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Analysis of a Hand-Foot-Mouth Disease Model with Standard Incidence Rate and Estimation for Basic Reproduction Number

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Abstract: Firstly, an SEIR mathematical model with standard incidence rate is established to describe the transmission of Hand-Foot-Mouth disease (HFMD). The equilibrium of the nondimensionalized model is calculated and the basic reproduction number of the model is defined. In addition, the local stability of the equilibrium is analyzed via the characteristic roots of the Jacobian matrix at the equilibrium, respectively. Numerical simulations are given to confirm the theoretical results. Secondly, a formula aimed to estimate the basic reproduction number of the transmission of HFMD is deduced. As examples to make use of the formula, the basic reproduction number of the HFMD transmission of Singapore of years 2015 and 2016 is estimated based on the newly infected cases notified by the surveillance organizations, respectively. The formula can realize real time estimation for the basic reproduction number and does not need to estimate the transmission efficiency of HFMD between individuals.

Keywords: Hand-Foot-Mouth disease; mathematical model; basic reproduction number; local asymptotical stability; estimation

1. Introduction

Hand-Foot-Mouth disease (HFMD) is caused by viruses that belong to the enterovirus genus (group), and it is a common disease of early childhood. The group of viruses includes polioviruses, coxsackieviruses, echoviruses and enteroviruses. Among these types of viruses, coxsackievirus A16 (COX A16) and enteroviruses 71 (EV71) are the most common ones to induce HFMD [1]. Since the first HFMD case was reported in New Zealand in 1957, this disease is now endemic worldwide. For example, there are 488,955 cases reported in mainland China in 2008 with an ill-death rate of 0.26/1000; and in 2009, there were 1,155,525 cases reported. Due to the high infectiousness of HFMD, it has been included into the Communicable Surveillance Network of China since 2008 [2].

Mathematical models are now applied to analyze the transmission and control of HFMD aimed to enable a better decision making for health policy makers [1,3–10]. For example, the following HFMD model was constructed in [1]

$$\begin{cases} \frac{dS}{dt} = \Lambda - \beta SI - \mu S + \delta R, \\ \frac{dE}{dt} = \beta SI - (\mu + k)E, \\ \frac{dI}{dt} = kE - (\mu + q + \epsilon_1 + \gamma_1)I, \\ \frac{dQ}{dt} = qI - (\mu + \epsilon_2 + \gamma_2)Q, \\ \frac{dR}{dt} = \gamma_1 I + \gamma_2 Q - (\mu + \delta)R. \end{cases}$$
(1)

and the model

$$\begin{cases}
\frac{dS}{dt} = \alpha N - \frac{(\beta_a + \beta_s)SI}{N} - \mu S, \\
\frac{dE}{dt} = \frac{\beta_a SI}{N} - (\mu + \lambda)E, \\
\frac{dI}{dt} = \frac{\beta_s SI}{N} - (\mu + \lambda)I, \\
\frac{dR}{dt} = \lambda (E + I) - \mu R,
\end{cases}$$
(2)

was established in [7]. In Models (1) and (2), the variable S, E, I, R, Q means the number of the susceptible, the exposed, the infective, the recovered, the quarantined individuals, respectively. The basic reproduction number, the local and global stability of the equilibrium, and the optimal control of Model (1) were analyzed in [1]. We note that only numerical simulations were given in [7] to analyze the transmission of HFMD.

The bilinear incidence rate was adopted in Model (1), while the standard incidence rate was introduced in Model (2). For infectious disease models to describe the transmission among individuals in a heterogeneous environment, the standard incidence rate is more reasonable [11]. Motivated by this, in this paper, we will establish a mathematical model with standard incidence rate to describe the transmission of HFMD diseases, and give the definition of the basic reproduction number of the model. Moreover, we will analyze the local stability of the equilibrium.

The basic reproduction number is related to the severity of communicable diseases. Usually, the disease is more urgent when the disease has a larger basic reproduction number. Different control measures are taken based on the severity of the disease that can be quantified by its basic reproduction number. Hence, the problem of how to estimate the value of the basic reproduction number during the prevalence of a disease with observable data received more and more attention [12,13]. The other aim of this paper is to provide a formula (method) to estimate the value of basic reproduction number of HFMD during its prevalence.

