



Article **DNA Methylation Is a Potential Biomarker for Cardiometabolic** Health in Mexican Children and Adolescents

Abeer A. Aljahdali ^{1,2}, Jaclyn M. Goodrich ^{3,*}, Dana C. Dolinoy ^{2,3}, Hyungjin M. Kim ⁴, Edward A. Ruiz-Narváez ², Ana Baylin ^{2,5}, Alejandra Cantoral ⁶, Libni A. Torres-Olascoaga ⁷, Martha M. Téllez-Rojo ⁷ and Karen E. Peterson ^{2,3}

- 1 Department of Clinical Nutrition, King Abdulaziz University, Jeddah 21589, Saudi Arabia 2
 - Department of Nutritional Sciences, University of Michigan, Ann Arbor, MI 48109, USA
- 3 Department of Environmental Health Sciences, University of Michigan, 1415 Washington Heights, Ann Arbor, MI 48109, USA
- 4 Center for Computing, Analytics and Research, University of Michigan, Ann Arbor, MI 48109, USA
- 5 Department of Epidemiology, University of Michigan, Ann Arbor, MI 48109, USA
- 6 Department of Health, Iberoamericana University, Mexico City 01219, Mexico
- 7 Center for Nutrition and Health Research, National Institute of Public Health, Cuernavaca 62100, Mexico
- Correspondence: gaydojac@umich.edu; Tel.: +1-(734)-647-4564

Abstract: DNA methylation (DNAm) is a plausible mechanism underlying cardiometabolic abnormalities, but evidence is limited among youth. This analysis included 410 offspring of the Early Life Exposure in Mexico to Environmental Toxicants (ELEMENT) birth cohort followed up to two time points in late childhood/adolescence. At Time 1, DNAm was quantified in blood leukocytes at long interspersed nuclear elements (LINE-1), H19, and 11β-hydroxysteroid dehydrogenase type 2 $(11\beta$ -HSD-2), and at Time 2 in peroxisome proliferator-activated receptor alpha (*PPAR*- α). At each time point, cardiometabolic risk factors were assessed including lipid profiles, glucose, blood pressure, and anthropometry. Linear mixed effects models were used for LINE-1, H19, and 11β-HSD-2 to account for the repeated-measure outcomes. Linear regression models were conducted for the cross-sectional association between $PPAR-\alpha$ with the outcomes. DNAm at LINE-1 was associated with log glucose at site 1 [$\beta = -0.029$, p = 0.0006] and with log high-density lipoprotein cholesterol at site 3 [$\beta = 0.063$, p = 0.0072]. 11 β -HSD-2 DNAm at site 4 was associated with log glucose ($\beta = -0.018$, p = 0.0018). DNAm at LINE-1 and 11β -HSD-2 was associated with few cardiometabolic risk factors among youth in a locus-specific manner. These findings underscore the potential for epigenetic biomarkers to increase our understanding of cardiometabolic risk earlier in life.

Keywords: cardiometabolic risk factors; population-based study; children and adolescents; Mexicans; biomarkers; epigenetics; DNA methylation

1. Introduction

Obesity is rising worldwide among children aged 5–19 years. In Latin America and the Caribbean region, prevalence rose over a 40-year period from 1.6% and 1.8% in 1975 to 10.4% and 13.4% in 2016 for girls and boys, respectively [1]. Obesity has been associated with increases in the risk and prevalence of cardiometabolic abnormalities among youth [2–4]. A cluster of cardiometabolic abnormalities, called metabolic syndrome [5,6], is considered a risk factor for cardiovascular disease (CVD) incidence, cardiovascular-related mortality, all-cause mortality [7,8], and other chronic diseases [9,10]. Rising prevalence of metabolic syndrome may be a driver of the CVD and type-2 diabetes epidemics [11]. Even though CVD outcomes are manifested in middle and late adulthood, cardiometabolic risk factors may become evident during childhood [12–17] and track into adulthood [4,18,19]. Identifying the determinants of cardiometabolic risk factors in youth could serve as a fundamental step for risk reduction and prevention [4,20].



Citation: Aljahdali, A.A.; Goodrich, J.M.; Dolinoy, D.C.; Kim, H.M.; Ruiz-Narváez, E.A.; Baylin, A.; Cantoral, A.; Torres-Olascoaga, L.A.; Téllez-Rojo, M.M.; Peterson, K.E. DNA Methylation Is a Potential Biomarker for Cardiometabolic Health in Mexican Children and Adolescents. Epigenomes 2023, 7, 4. https://doi.org/10.3390/ epigenomes7010004

Academic Editor: Frédéric Silvestre

Received: 13 January 2023 Revised: 28 January 2023 Accepted: 29 January 2023 Published: 3 February 2023



Copyright: © 2023 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/).

Epigenetic modification is one potential underlying mechanism in obesity, cardiometabolic abnormalities, and CVD [21–28]. Previous research highlighted the importance of epigenetics as a potential biomarker for screening, diagnosis, prognosis, and individualized treatment regimens [23,24,29–31]. DNA methylation (DNAm), a commonly studied epigenetic modification, has been associated with CVD [21–27,32,33] and cardiometabolic risk factors, mainly in adults [21,28,34–36]. Existing evidence showed that DNAm during early development was associated with obesity and CVD risk later in life [37]; early embryogenesis is a particularly sensitive time period for epigenetic alteration by environmental factors that may contribute to disease risk [38]. However, adolescence is also a susceptible period for the impact of environmental stimuli on DNAm [39,40]. Furthermore, adolescence is characterized by changes in body composition and hormonal milieu [41]—the hallmarks for cardiometabolic abnormalities [42]. Despite the importance of this milestone, scare evidence is available investigating the potential of using DNAm as biomarkers for

cardiometabolic health among children and adolescents. The current study will address this gap in knowledge by examining the association of DNAm in blood leukocytes with cardiometabolic risk factors among Mexican children and adolescents enrolled in the Early Life Exposures in Mexico to ENvironmental Toxicants (EL-EMENT) Cohort. Specifically, we quantified CpG site-specific DNAm at repetitive elements (long interspersed nuclear element-1, LINE-1), which comprises 15–17% of the human genome [43,44]. DNAm of LINE-1 is often used as a proxy measure for global DNAm [45], and it has been found to associate with CVD independent from well-established CVD risk factors in adults [46]. The other three genes are H19, 11β -hydroxysteroid dehydrogenase type 2 (11 β -HSD-2), and peroxisome proliferator-activated receptor alpha (*PPAR-* α), which were selected based on their associations with components of cardiometabolic health. H19 is an imprinted gene with a role in regulating cell formation and proliferation, body weight, adipogenesis, and brown adipose tissue thermogenesis [47–50], and abnormal fat partitioning is a crucial underlying factor in impaired cardiometabolic health [51]. 11β -HSD-2 converts cortisol to an inactive metabolite called cortisone [52,53]. Previous studies have associated 11β -HSD-2 regulation with blood pressure [54–56], insulin sensitivity [57], and type 2 diabetes [58]. Blood pressure and glucose hemostasis are cornerstones in assessing cardiometabolic health; however, the latter is of great interest for Hispanic youth as insulin resistance was reported among normal-weight Mexican youth [59]. Lastly, PPAR- α controls multiple lipid metabolism pathways [60,61], and it was associated with serum triglycerides [62]. PPAR- α dysregulation is thought to play a role in dyslipidemia, diabetes, and obesity [63]. Based on functions of the genes and results from related studies, we hypothesized that altered DNAm of these regions would associate with cardiometabolic health measures (waist circumference, blood pressure, and serum glucose, high-density lipoprotein cholesterol, and triglycerides) in children and adolescents.

2. Results

We assessed DNAm and cardiometabolic outcomes at one to two time points each in children from the ELEMENT cohort. The final sample sizes for LINE-1, 11β -HSD-2, and H19 were 242, 229, and 245 subjects, respectively, with DNAm at Time 1 and outcomes at Time 1 and/or Time 2. For *PPAR-a*, 345 subjects had DNAm and outcomes at Time 2 (Figure 1). Table 1 shows the demographic characteristics of the 410 children by time point. At Time 1, the mean (standard deviation (SD)) age of the sample was 10.34 (1.67) years and 53.25% were female. At Time 2, the mean age was 14.08 (2.03) years and 51.32% were female. Among cardiometabolic risk factors, only waist circumference and serum triglycerides values were greater at Time 2 than at Time 1. We examined the crude association between DNAm values across sites within each genomic region. We found that LINE-1, *H19*, and *PPAR-a* were moderately to strongly correlated with one another. *11β-HSD-2* sites were less correlated as we have observed in past studies with this same gene (Supplementary Tables S1–S4).

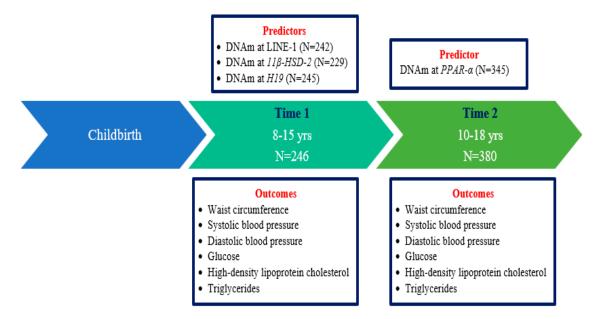


Figure 1. Summary of the Main Predictors and Outcomes for this Study and Number of Participants with the Data from the Early Life Exposures in Mexico to ENvironmental Toxicants (ELEMENT) Cohort. Abbreviations: DNAm = DNA methylation; Long interspersed nuclear elements (LINE-1); 11 β -hydroxysteroid dehydrogenase type 2 (*11\beta-HSD-2*); Peroxisome proliferator-activated receptor alpha (*PPAR-\alpha*).

Table 1. Descriptive Statistics of Mother and Child Characteristics of the Early Life Exposures in the Mexico to ENvironmental Toxicants (ELEMENT) Analytical Sample.

	Time 1 <i>n</i> = 246	Time 2 <i>n</i> = 380
Maternal Chara	cteristics (At Time of Child's Birth)	
Years of education, %		
<12 years	121 (49.19) 1	196 (51.58) ²
12 years	90 (36.59) ¹	131 (34.47) ²
>12 years	34 (13.82) ¹	52 (13.68) ²
Age at childbirth, (years)	26.86 (5.64) 1	26.47 (5.46) ²
Parity, %		
1	90 (36.59) ¹	144 (37.89) ²
2	89 (36.18) ¹	135 (35.53) ²
≥ 3	66 (26.83) ¹	100 (26.32) ²
Marital status, %		
Married	175 (71.14) ¹	274 (72.11) ²
Other	70 (28.46) 1	105 (27.63) ²
Enrollment in calcium supplementation study, %		
Not enrolled	152 (61.79) ¹	257 (67.63) ²
Enrolled	93 (37.80) ¹	122 (32.11) ²
Child	Characteristics (At birth)	
Female, %	131 (53.25)	195 (51.32)
Gestation age, (weeks)	38.85 (1.49) ³	38.79 (1.61) ⁴
Mode of delivery, %		
Vaginal delivery	140 (56.91) ⁵	220 (57.89) ⁶
C-Section	103 (41.87) ⁵	158 (41.58) ⁶
Birth weight, (kg)	3.15 (0.45) ⁷	3.15 (0.48) 6
Breastfeeding duration, (months)	8.15 (5.91) ¹	8.09 (6.07) ²
Child Char	acteristics (At follow-up visits)	
Age, (years)	10.34 (1.67)	14.08 (2.03)
Body mass index Z-score for age	0.85 (1.24)	0.53 (1.26) ⁶
Metabolic equivalents, (METs/week)	31.38 (19.97)	60.63 (38.76)
Total caloric intake, (kcal/day)	2636.32 (839.83)	2371.62 (936.82)
Pubertal onset, %	103 (41.87)	350 (92.11) ⁸

Tab	le 1	l. C	ont.

