

Review

The Effects of Community Home Visit and Peer Group Nutrition Intervention Delivery Platforms on Nutrition Outcomes in Low and Middle-Income Countries: A Systematic Review and Meta-Analysis

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Abstract: Weak delivery systems reduce the potential of evidence-supp orted interventions to improve nutrition. We synthesized the evidence for the effectiveness of nutrition-specific intervention delivery platforms for improving nutrition outcomes in low and middle-income countries (LMIC). A systematic literature search for studies published from 1997 to June 2018 resulted in the inclusion of 83 randomized controlled trials (RCTs), quasi-randomized, and controlled before-after studies across a variety of delivery platforms. In this paper, we report on meta-analysed outcomes for community health worker (CHW) home visits and mother/peer group delivery platforms. Compared to care as usual, CHW home visits increased early initiation of breastfeeding (EIBF) (OR: 1.50; 95% CI: 1.12, 1.99; n = 10 RCTs) and exclusive breastfeeding (EBF) (OR: 4.42; 95% CI: 2.28, 8.56; n = 9 RCTs) and mother/peer groups were effective for improving children's minimum dietary diversity (OR: 2.34; 95% CI: 1.17, 4.70; n = 4) and minimum meal frequency (OR: 2.31; 95% CI: 1.61, 3.31; n = 3). Pooled estimates from studies using both home visit and group platforms showed positive results for EIBF (OR: 2.13; 95% CI: 1.12, 4.05; n = 9), EBF (OR: 2.43; 95% CI: 1.70, 3.46; *n* = 12), and < 5 wasting (OR: 0.77; 95% CI: 0.67, 0.89; *n* = 4). Our findings underscore the importance of interpersonal community platforms for improving infant and young child feeding practices and children's nutritional status in LMICs.

Keywords: nutrition; intervention; delivery; systematic review; low and middle-income country

1. Introduction

Poor nutrition in the early years of life poses substantial, and potentially lifelong, costs to health, development, and socioeconomic well-being [1,2]. Preventing the adverse effects of undernutrition requires essential nutrition actions during the critical first 1000 days of life [3]. Despite evidence-supported interventions to improve maternal and child nutrition [4], their potential gains are often thwarted by weak delivery systems preventing maximum population reach and impact [5,6].

In most countries, a variety of intervention delivery platforms, or channels, are used to reduce access barriers to essential services. Notable examples are trained community health workers (CHWs) and mother/peer groups for health and nutrition promotion [7–9]. Though evidence supports CHW



home visits and women's groups for improving maternal and newborn health in low and middleincome countries (LMICs) [10–12], evidence from high-quality studies involving such community platforms for delivering nutrition-specific interventions is limited across diverse populations [5]. This is partly due to limited knowledge surrounding how best to integrate nutrition-focused interventions into these routine activities at the country level [13].

Emerging evidence indicates that advancements in nutrition require nutrition-sensitive programmes that enhance the coverage and effectiveness of direct nutrition interventions [14]. Cash transfers requiring compliance with specific health/nutrition behaviours are commonly used in social protection/safety net programmes in LMICs [15,16]. Though evidence suggests they are effective for poverty reduction [14,17,18], there is limited high-quality evidence on the effectiveness of conditional cash transfers (CCTs) for improving nutrition-specific outcomes among women and children [19]. Mobile health (mHealth) technologies (e.g., SMS messaging) are increasingly being used in global nutrition programming, yielding potentially cross-cutting benefits through linkages with community and financial platforms [20]. However, knowledge gaps remain surrounding the effectiveness of mHealth platforms for delivering nutrition-specific interventions [21,22].

We reviewed the available evidence to assess the effects of community-, financial- and technology-based nutrition intervention delivery platforms for improving nutrition outcomes in LMICs. In this paper, we report on meta-analysed outcomes for community health worker (CHW)/lay counselor home visits and mother/peer groups, which emerged as the most widely used community-based platforms for delivering nutrition-specific interventions in our review. While the multiplicity of health delivery platforms across LMICs is acknowledged, our review focused only on those delivering a nutrition-specific intervention or that integrated a direct nutrition component.

