

Article Financial Inclusion and Sustainable Growth in North African Firms: A Dynamic-Panel-Threshold Approach

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Abstract: This paper investigates the impact of financial inclusion on sustainable firm growth in Northern African countries (Egypt, Morocco, and Tunisia) during the period of 2007–2020. To this end, this study employs a dynamic panel threshold regression (DPTR) model. This model is a panel-data model that can capture different behaviors of data, depending on a threshold variable. The main results showed the existence of a threshold effect. This means that when financial inclusion is low (high), sustainable firm growth is limited (significant) due to the absence (presence) of appropriate financing, information, and financial tools. However, the levels of financial inclusion in North African countries are insufficient and require improvement. Therefore, it is essential for policymakers and managers to continue to promote the quality of financial inclusion by improving access to financial services and the regulatory environment to facilitate firms' access to financing and support their sustainability.

Keywords: financial inclusion; sustainable firm growth; DPTR; North African countries



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1. Introduction

Sustainable growth is essential for any firm, as it can not only increase revenues and profits, but also contribute to the long-term viability of firms. However, achieving sustainable growth may require significant investment, which can be financed in various ways, including debt. However, not all firms can use debt as a source of financing due to constraints related to information asymmetry, such as adverse selection and moral hazard. As a result, some firms are unable to access credit, which limits their ability to grow (Stiglitz and Weiss 1981; Carreira and Silva 2010).

In recent years, financial inclusion has been the subject of intense academic and policy debates, especially in the context of firm growth (e.g., Lee et al. 2020; Nizam et al. 2021). A lack of financial inclusion can limit the growth potential of small and medium-sized enterprises (SMEs) by limiting their access to financing, insurance, and other financial services. This, in turn, can stifle innovation and reduce economic growth (Abdul Karim et al. 2022). In addition, financial inclusion can improve firms' access to traditional financing sources, such as banks and financial markets. This is due to a reduction in information asymmetries between borrowers and lenders, as well as better financial transparency (Beck et al. 2012; Donati 2016). Thus, financial inclusion widens the range of financing choices available to firms and enhances their ability to obtain funds, which can stimulate their growth.

North African governments, especially those of Egypt, Tunisia, and Morocco, are focusing on strategies to encourage financial inclusion for SMEs. This is important because many SMEs struggle to access traditional bank loans, limiting their growth and

job-creation potential. Although progress has been made in recent years, access to financial services remains low in the region, with only 36% of adults having a bank account (Demirgüç-Kunt et al. 2022). Figure 1 illustrates a map of the number of deposit accounts with commercial banks per 1000 adults (in Tunisia, it is 37%, in Egypt, it is 27%, and in Morocco, it is 44%). However, mobile banking and other digital financial services are becoming increasingly popular, especially among young people and those in rural areas. Tunisia, Morocco, and Egypt have implemented successful financial-inclusion programs, including partnerships with microfinance institutions and national strategies to increase access to financial services for all citizens, especially SMEs and startups. The Central Bank of Egypt has launched a new financial-inclusion strategy for 2022–2025, with a focus on expanding access to financial services.



Figure 1. Map of the number of deposit accounts in North African countries. Source: IMF FAS and authors' calculation.

Despite the efforts made by these three countries to promote financial inclusion, it has been found that levels of financial inclusion remain relatively low. This raises concerns about the effectiveness of current policies and initiatives to support access to financial services for individuals and firms. In highlighting the low levels of financial inclusion in these countries, it is important to emphasize a critical issue that requires attention and appropriate action. In addition, by exploring the effects of financial inclusion on business growth in countries with low levels of inclusion, it is important to identify the specific barriers that limit access to financial services and hinder firms' development. This indepth understanding of the constraints encountered in these countries will enable the development of recommendations and targeted strategies to improve financial inclusion and stimulate sustainable firm growth.

Various theories have been set forth to explain the debt–firm-growth nexus, such as the agency theory (Jensen and Meckling 1976), the trade-off theory (Bradley et al. 1984), and the pecking-order theory (Myers and Majluf 1984). These theories focus on firms' financing choices by emphasizing the costs, benefits, and risks associated with each source of financing. However, in the era of financial inclusion, these theories are no longer sufficient to understand firms' financing choices and their impact on sustainable growth. Indeed, financial inclusion has a significant impact on firms' financing decisions by offering new sources of financing and improving forms' access to existing sources. Traditional

3 of 19

theories of capital structure do not take this new reality into account and need to be adapted to understand the effects of financial inclusion on firms' financing choices and their growth.

From an empirical viewpoint, the results of prior studies that focused on the relationship between financial inclusion and firm growth are mixed and inconclusive. On the one hand, the few studies that examined this relationship suggested a positive impact of financial inclusion on sales growth (e.g., Chauvet and Jacolin 2017; Lee et al. 2020). On the other hand, other studies indicated the existence of a nonlinear relationship between access to credit and sales growth (Nizam et al. 2021), and between leverage and sustainable firm growth (Akhtar et al. 2022). Hence, this study intends to fill this research gap by examining the threshold effect of financial inclusion on sustainable firm growth in North African countries. This threshold effect indicates two types of firm: those whose level of sustainable growth decreases as they should be more financially inclusive, but only up to a certain point (threshold), where lenders begin to extend credit to them because of high levels of financial inclusion; and those who exceed the optimal level, resulting in increase in their growth.

In this context, the main objective of this paper is to analyze the impact of financial inclusion on sustainable firm growth in the North African region. Specifically, this study investigates the threshold effect of financial inclusion on sustainable firm growth. It uses a sample of Northern African countries (Egypt, Morocco, and Tunisia) from the period of 2007–2020. It also uses the dynamic panel threshold regression (DPTR) model, run on a sample of listed North African firms. Furthermore, this study investigates the non-linear relationship between financial inclusion and sustainable firm growth by considering that managers-shareholders may react positively (negatively) to high (low) levels of financial inclusion. The main findings of this paper are that financial inclusion does not necessarily lead to increased sustainable firm growth, and that a certain threshold of financial inclusion must be reached by North African non-financial firms in order to enhance their growth. Most of the current studies indicate that an increase (decrease) in financial inclusion-level always leads to sustainable (unsustainable) firm growth. However, although several studies on financial inclusion and sustainable firm growth have been conducted in other countries, there are few studies on emerging regions such as North Africa (Egypt, Morocco, and Tunisia).

