

# Double shrinkage estimator

Following review about the double shrinkage estimator is due to Zhao [1].

Assume that there exists a statistic  $S_e^2$  independent of  $w_e$  that contains information of  $\sigma_e^2$ . It follows to assume in general that  $S_e^2|\sigma_e^2 \sim \sigma_e^2 \frac{\chi_{d_e}^2}{d_e}$  where  $d_e$  represents the degrees of freedom corresponding to the  $e^{th}$  statistic  $S_e^2$ . Now approximate  $\log\left(\frac{\chi_{d_e}^2}{d_e}\right)$  to follow  $N(m_e, \sigma_{ch,e}^2)$  distribution with mean  $m_e = E\left(\log\left(\frac{\chi_{d_e}^2}{d_e}\right)\right) = \psi\left(\frac{d_e}{2}\right) - \log\left(\frac{d_e}{2}\right)$  and variance  $\sigma_{ch,e}^2 = V\left(\log\frac{\chi_{d_e}^2}{d_e}\right) = \frac{d}{dx}\psi\left(\frac{d_e}{2}\right)$ , where  $\psi(x) = \frac{d}{dx}\log(\Gamma(x))$  is known as the digamma function.

Assumption of  $S_e^2|\sigma_e^2 \sim \sigma_e^2 \frac{\chi_{d_e}^2}{d_e}$  and the approximation of  $\log\left(\frac{\chi_{d_e}^2}{d_e}\right)$  to follow  $N(m_e, \sigma_{ch,e}^2)$  distribution follows that

$$\log(S_e^2)|\log(\sigma_e^2) \sim N(m_e + \log(\sigma_e^2), \sigma_{ch,e}^2). \quad (S1)$$

Furthermore, this model assumes that  $\log(\sigma_e^2)$  is a normal random variable with unknown mean  $\mu_v$  and variance  $\tau_v^2$ . Thus,

$$\log(\sigma_e^2) \sim N(\mu_v, \tau_v^2). \quad (S2)$$

Combining the information from equations 1 and 2 provides the following equation for  $\log(\sigma_e^2)|\log(S_e^2)$  as

$$\log(\sigma_e^2)|\log(S_e^2) \sim N\left(M_{v,e}(\log(S_e^2) - m_e) + (1 - M_{v,e})\mu_v, M_{v,e}\sigma_{ch,e}^2\right), \quad (S3)$$

where  $M_{v,e} = \tau_v^2/(\tau_v^2 + \sigma_{ch,e}^2)$ . Thus, the shrinkage variance estimate of  $\sigma_e^2$  can be obtained as the posterior mean from equation 3 as

$$\sigma_e^2 = \exp\left(M_{v,e}(\log(S_e^2) - m_e) + (1 - M_{v,e})\mu_v\right) \quad (S4)$$

under the assumption that both  $\mu_v$  and  $\tau_v^2$  are known. However,  $\mu_v$  and  $\tau_v^2$  are assumed unknown in the model.

Now focus on obtaining an empirical Bayes estimator  $\hat{\sigma}_e^2$  using the estimated  $\hat{\mu}_v$  and  $\hat{\tau}_v^2$  from the data. It can be obtained from equation 1 that

$$\log(S_e^2) - m_e|\log(\sigma_e^2) \sim N(\log(\sigma_e^2), \sigma_{ch,e}^2).$$

It is known from equation 2 that

$$\log(\sigma_e^2) \sim N(\mu_v, \tau_v^2).$$

Thus, using conditional expectation and conditional variance, it can be shown that  $E(\log(S_e^2) - m_e) = \mu_v$  and  $E(\log(S_e^2) - m_e)^2 = \mu_v^2 + \tau_v^2 + \sigma_{ch,e}^2$ . Hence, the model estimates  $\mu_v$  by  $\hat{\mu}_v = \frac{1}{n} \sum_{e=1}^n (\log(S_e^2) - m_e)$  and  $\tau_v^2$  by

$\hat{\tau}_v^2 = \left( \frac{1}{n} \left( \sum_{e=1}^n (\log(S_e^2) - m_e)^2 - \hat{\mu}_v^2 - \sigma_{ch,e}^2 \right) \right)_+$ , where  $+$  indicates the positive part estimator in standard notation. Hence,  $M_{v,e}$  can be estimated as  $\hat{M}_{v,e} = \hat{\tau}_v^2 / (\hat{\tau}_v^2 + \sigma_{ch,e}^2)$  and the empirical Bayes estimator of  $\sigma_e^2$  can be derived as

$$\hat{\sigma}_e^2 = \exp\left(\hat{M}_{v,e}(\log(S_e^2) - m_e) + (1 - \hat{M}_{v,e})\hat{\mu}_v\right).$$

Next step is to estimate  $\mu$  and  $\tau^2$  of the assumed distribution of  $\theta_e$  for  $e = 1, 2, 3, \dots, n$ . Revise that the model begins with assuming  $w_e | \theta_e, \sigma_e^2 \sim N(\theta_e, \sigma_e^2)$  and  $\theta_e \sim N(\mu, \tau^2)$  for  $e = 1, 2, 3, \dots, n$ . It can be shown following the same procedure discussed above for estimating  $\mu_v$  and  $\tau_v^2$  that

$$E(w_e | \sigma_e^2) = \mu \text{ and} \tag{S5}$$

$$E(w_e - \mu)^2 | \sigma_e^2 = \sigma_e^2 + \tau^2. \tag{S6}$$

From equation 5,  $\mu$  can be estimated by the weighted average as  $\hat{\mu} = \sum_{e=1}^n \frac{w_e / \hat{\sigma}_e^2}{\sum_{e=1}^n 1 / \hat{\sigma}_e^2}$ . Equation 6 can be used to estimate  $\tau^2$  as  $\hat{\tau}^2 = \left( \frac{\sum_{e=1}^n (w_e - \hat{\mu})^2 - \hat{\sigma}_e^2}{n} \right)_+$ . According to Zhao [1], this estimator  $\hat{\tau}^2$  can be inconsistent for  $\tau^2$  as  $n \rightarrow \infty$ . Therefore, following estimator  $\hat{\tau}^2 = \left( \frac{\sum_{e=1}^n (w_e - \hat{\mu})^2 - S_e^2 \exp(-m_e - \sigma_{ch,e}^2/2)}{n} \right)_+$  is suggested in the model. Two estimators of  $\hat{\tau}^2$  and  $\hat{\sigma}_e^2$  are used to derive double shrinkage estimator of  $\theta_e$  as discussed in the paper.

## Reference

- [1] Zhigen Zhao. "Double shrinkage empirical Bayesian estimation for unknown and unequal variances". In: *Statistics and Its Interface* 3.4 (2010), pp. 533–541.