2. Materials and Methods

2.1. Types of Studies

We conducted the systematic review and meta-analysis following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines [23] and according to the study protocol [24]. We searched the Cochrane Central Register of Controlled Trials, Embase, MEDLINE, Scopus, and Web of Science for English-language studies published from 1997 to June 2018 using Medical Subject Heading (MeSH) terms and keywords (MEDLINE search strategy included in Table S1). In addition, ClinicalTrials.gov and the WHO International Clinical Trials Registry Platform were searched for ongoing trials and we examined the grey literature for relevant studies. Studies were included if they: (i) occurred in a LMIC, as per the World Bank [25] definition at the time of publication; (ii) utilized a community, financial and/or technology platform for delivering a nutrition-specific intervention; (iii) utilized a randomized, cluster-randomized, quasi-randomized, stepped-wedge, controlled before-after, or interrupted time series design with a control group; and (iv) examined a relevant outcome. Excluded studies did not include a control group.

2.2. Outcomes

The primary study outcomes were: early initiation of breastfeeding (EIBF) (breastfeeding within one hour of birth), exclusive breastfeeding (EBF) (up to 6 months), minimum dietary diversity (MDD) and minimum meal frequency (MMF) (based on WHO [26] guidelines), compliance with micronutrient supplementation, anthropometric measures (height/length, weight, body mass index, stunting, underweight, wasting), hemoglobin/anemia, and micronutrient status. No secondary outcomes were assessed. Duration of exposure and timing of outcome measurement were not restrictions for exclusion.

2.3. Search Methods and Data Extraction

Two reviewers screened titles and abstracts to identify potentially eligible studies using Covidence (Veritas Health Innovation, Melbourne, Australia). Full texts of selected articles were reviewed independently by two reviewers for eligibility based on predefined inclusion criteria, with

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discrepancies resolved by discussion. Double data extraction using standard forms was conducted for studies that met the inclusion criteria. Extracted data consisted of: (i) author, publication year, country, area; (ii) study design; (iii) study population characteristics; (iv) type of delivery platform; (v) description and duration of intervention(s); (vi) control group description; (vii) outcome measurement(s); and (viii) effect measure(s). Reference lists from included studies and relevant systematic reviews were also examined.

2.4. Statistical Analysis

Pooled estimates of outcome measures were calculated using random-effects models (inversevariance method) in RevMan 5.3 (Cochrane Collaboration). Meta-analyses were conducted based on outcome and type of delivery platform and, when possible, were performed separately for randomized and non-randomized studies. Pooled estimates are presented as odds ratios (OR) with 95% confidence intervals (CI). Where necessary, data conversions were made using available study data. For study outcomes measured at multiple time points (e.g., anthropometry), we selected similar measurements across studies based on what we considered to be the most clinically relevant followup period for the particular outcome. Statistical heterogeneity for included studies was assessed by visually examining forest plots and quantitatively using Tau², I², and significance of the chi-squared test. Sensitivity analyses were performed based on removing studies with outlying effect sizes.

2.5. Quality Assessment

Two reviewers independently assessed the risk of bias for all included studies using the Cochrane risk-of-bias tool [27] and the Cochrane Effective Practice and Organisation of Care (EPOC) guidelines [28] for randomized and non-randomized studies, respectively. As the blinding of study participants and intervention providers is relatively uncommon in behavioural intervention trials, we concluded that non-blinding posed a low risk of bias for outcomes we considered to be minimally affected by a lack of blinding. We used the GRADE method [29] to evaluate the overall quality of evidence and produce a summary of findings table for each meta-analysed outcome.

