



Article Spatial Non-Stationarity-Based Landslide Susceptibility Assessment Using PCAMGWR Model

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Abstract: Landslide Susceptibility Assessment (LSA) is a fundamental component of landslide risk management and a substantial area of geospatial research. Previous researchers have considered the spatial non-stationarity relationship between landslide occurrences and Landslide Conditioning Factors (*LCFs*) as fixed effects. The fixed effects consider the spatial non-stationarity scale between different *LCFs* as an average value, which is represented by a single bandwidth in the Geographically Weighted Regression (GWR) model. The present study analyzes the non-stationarity scale effect of the spatial relationship between *LCFs* and landslides and explains the influence of factor correlation on the LSA. A Principal-Component-Analysis-based Multiscale GWR (PCAMGWR) model is proposed for landslide susceptibility mapping, in which hexagonal neighborhoods express spatial proximity and extract *LCFs* as the model input. The area under the receiver operating characteristic curve and other statistical indicators are used to compare the PCAMGWR model with other GWR-based models and global regression models, and the PCAMGWR model has the best prediction effect. Different spatial non-stationarity scales are obtained and improve the prediction accuracy of landslide susceptibility compared to a single spatial non-stationarity scale.

Keywords: landslide susceptibility; PCAMGWR model; spatial non-stationarity; factor correlation; spatial proximity; hexagonal neighborhoods

1. Introduction

Landslides are one of the most destructive and catastrophic geohazards worldwide and threaten the safety of humans and property in mountainous areas [1,2]. Located on the junction of the Asia–Europe plate, the Indian Ocean plate, and the Pacific plate, China has the characteristics of active and complex geological tectonic activities, numerous climate types, and frequent human activities, which leads to frequent geohazards. According to the "Chinese Geological Hazard Bulletin" [3], there were 166,828 geohazards in China from 2008 to 2020, of which 111,621 were landslide hazards, accounting for 66.9%. Landslide susceptibility refers to the occurrence possibility of a landslide under the combined effect of Landslide Conditioning Factors (*LCFs*) and predicts the location and probability of a landslide in a specific area [4]. Therefore, it is necessary to take measures to assess landslide susceptibility, and Landslide Susceptibility Mapping (LSM) is a common method for geohazard prevention.

Previous studies have investigated Landslide Susceptibility Assessment (LSA) in recent years, primarily relying on Geographic Information Systems (GIS) for qualitative analysis in the early stage [5–10]. Some methods of sorting and weighting parameters in qualitative research already pertain to the category of semiquantitative analysis [11,12]. The quantitative analysis method is based on the numerical expression of the relationship between *LCFs* and landslide occurrence [13], mainly including physics-based methods [14], statistical analysis [15,16], and artificial neural networks [17]. The statistical analysis



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Copyright: © 2022 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). includes bivariate and multivariate techniques [18,19]. Multivariate statistical analysis is mainly expressed in regression models, which consist of global regression models, such as logistic regression [20], and local regression models, such as Geographically Weighted Regression (GWR) [21]. The global regression models consider the influence of *LCFs* on landslides as stable for a region. The local regression models consider spatial variation in *LCFs* and landslides. The problem of spatial autocorrelation and spatial non-stationarity is well-known in LSA research. For example, Brenning estimates error rates for predicting "present" and "future" landslides based on a resampling approach that takes into account spatial autocorrelation [22]. The second law of geography states that there is widespread spatial heterogeneity in the relationship between given geospatial variables, namely spatial non-stationarity [23]. Since LSA is based on geospatial data analysis, the relationship between *LCFs* and landslides may be spatial non-stationarity [24]. The presence of spatial non-stationarity demonstrates that the traditional models are only applicable to the case of stationary spatial relations in the study area and cannot accurately fit the local relations [25]. The models which are widely used in LSA cannot thus express spatial non-stationarity.

The GWR model is a local regression model widely adopted in the research into spatial non-stationarity, which presumes an optimal spatial average based on a single bandwidth. Nowadays, GWR is widely used in research fields such as land use [26], housing price prediction [27], predicting forest-fire kernel density at multiple scales [28], and ecological environment protection [29]. In terms of geohazard susceptibility assessment, GWR is gradually becoming a key instrument in LSA. GWR considers the spatial variability of parameters to focus on the spatial non-stationarity of landslide predisposing factors [24,30–33]. The above studies of LSA are all based on the basic GWR model using a single optimal average scale to detect spatial non-stationarity, which leads to the fact that the spatial variation of all parameter estimates manifests the same scale characteristics. The basic GWR model is based on the assumption of a "best average", which ignores scale differences in local variation relationships and may produce unreliable results.

GWR performance depends on the spatial variation of the relationships between dependent and independent variables. The bandwidth is a constant distance in the fixed kernel to measure the spatial variation that plays a role in implementing the spatial effect of the observed points and neighbors [34]. However, the basic GWR model ignores the multiple spatial data relationships corresponding to spatial scale variations and uses the "best average" scale (single bandwidth) to reflect the spatial changes of all parameter estimates. Adopting multiple bandwidths could give Multiscale GWR (MGWR) the capability to potentially differentiate the scale of local, regional, and global processes by comparing the optimal bandwidths for different independent variables [35]. Researchers do not treat spatial non-stationarity in detail, and there is still a research gap regarding the study of the spatial non-stationarity scale effect of LSA. It is increasingly important to explore local, regional, and global processes in LSA.