The paper is organized as follows. In Section 2, we give the HFMD transmission model between individuals with standard incidence rate. The basic reproduction number of the model is defined and the stability of the equilibrium of the model is analyzed. Numerical simulations are given to verify the theoretical results. In Section 3, we deduce a formula to estimate the basic reproduction number and further this formula is applied to the transmission of HFMD in Singapore of years 2015 and 2016. Both yearly and real time estimation for the basic reproduction number are estimated. Brief discussions are given in the last section.

2. Mathematical Model and Analysis

2.1. The Model

In this section, we first establish the following model with a standard incidence rate in light of Models (1) and (2)

$$\begin{cases} \frac{dS}{dt} = b - \frac{\beta SI}{N} - \mu S + \delta R, \\ \frac{dE}{dt} = \frac{\beta SI}{N} - (\mu + \sigma) E, \\ \frac{dI}{dt} = \sigma E - (\mu + \gamma) I, \\ \frac{dR}{dt} = \gamma I - (\mu + \delta) R. \end{cases}$$
(3)

The biological implications of the parameters and variables of Model (3) are listed in Table 1. With the biological background of Model (3) concerned, it is natural to suppose that all the parameters of Model (3) are positive.

Since N = S + E + I + R, Model (3) can be rewritten as the following model

$$\begin{cases} \frac{dS}{dt} = b - \frac{\beta SI}{N} - \mu S + \delta(N - S - E - I), \\ \frac{dE}{dt} = \frac{\beta SI}{N} - (\mu + \sigma)E, \\ \frac{dI}{dt} = \sigma E - (\mu + \gamma)I, \\ \frac{dN}{dt} = b - \mu N. \end{cases}$$

$$(4)$$

We further introduce the following transformations of variables

$$s = \frac{S}{b/\mu}$$
, $e = \frac{E}{b/\mu}$, $i = \frac{I}{b/\mu}$, $n = \frac{N}{b/\mu}$

hence, Model (4) is transformed to the following nondimensionalized model

$$\begin{cases}
\frac{ds}{dt} = \mu - \frac{\beta si}{n} - \mu s + \delta(n - s - e - i), \\
\frac{de}{dt} = \frac{\beta si}{n} - (\mu + \sigma)e, \\
\frac{di}{dt} = \sigma e - (\mu + \gamma)i, \\
\frac{dn}{dt} = \mu - \mu N.
\end{cases} (5)$$

Table 1. The biological implications of the parameters and variables of Model (3).

Parameter or Variable	Biological Implications	Value
S	number of susceptible individuals	variable
E	number of exposed individuals	variable
I	number of infective individuals	variable
R	number of recovered individuals	variable
N	total number of individuals, $= S + E + I + R$	variable
b	recruitment rate	demographical
β	transmission efficiency rate	estimated
μ	natural death rate	demographical
δ	loss of immunity rate	0.07 [1,4]
$1/\sigma$	latent duration	estimated
γ	recovery rate	estimated

It is clear that the solutions of Model (5) with nonnegative initial values remain nonnegative for all t > 0. Moreover, we define the basic reproduction number R_0 of Model (5) as

$$R_0 = \frac{\beta \sigma}{(\gamma + \mu)(\sigma + \mu)}.$$
(6)

We have the following results on the existence of equilibrium for Model (5), which can be obtained via direct computation, we omit the details.

Theorem 1. (i) If $R_0 \le 1$, then Model (5) has only a disease free equilibrium $E_0(s_0, e_0, i_0, n_0)$ which satisfy $s_0 = 1, e_0 = 0, i_0 = 0, n_0 = 1$; (ii) If $R_0 > 1$, then Model (5) has a positive equilibrium $E^*(s^*, e^*, i^*, n^*)$ and the disease free equilibrium E_0 , where

$$s^* = \frac{1}{R_0}, e^* = \frac{(\gamma + \mu)(\mu + \delta)(R_0 - 1)}{\sigma(\beta + \frac{\beta\sigma\delta}{\sigma + \mu} + R_0\delta)}, i^* = \frac{(\mu + \delta)(R_0 - 1)}{\beta + \frac{\beta\sigma\delta}{\sigma + \mu} + R_0\delta}, n^* = 1.$$
 (7)

2.2. Local Stability Analysis

In this part, we will analyze the local asymptotical stability of the disease free equilibrium E_0 and the positive equilibrium E^* of Model (5). The method used is to analyze the sign of the real parts of characteristic roots of the Jacobian matrix at the corresponding equilibrium. If the real parts of all the characteristic roots of the Jacobian matrix are negative, then the respective equilibrium is locally asymptotically stable; if there is a characteristic root of the Jacobian matrix that has positive real part, then the respect equilibrium is not stable [1]. To obtain the sign of the characteristic roots of a matrix, the Routh-Hurwitz criteria are often applied [11].