3. Results

3.1. Search Results

We identified a total of 20,661 records in the search, of which 5277 were duplicates, and 15,164 were excluded at the title/abstract screening stage. We reviewed 219 full-text articles for eligibility and a total of 83 studies [30–85] were included in the full review (Figure A1). Table S2 summarizes the key characteristics of the CHW/lay counselor home visit and mother/peer group studies, which are the focus of this paper. These studies utilized one or both of these platforms to deliver nutrition education/behaviour change communication to mothers/caregivers of young children.

3.2. CHW/Peer Counselor Home Visit Platform

For the CHW/peer counselor home visit platform, we conducted meta-analyses for four outcomes: EIBF, EBF, <5 stunting and < 5 underweight (Table 1). The pooled estimate for EIBF based on 10 RCTs showed a positive effect (OR: 1.50; 95% CI: 1.12, 1.99) for home visits, compared to usual care (Figure 1). A smaller, though positive, effect was observed when excluding the Haider et al. study from the analysis (OR: 1.20; 95% CI: 1.02, 1.41). A pooled analysis of non-randomized studies (n = 4) also revealed a higher likelihood of EIBF among mothers receiving home visits, as compared to those receiving usual care (OR: 1.80; 95% CI: 1.17, 2.78).

| Study or Subgroup | log[Odds Ratio] | SE | Intervention Total | | Weight | Odds Ratio IV, Random, 95% CI | Odds Ratio IV, Random, 95% CI |
|-----------------------------------|---------------------|------------|-----------------------|-------|--------|----------------------------------|---|
| Ara 2018 | 0.8668 | 0.3119 | 165 | 164 | 7.8% | 2.38 [1.29, 4.38] | |
| Ayiasi 2016 | 0.3293 | 0.1102 | 601 | 741 | 11.4% | 1.39 [1.12, 1.73] | |
| Gabida 2015 | -0.0953 | 0.2152 | 182 | 175 | 9.6% | 0.91 [0.60, 1.39] | |
| Haider 2000 | 2.3019 | 0.1924 | 323 | 330 | 10.1% | 9.99 [6.85, 14.57] | |
| Hanson 2015 | 0.3402 | 0.0364 | 6562 | 6346 | 12.1% | 1.41 [1.31, 1.51] | 3 - |
| ljumba 2015 | 0.077 | 0.0763 | 1629 | 1865 | 11.8% | 1.08 [0.93, 1.25] | + |
| Lewycka 2013 | -0.2706 | 0.1234 | 677 | 715 | 11.2% | 0.76 [0.60, 0.97] | |
| Nayak 2016 | -0.087 | 0.5111 | 35 | 34 | 4.9% | 0.92 [0.34, 2.50] | |
| Nguyen 2017 | 0.2422 | 0.09 | 1000 | 1000 | 11.6% | 1.27 [1.07, 1.52] | |
| Penfold 2014 | 0.3365 | 0.2254 | 249 | 251 | 9.4% | 1.40 [0.90, 2.18] | |
| Total (95% CI) | | | 11423 | 11621 | 100.0% | 1.50 [1.12, 1.99] | • |
| Heterogeneity: Tau ² = | 0.17; Chi# = 147.16 | . df = 9 (| P < 0.00001); | = 94% | | | |
| Test for overall effect: | Z = 2.76 (P = 0.006 |) | | | | | 0.1 0.2 0.5 1 2 5 10 Favours [control] Favours [intervention |

Figure 1. Forest plot of the meta-analysis of early initiation of breastfeeding for home visits (randomized studies).

An analysis of nine RCTs examining the association between the home visit platform and EBF produced an overall effect of 4.42 (95% CI: 2.28, 8.56) (Figure 2). When removing the Haider et al. study, the odds of EBF remained significantly higher among mothers receiving home visits, compared to usual care (OR: 3.33; 95% CI: 1.94, 5.72). A pooled analysis of three non-randomized studies also revealed a positive effect (OR: 4.57; 95% CI: 1.36, 15.36) on EBF for mothers receiving home visits, compared to those receiving usual care.