Spatial non-stationarity prediction models and spatial relationships are crucial in LSA. The prediction model affects the accuracy of assessment results, and the spatial relationship impacts the input conditions of the spatial non-stationarity model. The spatial relation includes topological relation, metric relation, and azimuth relation. Spatial proximity relates to rhythmic connection to measure the distance between two neighborhoods in space, which is widely used in spatial data analysis and geographic information extraction [34,36]. The GWR model was established based on the spatial proximity of neighborhoods, with homogeneity within neighborhoods, and the heterogeneity between neighborhoods. Therefore, it is essential to choose an appropriate spatial proximity expression for expressing spatial proximity more precisely. Triangular tessellation uses triangles with two directions and is unpopular in studying spatial non-stationarity [31]. Administrative-district-based neighborhoods [37] and rectangle- or square-grid (Moore) boundaries [31] have been commonly used as the spatial proximity expressions of the GWR model in previous studies. Slope-unit-based neighborhoods are segmented to express the spatial proximity in land-slide susceptibility, which indicates an improvement in prediction accuracy compared

to grid units [24]. Slope-unit-based neighborhoods are considered to be more consistent with the topography [34]. Additionally, a hexagonal neighborhood is considered a better spatial structure for continuously dividing a two-dimensional space with an isotropic neighborhood and results in a better study of spatial non-stationarity [38,39].

In the previous assessments of landslide susceptibility considering spatial non-stationarity, there are gaps in the exploration of the spatial non-stationarity scale effect and the comparison of spatial proximity expressions. In the present study, the importance of assessing landslide susceptibility is analyzed from the spatial non-stationarity scale effect for the study area. Qingchuan county was the study area, where ten *LCFs* were selected. Moore neighborhoods, slope-unit-based neighborhoods, and hexagonal neighborhoods were established and compared. Factor analysis was then carried out to analyze the factor correction, and Exploratory Spatial Data Analysis (ESDA) was conducted to measure the spatial autocorrelation visualization. The PCAMGWR model was proposed to study the non-stationarity scale effect of the spatial relationship between *LCFs* and landslides, which employs PCA as a dimension reduction process and MGWR as a multiscale bandwidth acquisition method. The impact of the spatial non-stationarity effect on the LSA and different spatial scales of *LCFs* was obtained. Finally, the Receiver Operating Characteristic (ROC) curve was employed to verify the proposed model.

2. Study Area and Dataset

2.1. Overview of the Study Area

The southeastern area of China is a region with frequently occurring landslide hazards. The study area of Qingchuan County is a mountainous region on the northern edge of the Sichuan basin, in the southeast of China, adjacent to Shanxi province and Gansu province. The area is located at a latitude of 32°12′ N to 32°56′ N and a longitude of 104°36′ E to 105°38′ E, with an area of 3216 km². It has a complex topography mainly composed of mountains, hills, tablelands, valleys, and small flat dams, and the relative elevation is between 500 m and 3820 m.

The climate in the study area is a typically subtropical humid monsoonal climate that is hot and humid in summer and mild and arid in winter. The surface water system is developed, and the Bailong River and Qingzhu River run through the territory. The total water storage is above 15.7 billion m³, and the water energy reserves are larger than 1 million kW. The soil types are diverse, including yellow loam, yellow-brown loam, dark brown loam, and subalpine meadow soil. The dominant lithology is magmatic rocks, metamorphic rocks, and clastic rocks. Three fault zones are distributed in parallel, from northeast to southwest.

Qingchuan County is recognized as one of the most landslide-prone areas of China [40], and the landslide inventory map in Qingchuan is shown in Figure 1. Most landslide events are induced by natural environmental factors, such as rainfall and earthquakes, while only a few events are induced by human factors. Landslides are pivotal disturbances to the social development and socioeconomic growth of the region. LSM considering the spatial non-stationarity scale effect could significantly prevent the issue.



Figure 1. Landslide inventory map and location of Qingchuan County.

2.2. Dataset Preparation

In order to conduct the spatial non-stationarity exploration, the investigation of the spatial dataset was a continuing concern regarding the LSA. The main data included a vector map of contour lines, a geological map, settlement coordinates, aerial photographs, precipitation data, and vegetation coverage types, which were obtained from the Ministry of Land and Resources. After the field survey and aerial photo interpretation, 973 landslides were counted (Figure 1), and the *LCFs* played an extraordinary role in modeling the LSA [41]. Since there are no standard criteria for selecting *LCFs*, this study considered the general characteristics, working scale, and availability of the proper datasets [42–44]. Ten factors, namely elevation, slope, aspect, terrain relief, lithology, distance to fault zones, distance to stream, precipitation, vegetation coverage types, and distance to settlement, were obtained from GIS (Figure 2).