This study makes an important contribution to the literature by examining the non-linear relationship between financial inclusion and sustainable firm growth in North African countries. While many studies have addressed the relationship between financial inclusion and economic growth or sustainable economic growth (e.g., Ghosh 2011; Sethi and Acharya 2018; Emara and Said 2021; Pradhan et al. 2021; Akinrinola and Folorunso 2022; Ahmad et al. 2023; Adabor and Mishra 2023; Li and Wu 2023; Hussain et al. 2023), few have examined sustainable firm growth specifically (Yang and Zhang 2020). By focusing on the North African region, this study fills a research gap by providing new, region-specific knowledge. The findings may have important policy and practical implications for policy makers, financial institutions, and firms in these countries. Unlike previous research, which focused primarily on sales growth, this study uses the rate of sustainable firm growth as the dependent variable. Sustainable firm growth encompasses not only short-term financial performance, but also long-term financial performance (Khémiri and Noubbigh 2020, 2021; Nizam et al. 2021; Bui et al. 2021). By focusing on this more comprehensive measure of firm growth, this study seeks to assess how financial inclusion can influence long-term value creation, while considering one of the sustainable aspects of business development. This approach makes it possible to better assess the impact of financial inclusion on the overall performance and sustainability of firms. Finally, in contrast to previous studies, which focused on examining the linear relationship between financial inclusion and sustainable growth (Yang and Zhang 2020), this study uses a DPTR model to analyze the relationship between financial inclusion and sustainable firm growth. This model differs from previous approaches, which used curvilinear or static-panel-threshold regression models

(Nizam et al. 2021). The DPTR model is an advanced tool for identifying the threshold at which the level of change in financial inclusion has a significant impact on sustainable firm growth. By determining this threshold, this study provides valuable information on the specific conditions under which financial inclusion can have a significant effect on business performance. In addition, the econometric model used also makes it possible to identify differences in the threshold effect of financial inclusion on sustainable firm growth in different regimes (lower and higher). This provides a finer understanding of the complex relationship between financial inclusion and sustainable firm growth.

Although previous studies on the impact of financial inclusion on sustainable firm growth were conducted in other countries, there are few studies on developing countries, such as Egypt, Morocco, and Tunisia. Therefore, this research aims to answer the following research questions (RQ):

RQ1: Is there a critical threshold of financial inclusion beyond which firm sustainable growth in these countries is significantly stimulated?

RQ2: What are the specific barriers to financial inclusion faced by firms in these countries, and how can they be overcome?

The organization of this paper is as follows. Section 2 outlines the literature review and hypotheses, while Section 3 presents the methodology and data used in the study. Section 4 presents and discusses the findings, and Section 5 highlights the conclusion and implications, as well as the limitations of the study and future research directions.

2. Related Research and Hypothesis

According to Higgins (1977), the sustainable-growth rate is defined as the maximum speed at which a firm can increase its sales without exhausting its financial resources. This growth depends primarily on the expansion of its assets, which, in turn, is influenced by the rate of growth of its equity, provided that the firm maintains its operational efficiency and stable financial policies (Higgins 1977, 2003).

For more information on this concept, Van Horne (1987) offers a complementary perspective by emphasizing that the sustainable-growth rate is the maximum annual salesgrowth rate that a firm can achieve within predetermined operating, debt, and dividend ratios. These concepts are part of the financial theories on financial decisions and firm performance. More specifically, the relationship between financial inclusion and sustainable firm growth can be explained through the combination of sustainable-growth theory (Higgins 1977) and the literature related to leverage and firm performance. In fact, there is a strong relationship between the theories related to leverage and firm performance and the link between financial inclusion and sustainable firm growth. Both areas of research aim to optimize access to finance for firms. Traditional theories related to firm performance, such as the agency theory (Jensen and Meckling 1976), the trade-off theory (Bradley et al. 1984), and the pecking-order theory (Myers and Majluf 1984), seek to determine the most effective mechanisms for firms to optimize their financing structure and increase their performance. Similarly, financial inclusion aims to broaden the access of firms to a variety of traditional and alternative sources of financing to stimulate their growth. Therefore, although capitalstructure theories are more focused on firms' financing choices, while financial inclusion is more focused on the availability of financing sources, both areas share the common objective of improving firms' access to financing. Financial inclusion provides alternative sources of financing for firms due to information asymmetry (Beck et al. 2012). However, SMEs face external financing challenges due to the problem of asymmetric information (Stiglitz and Weiss 1981). According to the World Bank Group (2016), financial constraints, such as high concentration, low competition, and the predominance of public ownership in financial institutions, are major problems for SME financing. Financial inclusiveness, which encourages financial stability and provides opportunities for sound risk management and financial facilities, can address the financial-inclusion gap among SMEs and lead to significant growth. However, the very low rate of financial inclusion suggests that there is considerable untapped potential for developing better access to finance for SMEs.

In practice, regarding the use of credit to promote firm growth, there are two important factors to consider. The first is the ability to access credit, which several studies showed to be a significant factor in enhancing firm growth (e.g., Brown et al. 2005; Huynh and Petrunia 2010; Rahaman 2011; Ayaydın and Hayaloglu 2014; Akhtar et al. 2022). The second factor concerns lending relationships, which are another means through which to alleviate financing constraints. To protect themselves against the risk of default, lenders require restrictive clauses in the debt contract (Diamond 1984).

In this context, firm managers (especially in SMEs) always seek to improve lending relationships to receive credit (Rahaman 2011; Gopalan et al. 2011). Thus, firms evaluate their need for external financing based on the level of their sales. However, there is an opposing viewpoint, according to which access to credit is not beneficial for small firms (Oliveira and Fortunato 2006; Wu and Au Yeung 2012; Khémiri and Noubbigh 2020; Attia et al. 2023). In fact, small firms in particular face an increased risk of a lack of adequate repayment sources. Therefore, they seek to establish close lending relationships to improve their financing conditions. In addition, on the one hand, information asymmetries (adverse selection and moral hazard) limit the ability of firms to obtain bank loans (Stiglitz and Weiss 1981). On the other hand, non-performing loans or economic recessions may require bankers to limit the granting of loans to firms (Dimitrov and Tice 2006). This is likely to reduce firms' sales growth, mainly resulting from credit rationing (Carpenter and Petersen 2002; Carreira and Silva 2010; Beck et al. 2012).