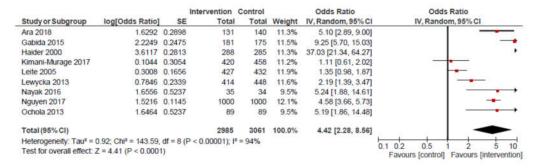


Figure 2. Forest plot of the meta-analysis of exclusive breastfeeding for home visits (randomized studies).

Meta-analyses revealed no significant impact of the home visit platform, compared to usual care, on < 5 stunting (OR: 2.12; 95% CI: 0.46, 9.74; n = 3) and a nearly significant 7% decrease in the likelihood of < 5 underweight (OR: 0.93; 95% CI: 0.82, 1.06; n = 4).

| | Anticipated Al | bsolute Effects | Relative | Nº of | Certainty of the Evidence (GRADE) | |
|---|----------------------------|--------------------------|------------------------------|---------------------------|---|--|
| Outcomes | Risk With Care as Usual | Risk With Home Visits | Effect (95% CI) | Participants (Studies) | | |
| Early initiation of breastfeeding RCTs ^a | 390 per 1000 | 490 per 1000 | OR 1.50 (1.12 to 1.99) | 23,044 (10) | ⊕⊕⊖⊖ LOW | |
| Early initiation of breastfeeding non RCTs ^a | 464 per 1000 | 609 per 1000 | OR 1.80 (1.17 to 2.78) | 9816 (4) | ⊕⊕⊖⊖ LOW | |
| Exclusive breastfeeding RCTs ^b | 350 per 1000 | 704 per 1000 | OR 4.42 (2.28 to 8.56) | 6046 (9) | ⊕⊕⊖⊖ LOW | |

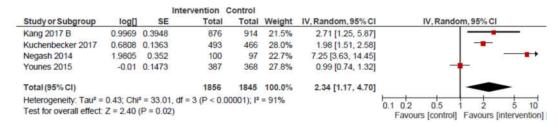
 Table 1. Summary of outcomes for community healthcare worker (CHW)/peer counselor home visit platform.

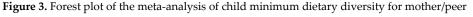
| | Anticipated Al | osolute Effects | Relative | Nº of | Certainty of the |
|--|--|-----------------|-------------------------------|---------------------------|---------------------|
| Outcomes | Risk With Risk With Care as Usual Home Visi | | Effect (95% CI) | Participants (Studies) | Evidence (GRADE) |
| Exclusive breastfeeding non RCTs ^b | 421 per 1000 | 768 per 1000 | OR 4.57 (1.36 to 15.36) | 2532 (3) | ⊕⊕⊖⊖ LOW |
| Stunting ^c | 431 per 1000 | 616 per 1000 | OR 2.12 (0.46 to 9.74) | 4436 (3) | ⊕○○○ VERY LOW |
| Underweight ^d | 379 per 1000 | 362 per 1000 | OR 0.93 (0.82 to 1.06) | 4515 (4) | ⊕⊕⊖⊖ Low |

^a Child is breastfed within one hour of birth. ^b Child is provided only breastmilk up to 4–6 months of age ^c Height-for-age is more than two standard deviations below the WHO Child Growth Standards median. ^d Weight-for-age is more than two standard deviations below the WHO Child Growth Standards median.

3.3. Mother/Peer Group Platform

For the mother/peer group platform, we conducted meta-analyses for seven outcomes: EIBF, EBF, MDD, MMF, and < 5 stunting, underweight and wasting (Table 2). Pooled estimates were nonsignificant for EIBF (OR: 1.10; 95% CI: 0.92, 1.32; n = 5), and EBF (OR: 1.78; 95% CI: 0.93, 3.41; n = 4), as compared to usual care. A pooled analysis of four studies revealed a higher likelihood of MDD (OR: 2.34; 95% CI: 1.17, 4.70) (Figure 3). However, removing the Negash et al. study resulted in a lower, and non-significant, effect on MDD (OR: 1.64; 95% CI: 0.92, 2.93) for children of mothers participating in group sessions, as compared to usual care. A similarly positive meta-analysed effect was observed for MMF (OR: 2.31; 95% CI: 1.61, 3.31, n = 3) (Figure 4). Removing the Kang et al. study reduced the likelihood of MMF to 2.13 (95% CI: 1.55, 2.92) for the mother/peer group platform, but remained significantly higher, as compared to usual care.





groups.