The digital elevation model (DEM) data at a spatial resolution of 10 m were first obtained by the vector map of contour lines in the GIS environment. The DEM data were used to extract the elevation, slope, aspect, terrain relief, and other conditioning factors describing the topography and geomorphology [45]. Precipitation data and vegetation coverage types were processed through interpolation, collected from different organizations and government departments. Distance to fault zones, distance to stream, and distance to settlement were calculated using a Euclidean distance tool in the GIS environment. In this study, the original data were used as the input of the model for continuous variables, and the unclassified variables were numbered according to their categories before being input into the model.



Figure 2. Zoning map of (a) elevation, (b) terrain relief, (c) slope, (d) aspect, (e) lithology, (f) distance to fault zones, (g) distance to stream, (h) distance to settlement, (i) vegetation coverage types, and (j) precipitation.

3. Materials and Methods

3.1. Flowchart

The methodological approach in this study was a methodology combining geospatial statistics and geohazard assessment and is shown in Figure 3. The research process consisted of the following five steps:

- 1. Three spatial neighborhood expressions were constructed in GIS—Moore neighborhoods, slope-unit-based neighborhoods, and hexagonal neighborhoods. The segmentation metric function proposed by Espindola [46] was then used for the prime spatial proximity expression and the extracted *LCF* was used as the input of the PCAMGWR model.
- 2. Based on the geoenvironmental condition of the study area, *LCFs* were selected, and thematic layers of *LCFs* were prepared. Then, the *LCFs* were analyzed using Pearson correlation analysis and multicollinearity test.
- 3. ESDA was used to investigate the validity of global regression, and the residual obtained by Ordinary Least Squares (OLS) was analyzed based on Moran's I autocorrelation.
- 4. PCAMGWR model was established for exploring the influence of spatial non-stationarity and factor correlation on LSM.
- 5. The accuracy of the proposed model was verified using statistical measures, and the spatial non-stationarity scale effect was analyzed and compared.



Figure 3. Methodology of research applied in this study.

3.2. Expression of Spatial Proximity Selection Method

In this study, three expressions of spatial proximity, namely Moore neighborhoods, slope-unit-based neighborhoods, and hexagonal neighborhoods, were established. The comparison and selection of spatial adjacency expression proceeded in three dimensions, which were homogeneity within a neighborhood, heterogeneity between neighborhoods, and RMSE value of neighborhood area. RMSE values were calculated for the area of the slope-unit-based neighborhoods and hexagonal neighborhoods. The smaller the RMSE value, the more stable the spatial proximity.

Espindola [46] proposed a function to ensure the expression of the spatial proximity of the study area. The function measures the quality of spatial adjacency expression by maximizing homogeneity within and heterogeneity between neighborhoods. The objective function combines the variance measure and the autocorrelation measure:

$$F(v, I) = F(v) + F(I)$$
(1)

where v is the intersegment variance within the neighborhood and I is the Moran's I to assess the intersegmental heterogeneity between neighborhoods.

3.3. Factor Analysis

3.3.1. Correlation Analysis

Pearson correlation analysis quantifies and interprets the correlation between *LCFs* [47], and the Pearson correlation coefficient is (-1, 1). If the coefficient is greater than 0, the larger the coefficient is, the stronger the positive correlation. If the coefficient is less than 0, the smaller the coefficient is, the stronger the negative correlation. The Pearson correlation coefficient is defined as follows [48]:

$$r = \frac{\left(N\sum LCF_i LCF_j - \sum LCF_i \sum LCF_j\right)}{\sqrt{NLCF_i^2 - \left(\sum LCF_i\right)^2}\sqrt{NLCF_j^2 - \left(\sum LCF_j\right)^2}}$$
(2)

where *r* is the correlation coefficient between two *LCFs* and *N* is the number of observations.

3.3.2. Multicollinearity Test

Multicollinearity refers to the fact that model estimates are distorted or difficult to estimate accurately due to the existence of accurate or highly correlated relationships between explanatory variables in a linear regression model. Multicollinearity between *LCFs* may reduce the prediction accuracy of linear regression models, such as GAM-style GWR methods, and the permutation importance of some models, such as (bagged or boosted) tree-based methods, may be faulty. Therefore, the multicollinearity test must be performed before GWR-based model prediction [49]. Variance Inflation Factor (*VIF*) and TOLerance (TOL) standards were applied for the multicollinearity test [50]. The threshold value of the TOL \leq 2 and *VIF* \geq 5 indicate the existence of multicollinearity between the *LCFs* [50]. The *VIF* can be obtained using Equation (3) [51]:

$$VIF = \frac{1}{1 - R_i^2} \tag{3}$$

where R_i is the negative correlation coefficient for regression analysis of the *i*th independent variable to other independent variables.

3.4. ESDA

ESDA denotes whether spatial autocorrelation exists and proves the explicit and quantitative spatial assessment of geographical change [52]. Spatial autocorrelation is a connatural peculiarity of geospatial datasets [53]. However, previous studies have seldom paid attention to the frequent occurrence of spatially autocorrelated residuals in regression models, which indicate a model misspecification problem and unreliable results [54]. To judge whether the global regression result was valid, the residual value of global regression was analyzed by the Moran index. If $P \leq 0.01$ and $Z \geq 2.58$, the data are positively aggregated.