However, recent studies examined the effects of financial inclusion on growth. These effects can be classified into two main levels: macroeconomic and microeconomic. At the macroeconomic level, previous studies investigated the linear and nonlinear links between financial inclusion, economic growth, and sustainable economic growth. Empirical studies on the linear correlation between financial inclusion and economic growth produced contradictory findings. Indeed, several previous studies showed that economic growth is driven by financial inclusion in both developed and developing countries (e.g., Ghosh 2011; Sethi and Acharya 2018; Emara and Said 2021; Ozturk and Ullah 2022; Akinrinola and Folorunso 2022; Ahmad et al. 2023; Adabor and Mishra 2023). In addition, other studies showed the importance of financial inclusion in determining sustainable economic growth in countries (e.g., Pradhan et al. 2021; Li and Wu 2023; Hussain et al. 2023). However, other studies revealed the opposite. In fact, these studies indicated that the inaccessibility or inadequacy of financial services is the main obstacle to the development of economies in developing countries, especially for financially limited firms or those located in remote regions, where financial services are not available (e.g., Abdul Karim et al. 2022; Chen et al. 2023). It should be noted that other recent studies explored the non-linearity between financial inclusion and economic growth, albeit using different econometric techniques (Abdul Karim et al. 2022; Daud and Ahmad 2023). These studies showed that increasing financial access for businesses and individuals leads to higher economic growth up to a certain threshold. However, beyond this threshold, this effect becomes negative in some economies. According to Rehman and Islam (2023), the financial system plays a crucial role in investment flows, as it facilitates the transfer of capital investments, creating value for multiple uses in the production process. Efficient financial systems use available resources to accelerate industrial productivity. Financial markets, which are integral parts of financial systems, stimulate income growth by reducing transaction costs and facilitating the distribution of resources. Consequently, financial development contributes to economic growth by encouraging capital accumulation and improving overall factor productivity.

The same finding is confirmed at the microeconomic level. More specifically, several studies have demonstrated a linear relationship between financial inclusion and firm growth, indicating that financial inclusion is a key determinant of firm growth (Donati 2016; Chauvet and Jacolin 2017; Lee et al. 2020; Yang and Zhang 2020; Marcelin et al. 2022). In addition, Nizam et al. (2021) demonstrated the existence of an optimal FI threshold (particularly credit access) below which FI stimulates sales growth, but beyond which it reduces it. Referring to previous research, it is noted that studies also found a non-

linear relationship between access to credit and sales growth (e.g., Coricelli et al. 2012; Molinari 2013; Khémiri and Noubbigh 2021; Akhtar et al. 2022). However, in the literature, the non-linear relationship between financial inclusion and sustainable firm growth has not been examined. This study aims to fill this gap while examining the threshold effect of financial inclusion on sustainable firm growth. Based on these different empirical studies, the hypotheses of this research are as follows:

H1. There is a threshold effect of financial inclusion on sustainable firm growth.

H1(a). Financial inclusion positively affects sustainable firm growth.

H1(b). Financial inclusion negatively affects sustainable firm growth.

3. Research Methodology

3.1. Data

The sample in our study included non-financial firms listed on North African stock markets, namely those of Tunisia, Egypt, and Morocco. There were several reasons for choosing these three countries. Firstly, these three countries have implemented ambitious digital transformation strategies, aiming to leverage the potential of digital technologies for industrialization, innovation, and competitiveness. They have invested in developing their digital infrastructure, skills, and ecosystems, as well as developing e-government, e-commerce, and e-learning (El Aynaoui et al. 2022). Secondly, the economies of other countries in the North African region (e.g., Algeria and Mauritania) are heavily dependent on natural resources, particularly the hydrocarbon sector. This can make their economies more vulnerable to fluctuations in oil and gas prices, and potentially less diversified than those of Egypt, Tunisia, and Morocco, which have sought to diversify their economic sectors. Finally, the selection of these three countries was influenced by the availability of data on financial inclusion, especially financial data on listed non-financial firms.

We used annual data from the period of 2007–2020, excluding financial institutions. To be included in the sample, firms needed to have at least three consecutive years of available data during the study period. After applying these criteria, we obtained a final sample of 155 firms, representing 11 different industries, as classified by Campbell (1996), with a total of 2170 firm–year observations. Table 1 presents the distribution of the firms in each country. To minimize the influence of outliers on our analysis, we winsorized all firm-level variables at the 1st and 99th percentiles.

 Table 1. Distribution of firms.

Countries	Number of Firms	Number of Observations	%
Egypt	91	1274	58.71%
Morocco	39	546	25.16%
Tunisia	25	350	16.13%
Total	155	2170	100%

Source: authors' calculation.

This choice is particularly important because our study period coincided with major changes that affected the economies of the North African region, as well as the business sector. Indeed, the subprime crisis of 2007 had negative effects on the region's economies, and the post-2010 period was marked by significant political and economic unrest, particularly with the Arab Spring, which hindered the growth of some countries in our sample, such as Egypt and Tunisia. In addition, the COVID-19 pandemic had a negative impact on the economies in the North African region, as well as on the global economy. Finally, we excluded insurance companies, financial institutions, and banks from our sample because they have different accounting methods, governance practices, and financial structures from non-financial firms. In this study, data were collected from different sources. The Refinitiv Eikon database was used to obtain information at the firm level, while macroeconomic data, such as gross domestic product (GDP)-growth rate, were extracted from the World Bank's World Development Indicators (WDI). Data related to the two dimensions of FI and the composite index were collected from International Monetary Fund (IMF)'s Financial Access Survey (FAS).

3.2. Econometric Model

To explore the nonlinear relationship between financial inclusion and sustainable firm growth in the North Africa region, we applied the DPTR model, developed by Seo and Shin (2016), considering potential endogeneity. To test the validity of our hypothesis, we used Equation (1), where sustainable firm growth (SFG) is considered as the dependent variable, while financial inclusion is considered as the variable of interest. We included several control variables in Equation (1) due to their potential effects on SFG. Since the growth in the current year is assumed to be the result of operations in the previous year (Khémiri and Noubbigh 2021), we established empirical model with a lag of one year for the variable of interest (FI). The empirical DPTR model's equation was then written as follows:

$$SFG_{cit} = (\beta_1 SFG_{cit-1} + \beta_2 FI_{ct} + \beta_3 SIZE_{cit} + \beta_4 TANG_{cit} + \beta_5 RISK_{cit} + \beta_6 NDTS_{cit} + \beta_7 LIQ_{cit} + \beta_8 GDP_{cit}) 1 \cdot (FI_{ct} \leq \gamma) + (\lambda_1 SGR_{cit-1} + \lambda_2 FI_{ct} + \lambda_3 SIZE_{cit} + \lambda_4 TANG_{cit} + \lambda_5 RISK_{cit} + \lambda_6 NDTS_{cit} + \lambda_7 LIQ_{cit4} + \lambda_8 GDP_{cit}) 1 \cdot (FI_{ct} > \gamma) + \varepsilon_{it}$$

$$(1)$$

where is SFG_{cit} the dependent variable, FI_{ct} is the time-varying regressor, $SIZE_{cit}$, $TANG_{cit}$, $RISK_{cit}$, $NDTS_{cit}$, LIQ_{cit} , and GDP_{cit} are the control variables, $1\{\cdot\}$ is an indicator function, and q_{it} is the threshold variable (FI_{ct}). The γ denotes the threshold parameter. The $\varepsilon_{it}(\varepsilon_{it} = \mu_i + \nu_{it})$ are the error components, where μ_i is the individual fixed effects and ν_{it} is the idiosyncratic random disturbance. The β and λ denote the coefficients of all exogenous variables for the lower and upper regimes, respectively. For more details, please refer to Seo and Shin (2016).