	Time 1 <i>n</i> = 246	Time 2 <i>n</i> = 380
Cardiometabo	lic risk factors (Outcomes)	
Waist circumference, (cm)	70.81 (10.71)	79.14 (11.42)
Systolic blood pressure, (mmHg)	102.74 (10.24)	97.23 (9.62)
Diastolic blood pressure, (mmHg)	65.58 (7.31)	62.24 (6.71)
Fasting glucose, (mg/dL)	86.98 (9.38)	77.48 (7.05) 9
High-density lipoprotein cholesterol, (mg/dL)	58.76 (11.92)	42.95 (8.87) ⁹
Triglycerides, (mg/dL)	87.89 (44.40)	105.81 (57.47) 9
DN	Am (Predictors)	
LINE-1 DNAm, % (averaged across 4 CpG sites)	78.49 (2.31) ⁵	N/A
11β-HSD-2 DNAm, % (averaged across 5 CpG sites) ^a	-0.85(1.34)	N/A
H19 DNAm, % (averaged across 4 CpG sites)	58.31 (5.16) ¹	N/A
<i>PPAR-</i> α DNAm, $\%$ (averaged across 2 CpG sites)	N/A	$10.62(2.09)^{10}$

Means (SD) or count (percentages) are presented for continuous or categorical variables, respectively. Number of missing values n = 245, 2 n = 379, 3 n = 242, 4 n = 377, 5 n = 243, 6 n = 378, 7 n = 244, 8 n = 373, 9 n = 342, n = 358. a Negative values appear for 11β -HSD-2 because values are standardized to controls included on each plate to reduce the impact of pyrosequencing batch effects. Abbreviations: DNAm = DNA methylation; Long interspersed nuclear elements (LINE-1); 11β -hydroxysteroid dehydrogenase type 2 (11β -HSD-2); Peroxisome proliferator-activated receptor alpha (*PPAR-a*).

2.1. Associations between the DNAm z-Score at LINE-1 and Repeated Measures of Cardiometabolic Risk Factors

In adjusted models, LINE-1 methylation levels were inversely associated with log serum fasting glucose at site 1 [$\beta = -0.029$, p = 0.0006]; for each one standard deviation increase in DNAm (i.e., +4%), there was an approximately 3% decrease in log fasting glucose. In addition, a positive association was detected between DNAm at site 3 and log serum fasting high-density lipoprotein cholesterol [$\beta = 0.063$, p = 0.0072], which means that for each one standard deviation increase in DNAm (i.e., +3%), there was an approximately 6% increase in log high-density lipoprotein cholesterol (Table 2). Sensitivity analysis (i.e., additionally adjusting for pubertal onset) did not attenuate the detected associations (Supplementary Table S5).

2.2. Associations between the DNAm z-Score at 11 β -HSD-2 and Repeated Measures of Cardiometabolic Risk Factors

DNAm at site 4 showed an inverse association with log serum fasting glucose (mg/dL) [$\beta = -0.018$, p = 0.0018] (Table 3). In sensitivity analysis (i.e., additionally adjusting for pubertal onset), associations found with fasting glucose maintained similar magnitude and significance (Supplementary Table S6).

2.3. Associations between the DNAm z-Score at H19 and Repeated Measures of Cardiometabolic Risk Factors

In the adjusted models, DNAm at none of the individual CpG sites was associated with the cardiometabolic outcomes (Supplementary Table S7). Results did not change in the two sensitivity analyses (Supplementary Tables S8 and S9).

2.4. Cross-Sectional Associations between the DNAm z-Score at PPAR- α and Cardiometabolic Risk Factors

In a cross-sectional analysis, DNAm was not associated with the cardiometabolic risk factors (Table 4). The sensitivity analysis (i.e., after adjusting for pubertal onset) showed the same result (Supplementary Table S10).

As an explanatory analysis, we assessed the crude association between DNAm values and gene expression for *PPAR-a*. RNA-seq data were available for a small subset of subjects at the same time point (i.e., Time 3) (n = 65). Weak non-significant positive correlations were identified between mRNA and DNAm (site 1: Spearman's correlation [rs] = 0.14, (p = 0.26); site 2: rs = 0.10, (p = 0.42); average of the two sites rs = 0.12, (p = 0.33)).

	LINE-1 z-Score	e at Site 1	LINE-1 z-Score	LINE-1 z-Score at Site 2		at Site 3	LINE-1 z-Score at Site 4		
	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	
	(Total num		Waist cir rvations = 441, of v d 199 (82.23%) sub		7.77%) subjects had	l one meası	ırement		
Model 1	-0.5960 (1.0435)	0.5684	1.1418 (1.4217)	0.4227	-0.4783 (1.1510)	0.6781	0.2997 (0.9013)	0.7398	
Model 2	0.5615 (1.0072)	0.5777	0.9837 (1.3686)	0.4730	-1.7757 (1.1106)	0.1111	0.3214 (0.8710)	0.7124	
	(Total num		Systolic bloo rvations = 441, of v d 199 (82.23%) sub	which 43 (12	7.77%) subjects had	l one meası	ırement		
Model 1	-0.4560 (0.8541)	0.5939	-0.1855 (1.1698)	0.8741	0.1632 (0.9435)	0.8628	0.9703 (0.7361)	0.1887	
Model 2	-0.9634 (0.8928)	0.2817	-0.00023 (1.2181)	0.9999	0.4640 (0.9898)	0.6397	0.8922 (0.7676)	0.2464	
	(Total num		Diastolic blo rvations = 441, of v d 199 (82.23%) sub	which 43 (12	7.77%) subjects had	l one meası	ırement		
Model 1	-0.5185 (0.5769)	0.3697	-0.1316 (0.7927)	0.8682	0.2271 (0.6379)	0.7221	0.3619 (0.4966)	0.4669	
Model 2	-0.6759 (0.5947)	0.2570	-0.04549 (0.8136)	0.9555	0.3404 (0.6613)	0.6072	0.3674 (0.5094)	0.4716	
	(Total num		Log-transformed rvations = 438, of v d 196 (80.99%) sub	which 46 (19	9.01%) subjects had	l one meası	ırement		
		uit		,	,				
Model 1	-0.01570 (0.007838)	0.0463	0.02427 (0.01086)	0.0263	-0.00357 (0.008708)	0.6825	-0.00361 (0.006726)	0.5917	
				0.0263 0.0160		0.6825 0.2162		0.5917	
	(0.007838) -0.02864 (0.008211)	0.0463 0.0006 * Log-transiber of observ	(0.01086) 0.02729	0.0160 sity lipopro which 46 (19	(0.008708) 0.01135 (0.009149) tein cholesterol (m 9.01%) subjects had	0.2162 g/dL)	(0.006726) -0.00142 (0.007028)		
Model 2	(0.007838) -0.02864 (0.008211)	0.0463 0.0006 * Log-transiber of observ	(0.01086) 0.02729 (0.01124) sformed high-dens rvations = 438, of v	0.0160 sity lipopro which 46 (19	(0.008708) 0.01135 (0.009149) tein cholesterol (m 9.01%) subjects had	0.2162 g/dL)	(0.006726) -0.00142 (0.007028)	0.8402	
Model 1 Model 2 Model 1 Model 2	(0.007838) 0.02864 (0.008211) (Total num 0.02078	0.0463 0.0006 * Log-trans iber of obser and	$(0.01086) \\ 0.02729 \\ (0.01124) \\ sformed high-dens \\ rvations = 438, of w \\ d 196 (80.99\%) sub \\ -0.02664$	0.0160 sity lipopro vhich 46 (19 jects had tv	(0.008708) 0.01135 (0.009149) tein cholesterol (m 9.01%) subjects had vo measurements) 0.01023	0.2162 g/dL) l one measu	(0.006726) -0.00142 (0.007028) urement -0.01677	0.8402	
Model 2 Model 1	(0.007838) -0.02864 (0.008211) (Total num 0.02078 (0.01893) -0.01466 (0.02111)	0.0463 0.0006 * Log-trans uber of obser and 0.2733 0.4881	(0.01086) 0.02729 (0.01124) sformed high-dens rvations = 438, of v d 196 (80.99%) sub -0.02664 (0.02610) -0.02801	0.0160 sity lipopro which 46 (19 jects had tv 0.3083 0.3306 d triglyceri which 46 (19	(0.008708) 0.01135 (0.009149) tein cholesterol (m 9.01%) subjects had vo measurements) 0.01023 (0.02099) 0.06331 (0.02334) des (mg/dL) 9.01%) subjects had	0.2162 g/dL) d one measu 0.6265 0.0072 *	(0.006726) -0.00142 (0.007028) urement -0.01677 (0.01627) -0.00571 (0.01822)		
Model 2 Model 1	(0.007838) -0.02864 (0.008211) (Total num 0.02078 (0.01893) -0.01466 (0.02111)	0.0463 0.0006 * Log-trans uber of obser and 0.2733 0.4881	(0.01086) 0.02729 (0.01124) sformed high-dens rvations = 438, of v d 196 (80.99%) sub -0.02664 (0.02610) -0.02801 (0.02873) Log-transforme rvations = 438, of v	0.0160 sity lipopro which 46 (19 jects had tv 0.3083 0.3306 d triglyceri which 46 (19	(0.008708) 0.01135 (0.009149) tein cholesterol (m 9.01%) subjects had vo measurements) 0.01023 (0.02099) 0.06331 (0.02334) des (mg/dL) 9.01%) subjects had	0.2162 g/dL) d one measu 0.6265 0.0072 *	(0.006726) -0.00142 (0.007028) urement -0.01677 (0.01627) -0.00571 (0.01822)	0.8402	

Table 2. Associations between the DNAm z-score at LINE-1 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models (n = 242).

Long interspersed nuclear elements (LINE-1). Model 1 included LINE-1 z-scores at CpG sites 1, 2, 3, and 4 as fixed effects and a compound symmetry matrix structure to model the covariance structure of the repeated measurements for each outcome. Model 2 was additionally adjusted for the following fixed effects: age, sex, and duration of breastfeeding. * p < 0.008.