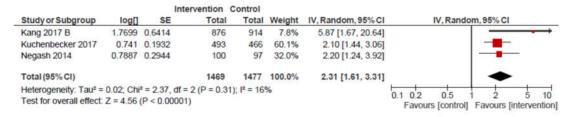


Figure 4. Forest plot of the meta-analysis of child minimum meal frequency for mother/peer groups.

In a pooled analysis of three studies that examined the effects of the group platform on < 5 stunting, underweight, and wasting, compared to care as usual, there was a nearly significant 15%

reduction in the odds of stunting (OR: 0.85; 95% CI: 0.68, 1.07) and a 24% reduction in the odds of underweight (OR: 0.76; 95% CI: 0.53, 1.07). For child wasting, the observed effect was 0.86 (95% CI: 0.65, 1.14).

| | Anticipated A | Absolute Effects | Relative | Nº of | Certainty of the | |
|--|----------------------------|-----------------------------|---------------------------|---------------------------|---------------------|--|
| Outcomes | Risk with Care as Usual | Risk with Group Sessions | Effect (95% CI) | Participants (Studies) | Evidence (GRADE) | |
| Early initiation of breastfeeding ^a | 651 per 1000 | 672 per 1000 | OR 1.10 (0.92 to 1.32) | 28,111 (5) | ⊕⊕⊕⊖ MODERATE | |
| Exclusive breastfeeding ^b | 413 per 1000 | 556 per 1000 | OR 1.78 (0.93 to 3.41) | 2900 (4) | ⊕⊕⊕⊖ Moderate | |
| Minimum dietary diversity ^c | 346 per 1000 | 554 per 1000 | OR 2.34 (1.17 to 4.70) | 3701 (4) | ⊕⊕⊖⊖ LOW | |
| Minimum meal frequency ^d | 808 per 1000 | 907 per 1000 | OR 2.31 (1.61 to 3.31) | 2946 (3) | ⊕⊕⊕⊕ HIGH | |
| Stunting ^e | 352 per 1000 | 315 per 1000 | OR 0.85 (0.68 to 1.07) | 6077 (3) | ⊕⊕⊕⊖ Moderate | |
| Underweight ^f | 382 per 1000 | 319 per 1000 | OR 0.76 (0.53 to 1.07) | 6019 (3) | ⊕⊕⊖⊖ LOW | |
| Wasting ^g | 231 per 1000 | 206 per 1000 | OR 0.86 (0.65 to 1.14) | 5924 (3) | ⊕⊕⊖⊖ LOW | |

| Table 2. Summary of outcomes for mother/peer gr | oup platform. |
|---|---------------|
|---|---------------|

^a Child is breastfed within one hour of birth. ^b Child is provided only breastmilk up to 4–6 months of age ^c Child aged 6–23 months is fed a daily minimum of four out of seven food groups according to WHO guidelines [25]. ^d Child aged 6–23 months is provided the appropriate number of daily solid/semi-solid/milk feeds (based on age and breastfeeding status) according to WHO guidelines [25]. ^e Height-for-age is more than two standard deviations below the WHO Child Growth Standards median. ^f Weight-for-age is more than two standard deviations below the WHO Child Growth Standards median. ^g Weight-for-height is more than two standard deviations below the WHO Child Growth Standards median.