3.3. Definition of Variables

3.3.1. Dependent Variable

The SFG rate represents the rate at which a firm can finance its growth using its own funds without borrowing from banks or other financial institutions. This rate is commonly used to forecast asset acquisitions, cash-flow projections, and borrowing strategies, measure long-term competitiveness and profitability, and evaluate the sustainability of long-term growth. According to Akhtar et al. (2022), the Higgins model (Higgins 1977) is used to calculate the SGR, which is defined as follows:

$$SFG = NPR \times ATR \times RR \times EM \tag{2}$$

where *NPR* is the net-profit ratio calculated by dividing the profit after tax by net sales, *ATR* is the asset-turnover ratio, calculated by dividing net sales by total assets, *RR* is the retention ratio, calculated by dividing retained earnings by net profit, and *EM* is the equity multiplier, calculated by dividing total assets by total equity.

3.3.2. Independent Variable

Following Wang and Luo (2022), we used two dimensions (access and usage) to measure financial inclusion. For the first dimension, we used the number of deposit accounts with commercial banks per 1000 adults, the number of ATMs per 100,000 adults (demographic penetration), and the number of ATMs per 1000 square kilometers (geographic penetration). For the second dimension, we also used two indicators, namely the outstanding deposits with commercial banks (% of GDP) and the outstanding loans from commercial banks (% of GDP). To construct the composite indicator, we followed Wang and Luo (2022) by using principal component analysis (PCA). Before performing the PCA, the indicators employed from each dimension must be normalized using the min-max normalization technique. Since the access dimension was composed of two indicators and the usage dimension was also composed of two indicators, we initially captured the common variation between the two indicators of each dimension using PCA and constructed the two dimensions of access and usage. We then applied PCA to extract principal component (PC) common to the two dimensions.

Two specific tests were performed in order to test the data's applicability for factor analysis: Bartlett's test of sphericity and the Kaiser-Meyer-Olkin (KMO) test. The results of these two tests are reported in Table A1. The *p*-values of the Bartlett sphericity test were all less than 0.05 (0.00), and the KMO values were equal to 0.50, which shows that the data were applicable to the PCA. Panel B in Table 2 shows the results of the PCA. Regarding the access dimension, the eigenvalues of the three PCs were 2.745, 0.217, and 0.038, respectively. These results reveal that the first PC explained about 91% of the variance of the corresponding sample. Except for the first PC, no other PC had an eigenvalue greater than one, indicating that it was suitable to take only the first PC and extract the access dimension using the weights (i.e., 0.698, 0.704, and -0.130) assigned to the first PC. Regarding the usage dimension, the eigenvalues of the two PCs were 1.196, and 0.804, respectively. These results show that the first PC explained about 96% of the variance of the corresponding sample. Apart from the first PC, no other PC had an eigenvalue greater than one, suggesting that it was crucial to take only the first PC and extract the usage dimension using the weights (i.e., 0.7071 and 0.7071) assigned to the first PC. Finally, for FI, the eigenvalues of the two PCs were 1.822 and 0.178, respectively, which indicated that the first PC explained about 91% of the variance of the corresponding sample. Furthermore, only the first PC had an eigenvalue greater than one, so we can suppose that the first PC appropriately explained the joint variation between the two dimensions.

 Table 2. Financial-inclusion indicators across North African countries.

Countries	Financial Inclusion Index	Access Dimension	Usage Dimension
Egypt	0.253	0.549	0.464
Morocco	0.418	0.447	0.511
Tunisia	0.294	0.483	0.323

Source: authors' calculation, Stata output.

Following Ahamed and Mallick (2019), we then constructed a multidimensional index for FI, as follows:

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$$I = \sum_{i=1}^{n} \tau_{ij} X_i \tag{3}$$

where τ_{ij} is the component weights and X_i is the original variables. We used the same weights (i.e., 0.7071) for both dimensions. To facilitate analysis, following Ahamed and Mallick (2019), we further normalized (using min–max normalization technique) this index on a scale of 0 to 1. Note that 0 indicates financial exclusion and 1 indicates financial inclusion.

Table 2 presents the financial-inclusion indicators across North African countries. The results suggest that the levels of financial inclusion in three countries are relatively low. However, it appears that the level of financial inclusion in Morocco, particularly access to credit, is higher than that of other countries in the region.

3.3.3. Control Variables

To assess the impact of financial inclusion on sustainable firm growth, we considered several control variables that were mentioned in previous studies. We employed firm size (size) to control financial constraints, determined by the natural logarithm of total assets. To assess the firm's ability to obtain external financing, we used the tangibility, calculated as the ratio of tangible assets to total assets (Khémiri and Noubbigh 2021). To control the

firm risk, we followed Akhtar et al. (2022) using Altman's Z-score (Altman 1983). This variable is calculated using the following formula:

$$Z = 1.2 \times \left(\frac{WC}{TA}\right) + 1.4 \times \left(\frac{NI}{TA}\right) + 3.3 \times \left(\frac{EBIT}{TA}\right) + 0.6 \times \left(\frac{BVE}{TBVL}\right) + 0.999 \times \left(\frac{SAL}{TA}\right)$$
(4)

where WC is working capital, TA is total assets, NI is net income, EBIT is earnings before interest and taxes, BE is book value of equity, TBVL is total book value of liabilities, and SAL is sales.

In addition, we included non-debt tax shields (NDTS), measured by the ratio of depreciation to total assets (Akhtar et al. 2022). To evaluate the firm's ability to meet its short-term commitments, we employed liquidity (LIQ) variable, measured as the current-assets-to-current-liabilities ratio (e.g., Khémiri and Noubbigh 2021; Akhtar et al. 2022). Finally, to control for macroeconomic stability and economic conditions, we employed GDP growth (Akhtar et al. 2022; Khémiri and Noubbigh 2021). All variables are presented in Table 3.

Table 3. Variable definition.

Variables	Acronyms	Definition	Source		
Dependent variable					
Sustainable firm growth	SFG	Net-profit ratio (profit after tax by net sales) multiplied by the asset-turnover ratio (net sales by total assets), multiplied by the retention ratio (retained earnings by net profit), multiplied by equity multiplier (total assets by total equity)	Refinitiv Eikon and authors' calculation		
Independent variable					
Financial-inclusion index	FI	Composite variable. PCA is applied to calculate the composite variable. FI represents FI indicator using min–max normalization.	IMF FAS and authors' calculation		
Control variables		-			
Firm size	SIZE	The natural logarithm of total assets	Refinitiv Eikon		
Asset tangibility	TANG	The ratio of fixed assets to total assets	Refinitiv Eikon		
Firm risk	RISK	Altman's Z-score	Refinitiv Eikon and authors' calculation		
Non-debt tax shields	NDTS	The ratio of depreciation to total assets	Refinitiv Eikon		
Liquidity	LIQ	The ratio of current assets to current liabilities	Refinitiv Eikon		
GDP growth	GDP	GDP growth rate (annual %)	WDI		
Noto: EAC is International Monotory Fund's Einspriel Assocs Currery WIPES is World Park Enternation Currery					

Note: FAS is International Monetary Fund's Financial Access Survey; WBES is World Bank Enterprise Survey; WDI is World Development Indicator. Source: authors' calculation, Stata output.