	11β-HSD-2 z-Score at Site 1		11β-HSD-2 z-Score at Site 2			11β-HSD-2 z-Score at Site 3		11β-HSD-2 z-Score at Site 4		11β-HSD-2 z-Score at Site 5	
	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	
(Total nu	mber of observ	vations = 415	, of which 43 (1		circumference cts had one me		nd 186 (81.22%	6) subjects ha	nd two measur	ements)	
Model 1	-0.3822 (1.0424)	0.7142	-0.08657 (0.7980)	0.9137	0.2635 (0.9701)	0.7862	0.5264 (0.7690)	0.4943	0.2132 (0.7252)	0.7690	
Model 2	-1.1319 (1.0012)	0.2595	0.2204 (0.7707)	0.7751	0.6173 (0.9303)	0.5076	0.5578 (0.7361)	0.4493	-0.1382 (0.6969)	0.8430	
(Total nu	mber of observ	vations = 415,	, of which 43 (1		lood pressure (cts had one me		nd 186 (81.22%	6) subjects ha	nd two measur	ements)	
Model 1	-1.6096 (0.8326)	0.0545	-0.7568 (0.6372)	0.2362	1.2770 (0.7754)	0.1010	0.3766 (0.6161)	0.5416	-0.4901 (0.5780)	0.3974	
Model 2	-1.4026 (0.8695)	0.1083	-0.7320 (0.6688)	0.2751	1.1599 (0.8074)	0.1524	0.3305 (0.6404)	0.6064	-0.3520 (0.6029)	0.5600	
(Total nu	mber of observ	vations = 415	, of which 43 (1		blood pressure cts had one me		nd 186 (81.22%	a) subjects hat	nd two measur	ements)	
Model 1	-0.9251 (0.5519)	0.0951	-0.8601 (0.4222)	0.0428	0.3540 (0.5143)	0.4920	0.4535 (0.4092)	0.2690	-0.01360 (0.3827)	0.9717	
Model 2	-0.8686 (0.5624)	0.1240	-0.8775 (0.4322)	0.0436	0.3201 (0.5221)	0.5404	0.4519 (0.4148)	0.2771	0.01427 (0.3888)	0.9708	
(Total nu	mber of observ	vations = 412			ed fasting glue cts had one me			6) subjects ha	ıd two measur	ements)	
Model 1	-0.00076 (0.007513)	0.9193	0.001955 (0.005764)	0.7348	0.006329 (0.006998)	0.3668	-0.01869 (0.005586)	0.0010 *	0.002692 (0.005216)	0.6064	
Model 2	0.009223 (0.007893)	0.2440	-0.00184 (0.006079)	0.7624	0.001102 (0.007320)	0.8805	-0.01837 (0.005817)	0.0018 *	0.007427 (0.005472)	0.1762	
(Total nu	mber of observ	vations = 412			ensity lipoprote cts had one me			6) subjects ha	nd two measur	ements)	
Model 1	0.002550 (0.01874)	0.8919	-0.00550 (0.01438)	0.7026	-0.00829 (0.01745)	0.6351	-0.01132 (0.01390)	0.4161	0.005434 (0.01303)	0.6770	
Model 2	0.02693 (0.02073)	0.1952	-0.02151 (0.01596)	0.1793	-0.02199 (0.01925)	0.2545	-0.01714 (0.01524)	0.2620	0.01880 (0.01442)	0.1938	
(Total 1	number of obse	ervations = 4			ned triglycerid ojects had one i		t and 183 (79.9	1%) subjects	had measuren	nents)	
Model 1	0.02425 (0.04126)	0.5572	0.03580 (0.03163)	0.2588	0.004623 (0.03838)	0.9042	0.01794 (0.03047)	0.5566	-0.00972 (0.02872)	0.7354	
	0.01469	0.7140	0.03065	0.3212	0.01000	0.7880	0.01977	0.5029	-0.01685	0.5453	

Table 3. Associations between the DNAm z-score at 11β -HSD-2 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models (n = 229).

3, 4, and 5 as fixed effects and a compound symmetry matrix structure to model the covariance structure of the repeated measurements for each outcome. Model 2 was additionally adjusted for the following fixed effects: age, and sex. * p < 0.008.

Table 4. Cross-sectional Associations between DNAm z-scores at *PPAR-* α and Cardiometabolic Risk Factors using Linear Regression (n = 345).

	PPAR-α z-Sco	re at Site 1	<i>PPAR-</i> α z-Score at Site 2		
	Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value	
	Waist circ	cumference (cm)	(n = 345)		
Model 1	0.71915 (0.71474)	0.3150	-1.70941 (0.65445)	0.0094	
Model 2	0.99917 (0.70529)	0.1575	-1.68127 (0.64618)	0.0097	

Systolic blood pressure (mmHg) ($n = 345$) Model 1 0.58582 (0.60305) 0.3320 -1.02922 (0.55218) 0.0632 Model 2 0.49623 (0.57982) 0.3927 -0.66490 (0.53123) 0.2116 Diastolic blood pressure (mmHg) ($n = 345$) Diastolic blood pressure (mmHg) ($n = 345$) 0.1668 -0.57466 (0.38679) 0.1383 Model 1 0.58530 (0.42242) 0.1668 -0.34026 (0.37311) 0.3624 Model 2 0.58072 (0.40724) 0.1548 -0.34026 (0.37311) 0.3624 Model 1 0.00598 (0.00614) 0.3305 0.00016627 (0.00790) 0.9779 Model 2 0.00282 (0.00609) 0.6443 0.00159 (0.00596) 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) Model 1 -0.00813 (0.01303) 0.5329 0.01206 (0.01273) 0.3445		PPAR-α z-Sco	PPAR-α z-Score at Site 1		<i>PPAR-</i> α z-Score at Site 2		
Model 1 $0.58582 (0.60305)$ 0.3320 $-1.02922 \\ (0.55218)$ 0.0632 Model 2 $0.49623 (0.57982)$ 0.3927 $-0.66490 \\ (0.53123)$ 0.2116 Diastolic blood pressure (mmHg) ($n = 345$) 0.668 $-0.57466 \\ (0.38679)$ 0.1383 Model 1 $0.58530 (0.42242)$ 0.1668 $-0.57466 \\ (0.38679)$ 0.1383 Model 2 $0.58072 (0.40724)$ 0.1548 $-0.34026 \\ (0.37311)$ 0.3624 Model 1 $0.00598 (0.00614)$ 0.3305 $0.00016627 \\ (0.00600)$ 0.9779 Model 2 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) Model 1 $-0.00813 \\ (0.01303)$ 0.5329 $0.01206 (0.01273)$ 0.3445		Estimate (SE)	<i>p</i> -Value	Estimate (SE)	<i>p</i> -Value		
Model 1 $0.58582 (0.60305)$ 0.3320 $0.051218)$ 0.0632 Model 2 $0.49623 (0.57982)$ 0.3927 -0.66490 (0.53123) 0.2116 Diastolic blood pressure (mmHg) ($n = 345$)Diastolic blood pressure (mmHg) ($n = 345$) 0.1668 -0.57466 (0.38679) 0.1383 Model 1 $0.58530 (0.42242)$ 0.1668 -0.57466 (0.38679) 0.1383 Model 2 $0.58072 (0.40724)$ 0.1548 -0.34026 (0.37311) 0.3624 Model 1 $0.00598 (0.00614)$ 0.3305 0.00016627 (0.00600) 0.9779 Model 1 $0.00598 (0.00614)$ 0.3305 0.00016627 (0.00600) 0.9779 Model 2 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$)Model 1 $-0.00813 \\ (0.01303)$ 0.5329 $0.01206 (0.01273)$ 0.3445 Model 2 -0.00419 0.7490 $0.00857 (0.01280)$ 0.5035		Systolic bloo	d pressure (mm	Hg) (<i>n</i> = 345)			
Model 2 $0.49623 (0.57982)$ 0.3927 (0.53123) 0.2116 Diastolic blood pressure (mmHg) ($n = 345$) 0.668 -0.57466 0.1383 Model 1 $0.58530 (0.42242)$ 0.1668 -0.57466 0.1383 Model 2 $0.58072 (0.40724)$ 0.1548 -0.34026 0.3624 Log-transformed fasting glucose (mg/dL) ($n = 310$) 0.00016627 0.9779 Model 1 $0.00598 (0.00614)$ 0.3305 0.00016627 0.9779 Model 2 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) 0.3445 Model 1 -0.00813 0.5329 $0.01206 (0.01273)$ 0.3445	Model 1	0.58582 (0.60305)	0.3320		0.0632		
Model 1 $0.58530 (0.42242)$ 0.1668 $-0.57466 \\ (0.38679)$ 0.1383 Model 2 $0.58072 (0.40724)$ 0.1548 $-0.34026 \\ (0.37311)$ 0.3624 Log-transformed fasting glucose (mg/dL) (n = 310) $0.00598 (0.00614)$ 0.3305 $0.00016627 \\ (0.00600)$ 0.9779 Model 1 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) (n = 310) 0.3445 Model 1 $-0.00813 \\ (0.01303)$ 0.5329 $0.01206 (0.01273)$ 0.3445	Model 2	0.49623 (0.57982)	0.3927		0.2116		
Model 1 $0.58530 (0.42242)$ 0.1668 (0.38679) 0.1383 Model 2 $0.58072 (0.40724)$ 0.1548 -0.34026 0.3624 Log-transformed fasting glucose (mg/dL) ($n = 310$) Log-transformed fasting glucose (mg/dL) ($n = 310$) 0.00016627 0.9779 Model 2 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) 0.00413 $0.00159 (0.00282)$ 0.3445 Model 1 -0.00813 0.5329 $0.01206 (0.01273)$ 0.3445		Diastolic bloc	d pressure (mn	hHg) (<i>n</i> = 345)			
Model 2 $0.58072 (0.40724)$ 0.1548 (0.37311) 0.3624 Log-transformed fasting glucose (mg/dL) ($n = 310$) $0.00598 (0.00614)$ 0.3305 $0.00016627 \\ (0.00600)$ 0.9779 Model 1 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) 0.00411 $-0.00813 \\ (0.01303)$ 0.5329 $0.01206 (0.01273)$ 0.3445 Acadel 2 $-0.00419 \\ 0.00857 (0.01280)$ 0.5035 0.5035	Model 1	0.58530 (0.42242)	0.1668	0.01 200	0.1383		
Model 1 $0.00598 (0.00614)$ 0.3305 $\begin{array}{c} 0.00016627 \\ (0.00600) \end{array}$ 0.9779 Model 2 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) $0.00813 \\ (0.01303)$ 0.5329 $0.01206 (0.01273)$ 0.3445 Model 1 $-0.00813 \\ (0.01303)$ $0.7490 \\ 0.7490 \\ 0.00857 (0.01280) \\ 0.00857 (0.01280) \\ 0.5035 \\ 0.5$	Model 2	0.58072 (0.40724)	0.1548		0.3624		
Model 1 $0.00598 (0.00614)$ 0.3305 (0.00600) 0.9779 Model 2 $0.00282 (0.00609)$ 0.6443 $0.00159 (0.00596)$ 0.7900 Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) $-0.00813 (0.01303)$ 0.5329 $0.01206 (0.01273)$ 0.3445 Model 1 $-0.00419 (0.00419)$ $0.7490 (0.00857 (0.01280))$ 0.5035		Log-transformed	fasting glucose	(mg/dL) (n = 310)			
Log-transformed high-density lipoprotein cholesterol (mg/dL) ($n = 310$) Model 1 -0.00813 (0.01303) 0.5329 0.01206 (0.01273) 0.3445 Apdel 2 -0.00419 0.7490 0.00857 (0.01280) 0.5035	Model 1	0.00598 (0.00614)	0.3305		0.9779		
-0.00813 0.5329 0.01206 0.01273 0.3445 $4 odel$ -0.00419 0.7490 0.00857 0.01280 0.5035	Model 2	0.00282 (0.00609)	0.6443	0.00159 (0.00596)	0.7900		
Model 1 (0.01303) 0.5329 $0.01206 (0.01273)$ 0.3445 $4 a d a l 2$ -0.00419 0.7490 $0.00857 (0.01280)$ 0.5035	Log	-transformed high-densi	ty lipoprotein c	holesterol (mg/dL) ($n =$	310)		
Model 2 0 7490 0 00857 (0 01280) 0 5035	Model 1		0.5329	0.01206 (0.01273)	0.3445		
	Model 2	0.00	0.7490	0.00857 (0.01280)	0.5035		
Log-transformed triglycerides (mg/dL) ($n = 310$)		Log-transformed	l triglycerides (1	ng/dL) (<i>n</i> = 310)			

Table 4. Cont.