3.5. Validation Method

The comprehensive performance of the proposed model in LSA was appraised. Several validation methods were employed for the model fitting degree and accuracy of model prediction, such as the Akaike Information Criterion (*AICc*). Area Under the ROC Curve (AUC) and the ROC curve were used to assess the accuracy of models [55]. The ROC curve took the false-positive rate specificity as the horizontal coordinate (the proportion of neighborhoods without landslide hazards that were correctly predicted) and the true-positive rate sensitivity as the vertical coordinate (the proportion of neighborhoods with landslide hazards that were correctly predicted). The AUC is the area under the ROC curve, which can more intuitively express the prediction accuracy of models. Its value range is (0.5, 1). The closer the value is to 1, the higher the model accuracy will be. It is generally considered that the prediction accuracy interval (0.5, 0.7) is relatively reasonable, (0.7, 0.8) is reasonable, and (0.8, 1) is very reasonable. If the AUC is extremely close to 1, the model reflects higher goodness-of-fit and consummate accuracy [56].

4. PCAMGWR Modeling

4.1. Principal Component Analysis (PCA)

PCA is an effective method for dimensionality reduction and the feature combination analysis of multivariate factors, which describes multivariate samples to identify spatial patterns [31]. PCA employs an orthogonal transformation to transform correlated variables into linearly unrelated variables, namely independent Principal Components (PCs) [57–59]. The first PC had the most significant contribution and contained as much information about the *LCFs* as possible, followed by the second PC and the third PC. The steps for eliminating factor correlation were as follows: (1) standardized processing of the original *LCFs*; (2) calculating the correlation coefficient matrix of the standardized matrix; (3) calculating the eigenvalues and eigenvectors of the coefficient matrix to determine the PCs; (4) computing the variance contribution rate and determining the number of PCs; (5) comprehensively evaluating the PCs.

4.2. MGWR

GWR is a linear regression but differs from traditional linear regressions in that GWR considers the influence of spatial relations on the model, namely spatial heterogeneity. GWR is one of the methods used in the exploration of spatial non-stationarity [60], and the regression coefficients change with the spatial location. However, there are limitations to this method, namely that it is useless for spatial multicollinearity [61] and neglects changes in the spatial scope of geographical units [62–64]. The MGWR model can solve these issues [65]. Consequently, the PCAMGWR model considering the spatial non-stationarity scale effect based on hexagonal neighborhoods was employed to assess landslide susceptibility.

Assuming that there are *n* observations, for the observation $i \in \{1, 2, ..., n\}$ at location (u_i, v_i) with *m* independent variables, where $j \in \{1, 2, ..., m\}$ at the *j*-th independent variable, the GWR model formulation is described as follows:

$$y_i = \sum_{j=0}^{m} \beta_j(u_i, v_i) x_{ij} + \varepsilon_i \tag{4}$$

where x_{ij} is the *j*-th independent variable; $\beta_j(u_i, v_i)$ is the *j*-th coefficient; ε_i is the error term; and y_i is the *i*-th dependent variable.

The MGWR model is mathematically expressed as follows [34,66]:

$$y_i = \sum_{j=0}^{m} \beta_{bw_j}(u_i, v_i) x_{ij} + \varepsilon_i$$
(5)

where bw_j in β_{bw_j} indicates the bandwidth used for calibration of the *j*-th spatial relationship. The selection of the bandwidth was relatively ordinary due to a single bandwidth required. By the trial selection of an initial bandwidth, the *AICc* was optimized to select the bandwidth, which is defined as:

$$AICc = 2nln\left(\frac{RSS}{n}\right) + nln2\pi + n\left\{\frac{n + tr(S)}{n - 2 - tr(S)}\right\}$$
(6)

where *RSS* is the sum of the error terms' square residual, and tr(S) is the trace of the hat matrix *S* and the Effective Number of Parameters (ENP) of the model. The bandwidth with the smallest *AICc* value was optimal. The Gaussian kernel data-borrowing scheme parameterized via *AICc* optimization was utilized throughout. However, the *AICc* cannot select bandwidths in MGWR on account of the great number of potential combinations of bandwidths, which may generate a different procedure required.

Model calibration for a Gaussian MGWR can be conducted by resorting to weighted least squares. The coefficient at the location (u_i, v_i) is estimated in Equation (7), where X is the design matrix and $W_{(u_i,v_i)}$ is the spatial weighting matrix for location (u_i, v_i) . $W_{(u_i,v_i)}$ is homogeneous for each relationship, with the identical bandwidth being adopted for all the relationships in the model.

$$\beta_{bw_j}(u_i, v_i) = \left[\left(X^T W_{ij} X \right) \right]^{-1} X^T W_{ij} y$$
(7)

The fixed Gaussian kernel function was used to calculate the spatial weights in MGWR on a par with GWR, which can be written as [48]:

$$W_{ij} = exp\left(\frac{-d_{ij}^2}{bw_j^2}\right) \tag{8}$$

where W_{ij} is the weight value of observation *j* for estimating the coefficient of observation *i*; d_{ij} is the straight-line distance between observations *i* and *j*; and bw_j is a constant bandwidth.