4. Results and Discussion

4.1. Summary Statistics and Correlation Matrix

Table 4 provides the descriptive statistics of all the variables for the period from 2007 to 2020. The average SFG was 0.30. The strong growth of financial markets and recent listings on the stock exchange partly explain this increase in the growth rates of firms in North Africa. Additionally, the financial-inclusion index had an average value of 0.301, ranging from 0.191 to 0.536. However, access to credit was still insufficient for the listed firms in the region, with an average value of 0.513, which is little higher than that of households for the use of financial services (0.543) in the USE. This situation suggests that firms in the North African region have limited access to formal financial services. Therefore, it is important that governments in the region improve financial-inclusion policies, particularly in relation to access to credit. The other variables showed steady growth from 2007 to 2020.

Variable	Obs.	Mean	Std. Dev.	Min	Max
SFG	2170	0.301	1.745	-48.938	18.785
FI	2170	0.301	0.088	0.191	0.536
ACC	2170	0.513	0.219	0.258	0.854
USE	2170	0.453	0.191	0.175	0.879
SIZE	2170	12.273	2.155	7.713	18.25
TANG	2170	0.275	0.212	0.011	0.874
RISK	2170	1.748	2.291	-2.713	12.609
NDTS	2170	0.033	0.031	0.003	0.233
LIQ	2170	0.520	0.270	0.005	0.929
GDP	2170	0.018	0.024	-0.097	0.056

Table 4. Descriptive statistics.

Note: This table summarizes the descriptive statistics, namely number of observations, mean, standard deviation, minimum value, and maximum value, for all variables used in this research. Source: authors' calculation, Stata output.

According to Gujarati (2004), if the correlation coefficients between the explanatory variables are less than 0.80, multicollinearity should not have a significant impact on the multiple-regression analysis. The results reported in Table 5 indicate the absence of a strong correlation between the independent variables in our study, reinforcing the validity of the results of our multiple-regression analysis. In addition, Table 5 also reveals the findings of a variance-inflation factor (VIF) test, which was employed to check the presence of multi-collinearity among all the independent variables. The values were clearly within the acceptable limit, i.e., VIF < 5, which confirmed that multicollinearity was absent among the variables (Gujarati 2004).

Table 5. Correlation matrix.

Variables	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	VIF
(1) SFG	1.000								
(2) FI	-0.061 *	1.000							1.18
(3) SIZE	0.063 *	0.055 *	1.000						1.01
(4) TANG	0.040	-0.075 *	0.022	1.000					1.37
(5) RISK	-0.055 *	-0.174 *	-0.057 *	-0.239 *	1.000				1.14
(6) NDTS	0.065 *	0.172 *	0.070 *	0.392 *	-0.069 *	1.000			1.27
(7) LIQ	0.006	0.117 *	0.036	-0.204 *	-0.062 *	-0.047 *	1.000		1.07
(8) GDP	-0.138 *	0.184 *	-0.035	0.011	0.088 *	-0.056 *	-0.072 *	1.000	1.08
Mean VIF									1.16

Notes: This table presents the correlation between all variables over the period between 2007 and 2020. * indicates significant correlation at the 5% level. Source: authors' calculation, Stata output.

4.2. Baseline Results

In this subsection, we examine the threshold effect of financial inclusion on the sustainable growth of the listed firms in North Africa. To this end, we adopted the DPTR model. The findings are reported in Table 6. The results of the bootstrap linearity test prove the existence of a threshold effect of FI on the SFG at the 1% significance level. Using financial inclusion as the transition variable, the threshold value was estimated at 0.284. In addition, of the observations, 51% fell in the lower regime and 49% in the upper regime.

The results presented in Table 6 indicate that the lagged variable coefficients of the SFG exhibited contrasting signs in different regimes (positive in the lower regime and negative in the upper regime), and that they were statistically significant at the 1% level. These results indicate that the current SFG is lower than the previous SFG, suggesting a lack of accelerated growth, and vice versa.

VARIABLES	Lower FI Regime	Upper FI Regime	Difference ($\delta = \lambda - \beta$)
SFG _{t-1}	0.884 ***	-0.748 ***	-1.632 ***
	(0.006)	(0.001)	(0.005)
FI	5.754 ***	-4.958 ***	-10.712 ***
	(0.066)	(0.050)	(0.016)
Size	-0.008 ***	0.111 ***	0.119 ***
	(0.001)	(0.003)	(0.002)
TANG	-0.202 ***	0.886 ***	1.088 ***
	(0.025)	(0.030)	(0.005)
RISK	0.023 ***	-0.070 ***	-0.093 ***
	(0.001)	(0.002)	(0.001)
NDTS	-3.215 ***	-0.534 ***	2.681 ***
	(0.177)	(0.197)	(0.020)
LIQ	-0.401 ***	0.577 ***	0.978 ***
	(0.012)	(0.018)	(0.006)
GDPG	-8.064 ***	10.445 ***	18.509 ***
	(0.111)	(0.098)	(0.013)
constant			-0.417 ***
			(0.035)
Threshold value $(\hat{\gamma})$		0.284 *** [0.280, 0.288]	
Percentage (%)	51%	49%	
Bootstrap (<i>p</i> -value)		0.000	
Observations			2170
Number of firms			155

Table 6. Baseline results.

Notes: this table shows the outcomes of the nonlinear relationship between FI and SFG using Seo and Shin's (2016) DPTR model. Standard errors are displayed in parentheses. *** denotes statistical significance at the 1% level. Source: authors' calculation, Stata output.

The findings show that the coefficients of the FI variable were significant for both regimes (lower and upper). In the lower regime, which demonstrated a low FI level, the estimated coefficient was positive and statistically significant, supporting the existence of a positive relationship between financial inclusion and SFG. However, in the upper regime, which had a high FI level, the estimated coefficient was negative and statistically significant, indicating the presence of a negative relationship between FI and SFG. Based on these results, the lower regime appears to be the optimal regime, since the coefficient is greater than the upper-regime coefficient. Specifically, a 1.0% FI leads to a 5.75% increase in SFG.

Regarding the control variables, we observed that the size variable had a negative correlation with sustainable firm growth in the lower FI regime, but a positive correlation in the upper regime. These results indicate that small firms have higher (lower) growth than large (small) firms. Although large (small) firms may benefit from considerable economies of scale, they often reach a minimum cost threshold, beyond which they cannot continue to grow. The negative correlation suggests that Gibrat's law does not apply, and vice versa. These results are consistent with those obtained by Akhtar et al. (2022) and Khémiri and Noubbigh (2021).