Peroxisome proliferator-activated receptor alpha (*PPAR-* α). Model 1 included PPAR- α z-scores for CpG sites 1 and 2. Model 2 was additionally adjusted for age, and sex.

0.6873

0.4956

0.00118 (0.02989)

-0.01116

(0.02989)

0.9684

0.7092

3. Discussion

Model 1

Model 2

0.01232 (0.03058)

0.02086 (0.03057)

In this study, the relationships between DNAm at LINE-1, 11β -HSD-2, H19, and *PPAR-a* with cardiometabolic risk factors were investigated among Mexican children and adolescents enrolled in a well-characterized birth cohort from Mexico City. Among cardiometabolic components, fasting glucose and high-density lipoprotein cholesterol were associated with DNAm of at least one genomic region. To the best of our knowledge, this is the first study investigating the potential of DNAm as a biomarker for cardiometabolic risk factors among Mexican youth using hypothesis-driven genomic regions.

The inverse and positive associations between LINE-1 DNAm and glucose and highdensity lipoprotein cholesterol are in line with current evidence linking LINE-1 hypomethylation with genomic instability and CVD [46,64–66]. Furthermore, few studies conducted on adult populations showed inverse relationships between LINE-1 DNAm and impaired carbohydrate metabolism [67] and fasting glucose [62,68]. Scare and inconsistent evidence is available among pediatric populations with regard to cardiometabolic health and LINE-1 DNAm [69,70], where an inverse association detected with the waist circumference z-score [69] and null associations were reported with adiposity markers [70]. We acknowledge the complexity of crude comparisons across the studies because of the mismatch in the study endpoints and sample characteristics; therefore, future prospective studies are needed to strengthen the use of LINE-1 DNAm as a proxy for cardiometabolic health among youth.

We found that a one standard deviation increase in 11β -HSD-2 DNAm at site 4 (i.e., +2%) was associated with a decrease of 2% in fasting glucose. Our results could be explained

in light of the limited studies that investigated the connection between 11β -HSD-2 and glucose metabolism in adult populations [57,58]. Müssig et al. reported inverse association between 1 β -HSD2 activity and insulin sensitivity [57], and Jang and colleagues found higher 11 β -HSD2 enzyme activity among subjects with type 2 diabetes [58]. It is worth noting that not only is 11 β -HSD-2 expression regulated by other epigenetic modifications [71], age [72], and lifestyle factors [73], but a lack of association was also documented earlier between 11 β -HSD2 enzyme activity and mRNA expression [58]. As our results showed the potential of 11β -HSD-2 DNAm as a cardiometabolic biomarker among youth, future studies are needed combining DNAm, gene expression, and enzyme activity assessment to strengthen the evidence for the role of 11β -HSD-2 in cardiometabolic risk.

The present study has multiple strengths, including the prospective assessment of the association between DNAm at four genomic regions and up to two repeated measures of cardiometabolic risk factors during a sensitive period of growth, development, and maturation. We used a robust statistical model to account for the longitudinal data structure and conducted site-specific analyses for examining the association between the DNAm of each region with cardiometabolic risk factors. Site-specific approaches may be better when the data are not as correlated or when some CpG sites are much more variable than others in order to capture the complexity of the data. Our data come from a well-characterized birth cohort, ELEMENT, which allowed for assessing whether any of the mother's sociodemographic and reproductive characteristics would be potential confounding factors to account for. Furthermore, peripheral blood was used to quantify the DNAm because blood is an accessible tissue and commonly collected in clinical setting and epidemiological studies [31], which is a strength for investigating potential biomarkers for cardiometabolic risk factors among children.

With regard to the study limitations, the use of bisulfite treatment to measure DNAm does not distinguish between cytosine methylation (5mC) and cytosine hydroxymethylation (5hmC) [74], and 5hmC has its own distinct impact on gene regulation, which was not captured by our method. Therefore, the DNAm values might be confounded by hydroxymethylation because both 5hmC and 5mC are captured in the total DNAm percentage. Future studies should apply laboratory techniques that allow for distinguishing between 5hmC and 5mC. Additionally, our work has the limitation of including only DNAm without gene expression data for three of the four regions assessed. Because gene expression could be influenced by multiple factors, including other epigenetic modifications, physiological conditions, and lifestyle factors, we recommend future studies supplement the assessment of DNAm with gene expression and carefully take into account the other potential factors that influence gene expression. Such evidence will strengthen the use of DNAm as a clinical biomarker for cardiometabolic health if clinical validation studies confirm its utility.

We acknowledge the age heterogeneity as our analysis includes pre-teenagers and teenagers; given our small sample size, we did not explore the relationships stratified by age groups. Thus, future studies are needed to investigate the potential role of age in modifying the association between DNAm and cardiometabolic risk factors during pubertal transition. Additionally, the magnitude of detected associations was small, which might not be of clinical significance. However, small effect sizes are typically reported in epigenetic studies [62,65,67,68,75]. Because small effects may still have relevance for children's health outcomes [75], further studies are needed to enhance our understanding of the cause-andeffect relationship between DNAm and cardiometabolic health by validating our results in independent large-scale population-based youth populations with objective assessment of lifestyle patterns known to influence DNAm. Such evidence will facilitate the progress toward increasing the reproducibility and strengthening the biological relevance of DNAm biomarkers. Additionally, despite our consideration for addressing multiple testing, we still acknowledge the possibility of reporting false positive results due to chance. Lastly, the possibility of residual confounding—such as smoking status and genetic variants—and reverse causation between DNAm and cardiometabolic outcomes cannot be ruled out.

4. Materials and Methods

4.1. Study Population

The analytical sample consisted of offspring who participated in two of three sequentially enrolled birth cohorts of the ELEMENT project in Mexico City, Mexico. A comprehensive description of the ELEMENT project and the eligibility and exclusion criteria are available elsewhere [76]. Briefly, the ELEMENT project included mother–child dyads recruited from maternity hospitals representing women from low- to middle-income population groups from 1997 to 2005 [77]. Mothers recruited for one of the birth cohorts were enrolled in a randomized controlled trial (RCT) that examined the role of daily calcium supplementation during pregnancy (1200 mg/day) in mitigating the effect of lead exposure on the neurobehavioral and physical developmental outcomes in offspring [76]. Offspring were followed at multiple time points in childhood and through adolescence; the aim of the follow-up visits was to follow as many children from the original birth cohort as possible, prioritizing younger ages at specific time points. The sample size for each follow-up visit was determined by the aims for the original grant-funded visit.

We utilized available data from two follow-up visits. In the first follow-up visit, herein called Time 1, we planned to follow 250 children aged between 8 and 15 years. We prioritized children according to availability of prenatal biological samples for offspring from the original birth cohorts [76]. The second follow up visit, Time 2, was conducted on average 2 years later (maximum time to follow-up was 4.6 years). We planned to follow >500 children from the original birth cohorts. We prioritized the 250 subjects from the Time 1 visit (of which a large majority (~90%) returned) and added additional ELEMENT children who were not included in the Time 1 visit. Based on a statistical power calculation and available funds, we selected a sub-sample of these for epigenetic analysis (all children at Time 1 and >350 at Time 2). Children were 10–18 years of age at Time 2.

The analytical sample for the genomic regions LINE-1, *H19*, and *11* β -*HSD-2* included children and adolescents who had DNAm data at Time 1 and had data for at least one of the six cardiometabolic risk factors (i.e., waist circumference, systolic and diastolic blood pressure, fasting glucose, triglycerides, high density lipoprotein cholesterol) at Time 1 and/or Time 2. DNAm at *PPAR-* α was measured only at Time 2; subjects with these data and at least one of the six cardiometabolic risk factors were included for the analytical sample for *PPAR-* α models. The National Institute of Public Health of Mexico and the University of Michigan institutional review boards approved the research protocols. Written informed consents were collected from mothers upon their enrollments and assent from adolescents.

4.2. Laboratory Measurements and Outcomes

4.2.1. DNA Methylation Analysis

The current study limits its focus to four genomic regions, which have previously been associated with cardiometabolic risk factors. Whole blood samples were collected via venipuncture into tubes containing ethylenediaminetetraacetic acid (EDTA) preservative (Paxgene and BD Vacutainer) by trained staff following standard protocols. Highmolecular-weight DNA was extracted from blood leukocytes with the PAXgene Blood DNA kit (PreAnalytix, Switzerland) or the Flexigene kit (Qiagen). The extracted DNA samples were treated with sodium bisulfite using Epitect (Qiagen, Valencia, CA, USA) or EZ DNA Methylation kits (Zymo Research, Irvine, CA, USA) following the standard methods previously published [78]. The purpose of bisulfite treatment was to convert the un-methylated cytosines to uracil and to preserve the methylated cytosines. The bisulfite-treated DNA samples were amplified using HotStarTaq Master Mix (Qiagen), and primers designed to amplify each region of interest. Pyrosequencing was performed using either PyroMark Q96 MD (Qiagen) or PyroMark Q96 ID (Qiagen). Pyro Q-CpG Software calculated the percent methylation and performed internal quality control checks. At Time 1, DNAm was quantified for H19 (4 CpG sites in the imprinting control region), for LINE-1 (4 CpG sites in a conserved region across many LINE-1s), and for 11β -HSD-2 (5 CpG sites in the

promoter region) and at Time 2 for *PPAR-* α (2 CpG sites in the promoter region) following the protocols published previously [79–81]. Information on these genomic regions and the primer sequences is presented in Supplementary Table S11 [77]. More than 10% of all samples and controls of human DNA with known percentages of DNAm (0%, 25%, 50%, 75%, and 100%) were run in duplicate and included in each pyrosequencing batch (96-well plate). The average of duplicate samples was used when applicable [82]. DNAm data from LINE-1, *11* β -*HSD-2*, and *H19* suggested a batch effect, and the methylation percentages were standardized to adjust for the batch effects as described previously [82]. We then standardized DNAm values for each region to have mean 0 standard deviation 1 based on the sample's mean and standard deviation values to express the DNAm as a z-score, and these z-scores were used in statistical analysis.

Samples collected at Time 1 were not preserved for downstream RNA isolation. At Time 2, blood leukocytes preserved for RNA isolation were collected from all participants and archived. Of these, 72 were selected for next-generation sequencing of RNA ('RNA-Seq'). Samples were prioritized for selection that had the highest quality and quantity of RNA and had complete datasets needed for previous questions of interest [83]. Of those, 65 were from participants included in this manuscript. The read count of *PPAR-* α from the RNA-seq was used to assess the relationship between DNAm and gene expression for *PPAR-* α . The RNA-seq protocol followed was previously described [83].

4.2.2. Cardiometabolic Risk Factors

Anthropometric Measures

Duplicate measurements were collected by trained research staff for body weight to the nearest 0.1 kg using a digital scale (BAME Model 420; Catálogo Médico) and In-Body 230 (Biospace Co, Ltd, Seoul, Republic of Korea), height to the nearest 0.5 cm, and waist circumference to the nearest 0.1 cm using a non-stretchable measuring tape SECA (model 201, Hamburg, Germany) [84]. The average of the two measurements was used for the analysis [85]. These measurements were conducted at Time 1 and Time 2.