Denote the right side of each equation of Model (5) as f_1 , f_2 , f_3 , f_4 , respectively. The Jacobian matrix of Model (5) is as follows

$$\mathbf{J} = \begin{pmatrix} \frac{\partial f_1}{\partial s} & \frac{\partial f_1}{\partial e} & \frac{\partial f_1}{\partial i} & \frac{\partial f_1}{\partial i} & \frac{\partial f_1}{\partial n} \\ \frac{\partial f_2}{\partial s} & \frac{\partial f_2}{\partial e} & \frac{\partial f_2}{\partial i} & \frac{\partial f_2}{\partial n} \\ \frac{\partial f_3}{\partial s} & \frac{\partial f_3}{\partial e} & \frac{\partial f_3}{\partial i} & \frac{\partial f_3}{\partial n} \\ \frac{\partial f_4}{\partial e} & \frac{\partial f_4}{\partial e} & \frac{\partial f_4}{\partial i} & \frac{\partial f_4}{\partial n} \end{pmatrix} = \begin{pmatrix} -\frac{\beta si}{n} - \mu - \delta & -\delta & -\frac{\beta si}{n} - \delta & \frac{\beta si}{n^2} \\ \frac{\beta si}{n} & -(\mu + \sigma) & \frac{\beta s}{n} & -\frac{\beta si}{n^2} \\ 0 & \sigma & -(\gamma + \mu) & 0 \\ 0 & 0 & 0 & -\mu \end{pmatrix}$$
(8)

Theorem 2. If $R_0 < 1$, then the disease free equilibrium E_0 is locally asymptotically stable; If $R_0 > 1$, then the disease free equilibrium E_0 is unstable.

Proof. In view of Equation (8), the Jacobian matrix, $J(E_0)$, at the disease free equilibrium E_0 is

$$\mathbf{J}(E_0) = \begin{pmatrix} -\mu - \delta & -\mu & -\beta - \delta & 0 \\ 0 & -(\sigma + \mu) & \beta & 0 \\ 0 & \sigma & -(\gamma + \mu) & 0 \\ 0 & 0 & 0 & -\mu \end{pmatrix}.$$

Hence, the characteristic equation of $J(E_0)$ is

$$|\lambda \mathbf{I} - \mathbf{J}(E_0)| = (\lambda + \mu)(\lambda + \mu + \delta)\{\lambda^2 + [(\sigma + \mu) + (\gamma + \mu)]\lambda + (\sigma + \mu)(\gamma + \mu) - \beta\sigma\} = 0, \quad (9)$$

where **I** is the unity matrix. From Equation (9), it is clear that the two characteristic roots of $J(E_0)$ is $\lambda_1 = -\mu$, $\lambda_2 = -(\mu + \delta)$, that are both negative. The other two characteristic roots are determined by the following equation

$$\lambda^2 + [(\sigma + \mu) + (\gamma + \mu)]\lambda + (\sigma + \mu)(\gamma + \mu) - \beta\sigma = 0. \tag{10}$$

If we denote the two roots of (10) as λ_3 , λ_4 , respectively, then

$$\lambda_3 + \lambda_4 = -[(\sigma + \mu) + (\gamma + \mu)] < 0,$$
 (11)

$$\lambda_3 \lambda_4 = (\sigma + \mu)(\gamma + \mu) - \beta \sigma. \tag{12}$$

From Equation (6), the definition of R_0 , we have $\lambda_3\lambda_4>0$ provided that $R_0<1$. Thus, the real parts of λ_3 and λ_4 are both negative. From above, all the roots of Equation (9) have negative real parts if $R_0<1$. It follows that the disease free equilibrium E_0 is locally asymptotically stable if $R_0<1$.