In addition, the tangibility (TANG) had negative and positive effects in the lower and upper regimes, respectively. From an economic viewpoint, the negative effect of the TANG on sustainable firm growth suggests that firms in North African countries do not have sufficient tangible assets to guarantee their sustainability, and vice versa. Specifically, a lack of guarantees prevents firms from accessing external financing, discourages lenders from granting loans to finance their investment projects, and ensures their sustainability.

However, firm risk (risk) had a positive correlation with sustainable firm growth in the lower FI regime, but a negative correlation in the upper FI regime. These results indicate that North African firms that are high in risk may find it more difficult to attract investors or secure financing, which can limit their ability to invest in new projects or expand their operations, and vice versa. Furthermore, the NDTS variable had a negative and significant effect on sustainable firm growth in all the regimes. From an economic perspective, this negative correlation indicates that these North African firms with higher levels of NDTS may have lower taxable income, which may limit their ability to reinvest profits into growth opportunities. This is because they may have less cash available for investment after paying taxes.

However, the liquidity (LIQ) had a negative correlation with sustainable firm growth in the lower FI regime, but a positive correlation in the upper FI regime. Indeed, North African firms that maintain high levels of liquidity may be more conservative in their investment decisions, as they prioritize maintaining a cash cushion over pursuing growth opportunities, and vice versa. This can limit their ability to invest in new projects, expand their operations, or pursue new markets.

The macroeconomic variables had a considerable influence on sustainable firm growth. Indeed, GDP had a negative correlation with sustainable firm growth in the lower FI regime, but a positive correlation in the upper FI regime. This suggests that sustained economic growth in the region promotes firm growth, and vice versa. This is because strong economic growth stimulates demand for goods and services, resulting in increased sales and profits for listed firms.

4.3. Robustness Checks

4.3.1. Alternative Measure of Dependent Variable

To ensure the robustness of our findings, we conducted a further test by changing the SFG measure. Following Zhang and Chen (2017) and Akhtar et al. (2022), we employed Van Horne's static model of the SFG denoted by SFGA. The SFGA is measured as follows: retained profits × net profit rate × (1 + debt/equity ratio) × {1/(total assets/total sale) - 1}. The DPTR results are reported in Table 7. Our findings were consistent with those in Table 6, indicating the presence of a threshold effect of financial inclusion on sustainable firm growth. In addition, the estimated threshold value ($\hat{\gamma}$) was 0.252, which was close to the threshold value of 0.284 in the baseline results. Furthermore, the sign and significance of the coefficients of FI ($\hat{\beta}_2$ and $\hat{\lambda}_2$) were consistent with the outcomes demonstrated in Table 6. This suggests that the previous results were robust.

Table 7. Robustness check: alternative measurement of SFG.

VARIABLES	Lower FI Regime	Upper FI Regime	Difference
SFGA _{t-1}	1.014 ***	-0.884 ***	-1.898 ***
	(0.001)	(0.0001)	(0.0011)
FI	0.262 ***	-0.451 ***	-0.713 ***
	(0.019)	(0.020)	(0.001)
Size	0.012 ***	-0.009 ***	-0.021 ***
	(0.000)	(0.001)	(0.001)
TANG	-0.276 ***	0.030 ***	0.306 ***
	(0.004)	(0.005)	(0.001)
RISK	0.003 ***	-0.004 ***	-0.007 ***
	(0.0001)	(0.0002)	(0.001)
NDTS	-1.728 ***	1.582 ***	3.31 ***
	(0.029)	(0.036)	(0.007)
LIO	0.068 ***	-0.079 ***	-0.147 ***
~	(0.001)	(0.002)	(0.001)
GDPG	-0.651 ***	0.434 ***	1.085 ***
	(0.025)	(0.032)	(0.007)
constant			0.261 ***
			(0.009)
Threshold value ($\hat{\gamma}$)		0.252 *** [0.251, 0.253]	
Percentage (%)	32%	68%	
Bootstrap (p-value)		0.000	
Observations			2170
Number of firms			155

Notes: this table shows the outcomes of the nonlinear relationship between FI and SFGA using Seo and Shin's (2016) DPTR model. Standard errors are displayed in parentheses. *** denotes statistical significance at the 1% level. Source: authors' calculation, Stata output.

4.3.2. Decomposition of FI

Next, we performed an additional robustness test by examining the impact of each of the financial-inclusion components (ACC and USE) on the sustainable firm growth. To this end, we also employed the DPTR model. Table 8 presents the results, which suggest how the various components of financial inclusion affect North African sustainable firm growth. Compared with the results obtained in the previous regressions (Table 6), there were some differences, indicating the robustness of our results.

Dependent Variable: SFG		Access			Usage	
VARIABLES	Lower FI Regime	Upper FI Regime	Difference	Lower FI Regime	Upper FI Regime	Difference
SFG _{t-1}	1.143 ***	-1.051 ***	-2.194 ***	0.058 ***	0.765 ***	0.707 ***
	(0.018)	(0.010)	(0.008)	(0.004)	(0.001)	(0.003)
FI	-3.371 ***	6.463 ***	9.834 ***	-4.194 ***	4.317 ***	8.511 ***
	(0.235)	(0.233)	(0.002)	(0.135)	(0.133)	(0.002)
Size	0.045 ***	-0.180 ***	-0.225 ***	-0.063 ***	0.004	0.067 ***
	(0.003)	(0.004)	(0.001)	(0.003)	(0.004)	(0.001)
TANG	-0.453 ***	0.582 ***	1.035 ***	0.520 ***	-1.639 ***	-2.159 ***
	(0.025)	(0.068)	(0.043)	(0.037)	(0.062)	(0.025)
RISK	0.040 ***	-0.143 ***	-0.183 ***	-0.026 ***	0.040 ***	0.066 ***
	(0.003)	(0.004)	(0.001)	(0.003)	(0.004)	(0.001)
NDTS	-4.698 ***	7.219 ***	11.917 ***	10.654 ***	-10.739 ***	-21.393 ***
	(0.243)	(0.348)	(0.105)	(0.336)	(0.315)	(0.021)
LIQ	0.105 ***	0.226 ***	0.121 ***	0.055 *	-0.227 ***	-0.282 ***
	(0.013)	(0.023)	(0.010)	(0.032)	(0.036)	(0.004)
GDPG	0.391 ***	-2.082 ***	-2.473 ***	2.683 ***	-4.295 ***	-6.978 ***
	(0.037)	(0.042)	(0.005)	(0.116)	(0.114)	(0.002)
constant		0.899 ***				-0.127 *
		(0.146)				(0.066)
			0.392 ***			0.293 ***
Threshold value ($\hat{\gamma}$)			[0.371, 0.414]			[0.287, 0.298]
Percentage (%)	49%	51%		22%	78%	[
Bootstrap (<i>p</i> -value)			0.000			0.000
Observations			2170			2170
Number of firms			155			155

Table 8. Robustness checks: decomposition of FI.