Blood Pressure Measurements

Duplicate readings of systolic and diastolic blood pressure were recorded in a seated position using a mercury sphygmomanometer (TXJ-10 MD 3000 model, Homecare, Nanjing, China), and the average of the two measurements was used for the analysis. These measurements were conducted at Time 1 and Time 2.

Fasting Biomarkers

At each follow-up visit (T1 and T2), trained research staff collected blood samples from children after an 8 h overnight fast. Fasting glucose and lipids were measured in serum at the Michigan Diabetes Research Center Chemistry Laboratory. Specifically, fasting glucose was assessed via automated chemiluminescence immunoassay (Immulite 1000; Siemens Medical Solutions). Triglycerides were quantified via an enzymatic colorimetric method using a Cobas Mira automated chemistry analyzer (Roche Diagnostics, Indianapolis, IN, USA). The level of high-density lipoprotein cholesterol was obtained by using direct high-density lipoprotein cholesterol (Roche Diagnostics) [85]. All serum markers were above the limit of detection (LOD).

4.3. Covariates

Based on prior knowledge of cardiovascular and metabolic health, covariates assessed for this research were classified as (1) maternal and child characteristics around the time of birth (sex, birth weight, gestation age, mode of delivery, duration of breastfeeding, and mothers' age, marital status, parity, years of education, and enrollment in the calcium supplementation study during pregnancy) and (2) follow-up characteristics for the children, which were measured at the baseline visit for each exposure, e.g., child's age, total caloric intake, physical activity measured as metabolic equivalents, and pubertal onset. In our statistical analysis section, we explained our rationale for selecting covariates in each adjusted model.

After childbirth, mothers reported household and demographic information, including their ages, marital status (married compared to any other status), parity status $(1, 2, \ge 3)$, and years of education (<12 yrs, 12 yrs, or >12 yrs), gestational age estimated by a registered nurse, and mode of delivery (vaginal, or C-section childbirth). The newborns were followed until 5 years of age, and information about self-reported breastfeeding duration was estimated [86]. Since cohort 3 was an RCT for daily calcium supplementation during the first trimester of pregnancy until 1-year postpartum and cohort 2 participants were not part of a trial, we created a binary indicator for mothers who received the calcium treatment (yes/no) with all mothers from cohort 2 falling into the 'no' category [76,87].

During each of two follow-up visits, total caloric intake was quantified using a semiquantitative food frequency questionnaire (FFQ) that captured the intake over the previous week [84,88]. The FFQ was adapted from the Mexican National Health and Nutrition Survey, and FFQs were analyzed using food composition software developed by the National Institute of Public Health, Mexico [89]. A physical activity questionnaire was developed based on the Youth Activity Questionnaire (YAQ) and validated relative to 24 h physical activity recall among Mexican school-children aged 10 to 14 years in Mexico City [90]. For each self-reported physical activity, the corresponding metabolic equivalent was multiplied by the activity intensity [91]. The total metabolic equivalents per week were calculated by summing the metabolic equivalents for all activities. Puberty was assessed through Tanner staging for breast and pubic hair (for girls) or genitalia and pubic hair (for boys) [92,93] by trained physicians [94]. Consistent with previous ELEMENT publications in which pubertal onset was a covariate, we classified children as having pubertal onset when the Tanner Stage for either or both of pubic hair and genital development (boys) or pubic hair and breast development (girls) was greater than one [95–97].

4.4. Statistical Analysis

Outcomes were cardiometabolic risk factors: waist circumference, systolic blood pressure, diastolic blood pressure, glucose, high-density lipoprotein cholesterol, and triglycerides. Dependent variables of interest were DNAm z-scores at LINE-1, *11β-HSD-2*, *H19*, and *PPAR-α* after standardizing the values based on the sample's mean and standard deviation for each site. Outcomes and exposures were treated as continuous in our models. The demographic characteristics of the study participants were presented as the mean (SD) and counts (proportions) for continuous and categorical variables, respectively.

DNAm percentages were quantified at multiple loci (CpG sites) within the same genomic region (i.e., *H19*: 4 CpG sites, LINE-1: 4 CpG sites, *11β-HSD-2*: 5 CpG sites, and *PPAR-α*:2 CpG sites). For each genomic region, the DNAm percentages at all CpG sites were included as repeated measures of the same variable in models of each outcome. To illustrate, the LINE-1 z-scores at CpG site 1, 2, 3, and 4 were included as four fixed effects in our models, and the same strategy was applied for other genes. This analytical approach was used in previous publications [98].

To examine the relationship between DNAm at Time 1 for LINE-1, 11β -HSD-2, and H19 and each cardiovascular risk factor outcome, separate linear mixed-effects models with a compound symmetry covariance structure were used to model the covariance structure of the repeated outcome assessed at Time 1 and 2. We used linear regression to assess the cross-sectional association between DNAm at *PPAR-α* and the outcomes because this gene was only measured at Time 2. For each exposure, the crude model included only DNAm z-scores at multiple CpG sites for a genomic region. Due to the biological plausibility for the sex and age difference in DNAm, we considered age and sex as mandatory covariates in any fully adjusted model. For the other covariates, we followed a parsimonious approach. Therefore, covariates were adjusted for only if they were potential confounders among our study population based on the significance of their statistical association with each gene of interest (i.e., p < 0.05). We investigated the confounding factors for each genomic region

by examining the distribution of childbirth and follow-up characteristics across quartiles of average DNAm z-scores of all loci within the region using either analysis of variance or Kruskal-Wallis H tests for continuous covariates that were normally and non-normally distributed, respectively, and a chi-squared test for categorical covariates. Based on these investigations to select the confounding factors, only LINE-1 DNAm was associated with breastfeeding duration. Therefore, LINE-1 models included breastfeeding duration, in addition to age and sex (Supplementary Tables S12–S15).

Our mixed-effects models' tables show information about the total sample size (i.e., number of unique subjects), total number of observations used in each model, and number of subjects with repeated measures for each outcome. Our linear regression models' tables show information about the total sample size for each outcome. Collinearity was assessed in the linear regression models using variance inflation factors. We conducted sensitivity analyses. First, we adjusted for the pubertal onset at Time 1 for LINE-1, *11β*-HSD-2, and *H19* and at Time 2 for *PPAR-α* because puberty has been associated with DNAm [39]. We also repeated the analysis after excluding one outlier value in DNAm for *H19*. The SAS statistical software package, version 9.4, was used for analyses (SAS Corp, Cary, NC), and a p < 0.008 was considered a statistically significant association following correction for multiple testing of six outcomes (p < 0.008 or 0.05/6).

5. Conclusions

In conclusion, we observed associations between DNAm at specific CpG sites for LINE-1 and glucose and high-density lipoprotein cholesterol and for 11β -HSD-2 and glucose in a sample of Mexican youth. Our finding supplemented existing knowledge on the potential of epigenetics to identify the molecular mechanism underlying cardiometabolic abnormalities, and it could open the door for targeted interventions among youth. Nevertheless, our results merit further investigation to replicate, validate, and expand on the use of DNAm though carefully designed prospective studies in multiple independent pediatric populations. Moreover, since our study only focused on four genomic regions, we recommend future studies employ epigenome-wide approaches to identify all important genes for these outcomes in youth.

Supplementary Materials: The following supporting information can be downloaded at: https://www.mdpi.com/article/10.3390/epigenomes7010004/s1, Table S1: Spearman's rank correlation coefficients between the DNAm z-scores at LINE-1 CpG sites; Table S2: Spearman's rank correlation coefficients between the DNAm z-scores at 11β -HSD-2 CpG sites; Table S3: Spearman's rank correlation coefficients between the DNAm z-scores at H19 CpG sites; Table S4: Spearman's rank correlation coefficients between the DNAm z-scores at PPAR-α CpG sites; Table S5: Associations between the DNAm z-score at LINE-1 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models Adjusted for Pubertal Onset (n = 242); Table S6: Associations between the DNAm z-score at 11β -HSD-2 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models Adjusted for Pubertal Onset (n = 229); Table S7: Associations between the DNAm z-score at H19 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models (n = 245); Table S8: Associations between the DNAm z-score at H19 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models after the Removal of Outlier DNAm Values (n = 244); Table S9: Associations between the DNAm z-score at H19 and Repeated Measures of Cardiometabolic Risk Factors using Mixed-effects Models Adjusted for Pubertal Onset (n = 245); Table S10: Cross-sectional Associations between the DNAm z-score at *PPAR-\alpha* and Cardiometabolic Risk Factors using Linear Regression Adjusted for Pubertal Onset (n = 345); Table S11: Primer Sequences and Details of CpG Sites Assessed; Table S12: Average DNAm z-score at LINE-1 and Confounder Selection; Table S13: Average DNAm z-score at 11β -HSD-2 and Confounder Selection; Table S14: Average DNAm z-score at H19 and Confounder Selection; Table S15: Average DNAm z-score at *PPAR-* α and Confounder Selection.

Author Contributions: Data provision: J.M.G., D.C.D., K.E.P., A.C., M.M.T.-R. and L.A.T.-O.; conceptualization: A.A.A., J.M.G., K.E.P., H.M.K., E.A.R.-N. and A.B.; data analysis: A.A.A.; writing—original draft: A.A.A.; writing—review and editing: J.M.G. and K.E.P.; supervision: J.M.G. and K.E.P.

All authors provided critically important intellectual contribution to the content and have read. All authors have read and agreed to the published version of the manuscript.

Funding: Funding for the Early Life Exposure in Mexico to ENvironmental Toxicants (ELEMENT) was provided by the U.S. Environmental Protection Agency (US EPA) (RD83480019, RD83543601) and the National Institute for Environmental Health Sciences (NIEHS) (P20 ES018171, P01 ES02284401, and R35 ES031686). No additional financial support was received.

Institutional Review Board Statement: This study was conducted according to the guidelines of the Declaration of Helsinki and approved by the Institutional Review Boards of the University of Michigan and the National Institute of Public Health of Mexico.

Informed Consent Statement: Written informed consents were collected from mothers upon their enrollments in the ELEMENT project with assent from adolescents. No additional consent was required for publication.

Data Availability Statement: The datasets generated during and/or analyzed during the current study are not publicly available due to human subjects' rights, but the data are available upon reasonable request to the ELEMENT PI, Karen Peterson (karenep@umich.edu) for review by the ELEMENT committee.

Acknowledgments: We gratefully acknowledge the mothers and children who participated in the Early Life Exposure in Mexico to Environmental Toxicants (ELEMENT) study and American British Cowdray Medical Center (ABC) for providing facilities for this research.

Conflicts of Interest: The authors declare that there is no conflict of interest.