If $R_0 > 1$, then $(\sigma + \mu)(\gamma + \mu) - \beta \sigma < 0$. Consider the function

$$F(\lambda) = \lambda^2 + [(\sigma + \mu) + (\gamma + \mu)]\lambda + (\sigma + \mu)(\gamma + \mu) - \beta\sigma,$$

we have $F(0) = (\sigma + \mu)(\gamma + \mu) - \beta \sigma < 0$ and $\lim_{\lambda \to +\infty} F(\lambda) = +\infty$. Therefore, there is a positive root of $F(\lambda)$. That is, there is characteristic root with positive real part for $J(E_0)$ if $R_0 > 1$, hence, the disease free equilibrium E_0 is unstable when $R_0 > 1$. \square

Theorem 3. If $R_0 > 1$, then the positive equilibrium E^* of Model (5) is locally asymptotically stable.

Proof. The Jacobian matrix $J(E^*)$ at the disease free equilibrium E^* is

$$\mathbf{J}(E^*) = \begin{pmatrix} -\beta i^* - \mu - \delta & -\delta & -\beta s^* - \delta & \beta s^* i^* \\ \beta i^* & -(\sigma + \mu) & \beta s^* & -\beta s^* i^* \\ 0 & \sigma & -(\gamma + \mu) & 0 \\ 0 & 0 & 0 & -\mu \end{pmatrix}$$

which follows

$$|\lambda \mathbf{I} - \mathbf{J}(E^*)| = (\lambda + \mu)(a_3\lambda^3 + a_2\lambda^2 + a_1\lambda + a_0) = 0,$$
(13)

where

$$a_{3} = 1, a_{2} = \delta + \gamma + 3\mu + \sigma + \beta i^{*},$$

$$a_{1} = 3\mu^{3} + \delta\gamma + 2\delta\mu + \delta\sigma + \gamma\sigma + 2\mu\sigma + 2\beta i^{*}\mu + \beta i^{*}\sigma - \beta s^{*}\sigma + \beta\delta i^{*} + \beta\gamma i^{*},$$

$$a_{0} = \gamma\mu^{2} + \mu^{2}\sigma + \mu^{3} + \delta\gamma\mu + \delta\mu\sigma + \gamma\mu\sigma + \beta i^{*}\mu^{2} + \beta\delta i^{*}\sigma + \beta\gamma i^{*}\delta + \beta i^{*}\mu\sigma - \beta\mu s^{*}\sigma + \beta\delta\gamma i^{*} + \beta\gamma i^{*}\mu,$$

and s^* , i^* are given in Equation (7).

Noting that

$$\beta s^* \sigma = \sigma \gamma + \mu \gamma + \mu^2 + \sigma \mu, \ \beta \mu s^* \sigma = \mu \sigma \gamma + \mu^2 \gamma + \mu^3 + \sigma \mu^2,$$

it is consequent that $a_k > 0$, k = 0, 1, 2, 3. Further, $a_1 a_2 - a_0 a_3 > 0$ can also be verified directly. In fact,

$$\begin{split} a_1a_2 - a_0a_3 &= 2\mu^2\delta + \delta^2\gamma + 2\delta^2\mu + \gamma\delta\mu + \delta^2\sigma + 2\beta\delta i^*\mu + \beta\delta i^*\sigma + \beta\delta^2 i^* \\ &+ \beta\delta\gamma i^* + 2\mu^2\gamma + \delta\gamma^2 + 2\delta\mu\gamma + \gamma^2\mu + \gamma\mu\sigma + \beta\gamma\mu i^* + \beta\gamma\delta i^* \\ &+ \beta\gamma^2 i^* + 6\mu^3 + 2\delta\gamma\mu + 6\mu^2\delta + 3\mu^2\gamma + 2\delta\gamma\mu + 6\mu^2\delta + 3\mu^2\gamma \\ &+ 3\mu\delta\sigma + 3\mu^2\delta + 6\mu^2\beta i^* + 3\beta i^*\sigma\mu + 3\beta\mu\delta i^* + 2\beta\mu\gamma i^* + 2\mu^2\sigma \\ &+ \delta\gamma\sigma + 2\delta\mu\sigma + \gamma\mu\sigma + \delta\sigma^2 + \mu\sigma^2 + 2\beta\sigma\mu i^* + \beta i^*\sigma^2 + \beta\delta\sigma \\ &+ \beta\gamma\sigma i^* + \mu^2\beta i^* + 2\delta\mu\beta i^* + \beta i^*\mu\sigma + 2\beta^2 i^{*2}\mu + \beta^2 i^{*2}\sigma \\ &+ \beta^2 i^{*2}\delta + \beta^2 i^{*2}\gamma. \end{split}$$

Hence, all the solutions of (13) have negative real parts by the Routh-Hurwitz criteria. Thus, E^* is locally asymptotically stable. \Box

2.3. Numerical Simulation of Stability

In this part, we give two numerical examples to verify the stability results. Firstly, we set $\beta = 0.1$, $\gamma = 0.117$, $\sigma = 0.15$, $\mu = 0.000039$, $\delta = 0.07$. From Equation (6), we have $R_0 = 0.8542 < 1$. Hence, from Theorem 2, the disease free equilibrium of Model (5) is locally asymptotically stable. This can be observed in Figure 1a since the number of the infected individuals tends to 0 as $t \to +\infty$.