3.4. Combined Home Visit and Group Platforms

Through examining studies that used both home visit and group platforms for delivery of nutrition interventions, we explored the synergistic benefits of using both modalities for improving nutrition outcomes. Meta-analyses were conducted for EIBF, EBF and stunting underweight and wasting (Table 3). Pooled results showed a significantly higher likelihood of EIBF (OR: 2.13; 95% CI: 1.12, 4.05, n = 9 among mothers receiving home visits and participating in group sessions, compared to usual care (Figure 5). Excluding the Bhutta et al., Crookston et al., and Lamstein et al. studies reduced the likelihood of EIBF to a non-significant 1.15 (95% CI: 0.82, 1.62). For EBF, the observed effect was 2.43 (95% CI: 1.70, 3.46) across 12 studies (Figure 6). Removing the Bhandari et al. and Crookston et al. studies revealed a lower, though still significantly higher than in usual care, likelihood of EBF among mothers benefiting from home visits and group sessions (OR: 1.79; 95% CI: 1.38, 2.32).

| | | | Intervention | Control | | Odds Ratio | | | Odds | Ratio | | |
|---------------------------------------|----------------------------------|----------|------------------------------|---------|--------|--------------------|-----|----------------|-----------------------|---------------------|-------|-----|
| Study or Subgroup | log[Odds Ratio] | SE | Total | Total | Weight | IV, Random, 95% C | Ľ | | IV, Rando | om, 95% Cl | | |
| Bhutta 2008 | 1.9876 | 0.1653 | 395 | 375 | 11.1% | 7.30 [5.28, 10.09] | | | | | - | • • |
| Crookston 2007 | 2.0812 | 0.2257 | 222 | 245 | 10.8% | 8.01 [5.15, 12.47] | | | | | - | -++ |
| Kung'u 2018 A Ethiopia | 0.5306 | 0.1248 | 780 | 680 | 11.2% | 1.70 [1.33, 2.17] | | | | | | |
| Kung'u 2018 B- Kenya | 0.2776 | 0.2635 | 354 | 381 | 10.6% | 1.32 [0.79, 2.21] | | | | • | | |
| Kung'u 2018 C Senegal | -0.4943 | 0.105 | 1436 | 559 | 11.3% | 0.61 [0.50, 0.75] | | | | | | |
| Lamstein 2017 | 1.8646 | 0.0623 | 2249 | 2951 | 11.3% | 6.45 [5.71, 7.29] | | | | | - | - |
| Lewycka 2013 | -0.2695 | 0.1247 | 649 | 715 | 11.2% | 0.76 [0.60, 0.98] | | | | 1 | | |
| Liang 2018 | 0.2624 | 0.0451 | 2693 | 8425 | 11.4% | 1.30 [1.19, 1.42] | | | | • | | |
| Memon 2015 | 0.6133 | 0.156 | 316 | 361 | 11.1% | 1.85 [1.36, 2.51] | | | | - | | |
| Total (95% Ci) | | | 9094 | 14692 | 100.0% | 2.13 [1.12, 4.05] | | | | - | - | |
| Heterogeneity: Tau ² = 0.9 | 4; Chi ² = 735.31, df | = 8 (P - | < 0.00001); I ² = | 99% | | | - | - | | | - 1 | |
| Test for overall effect: Z = | 2.32 (P = 0.02) | | | | | | 0.1 | 0.2 Favours | 0.5 (Intervention) | 1 2 Favours (con | trol] | 10 |

Figure 5. Forest plot of the meta-analysis of early initiation of breastfeeding for home visits and mother/peer groups.