References

- Di Cesare, M.; Soric, M.; Bovet, P.; Miranda, J.J.; Bhutta, Z.; Stevens, G.A.; Laxmaiah, A.; Kengne, A.P.; Bentham, J. The epidemiological burden of obesity in childhood: A worldwide epidemic requiring urgent action. *BMC Med.* 2019, *17*, 212. [CrossRef] [PubMed]
- Tavares Giannini, D.; Caetano Kuschnir, M.C.; Szklo, M. Metabolic syndrome in overweight and obese adolescents: A comparison of two different diagnostic criteria. *Ann. Nutr. Metab.* 2014, 64, 71–79. [CrossRef] [PubMed]
- 3. Reinehr, T.; de Sousa, G.; Toschke, A.M.; Andler, W. Comparison of metabolic syndrome prevalence using eight different definitions: A critical approach. *Arch. Dis. Child.* **2007**, *92*, 1067–1072. [CrossRef] [PubMed]
- Flouris, A.D.; Bouziotas, C.; Christodoulos, A.D.; Koutedakis, Y. Longitudinal preventive-screening cutoffs for metabolic syndrome in adolescents. *Int. J. Obes.* 2008, 32, 1506–1512. [CrossRef] [PubMed]
- Alberti, K.G.; Eckel, R.H.; Grundy, S.M.; Zimmet, P.Z.; Cleeman, J.I.; Donato, K.A.; Fruchart, J.C.; James, W.P.; Loria, C.M.; Smith, S.C., Jr.; et al. Harmonizing the metabolic syndrome: A joint interim statement of the International Diabetes Federation Task Force on Epidemiology and Prevention; National Heart, Lung, and Blood Institute; American Heart Association; World Heart Federation; International Atherosclerosis Society; and International Association for the Study of Obesity. *Circulation* 2009, 120, 1640–1645. [CrossRef]
- 6. Handelsman, Y. Metabolic syndrome pathophysiology and clinical presentation. *Toxicol. Pathol.* 2009, 37, 18–20. [CrossRef]
- Galassi, A.; Reynolds, K.; He, J. Metabolic syndrome and risk of cardiovascular disease: A meta-analysis. Am. J. Med. 2006, 119, 812–819. [CrossRef] [PubMed]
- Gami, A.S.; Witt, B.J.; Howard, D.E.; Erwin, P.J.; Gami, L.A.; Somers, V.K.; Montori, V.M. Metabolic syndrome and risk of incident cardiovascular events and death: A systematic review and meta-analysis of longitudinal studies. *J. Am. Coll. Cardiol.* 2007, 49, 403–414. [CrossRef]
- Ford, E.S. Risks for all-cause mortality, cardiovascular disease, and diabetes associated with the metabolic syndrome: A summary
 of the evidence. *Diabetes Care* 2005, 28, 1769–1778. [CrossRef]
- 10. Stocks, T.; Bjorge, T.; Ulmer, H.; Manjer, J.; Haggstrom, C.; Nagel, G.; Engeland, A.; Johansen, D.; Hallmans, G.; Selmer, R.; et al. Metabolic risk score and cancer risk: Pooled analysis of seven cohorts. *Int. J. Epidemiol.* **2015**, *44*, 1353–1363. [CrossRef]
- 11. Zimmet, P.; Magliano, D.; Matsuzawa, Y.; Alberti, G.; Shaw, J. The metabolic syndrome: A global public health problem and a new definition. *J. Atheroscler. Thromb.* **2005**, *12*, 295–300. [CrossRef] [PubMed]
- 12. Hong, Y.M. Atherosclerotic cardiovascular disease beginning in childhood. Korean Circ. J. 2010, 40, 1–9. [CrossRef] [PubMed]
- 13. Berenson, G.S.; Srinivasan, S.R.; Bao, W.; Newman, W.P., 3rd; Tracy, R.E.; Wattigney, W.A. Association between multiple cardiovascular risk factors and atherosclerosis in children and young adults. The Bogalusa Heart Study. *N. Engl. J. Med.* **1998**, *338*, 1650–1656. [CrossRef] [PubMed]
- Sinha, R.; Fisch, G.; Teague, B.; Tamborlane, W.V.; Banyas, B.; Allen, K.; Savoye, M.; Rieger, V.; Taksali, S.; Barbetta, G.; et al. Prevalence of impaired glucose tolerance among children and adolescents with marked obesity. *N. Engl. J. Med.* 2002, 346, 802–810. [CrossRef] [PubMed]

- 15. Marcovecchio, M.L.; Patricelli, L.; Zito, M.; Capanna, R.; Ciampani, M.; Chiarelli, F.; Mohn, A. Ambulatory blood pressure monitoring in obese children: Role of insulin resistance. *J. Hypertens.* **2006**, *24*, 2431–2436. [CrossRef] [PubMed]
- 16. D'Adamo, E.; Impicciatore, M.; Capanna, R.; Loredana Marcovecchio, M.; Masuccio, F.G.; Chiarelli, F.; Mohn, A.A. Liver steatosis in obese prepubertal children: A possible role of insulin resistance. *Obesity (Silver Spring)* **2008**, *16*, 677–683. [CrossRef]
- Giannini, C.; Diesse, L.; D'Adamo, E.; Chiavaroli, V.; de Giorgis, T.; Di Iorio, C.; Chiarelli, F.; Mohn, A. Influence of the Mediterranean diet on carotid intima-media thickness in hypercholesterolaemic children: A 12-month intervention study. *Nutr. Metab. Cardiovasc. Dis.* 2014, 24, 75–82. [CrossRef]
- 18. Nicklas, T.A.; von Duvillard, S.P.; Berenson, G.S. Tracking of serum lipids and lipoproteins from childhood to dyslipidemia in adults: The Bogalusa Heart Study. *Int. J. Sports Med.* 2002, 23 (Suppl. 1), S39–S43. [CrossRef]
- 19. Morrison, J.A.; Friedman, L.A.; Wang, P.; Glueck, C.J. Metabolic syndrome in childhood predicts adult metabolic syndrome and type 2 diabetes mellitus 25 to 30 years later. *J. Pediatr.* **2008**, *152*, 201–206. [CrossRef]
- Eloranta, A.M.; Schwab, U.; Venalainen, T.; Kiiskinen, S.; Lakka, H.M.; Laaksonen, D.E.; Lakka, T.A.; Lindi, V. Dietary quality indices in relation to cardiometabolic risk among Finnish children aged 6–8 years—The PANIC study. *Nutr. Metab. Cardiovasc. Dis.* 2016, 26, 833–841. [CrossRef]
- 21. Izquierdo, A.G.; Crujeiras, A.B. Epigenetic biomarkers in metabolic syndrome and obesity. In *Prognostic Epigenetics*; Sharma, S., Ed.; Elsevier: Amsterdam, The Netherlands, 2019; pp. 269–287.
- 22. Zhang, Y.; Zeng, C. Role of DNA methylation in cardiovascular diseases. Clin. Exp. Hypertens. 2016, 38, 261–267. [CrossRef]
- 23. Costantino, S.; Libby, P.; Kishore, R.; Tardif, J.C.; El-Osta, A.; Paneni, F. Epigenetics and precision medicine in cardiovascular patients: From basic concepts to the clinical arena. *Eur. Heart J.* **2018**, *39*, 4150–4158. [CrossRef]
- 24. Prasher, D.; Greenway, S.C.; Singh, R.B. The impact of epigenetics on cardiovascular disease. *Biochem. Cell Biol.* 2020, *98*, 12–22. [CrossRef] [PubMed]
- Agha, G.; Mendelson, M.M.; Ward-Caviness, C.K.; Joehanes, R.; Huan, T.; Gondalia, R.; Salfati, E.; Brody, J.A.; Fiorito, G.; Bressler, J. Blood leukocyte DNA methylation predicts risk of future myocardial infarction and coronary heart disease: A longitudinal study of 11 461 participants from population-based cohorts. *Circulation* 2019, 140, 645–657. [CrossRef] [PubMed]
- Duan, L.; Liu, C.; Hu, J.; Liu, Y.; Wang, J.; Chen, G.; Li, Z.; Chen, H. Epigenetic mechanisms in coronary artery disease: The current state and prospects. *Trends Cardiovasc. Med.* 2018, 28, 311–319. [CrossRef] [PubMed]
- 27. van der Harst, P.; de Windt, L.J.; Chambers, J.C. Translational Perspective on Epigenetics in Cardiovascular Disease. J. Am. Coll. Cardiol. 2017, 70, 590–606. [CrossRef]
- Samblas, M.; Milagro, F.I.; Martinez, A. DNA methylation markers in obesity, metabolic syndrome, and weight loss. *Epigenetics* 2019, 14, 421–444. [CrossRef]
- 29. Soler-Botija, C.; Galvez-Monton, C.; Bayes-Genis, A. Epigenetic Biomarkers in Cardiovascular Diseases. *Front. Genet.* **2019**, 10, 950. [CrossRef]
- 30. Chen, J.; Sun, H.; Tang, W.; Zhou, L.; Xie, X.; Qu, Z.; Chen, M.; Wang, S.; Yang, T.; Dai, Y.; et al. DNA methylation biomarkers in stool for early screening of colorectal cancer. *J. Cancer* 2019, *10*, 5264–5271. [CrossRef]
- 31. Crujeiras, A.B.; Diaz-Lagares, A. DNA methylation in obesity and associated diseases. In *Epigenetic Biomarkers and Diagnostics*; Elsevier: Amsterdam, The Netherlands, 2016; pp. 313–329.
- 32. Kim, M.; Long, T.I.; Arakawa, K.; Wang, R.; Yu, M.C.; Laird, P.W. DNA methylation as a biomarker for cardiovascular disease risk. *PLoS ONE* **2010**, *5*, e9692. [CrossRef]
- 33. Westerman, K.; Sebastiani, P.; Jacques, P.; Liu, S.; DeMeo, D.; Ordovas, J.M. DNA methylation modules associate with incident cardiovascular disease and cumulative risk factor exposure. *Clin. Epigenetics* **2019**, *11*, 142. [CrossRef] [PubMed]
- 34. Antoun, E.; Issarapu, P.; di Gravio, C.; Shrestha, S.; Betts, M.; Saffari, A.; Sahariah, S.A.; Sankareswaran, A.; Arumalla, M.; Prentice, A.M.; et al. DNA methylation signatures associated with cardiometabolic risk factors in children from India and The Gambia: Results from the EMPHASIS study. *Clin. Epigenetics* **2022**, *14*, 6. [CrossRef]
- Day, S.E.; Coletta, R.L.; Kim, J.Y.; Garcia, L.A.; Campbell, L.E.; Benjamin, T.R.; Roust, L.R.; De Filippis, E.A.; Mandarino, L.J.; Coletta, D.K. Potential epigenetic biomarkers of obesity-related insulin resistance in human whole-blood. *Epigenetics* 2017, 12, 254–263. [CrossRef]
- Costantino, S.; Mohammed, S.A.; Ambrosini, S.; Paneni, F. Epigenetic processing in cardiometabolic disease. *Atherosclerosis* 2019, 281, 150–158. [CrossRef] [PubMed]
- 37. Li, E.; Zhang, Y. DNA methylation in mammals. Cold Spring Harb. Perspect. Biol. 2014, 6, a019133. [CrossRef]
- Han, L.; Zhang, H.; Kaushal, A.; Rezwan, F.I.; Kadalayil, L.; Karmaus, W.; Henderson, A.J.; Relton, C.L.; Ring, S.; Arshad, S.H.; et al. Changes in DNA methylation from pre- to post-adolescence are associated with pubertal exposures. *Clin. Epigenetics* 2019, 11, 176. [CrossRef] [PubMed]
- Wu, Y.; Peterson, K.E.; Sanchez, B.N.; Dolinoy, D.C.; Mercado-Garcia, A.; Tellez-Rojo, M.M.; Goodrich, J.M. Association of blood leukocyte DNA methylation at LINE-1 and growth-related candidate genes with pubertal onset and progression. *Epigenetics* 2018, 13, 1222–1233. [CrossRef]
- 40. Goran, M.I.; Gower, B.A. Longitudinal study on pubertal insulin resistance. Diabetes 2001, 50, 2444–2450. [CrossRef]
- Magge, S.N.; Goodman, E.; Armstrong, S.C.; Committee On, N.; Section On, E.; Section On, O. The Metabolic Syndrome in Children and Adolescents: Shifting the Focus to Cardiometabolic Risk Factor Clustering. *Pediatrics* 2017, 140, e20171603. [CrossRef]