Notes: this table shows the outcomes of the nonlinear relationship between decompositions of FI and SFG using Seo and Shin's (2016) DPTR model. Standard errors are displayed in parentheses. *** and * denote statistical significance at the 1%, and 10% levels, respectively. Source: authors' calculation, Stata output.

For the first dimension (ACC), the bootstrap test for linearity was significant at the 1% level, which confirmed the existence of a threshold effect of ACC on sustainable firm growth. The threshold value $\hat{\gamma}$ was estimated at 0.392. In the lower regime, with low ACC, the estimated coefficient ($\hat{\beta}_2$) was significantly negative, suggesting that the relationship between ACC and sustainable firm growth was negative. However, in the upper regime, with high ACC, the estimated coefficient ($\hat{\lambda}_2$) was significantly positive, suggesting that the relationship between ACC and sustainable firm growth was negative. However, in the upper regime, with high ACC, the estimated coefficient ($\hat{\lambda}_2$) was significantly positive, suggesting that the relationship between ACC and sustainable firm growth was positive. These results can be explained by the fact that when credit access is very high, it is more likely that some of the available credit is misallocated towards less productive uses that do not boost sustainable firm growth. This diminishes the benefits of increased access at higher levels, and vice versa.

For the second dimension (USE), the bootstrap test for linearity was significant at the 1% level, which also confirmed the existence of a threshold effect of the USE on the SFG. The threshold value $\hat{\gamma}$ was estimated at 0.293. The sign and significance of the coefficients of the USE ($\hat{\beta}_2$ and $\hat{\lambda}_2$) were consistent with the results shown in column (2). These results

can also be explained by the fact that at low levels of financial-service usage (USE), firms may face capital constraints that limit their ability to grow and vice versa. As they gain more access to financing, these constraints ease, and growth accelerates.

4.3.3. Alternative Econometric Techniques

To assess the robustness of the empirical findings, this study employs three alternative econometric techniques, namely Seo and Shin's (2016) DPTR model, with a kink, Kremer et al.'s (2013) DPTR model, and Seo and Shin's (2016) Static PTR model to estimate Equation (2).

According to Seo et al. (2019), the presence of a model kink depends on this condition, i.e., if $(1, X'_{it})\delta = k(q_{it} - \gamma)$ for some *k*. The results are reported in column (1) in Table 9. The estimated threshold value ($\hat{\gamma}$) was 0.278, which was close to the threshold value of 0.284 mentioned in the baseline results. These results confirm the previous findings when the presence of a kink is considered.

Table 9. Robustness check: changing the econometric method using DPTR m	nodel.
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Dependent Variable: SFG	(1)	(2)
VARIABLES	Seo and Shin's (2016) DPTR Model with a Kink	Kremer et al.'s (2013) DPTR
L.SGR	0.286 ***	0.686 **
	(0.000)	(0.305)
FI	4.814 ***	
	(0.013)	
Size	-0.027 ***	-0.261
	(0.000)	(0.542)
TANG	-0.373 ***	-2.033
	(0.009)	(3.054)
RISK	-0.015 ***	-0.236
	(0.001)	(0.399)
NDTS	2.423 ***	12.61
	(0.062)	(30.18)
LIQ	-0.385 ***	1.194
	(0.005)	(1.385)
GDPG	-1.673 ***	-20.71 **
	(0.018)	(10.39)
Threshold value $(\hat{\gamma})$	0.278 *** [0.333, 0.339]	0.2493
		[0.2493, 0.2494]
Below the threshold $(\hat{\beta}_2)$		6.127 **
·		(2.593)
Above the threshold $(\hat{\lambda}_2)$		-18.00 **
× _/		(7.832)
kink_slope	-4.749 ***	× ,
1.	(0.020)	
Sup Wstar		95.415 ***
-		(12.810)
Observations	2170	2015
Number of firms	155	155
Time fixed effects	No	Yes

Notes: this table shows the outcomes of the nonlinear relationship between FI and SFG using Seo and Shin's (2016) DPTR model with a kink (column 1) and Kremer et al.'s (2013) DPTR model. Standard errors are displayed in parentheses. ***, and ** denote statistical significance at the 1% and 5%, levels, respectively. Source: authors' calculation, Stata output.

For Kremer et al.'s (2013) model, we also considered financial inclusion as the threshold variable and the core independent variable. The findings are reported in column (2) in Table 9. We observed that the estimated threshold value ($\hat{\gamma}$) was 0.249, and the sign and

significance of the coefficients of financial inclusion ($\hat{\beta}_2$ and $\hat{\lambda}_2$) were in line with the findings reported in Table 6. This suggests that the results discussed above are robust.

The results from Seo and Shin's (2016) static PTR model are reported in Table 10. The estimated threshold value ($\hat{\gamma}$) was 0.336. In addition, the coefficients of FI in the two regimes ($\hat{\beta}_2$ and $\hat{\lambda}_2$) were 0.432 and -0.187, respectively, and their sign and significance were in line with the findings reported in Table 6, suggesting that the main results were robust.

VARIABLES	Lower FI Regime	Upper FI Regime	Difference
FI	0.432 ***	-0.187 ***	-0.619 ***
	(0.013)	(0.026)	(0.013)
Size	-0.006 ***	0.153 ***	0.159 ***
	(0.001)	(0.005)	(0.004)
TANG	-0.539 ***	1.414 ***	1.953 ***
	(0.011)	(0.062)	(0.051)
RISK	0.024 ***	-0.137 ***	-0.161 ***
	(0.001)	(0.009)	(0.008)
NDTS	-1.965 ***	2.370 ***	4.335 ***
	(0.087)	(0.263)	(0.176)
LIQ	-0.281 ***	0.424 ***	0.705 ***
	(0.005)	(0.020)	(0.015)
GDPG	-2.076 ***	3.798 ***	5.874 ***
	(0.040)	(0.081)	(0.041)
constant			-2.506 ***
			(0.085)
Threshold value ($\hat{\gamma}$)		0.336 *** [0.333, 0.339]	
Percentage (%)	78%	22%	
Bootstrap (<i>p</i> -value)		0.000	
Observations			2170
Number of firms			155

Table 10. Robustness check: results of Seo and Shin's (2016) static PTR model.

Notes: this table shows the outcomes of the nonlinear relationship between FI and SFG using Seo and Shin's (2016) static PTR model. Standard errors are displayed in parentheses. *** denotes statistical significance at the 1% levels. Source: authors' calculation, Stata output.