| Study or Subgroup | log[Odds Ratio] | SE | Intervention Total | | Weight | Odds Ratio IV, Random, 95% Cl | Odds Ratio IV, Random, 95% Cl |
|--------------------------------------|----------------------|---------|-----------------------|-------|--------|----------------------------------|--|
| Bhandari 2003 | 2.8679 | 0.5004 | 464 | 411 | 5.4% | 17.60 [6.60, 46.93] | \rightarrow |
| Bhutta 2008 | 0.7148 | 0.1502 | 395 | 375 | 8.7% | 2.04 [1.52, 2.74] | |
| Brasington 2016 A- upper Egypt | 0.1242 | 0.0957 | 877 | 881 | 9.0% | 1.13 [0.94, 1.37] | + |
| Brasington 2016 B lower Egypt | 0.0325 | 0.0986 | 802 | 885 | 9.0% | 1.03 [0.85, 1.25] | - |
| Crookston 2007 | 2.7293 | 0.2349 | 222 | 245 | 8.0% | 15.32 [9.67, 24.28] | , |
| Kung'u 2018 A Ethiopia | 0.9632 | 0.2018 | 780 | 680 | 8.3% | 2.62 [1.76, 3.89] | |
| Kung'u 2018 B- Kenya | 0.8154 | 0.159 | 354 | 381 | 8.6% | 2.26 [1.65, 3.09] | |
| Kung'u 2018 C Senegal | 0.3148 | 0.12 | 1436 | 569 | 8.9% | 1.37 [1.08, 1.73] | |
| Lamstein 2017 | 0.6178 | 0.1018 | 719 | 911 | 9.0% | 1.85 [1.52, 2.26] | |
| Lewycka 2013 | 1.5582 | 0.2192 | 387 | 448 | 8.1% | 4.75 [3.09, 7.30] | |
| Liang 2018 | 0.8416 | 0.1527 | 2693 | 8425 | 8.7% | 2.32 [1.72, 3.13] | |
| Nair 2017 | 0.0583 | 0.1834 | 1315 | 1368 | 8.4% | 1.06 [0.74, 1.52] | |
| Total (95% CI) | | | 10444 | 15569 | 100.0% | 2.43 [1.70, 3.46] | • |
| Heterogeneity: Tau? = 0.36; Chi? = | = 199.01, df = 11 (P | < 0.000 | 01); 12 = 94% | | | E E | |
| Test for overall effect: Z = 4.90 (P | | | | | | 0. | 1 0.2 0.5 1 2 5 10 Favours [control] Favours [Intervention] |

Figure 6. Forest plot of the meta-analysis of exclusive breastfeeding for home visits and mother/peer groups.

The impact of the home visit and group session platform on stunting among children < 5 years was 0.97 (95% CI: 0.72, 1.30) across a pooled analysis of eight studies. The impact on < 5 underweight was 0.84 (95% CI: 0.61, 1.14; n = 6). A 23% reduction in the odds of < 5 wasting (OR: 0.77; 95% CI: 0.67, 0.89) was observed across four studies that used the home visit and group platform, compared to usual care (Figure 7).

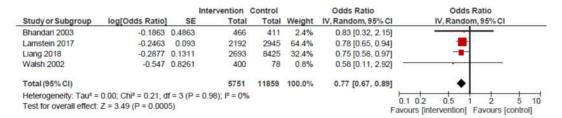


Figure 7. Forest plot of the meta-analysis of child wasting for home visits and mother/peer groups.

| | Anticipat | ed Absolute Effects | Relative | Nº of | Certainty of the |
|--|----------------------------|---|------------------------------|---------------------------|---------------------|
| Outcomes | Risk with Care as Usual | Risk with Home Visits and Group Sessions | Effect (95% CI) | Participants (Studies) | Evidence (GRADE) |
| Early initiation of breastfeeding ^a | 443 per 1000 | 628 per 1000 | OR 2.13 (1.12 to 4.05) | 23,786 (9) | ⊕○○○ VERY LOW |
| Exclusive breastfeeding ^b | 302 per 1000 | 513 per 1000 | OR 2.43 (1.70 to 3.46) | 26,013 (12) | ⊕○○○ VERY LOW |
| Stunting ^c | 315 per 1000 | 309 per 1000 | OR 0.97 (0.72 to 1.30) | 23,000 (8) | ⊕⊕⊖⊖ LOW |
| Underweight ^d | 146 per 1000 | 125 per 1000 | OR 0.84 (0.61 to 1.14) | 19,108 (6) | ⊕⊕⊖⊖ LOW |
| Wasting ^e | 59 per 1000 | 46 per 1000 | OR 0.77 (0.67 to 0.89) | 17,610 (4) | ⊕⊕⊖⊖ LOW |

Table 3. Summary of outcomes for using both CHW/peer counselor home visit and mother/peer group platforms.