- 42. Ardeljan, D.; Taylor, M.S.; Ting, D.T.; Burns, K.H. The Human Long Interspersed Element-1 Retrotransposon: An Emerging Biomarker of Neoplasia. *Clin. Chem.* 2017, *63*, 816–822. [CrossRef]
- 43. Beck, C.R.; Garcia-Perez, J.L.; Badge, R.M.; Moran, J.V. LINE-1 elements in structural variation and disease. *Annu. Rev. Genom. Hum. Genet.* **2011**, *12*, 187–215. [CrossRef] [PubMed]
- 44. Sant, K.E.; Goodrich, J.M. Methods for Analysis of DNA Methylation. In *Toxicoepigenetics*; McCullough, S.D., Dolinoy, D.C., Eds.; Elsevier: Amsterdam, The Netherlands, 2019; pp. 347–377.
- 45. Muka, T.; Koromani, F.; Portilla, E.; O'Connor, A.; Bramer, W.M.; Troup, J.; Chowdhury, R.; Dehghan, A.; Franco, O.H. The role of epigenetic modifications in cardiovascular disease: A systematic review. *Int. J. Cardiol.* **2016**, *212*, 174–183. [CrossRef] [PubMed]
- Lai, S.; Du, K.; Shi, Y.; Li, C.; Wang, G.; Hu, S.; Jia, X.; Wang, J.; Chen, S. Long Non-Coding RNAs in Brown Adipose Tissue. Diabetes Metab. Syndr. Obes. 2020, 13, 3193–3204. [CrossRef] [PubMed]
- Schmidt, E.; Dhaouadi, I.; Gaziano, I.; Oliverio, M.; Klemm, P.; Awazawa, M.; Mitterer, G.; Fernandez-Rebollo, E.; Pradas-Juni, M.; Wagner, W.; et al. LincRNA H19 protects from dietary obesity by constraining expression of monoallelic genes in brown fat. *Nat. Commun.* 2018, *9*, 3622. [CrossRef] [PubMed]
- Huang, R.C.; Galati, J.C.; Burrows, S.; Beilin, L.J.; Li, X.; Pennell, C.E.; van Eekelen, J.; Mori, T.A.; Adams, L.A.; Craig, J.M. DNA methylation of the IGF2/H19 imprinting control region and adiposity distribution in young adults. *Clin. Epigenetics* 2012, *4*, 21. [CrossRef]
- Bowman, A.; Peterson, K.E.; Dolinoy, D.C.; Meeker, J.D.; Sanchez, B.N.; Mercado-Garcia, A.; Tellez-Rojo, M.M.; Goodrich, J.M. Phthalate Exposures, DNA Methylation and Adiposity in Mexican Children Through Adolescence. *Front. Public Health* 2019, 7, 162. [CrossRef]
- 50. Bray, G.A.; Heisel, W.E.; Afshin, A.; Jensen, M.D.; Dietz, W.H.; Long, M.; Kushner, R.F.; Daniels, S.R.; Wadden, T.A.; Tsai, A.G.; et al. The Science of Obesity Management: An Endocrine Society Scientific Statement. *Endocr. Rev.* **2018**, *39*, 79–132. [CrossRef]
- 51. Patel, H.; Dhangar, K.; Sonawane, Y.; Surana, S.; Karpoormath, R.; Thapliyal, N.; Shaikh, M.; Noolvi, M.; Jagtap, R. In search of selective 11β-HSD type 1 inhibitors without nephrotoxicity: An approach to resolve the metabolic syndrome by virtual based screening. *Arab. J. Chem.* 2018, *11*, 221–232. [CrossRef]
- 52. Hintzpeter, J.; Stapelfeld, C.; Loerz, C.; Martin, H.J.; Maser, E. Green tea and one of its constituents, Epigallocatechine-3-gallate, are potent inhibitors of human 11beta-hydroxysteroid dehydrogenase type 1. *PLoS ONE* **2014**, *9*, e84468. [CrossRef]
- Friso, S.; Pizzolo, F.; Choi, S.W.; Guarini, P.; Castagna, A.; Ravagnani, V.; Carletto, A.; Pattini, P.; Corrocher, R.; Olivieri, O. Epigenetic control of 11 beta-hydroxysteroid dehydrogenase 2 gene promoter is related to human hypertension. *Atherosclerosis* 2008, 199, 323–327. [CrossRef]
- 54. Drake, A.J.; McPherson, R.C.; Godfrey, K.M.; Cooper, C.; Lillycrop, K.A.; Hanson, M.A.; Meehan, R.R.; Seckl, J.R.; Reynolds, R.M. An unbalanced maternal diet in pregnancy associates with offspring epigenetic changes in genes controlling glucocorticoid action and foetal growth. *Clin. Endocrinol.* **2012**, *77*, 808–815. [CrossRef] [PubMed]
- Krupp, D.; Shi, L.; Maser-Gluth, C.; Pietzarka, M.; Remer, T. 11beta Hydroxysteroid dehydrogenase type 2 and dietary acid load are independently associated with blood pressure in healthy children and adolescents. *Am. J. Clin. Nutr.* 2013, 97, 612–620. [CrossRef]
- Mussig, K.; Remer, T.; Haupt, A.; Gallwitz, B.; Fritsche, A.; Haring, H.U.; Maser-Gluth, C. 11beta-hydroxysteroid dehydrogenase 2 activity is elevated in severe obesity and negatively associated with insulin sensitivity. *Obesity (Silver Spring)* 2008, 16, 1256–1260. [CrossRef]
- 57. Jang, C.; Obeyesekere, V.R.; Dilley, R.J.; Krozowski, Z.; Inder, W.J.; Alford, F.P. Altered activity of 11beta-hydroxysteroid dehydrogenase types 1 and 2 in skeletal muscle confers metabolic protection in subjects with type 2 diabetes. *J. Clin. Endocrinol. Metab.* 2007, *92*, 3314–3320. [CrossRef]
- 58. Barbosa-Cortes, L.; Villasis-Keever, M.A.; Del Prado-Manriquez, M.; Lopez-Alarcon, M. Adiposity and Insulin Resistance in Children from a Rural Community in Mexico. *Arch. Med. Res.* **2015**, *46*, 214–220. [CrossRef] [PubMed]
- 59. Bougarne, N.; Weyers, B.; Desmet, S.J.; Deckers, J.; Ray, D.W.; Staels, B.; De Bosscher, K. Molecular Actions of PPARalpha in Lipid Metabolism and Inflammation. *Endocr. Rev.* 2018, *39*, 760–802. [CrossRef]
- Burri, L.; Thoresen, G.H.; Berge, R.K. The Role of PPARalpha Activation in Liver and Muscle. *PPAR Res.* 2010, 2010, 542359. [CrossRef]
- Castellano-Castillo, D.; Moreno-Indias, I.; Sanchez-Alcoholado, L.; Ramos-Molina, B.; Alcaide-Torres, J.; Morcillo, S.; Ocana-Wilhelmi, L.; Tinahones, F.; Queipo-Ortuno, M.I.; Cardona, F. Altered Adipose Tissue DNA Methylation Status in Metabolic Syndrome: Relationships Between Global DNA Methylation and Specific Methylation at Adipogenic, Lipid Metabolism and Inflammatory Candidate Genes and Metabolic Variables. *J. Clin. Med.* 2019, *8*, 87. [CrossRef] [PubMed]
- 62. Contreras, A.V.; Torres, N.; Tovar, A.R. PPAR-alpha as a key nutritional and environmental sensor for metabolic adaptation. *Adv. Nutr.* **2013**, *4*, 439–452. [CrossRef]
- 63. Guarrera, S.; Fiorito, G.; Onland-Moret, N.C.; Russo, A.; Agnoli, C.; Allione, A.; Di Gaetano, C.; Mattiello, A.; Ricceri, F.; Chiodini, P.; et al. Gene-specific DNA methylation profiles and LINE-1 hypomethylation are associated with myocardial infarction risk. *Clin. Epigenetics* **2015**, *7*, 133. [CrossRef]
- 64. Wei, L.; Liu, S.; Su, Z.; Cheng, R.; Bai, X.; Li, X. LINE-1 hypomethylation is associated with the risk of coronary heart disease in Chinese population. *Arq. Bras. Cardiol.* **2014**, *102*, 481–488. [CrossRef]