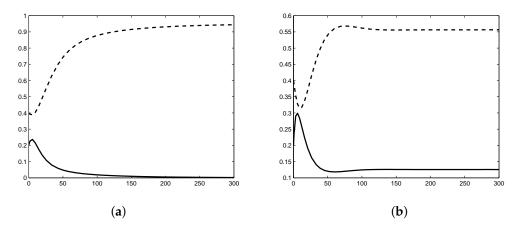


Figure 1. The dash line is the variation of s, while the solid line is the variation of i. (a) The local stability of the disease free equilibrium when $R_0 < 1$. s tends to 1 and i tends to 0 as $t \to +\infty$; (b) The local stability of the positive equilibrium when $R_0 > 1$. s tends to $s^* = 1/R_0 = 0.5853$ and i tends to $i^* \approx 0.1488 > 0$ as $t \to +\infty$.

Secondly, we set $\beta=0.2$, $\gamma=0.117$, $\sigma=0.25$, $\mu=0.000039$, $\delta=0.07$. From Equation (6), we have $R_0=1.7086>1$. Hence, from Theorem 3, the positive equilibrium of Model (5) is locally asymptotically stable. It can be observed in Figure 1b since the number of the infected individuals tends to 0.1488 as $t\to +\infty$.

3. Estimation of Basic Reproduction Number

3.1. Estimation Formula

In this part, we will derive a formula that can be applied for the estimation of the basic reproduction number with the infected cases periodically notified by the surveillance organizations. We begin with Model (4). Noting that the number of the infected I is much smaller than the total population size N in a region, we set $S \approx N$, thus we rewrite the second and the third equation of Model (4) as follows

$$\begin{cases}
\frac{dE}{dt} = -(\sigma + \mu)E + \beta I, \\
\frac{dI}{dt} = \sigma E - (\gamma + \mu)I.
\end{cases}$$
(14)

Further, we suppose the increasing of the accumulation of the exposed and the infected is exponential [14], that is

$$E = a \exp(\lambda t), I = b \exp(\lambda t), \tag{15}$$

where λ is the force of infection and a and b are constants determined by the initial values of *E* and *I*. Since Model (14) is linear, λ is the principal eigenvalue of the matrix

$$A = \begin{pmatrix} -(\sigma + \mu) & \beta \\ \sigma & -\gamma + \mu \end{pmatrix}.$$

It is direct to verify that one eigenvalue of A is positive and the other one of A is negative when $R_0 > 1$, hence, the value of λ is given by

$$\lambda = \frac{-[(\sigma + \mu) + (\gamma + \mu)] + \sqrt{[(\sigma + \mu) + (\gamma + \mu)]^2 - 4[(\sigma + \mu)(\gamma + \mu) - \beta\sigma]}}{2},$$
 (16)

and from Equation (6) it can be further written as the following

$$\lambda = \frac{-[(\sigma + \mu) + (\gamma + \mu)] + \sqrt{[(\sigma + \mu) + (\gamma + \mu)]^2 - 4(\sigma + \mu)(\gamma + \mu)(1 - R_0)}}{2}.$$
 (17)

Thus, from Equation (17) we have

$$R_0 = 1 + \frac{\lambda^2 + \lambda(\sigma + \gamma + 2\mu)}{(\sigma + \mu)(\gamma + \mu)}.$$
 (18)

The Equation (18) is the formula can be used for the estimation for the basic reproduction number of the transmission of HFMD during the prevalent stage of the disease. The parameter λ in Equation (18) can be evaluated via the newly infected cases notified by the surveillance organizations periodically. We add these data to get the accumulation of the infected cases, followed λ is estimated via curve fitting. The other parameters, σ , μ , γ , in Equation (18) can be obtained via demographical or pathological statistics.