4.4. Discussion

The objective of this study was to theoretically clarify and empirically evaluate how financial inclusion affects sustainable firm growth, addressing the mixed findings in the literature (Nizam et al. 2021). The literature indicates that financial inclusion is strictly related to firm sustainability, and that access to formal financial services is beneficial for sustainable firm growth (Yang and Zhang 2020). On the other hand, previous findings also suggested that high levels of financial inclusion (access to credit) are detrimental to sales growth, as they involve costs that generate risk (Ahmad et al. 2023). To verify whether positive and negative aspects coexist in the relationship between financial inclusion and sustainable firm growth, we examined this relationship by considering both the positive and the negative effects of financial inclusion, testing the threshold effect of financial inclusion on sustainable firm growth. To this end, we ran the DPTR model, using recent data covering the period of 2007–2020. Our results provide empirical evidence that supports hypothesis H1, that financial inclusion exerts a threshold effect on sustainable firm growth.

Furthermore, the estimated threshold indicates that there is an optimal level below which financial inclusion is beneficial for the sustainable growth of North African non-financial firms. However, above this optimal level, financial inclusion hinders the sustainable growth of these firms. This leads to lower and upper regimes (Nizam et al. 2021; Abdul Karim et al. 2022; Daud and Ahmad 2023). More specifically, in the lower regime, when financial inclusion is low, a positive effect is observed on sustainable firm growth. This means that when firms have limited access to formal financial services, an increase in this access has a positive impact on their sustainable growth. This may be explained by the fact that financial inclusion improves firms' access to the financial resources they need to invest, develop, and innovate, which stimulates their long-term growth. In this case, agency theory and Hypothesis H1(a) are accepted. North African firms are likely to have alternative sources of finance and access to credit to support their growth. Based on these results, it appears that the levels of financial inclusion in North African countries are still increasing. Therefore, it is important for governments and policy makers in this region to adopt appropriate policies to reduce the problems related to difficulties in accessing formal financial services. These results are like those found in other studies (macroeconomic (Akinrinola and Folorunso 2022; Ahmad et al. 2023; Adabor and Mishra 2023) and microeconomic (Chauvet and Jacolin 2017; Lee et al. 2020; Yang and Zhang 2020; Marcelin et al. 2022).

However, in the upper regime, when financial inclusion is high, there is a negative effect on sustainable firm growth. This suggests that high levels of financial inclusion may lead to additional costs and risks for firms. These costs and risks may limit the ability of firms to maintain sustainable growth, as they face greater competitive pressures and increased financial constraints. These results prove the validity of the sustainable growth and trade-off theories, supporting Hypothesis H1(b). Furthermore, this negative impact reveals that North African firms are unable to access formal financial services to grow because they are weakly inclusive. They therefore face external financing constraints. Certainly, these results are not surprising, since previous analyses showed that FI levels in North African countries are low. These results are like those found by Abdul Karim et al. (2022) and Chen et al. (2023). Finally, our results provide support for the agency, trade-off, and sustainable-growth theories. Specifically, an optimal financial-inclusion threshold was detected, indicating that at low levels of financial inclusion, non-financial firms in North Africa (particularly Egypt, Morocco, and Tunisia) can count on formal financial services to maintain their sustainability. This result was in line with the predictions of agency theory. However, beyond this level, firms must modify their financing policies, limiting their access to financial services to avoid costs and risks that would jeopardize their sustainability. Under these conditions, firms can reinvest their own capital and avoid additional external financing. This is consistent with the sustainable-growth and trade-off theories. In conclusion, the validity of our findings was reinforced by the thorough robustness test.

5. Conclusions and Implications

This study attempted to examine the nonlinear relationship between financial inclusion and SFG. To this end, we used the DPTR method. The findings showed that there is a threshold effect of financial inclusion on sustainable firm growth. This means that when financial inclusion is high (low), the growth of firms is significant (limited) due to the presence (absence) of appropriate financing, information, and financial tools. However, at an optimal level, when financial inclusion decreases (increases), firms can limit (access) financing and additional resources to invest in their growth (to avoid risks), which can lead to a decrease (increase) in their growth. In this case, the agency and trade-off theories are partially accepted, as they predict the existence of a threshold effect of financial inclusion. More specifically, they are accepted for the case of the nonlinear relationship between access and sustainable growth. Previous sustainable growth, as well as financial and macroeconomic variables, are also considered significant.

5.1. Implications

These findings may be useful for managers and policy makers. Managers should focus on finding the optimal level of financial inclusion (ACC and USE) for optimizing growth. Limited access to finance constrains growth, but excessive access can also be detrimental. In addition, firms need to improve their learning and expertise in managing leverage and financial resources and closely monitor the costs of capital and interest payments. If these start to overwhelm real investment and constrain growth, excessive financial inclusion may have occurred. Policymakers should strive to expand access to finance for North African SMEs and encourage financial inclusion. However, it must be recognized that more is not always better—extreme over-indebtedness must also be avoided. Policies should be designed to ensure optimal and moderate financial-inclusion levels. In addition, macroprudential policies and countercyclical buffers that can help to curb excess leverage during economic crises should be considered. Finally, financial education and resources for North African SMEs should be improved to help them maximize the benefits of increased financial inclusion. This includes education on the prudent use of leverage, managing costs of financing, and preventing over-indebtedness.

5.2. Limitations and Future Research

This study has some limitations that open up many opportunities for future research. To ensure the applicability of the results, future studies should have a broader scope and include other African countries. In addition, future research could focus on the methods that firms can adopt to optimize the use of financial services and improve their sustainable growth. Furthermore, our sample was limited to listed firms, so we urge scholars to perform studies that also include non-listed firms in the future. Finally, this research on sustainable firm growth reflects the context of North African countries. We encourage scholars to conduct further research by examining other African countries and performing comparative analyses.

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Appendix A

Table A1. Results of Bartlett test of sphericity, KMO measure of sampling adequacy, and PCA (total variance explained).

		Panel A: Tests of	Applicability			
		Bartl	ett Test of Spheric	ity	КМО	
		Chi-Square	Degrees of Freedom	<i>p</i> -Value		
Access dimension		95.554	3	0.000	0.696	
Usage dimension		38.103	1	0.000	0.500	
Financial inclusion		28.663	1	0.000	0.500	
	Pa	nnel B: PCA (Total Va	ariance Explained))		
Component	Component Eigenvalue Difference % of Cumula Variance Variance					
Access dimension	1	2.745	2.529	0.915	0.915	
	2 3	0.217 0.038	0.179	0.072 0.013	0.987 1.000	
Usage dimension	1	1.196	0.392	0.958	0.598	
	2	0.804	-	0.402	1.000	
Financial inclusion	1	1.822	1.643	0.911	0.911	
	2	0.178	-	0.089	1.000	

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