- 65. Baccarelli, A.; Wright, R.; Bollati, V.; Litonjua, A.; Zanobetti, A.; Tarantini, L.; Sparrow, D.; Vokonas, P.; Schwartz, J. Ischemic heart disease and stroke in relation to blood DNA methylation. *Epidemiology* **2010**, *21*, 819–828. [CrossRef] [PubMed]
- María Martín-Núñez, G.; Rubio-Martín, E.; Cabrera-Mulero, R.; Rojo-Martínez, G.; Olveira, G.; Valdés, S.; Soriguer, F.; Castano, L.; Morcillo, S. Type 2 diabetes mellitus in relation to global LINE-1 DNA methylation in peripheral blood: A cohort study. *Epigenetics* 2014, 9, 1322–1328. [CrossRef] [PubMed]
- 67. Turcot, V.; Tchernof, A.; Deshaies, Y.; Perusse, L.; Belisle, A.; Marceau, S.; Biron, S.; Lescelleur, O.; Biertho, L.; Vohl, M.C. LINE-1 methylation in visceral adipose tissue of severely obese individuals is associated with metabolic syndrome status and related phenotypes. *Clin. Epigenetics* **2012**, *4*, 10. [CrossRef] [PubMed]
- 68. Perng, W.; Mora-Plazas, M.; Marin, C.; Rozek, L.S.; Baylin, A.; Villamor, E. A prospective study of LINE-1DNA methylation and development of adiposity in school-age children. *PLoS ONE* **2013**, *8*, e62587. [CrossRef]
- Dunstan, J.; Bressler, J.P.; Moran, T.H.; Pollak, J.S.; Hirsch, A.G.; Bailey-Davis, L.; Glass, T.A.; Schwartz, B.S. Associations of LEP, CRH, ICAM-1, and LINE-1 methylation, measured in saliva, with waist circumference, body mass index, and percent body fat in mid-childhood. *Clin. Epigenetics* 2017, *9*, 29. [CrossRef]
- 70. Rezaei, M.; Andrieu, T.; Neuenschwander, S.; Bruggmann, R.; Mordasini, D.; Frey, F.J.; Vogt, B.; Frey, B.M. Regulation of 11beta-hydroxysteroid dehydrogenase type 2 by microRNA. *Hypertension* **2014**, *64*, 860–866. [CrossRef]
- Campino, C.; Martinez-Aguayo, A.; Baudrand, R.; Carvajal, C.A.; Aglony, M.; Garcia, H.; Padilla, O.; Kalergis, A.M.; Fardella, C.E. Age-related changes in 11beta-hydroxysteroid dehydrogenase type 2 activity in normotensive subjects. *Am. J. Hypertens.* 2013, 26, 481–487. [CrossRef]
- 72. Kargl, C.; Arshad, M.; Salman, F.; Schurman, R.C.; Del Corral, P. 11beta-hydroxysteroid dehydrogenase type-II activity is affected by grapefruit juice and intense muscular work. *Arch. Endocrinol. Metab.* **2017**, *61*, 556–561. [CrossRef] [PubMed]
- 73. Huang, Y.; Pastor, W.A.; Shen, Y.; Tahiliani, M.; Liu, D.R.; Rao, A. The behaviour of 5-hydroxymethylcytosine in bisulfite sequencing. *PLoS ONE* **2010**, *5*, e8888. [CrossRef]
- 74. Breton, C.V.; Marsit, C.J.; Faustman, E.; Nadeau, K.; Goodrich, J.M.; Dolinoy, D.C.; Herbstman, J.; Holland, N.; LaSalle, J.M.; Schmidt, R.; et al. Small-Magnitude Effect Sizes in Epigenetic End Points are Important in Children's Environmental Health Studies: The Children's Environmental Health and Disease Prevention Research Center's Epigenetics Working Group. *Environ. Health Perspect.* 2017, 125, 511–526. [CrossRef] [PubMed]
- Perng, W.; Tamayo-Ortiz, M.; Tang, L.; Sanchez, B.N.; Cantoral, A.; Meeker, J.D.; Dolinoy, D.C.; Roberts, E.F.; Martinez-Mier, E.A.; Lamadrid-Figueroa, H.; et al. Early Life Exposure in Mexico to ENvironmental Toxicants (ELEMENT) Project. *BMJ Open* 2019, 9, e030427. [CrossRef]
- 76. Wu, Y.; Goodrich, J.M.; Dolinoy, D.C.; Sanchez, B.N.; Ruiz-Narvaez, E.A.; Banker, M.; Cantoral, A.; Mercado-Garcia, A.; Tellez-Rojo, M.M.; Peterson, K.E. Accelerometer-measured Physical Activity, Reproductive Hormones, and DNA Methylation. *Med. Sci. Sports Exerc.* 2020, 52, 598–607. [CrossRef] [PubMed]
- 77. Grunau, C.; Clark, S.J.; Rosenthal, A. Bisulfite genomic sequencing: Systematic investigation of critical experimental parameters. *Nucleic Acids Res.* **2001**, *29*, e65. [CrossRef]
- Goodrich, J.M.; Sanchez, B.N.; Dolinoy, D.C.; Zhang, Z.; Hernandez-Avila, M.; Hu, H.; Peterson, K.E.; Tellez-Rojo, M.M. Quality control and statistical modeling for environmental epigenetics: A study on in utero lead exposure and DNA methylation at birth. *Epigenetics* 2015, 10, 19–30. [CrossRef] [PubMed]
- 79. Virani, S.; Dolinoy, D.C.; Halubai, S.; Jones, T.R.; Domino, S.E.; Rozek, L.S.; Nahar, M.S.; Padmanabhan, V. Delivery type not associated with global methylation at birth. *Clin. Epigenetics* **2012**, *4*, 8. [CrossRef] [PubMed]
- Hoyo, C.; Murtha, A.P.; Schildkraut, J.M.; Jirtle, R.L.; Demark-Wahnefried, W.; Forman, M.R.; Iversen, E.S.; Kurtzberg, J.; Overcash, F.; Huang, Z.; et al. Methylation variation at IGF2 differentially methylated regions and maternal folic acid use before and during pregnancy. *Epigenetics* 2011, 6, 928–936. [CrossRef]
- Goodrich, J.M.; Dolinoy, D.C.; Sanchez, B.N.; Zhang, Z.; Meeker, J.D.; Mercado-Garcia, A.; Solano-Gonzalez, M.; Hu, H.; Tellez-Rojo, M.M.; Peterson, K.E. Adolescent epigenetic profiles and environmental exposures from early life through peri-adolescence. *Environ. Epigenet.* 2016, 2, dvw018. [CrossRef]
- Jansen, E.C.; Dolinoy, D.; Peterson, K.E.; O'Brien, L.M.; Chervin, R.D.; Cantoral, A.; Tellez-Rojo, M.M.; Solano-Gonzalez, M.; Goodrich, J. Adolescent sleep timing and dietary patterns in relation to DNA methylation of core circadian genes: A pilot study of Mexican youth. *Epigenetics* 2021, 16, 894–907. [CrossRef]
- 83. Betanzos-Robledo, L.; Rodriguez-Carmona, Y.; Contreras-Manzano, A.; Lamadrid-Figueroa, H.; Jansen, E.; Tellez-Rojo, M.M.; Perng, W.; Peterson, K.; Hebert, J.R.; Shivappa, N.; et al. Greater cumulative exposure to a pro-inflammatory diet is associated with higher metabolic syndrome score and blood pressure in young Mexican adults. *Nutr. Res.* **2020**, *81*, 81–89. [CrossRef]
- 84. Perng, W.; Fernandez, C.; Peterson, K.E.; Zhang, Z.; Cantoral, A.; Sanchez, B.N.; Solano-Gonzalez, M.; Tellez-Rojo, M.M.; Baylin, A. Dietary Patterns Exhibit Sex-Specific Associations with Adiposity and Metabolic Risk in a Cross-Sectional Study in Urban Mexican Adolescents. *J. Nutr.* **2017**, *147*, 1977–1985. [CrossRef]
- 85. Kasper, N.; Peterson, K.E.; Zhang, Z.; Ferguson, K.K.; Sanchez, B.N.; Cantoral, A.; Meeker, J.D.; Tellez-Rojo, M.M.; Pawlowski, C.M.; Ettinger, A.S. Association of Bisphenol A Exposure with Breastfeeding and Perceived Insufficient Milk Supply in Mexican Women. *Matern. Child Health J.* **2016**, *20*, 1713–1719. [CrossRef]

- Ettinger, A.S.; Lamadrid-Figueroa, H.; Mercado-Garcia, A.; Kordas, K.; Wood, R.J.; Peterson, K.E.; Hu, H.; Hernandez-Avila, M.; Tellez-Rojo, M.M. Effect of calcium supplementation on bone resorption in pregnancy and the early postpartum: A randomized controlled trial in Mexican women. *Nutr. J.* 2014, *13*, 116. [CrossRef]
- Rodriguez-Ramirez, S.; Mundo-Rosas, V.; Jimenez-Aguilar, A.; Shamah-Levy, T. Methodology for the analysis of dietary data from the Mexican National Health and Nutrition Survey 2006. *Salud Publica Mex.* 2009, *51* (Suppl. 4), S523–S529. [CrossRef] [PubMed]
- 88. Ramírez Silva, I.; Barragán-Vázquez, S.; Rodríguez-Ramírez, S.; Rivera-Dommarco, J.A.; Mejía-Rodríguez, F.; Barquera-Cervera, S.; Tolentino-Mayo, L.; Flores-Aldana, M.; Villalpando-Hernández, S.; Ancira-Moreno, M. Base de Alimentos de México (BAM): Compilación de la Composición de los Alimentos Frecuentemente Consumidos en el país, Version 18.1.1. 2021. Available online: https://insp.mx/informacion-relevante/bam-bienvenida (accessed on 1 January 2023).
- Hernández, B.; Gortmaker, S.L.; Laird, N.M.; Colditz, G.A.; Parra-Cabrera, S.; Peterson, K.E. Validez y reproducibilidad de un cuestionario de actividad e inactividad física para escolares de la ciudad de México. *Salud Pública México* 2000, 42, 315–323. [CrossRef]
- Ainsworth, B.E.; Haskell, W.L.; Whitt, M.C.; Irwin, M.L.; Swartz, A.M.; Strath, S.J.; O'Brien, W.L.; Bassett, D.R., Jr.; Schmitz, K.H.; Emplaincourt, P.O.; et al. Compendium of physical activities: An update of activity codes and MET intensities. *Med. Sci. Sports Exerc.* 2000, 32, S498–S504. [CrossRef] [PubMed]
- 91. Marshall, W.A.; Tanner, J.M. Variations in pattern of pubertal changes in girls. Arch. Dis. Child. 1969, 44, 291–303. [CrossRef]
- 92. Marshall, W.A.; Tanner, J.M. Variations in the pattern of pubertal changes in boys. Arch. Dis. Child. 1970, 45, 13–23. [CrossRef]
- Chavarro, J.E.; Watkins, D.J.; Afeiche, M.C.; Zhang, Z.; Sanchez, B.N.; Cantonwine, D.; Mercado-Garcia, A.; Blank-Goldenberg, C.; Meeker, J.D.; Tellez-Rojo, M.M.; et al. Validity of Self-Assessed Sexual Maturation Against Physician Assessments and Hormone Levels. J. Pediatr. 2017, 186, 172–178.e173. [CrossRef]
- LaBarre, J.L.; Peterson, K.E.; Kachman, M.T.; Perng, W.; Tang, L.; Hao, W.; Zhou, L.; Karnovsky, A.; Cantoral, A.; Tellez-Rojo, M.M.; et al. Mitochondrial Nutrient Utilization Underlying the Association Between Metabolites and Insulin Resistance in Adolescents. J. Clin. Endocrinol. Metab. 2020, 105, 2442–2455. [CrossRef]
- Aljahdali, A.A.; Peterson, K.E.; Cantoral, A.; Ruiz-Narvaez, E.; Tellez-Rojo, M.M.; Kim, H.M.; Hebert, J.R.; Wirth, M.D.; Torres-Olascoaga, L.A.; Shivappa, N.; et al. Diet Quality Scores and Cardiometabolic Risk Factors in Mexican Children and Adolescents: A Longitudinal Analysis. *Nutrients* 2022, 14, 896. [CrossRef] [PubMed]
- 96. Aljahdali, A.A.; Baylin, A.; Ruiz-Narvaez, E.A.; Kim, H.M.; Cantoral, A.; Tellez-Rojo, M.M.; Banker, M.; Peterson, K.E. Sedentary patterns and cardiometabolic risk factors in Mexican children and adolescents: Analysis of longitudinal data. *Int. J. Behav. Nutr. Phys. Act.* **2022**, *19*, 143. [CrossRef] [PubMed]
- Needham, B.L.; Smith, J.A.; Zhao, W.; Wang, X.; Mukherjee, B.; Kardia, S.L.; Shively, C.A.; Seeman, T.E.; Liu, Y.; Diez Roux, A.V. Life course socioeconomic status and DNA methylation in genes related to stress reactivity and inflammation: The multi-ethnic study of atherosclerosis. *Epigenetics* 2015, *10*, 958–969. [CrossRef] [PubMed]
- Smith, J.A.; Zhao, W.; Wang, X.; Ratliff, S.M.; Mukherjee, B.; Kardia, S.L.R.; Liu, Y.; Roux, A.V.D.; Needham, B.L. Neighborhood characteristics influence DNA methylation of genes involved in stress response and inflammation: The Multi-Ethnic Study of Atherosclerosis. *Epigenetics* 2017, 12, 662–673. [CrossRef] [PubMed]

Disclaimer/Publisher's Note: The statements, opinions and data contained in all publications are solely those of the individual author(s) and contributor(s) and not of MDPI and/or the editor(s). MDPI and/or the editor(s) disclaim responsibility for any injury to people or property resulting from any ideas, methods, instructions or products referred to in the content.