The back-fitting algorithm was adopted to calibrate the MGWR model, which maximizes the expected log-likelihood and is generally used to calibrate generalized additive models (GAMs) [67]. The logic of GAM, $\beta_{bw_j}x_j$ in MGWR is defined as the *j*-th additive term f_j resulting in the GAM-style MGWR:

$$y = \sum_{j=0}^{m} f_j + \varepsilon$$
(9)

The back-fitting algorithm was a smoother method for calibrating the model, and the specific process was as follows [65]. Firstly, all additive terms were initialized, the dependent variables set, and the errors calculated. These errors, plus the "current" value of the first term f_0 , were then regressed on x_0 using GWR, which produced an optimal bandwidth bw_0 for the relationship between y and x_0 , as well as a new set of local estimates for the relationship between y and x_0 that was used to update the value of the first term f_0 . The second variable x_1 followed the same procedure as x_0 above, and the process was repeated until the first iteration was completed. The iteration continued until the change of all terms in successive iterations was less than the score of change (SOC).

Two decisions from the user are involved in the algorithm. The first concern is initializing the local coefficient estimates, which might affect the number of iterations needed to reach convergence instead of the selection of the optimal bandwidth. The GWR was thus used to estimate the initial MGWR. The second decision is the choice of the termination criterion—the value of the differential between successive iterations, namely the SOC by which the process is converged. *SOC-f* focuses on the relative changes of the

additive terms rather than on overall model fitting. The calculation formula of *SOC-f* is as follows:

$$SOC_{f} = \sqrt{\frac{\sum_{j=1}^{p} \frac{\sum_{i=1}^{n} \left(\hat{f}_{ij}^{new} - \hat{f}_{ij}^{old}\right)^{2}}{n}}{\sum_{i=1}^{n} \left(\sum_{i=1}^{p} \hat{f}_{ij}^{new}\right)^{2}}}$$
(10)

5. Results 5.1. Expression of Spatial Proximity

The spatial proximity of the study region was expressed in the GIS environment, which incorporated Moore neighborhoods (Figure 4a), slope-unit-based neighborhoods (Figure 4b), and hexagonal neighborhoods (Figure 4c).



Figure 4. Schematic diagram of spatial proximity in Qingchuan county: (**a**) Moore neighborhoods, (**b**) slope-unit-based neighborhoods, (**c**) hexagonal neighborhoods.

The present research primarily explored the homogeneity within the neighborhood and the heterogeneity between the neighborhoods, as shown in Figure 5. The hexagonal neighborhoods are shown in orange, the Moore neighborhoods in green, and the slope-unitbased neighborhoods in purple. The solid lattice of diamonds on the left represents the distribution of the F values in each type of neighborhood. On the right is a box plot to figure out the distribution characteristics of the F value. The black horizontal line is the middle line of the F value, and the box represents the range of 1/4 to 3/4. The F value ranges of the hexagonal neighborhoods, Moore neighborhoods, and slope-unit-based neighborhoods were, respectively, (0.47, 1.92), (0, 5), and (0.1, 2). The F value distribution in the hexagonal neighborhoods was relatively concentrated and the spectrum was small. The medians were ordered as follows: slope-unit-based neighborhoods, hexagonal neighborhoods, Moore neighborhoods. From the distribution of the F value, slope units and hexagons performed better. To further select the most suitable expression of spatial proximity, the RMSE values of the slope elements and hexagons were calculated, and the hexagonal neighborhoods expressed the spatial proximity of the study area with a greater RMSE.



Figure 5. Boxplot of F value distribution of spatial proximity expressions.

5.2. Correlation Analysis and Multicollinearity Test

Pearson correlation analysis of the *LCFs* was conducted as a preliminary analysis, as shown in Figure 6. The darker the red is, the stronger the positive correlation; the darker the blue is, the stronger the negative correlation. The maximum correlation coefficient was 0.72 for the distance to settlement and elevation, revealing that the higher the elevation is, the fewer the settlements. The correlation coefficients between the distance to the fault zones and distance to settlement, the terrain relief and slope, the elevation and distance to the settlement, and the elevation and distance to fault zones were between 0.5 and 0.8, showing a moderate positive correlation. The other factors were weakly correlated or irrelevant. The symbol \times in Figure 6 indicates no statistically significant correlation between factors. There were no highly correlated factors among the *LCFs* selected in this study.



Figure 6. Elliptical diagram of Pearson correlation of *LCFs*. L—lithology; DSt—distance to stream; DSe—distance to settlement; DFZ—distance to fault zones; A—aspect; S—slope; TR—terrain relief; VCT—vegetation cover type; Pre—precipitation; E—elevation.

Regression analysis is prone to producing high multicollinearity between independent variables. To test the global multicollinearity, the *VIFs* of various *LCFs* were calculated, as shown in Table 1. A *VIF* value greater than 10 denotes the existence of global multicollinearity. In this study, the maximum *VIF* value was 2.861, lower than 5, indicating that all *LCFs* passed the global multicollinearity test.