3.2. Application of the Estimation Formula

In this part, we give an example to show how to apply Equation (18) to obtain the estimation value of the basic reproduction number of the transmission of HFMD. The newly infected cases of HFMD in Singapore are periodically notified by Ministry of Health of Singapore that can be accessed

from the website [15]. The newly infected cases of HFMD of years 2015 and 2016 of each e-weeks (epidemological weeks, consisting of 7 days) are depicted in Figure 2.

We first show how to estimate the value of λ in Equation (18). Noting that λ is the force of infection, which is the exponential increasing rate of the accumulation of the newly infected cases, we add the newly infected cases from the 1st e-week to the j^{th} e-week of one year, where $j=2,\cdots,m$ and m is the last e-week of the year. Thus, we obtain the accumulation vector of the newly infected cases with respect to each e-week of that year consisting of m dimensions. We denote this vector as y. We define an m dimensional vector, $x=(7,14,21,\cdots,m\times7)$, as the time vector whose unit is days. Followed the MATLAB curve fitting toolbox is applied to fit y with respect to x as $y=a\exp(bx)$. Therefore, the value of b is the estimation of b. An b0 and 1 and the curve fitting is better if it merits a larger b1 and b2 are value. It is clear that the values of b3 and b4 are different if we apply curve fitting in different time intervals from the b4 e-week to the b5 e-week of one year. Hence we select the value of b5 that has the largest b6 as the estimation of b6 of that year.

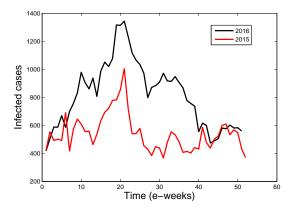


Figure 2. The newly infected cases of years 2015 and 2016 in Singapore notified by Ministry of Health of Singapore. The red line is the variations of the newly infected cases of year 2015, and the black line is that of 2016.

For the transmission of HFMD in Singapore of year 2015, we obtain the largest R – square value of 0.9692 when j = 22 and the corresponding value of b is 0.1317 (mean value of the 95% Confidence Interval). The curve fitting of y with respect to x from the 1st e-week to the 22nd e-week of year 2015 is depicted in Figure 3a. Thus, the estimation value of λ in year 2015 is

$$\lambda_{2015} = 0.1317. \tag{19}$$

Similarly, we can also obtain the largest R-square value of 0.9720 when j=22 for the transmission of HFMD in Singapore of year 2016. The curve fitting of y with respect to x from the 1st e-week to the 22nd e-week of year 2016 is depicted in Figure 3b. And the estimation value of λ (mean value of the 95% Confidence Interval) in year 2016 is

$$\lambda_{2016} = 0.1462. \tag{20}$$

Next, we determine the other parameters in Equation (18). The latent period of HFMD is 4–7 days [16]. Thus we set the average latent period as $1/\sigma = 5.5$ days, that is,

$$\sigma = \frac{1}{5.5} = 0.1818. \tag{21}$$

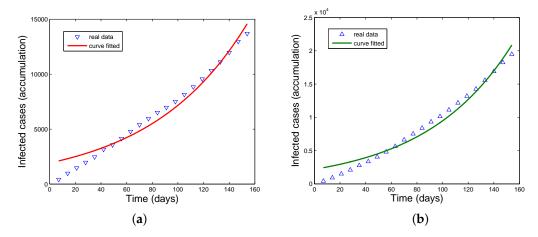


Figure 3. The accumulative infected cases fitted by the curve $y = a \exp(bx)$. (a) The year 2015; (b) The year 2016.

The period for an infected individual to recover from HFMD is 7–10 days [7]. We also set the average recovery period as $1/\gamma = 8.5$ days, that is,

$$\gamma = \frac{1}{8.5} = 0.1176. \tag{22}$$

The life expectancy at birth of Singapore of year 2015 is 82.7 years [17]. Thus, we set

$$\mu = \frac{1}{82.7 \times 365} = 0.00003313. \tag{23}$$

Now the parameters appeared in Equation (18) are all estimated. Substituting Equations (19), (21)–(23) into Equation (18), we obtain that the basic reproduction number of the transmission of HFMD of year 2015 in Singapore is

$$R_0(2015) = 1.1924. (24)$$

And from Equations (18), (20)–(23), the basic reproduction number of the transmission of HFMD of year 2016 in Singapore is

$$R_0(2016) = 1.2146. (25)$$

3.3. Real Time Estimation of R_0

In Section 3.2, we estimate the basic reproduction number R_0 of the transmission of HFMD in Singapore of years 2015 and 2016, respectively. We call this the yearly estimation of R_0 . From the estimation process, we know that the basic reproduction number can be estimated till we have the newly infected cases of each e-week in the whole year. However, it is more appropriate to estimate the basic reproduction number of the present e-week with the previous data used, that is, we should estimate the real time basic reproduction number. In this part, we elaborate on this topic.