Table 1. Multicollinearity test of *LCFs*.

LCFs	VIF	TOL
Lithology	1.328	0.753
Distance to stream	1.109	0.901
Distance to settlement	2.255	0.443
Distance to fault zones	2.195	0.456
Aspect	1.025	0.975
Slope	1.540	0.650
Terrain relief	1.719	0.582
Vegetation cover type	1.019	0.981
Precipitation	1.052	0.950
Elevation	2.861	0.350

5.3. ESDA

ESDA is an essential method that was used to examine the spatial associations among the *LCFs* and aid with the model development. A classic spatial autocorrelation index, namely the Moran's I, was computed using the residual of global regression and was implemented to explore the spatial non-stationarity effects [68].

The landslide susceptibility prediction based on OLS was carried out, and the residual distribution is shown in Figure 7a. The red area is highly consistent with the location of the landslides. The Moran's index analysis based on OLS residual values judges the effectiveness of global regression results. If $P \leq 0.01$ and $Z \geq 2.58$, it is a positive aggregation. The Moran's index analysis gave a result of $P \leq 0.0001$ and the *Z* value was 18.292, indicating that the global regression was invalid. The residual was spatially autocorrelated to a level of statistical significance, which was brought about by a non-stationary spatial process (Figure 7b). Therefore, a model with an interpretation of spatial non-stationarity relationships is necessary for LSA.



Figure 7. Diagram of ESDA: (**a**) residual distribution diagram based on OLS, (**b**) Moran's index analysis of residual values.

5.4. LSMs Based on PCAMGWR Model

Six principal components were obtained by principal component analysis and are shown in Table 2. The overall contribution rate was 82.704% to represent the entirety elementarily. PC1 was interpreted as a comprehensive factor, since it has a high favorable loading of elevation, distance to fault zones, distance to settlement, and terrain relief. PC2 has a high favorable loading of the slope. PC3 has a high positive loading of precipitation. PC4 has high favorable loading of vegetation cover type. PC5 has a high negative loading of aspect. PC5 has a high negative loading of distance to stream.

Table 2. Component matrix of LCFs.

LCFs	PC1	PC2	PC3	PC4	PC5	PC6
Lithology	0.500	-0.448	-0.160	0.109	-0.211	0.176
Distance to stream	0.272	0.322	0.542	0.156	0.071	-0.621
Distance to settlement	0.782	-0.267	0.120	-0.018	-0.012	-0.065
Distance to fault zones	0.803	-0.291	0.097	0.003	-0.084	0.163
Aspect	0.058	0.396	0.031	0.509	-0.740	0.088
Slope	0.483	0.630	-0.337	0.027	0.182	0.171
Terrain relief	0.604	0.528	-0.286	-0.042	0.227	0.075
Vegetation cover type	-0.034	-0.277	-0.190	0.819	0.450	-0.058
Precipitation	-0.012	0.164	0.740	0.109	0.232	0.592
Elevation	0.865	-0.086	0.112	-0.084	0.000	-0.147

Normal distribution and axle-whiskers were employed to analyze the differences between the PCs and *LCFs*, respectively, as shown in Figure 8. Although the original *LCFs* obeyed normal distribution, there was a significant difference between the factors, and the data were more discrete. The distribution of the principal components was more uniform and closer to the standard normal distribution using the principal component analysis.



Figure 8. Normal distribution and axial-whisker plots: (a) LCFs; (b) PCs.

The LSMs of the study area were independently obtained using the PCAMGWR, MGWR, PCAGWR, and GWR models, as show in Figure 9. The LSMs were divided into four levels using the quantile method, namely low, moderate, high, and very high susceptibility. Figure 10 shows the area percentage of very high, high, moderate, and low susceptibility.



Figure 9. LSMs generated by models: (a) PCAMGWR, (b) MGWR, (c) PCAGWR, and (d) GWR.



Figure 10. Area percentage of various landslide susceptibility levels produced in LSMs.

The PCAMGWR and MGWR models exhibited more veracious and reasonable predictions. The high and very high susceptibility areas were more consistent with investigated landslide locations due to the consideration of the scale difference effect of spatial nonstationarity. Apart from the significant variances in the land area composition with the various landslide susceptibility levels, the landslide susceptibility zoning results of the two models considering the spatial non-stationarity scale differences were better combined with topographic factors. The PCAMGWR and MGWR models comprehensively considered topography, streams, and settlements. Especially in the western mountainous area, the landslide susceptibility was low in the area with very high altitude but far from the river and with rare human activities, while the landslide susceptibility was medium in the area with high altitude and close to the river and human activities. It can be seen that the PCAMGWR model proposed in this study can reflect the non-stationarity scale difference in the spatial relationship between *LCFs* and landslides. Moreover, the LSMs based on the PCAMGWR model and MGWR model were similar, indicating that MGWR can deal with factor correlation and multicollinearity. There was a large gap between the LSMs based on the MGWR-based models and GWR-based models, which can be seen considering the spatial non-stationarity scale difference, making the zoning results tend towards the actual situation. In addition, there were many similarities between the four models. There were southwest-northeast spatial distributions along fault zones in the LSMs obtained using the four models, which may be attributed to the fault zones' influences on slope and rock mass stability.

5.5. Analysis and Comparison of Spatial Non-Stationarity Scale Effect

Compared with the GWR model, the MGWR model showed an improvement in the research of spatial non-stationarity. MGWR allowed the parameter estimates to vary spatially and generated a single optimal scale (bandwidth) for the non-stationary spatial relationship between landslides and each independent variable. The spatial variation of different processes was modeled at different spatial scales. The optimal bandwidths deduced by the GWR, PCAGWR, MGWR, and PCAMGWR models were direct indicators of spatial scale, indicating the individual spatial relationship between landslide and independent variables. Figure 11 indicates the bandwidth search process for each independent variable generated by the MGWR model (Figure 11a) and PCAMGWR model (Figure 11b), a process that was observed to operate at different spatial scales.

A variable with a large bandwidth affects the dependent variable at a large scale, so the standard deviation of the parameter estimates is slight. In contrast, a variable with a small bandwidth affects the dependent variable at a local scale, so the standard deviation of the local parameter estimates is significant. The optimal bandwidth of terrain relief and interpolation was 9018.89 and 784.84 in the MGWR model, with the number of iterations being 59. Additionally, the variable affected landslide susceptibility at the local scale, and its parameter estimate had significant variances over space. The relationships between the other nine *LCFs* and landslides exhibited spatial non-stationarity, but the processes varied at regional spatial scales. In the PCAMGWR model, PC2 and interpolation demonstrated solid locally spatial non-stationarity scale effects, and the optimal bandwidth was 8962.5 and 784.32. The number of iterations of the PCAMGWR model was 40. The spatial non-stationarity scale effects between the other principal components and landslides were at the regional scale. Bandwidth selection for different parameters may stop at different steps depending on the properties of the spatial non-stationarity scale effect.

However, the optimal bandwidths generated using the GWR model and PCAGWR model were 4127.02 and 2868.55, which implies that all variables affected landslides with the same spatial non-stationarity scale effect of extreme restriction. The bandwidth produced by GWR or PCAGWR was the weighted average of the independent spatial processes of each factor and landslides, with varying degrees of spatial non-stationarity, as shown in Figure 12. The convergence rate of the two models was similar under the same scale effect,



and there was no continuous decline in *AICc* values. The weighting is a function of the explanatory ability of each relationship in the local model.

Figure 11. Iterative process of bandwidths based on the back-fitting algorithm: (**a**) *LCFs* for MGWR, (**b**) PCs for PCAMGWR.



Figure 12. Bandwidth iteration process diagram based on a single spatial non-stationarity scale.

5.6. Validation and Accuracy Assessment

The whole dataset was input into the ROC calculation tool of origin software to obtain the ROC curve and AUC value. The accuracy of the model prediction was validated by the AUC and other statistical indicators, shown in Table 3. The bandwidth selection criterion of the basic GWR model was the minimum AICc value, and the bandwidth selection criterion of the MGWR and PCAMGWR models was SOC-f dissimilarity, but the AICc value could still be enumerated. Therefore, The AICc value served as a valid indicator of the model prediction. The PCAMGWR model had the minimum AICc value of 78,291.042 and indicated the maximum accuracy of bandwidth selection and model prediction, with the AIC value also being the minimum, while the BIC value was inferior to GWR. The ROC curves of various models are drawn in Figure 13, and the AUCs of Global Linear Regression (GLR), Logistic Regression (LR), GWR, PCAGWR, MGWR, and PCAMGWR were, respectively, 0.69263, 0.7458, 0.82128, 0.83707, 0.90352, and 0.90355. The MGWR model considering multiscale bandwidth showed a significant improvement in the accuracy of prediction results compared with the GWR model. Consequently, spatial non-stationarity scale variances were subsistent between *LCFs* in the study area, and it is significant to identify spatial non-stationarity scale effects in LSA.

Table 3. Accuracy statistics of models.

_						_
	Model	AIC	AICc	BIC	AUC	
	PCAMGWR	78,228.039	78,291.042	85,829.127	0.89773	
	MGWR	78,232.004	78,295.213	85,845.297	0.89771	
	PCAGWR	78,682.364	78,696.459	82,307.355	0.83198	
	GWR	78,785,304	78,794,218	81.672.072	0.81701	



Figure 13. Area under the ROC for different models. Figure 14a shows the convergence of *SOC-f* during the fitting of the back-fitting algorithm for the MGWR model and PCAMGWR model. The speedy convergence rate means that bandwidth was not chosen at each iteration step, and the optimization stopped at convergence inversely. It can be seen from Figure 14b that the optimal bandwidth was selected based on *AICc* at a slow convergence, and the *AICc* value did not continue to decline. It is hard to differentiate the *SOC-f* of PCAMGWR and MGWR models in detail, and the PCAGWR model represented by the black dot plot was better than the GWR model regarding the convergence of *AICc* values.



Figure 14. Parameter variation diagram of bandwidth judgment criterion: (**a**) *SOC-f* for MGWR and PCAMGWR, (**b**) *AICc* for GWR and PCAGWR.