In Section 3.2, we have supposed that the accumulation of the newly infected cases is exponential increasing (refer to Equation (15)) and they are solutions of the simplified Model (14). Theoretically, the accumulations of the newly infected cases with respect to time are corresponding to the integration of the newly infected cases. Hence, we also suppose that the increasing (or decreasing) of the newly infected cases is exponential since the derivation of the exponential function is still an exponential one. Based on this observation, we can realize the real time estimation of the basic reproduction number.

The details to realize the real time estimation is as follows. Suppose that we want to estimate the R_0 of the h^{th} e-week. We collect the newly infected cases, $y_{h-j+1}, y_{h-j+2}, \cdots, y_h$, of the

 $(h-j+1)^{th}$, $(h-j+2)^{th}$, \cdots , h^{th} e-week, and denote the data as a vector y. Denote x as the vector $(7,14,\cdots,j\times7)$ since the time unit is days. Followed, we fit y with respect to x as $y=a\exp(bx)$ to obtain $\lambda=b$. Substituting the estimated λ together with Equations (21)–(23) into Equation (18), we obtain the estimation of the basic reproduction number at the h^{th} e-week. With this process applied to different h, we obtain the real time estimation of R_0 . Figure 4 depicts the variations of R_0 of the transmission of HFMD in Singapore of years 2015 and 2016. We note that in Figure 4, we set j=6.

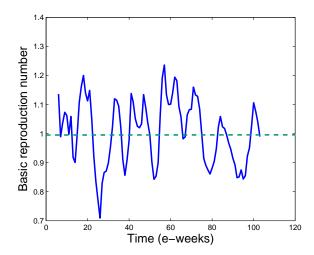


Figure 4. The real time estimation of the basic reproduction number of hand-foot-mouth disease (HFMD) in Singapore of years 2015 and 2016. The time zone is from the 6th e-week of 2015 to the last e-week of 2016.

4. Discussion

Mathematical models are now widely applied to analyze the transmission of communicable diseases. The basic reproduction number is the key parameter in theoretical analysis of the models. Usually, the disease free equilibrium is stable when the basic reproduction number is less than 1, and the endemic (positive) equilibrium exists which is stable when the basic reproduction number is larger than 1. In this paper, the HFMD model with standard incidence rate established has been proven to be consistent with this threshold property of the basic reproduction number.

The basic reproduction number is also important in practice. The value of this parameter is related to the severity of the transmission of the infectious diseases, can provide references for the strength of control measures should be taken for health policy makers. Hence, how to estimate the value of the basic reproduction number for communicable disease receives more and more attention by researchers [1,12–14,18]. In this paper, we provide a formula to estimate the basic reproduction number both yearly and real time for the transmission of HFMD. And the transmission of HFMD in Singapore of years 2015 and 2016 is as an example to show that the formula can be used in practice. It is estimated that the yearly basic reproduction number of HFMD in Singapore is 1.1924 of year 2015 and 1.2146 of year 2016. Both of them are greater than 1, which implies that HFMD was prevalent in Singapore during these two years.

In Section 3, we take the mean values of the parameters to estimate the basic reproduction number. However, in fact, some parameters in Equation (18) vary their values in intervals. That is, the latent period is 4–7 days, the recovery period is 7–10 days and the 95% Confidence Interval of λ of year 2015 is [0.01182, 0.01451], of year 2016 is [0.01316, 0.01608]. From Equation (18), we can obtain the yearly basic reproduction number of year 2015 is in the interval [1.1339, 1.2613] and of year 2016 is in the interval [1.1496, 1.2914].

We note that one should estimate the value of β to apply Equation (6) directly to estimate the basic reproduction number [1]. The estimation Formula (18) here does not need to estimate the value of β . The most urgent need to apply Equation (18) is in the newly-infected cases periodically notified by surveillance organizations.

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