Meanwhile, the assessment results of the PCAGWR compared with GWR provided a significant priority ranking. The assessment results of the PCAGWR model were more accurate than those of GWR regarding the elimination of local collinearity. The AUC of PCAMGWR increased less than that of MGWR, and the PCA had a minor role in collinearity elimination for the PCAMGWR model, which is reflective of the fact that the MGWR model could eliminate most of the local collinearity of *LCFs* to improve the prediction accuracy of landslide susceptibility.

6. Discussion

Landslides are a major geohazard, causing casualties and economic losses. Therefore, it is necessary to establish a suitable model to assess landslide susceptibility and make a landslide zoning map. The assessment methods and models of landslide susceptibility have been discussed in many studies. For example, a comparative study of WOE, AHP, ANN, and GLR procedures for landslide susceptibility zonation is presented in [69]. WOE can assess the impact of different classes of each *LCF*, but neglects the correlation between *LCFs*. The logistic regression (LR) method is a static susceptibility model that has limited application for predicting future landslide probability under potential rainfall events [70]. Moreover, LR is capable of analyzing the relationship among the *LCFs*, but it is not able to evaluate the impact of different classes [71]. A support vector machine (SVM) is a machine learning algorithm that uses a small number of samples, but a high-quality informative database is essential to improve model performance [72].

Compared with the assessment method of LSA, the study of the essential attribute, namely the spatial non-stationarity, of the landslide as a geospatial phenomenon is insufficient. At present, a small number of scholars have carried out studies on the consideration of spatial non-stationarity and the application of the GWR idea in LSA, and have proved that the GWR model is superior to some traditional models, such as global linear regression (GLR) [21,33], ANN and OLS [73], SVM [74], and SR [30]. However, there is a research gap in the study of the non-stationarity scale of the spatial relationship between landslides and *LCFs*.

This study mainly discusses the influence of spatial non-stationarity on LSA results and the difference of spatial non-stationarity scale among different factor combinations. In addition, spatiotemporal non-stationarity is rarely considered in the study of geospatial data, especially in the field of geological hazard assessment. In this study, only spatial non-stationarity was considered. Therefore, time may be introduced into non-stationarity in geospatial data studies. The analysis of the impact of scale variance on model prediction accuracy is a main area of focus. The impact of such variance is more directly reflected in the spatial differentiation of regression coefficients, which is a shortcoming in the present study. In addition, the studies regarding GWR and MGWR models take full datasets as the model input without considering the division of training set and testing set [21,75,76]. The indicators which are generated during the model run, such as AICc, are generally used for model performance testing, and the present study also verified the assessment accuracy using these indicators. However, most LSA studies divide the training set and testing sets, and future studies can thus focus on the comparison between analysis with and without division methods. On the other hand, the present study analyzed the spatial non-stationarity scale of landslide susceptibility using GAM-style MGWR. Thus, future studies can explore the interaction effects and nonlinear relationships. Moreover, the LSMs were related to zoning methods, and the quantile method was selected for the zoning of the landslide susceptibility map, which may have reduced the accuracy. All data used for model input were normalized, and the influence of the normalization process on a non-stationarity scale is uncertain, which affects the explanatory power of the model. The degree of elimination of local collinearity by the MGWR model may have been related to the composition of LCFs, which needs to be further studied. These limitations may further give rise to uncertainties in LSA.

7. Conclusions

Geospatial data may lead to the spatial non-stationarity process, and the scale at which each independent variable affects a dependent variable may vary according to the independent variables. The present study thus considered spatial non-stationarity and scale variations in LSA using a PCAMGWR model. The results indicated that the PCAMGWR model provides more reliable information for LSA than other GWR models and achieves a higher accuracy in LSM by performing better at alleviating residual autocorrelation.

The present study determined the respective bandwidths for each independent variable and revealed the association between the independent variables and landslide susceptibility using the PCAMGWR model. The model relaxes the single-bandwidth assumption of the basic GWR model and allows independent variable-specific bandwidths to be optimized. The results demonstrated that there are scale variations in LSA. For example, PC2 affected the landslide susceptibility at a local scale, namely the local parameters associated with the variable varied across space. The basic GWR model was outperformed at differentiating such scale variations and can be substituted by PCAMGWR.

Moreover, according to the AUCs, compared to GLR and LR, the PCAMGWR and MGWR models can better analyze the impacts of non-stationarity scale variation and factor correlation on LSM than GWR and PCAGWR. The four models with and without the elimination of factor correlations were compared, and it was indicated that the PCAGWR and PCAMGWR models benefited from the elimination of factor correlations by PCA, and the PCAMGWR model was preferred to PCAGWR because the scale variation of spatial non-stationarity had a greater impact than factor correlation. Meanwhile, the MGWR and PCAMGWR models benefited from the consideration of the scale variation, and PCAMGWR was preferred to MGWR. Spatial statistical models are useful for analyzing the determinants of landslide susceptibility by considering spatial dependency and spatial heterogeneity. The present study reveals the superiority of a new approach, namely the PCAMGWR model, to consider the spatial characteristics, non-stationarity scale variations, and factor correlations.

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