^a Child is breastfed within one hour of birth. ^b Child is provided only breastmilk up to 4–6 months of age ^c Height-for-age is more than two standard deviations below the WHO Child Growth Standards median. ^d Weight-for-age is more than two standard deviations below the WHO Child Growth Standards median. ^e Weight-for-height is more than two standard deviations below the WHO Child Growth Standards median.

4. Discussion

In this paper, we have summarized the available evidence for the effectiveness of CHW/peer counselor home visits and mother/peer groups for improving child nutrition outcomes in LMICs. These syntheses comprise a subset of results from a systematic review that examined various nutrition intervention delivery platforms in LMICs. Our findings reveal that one-to-one home visits by CHWs/lay peer counselors are effective for improving breastfeeding practices, which is consistent with prior evidence. Lewin et al. [86] showed that using lay health workers resulted in a higher likelihood of EIBF (RR: 1.36; 95% CI: 1.14, 1.61) and EBF (RR: 2.78; 95% CI: 1.74, 4.44), when compared to care as usual. Other reviews also concluded that peer support, in one-to-one or group settings, is effective for improving EBF practices [87–89]. The lack of effect of the mother/peer group platform on breastfeeding in our study was, therefore, unexpected and may have been due to differences in programme factors across studies, e.g., the frequency, content, and facilitation of group sessions. We were not able to separate the effects of using trained CHWs versus peer volunteer facilitators for conducting group sessions. Further research should explore whether facilitator characteristics may affect programme targets. Further, the higher likelihood of EIBF and EBF in pooled analyses of studies involving both home visit and mother/peer group platforms suggests home visits are important for breastfeeding promotion. We speculate that this is due to the one-to-one reinforcement of knowledge and skills acquired in group sessions and/or mothers feeling more comfortable discussing and practicing breastfeeding in the privacy of home settings. Household service delivery is emphasized in the 2013 Lancet nutrition series [4] as a means for increasing the coverage of nutrition intervention at the community level.

However, our results suggest that the mother/peer group platform is effective for improving the quantity and quality of young children's diets. This is likely due, in part, to participatory cooking

demonstrations that frequently occur alongside counseling/coaching during group sessions and which have been shown to be effective for improving complementary feeding practices [90]. Our findings are consistent with a review of 36 studies that showed that women's groups focused on behaviour change are beneficial for improving child feeding practices in South Asia [13].

The absence of an observed effect of the home visit and/or group session platform on child stunting and underweight in our review was not unexpected, as effects of behavioural interventions on anthropometric outcomes are generally small, due to insufficient time periods to affect changes in child growth. Moreover, as intervention–impact pathways are influenced by a complex interplay of factors at all levels, demonstrating an attributable impact of behaviour-focused interventions on nutritional status is often challenging. Nonetheless, the observed 23% reduction in the odds of < 5 wasting using the combined home visit and group platform is encouraging, and suggests acute thinness may be more responsive to education/behaviour change interventions than the long-term nutritional deficits underlying stunting.

The moderate to high study heterogeneity in our review was expected. Though too few studies prevented subgroup investigations of sources of heterogeneity, we speculate much of the heterogeneity is attributable to underlying differences in implementation-related factors affecting coverage levels achieved. Differences in coverage varied by platform, with home visits achieving higher average coverage than peer groups (59.5% vs. 50.5%) among included studies. In a study in Bangladesh, Owais et al. [81] reported that 83% of mothers received a CHW home visit within the previous month, while only 13% reported attending a mother-to-mother support group for nutrition education. In a Lady Health Worker study in Pakistan, Yousafzai et al. [85] reported that 75% of households received monthly home visits, while 31% of caregivers participated in monthly group meetings. Lastly, Nair et al. [80] reported that 80% of mothers received a home visit and 56% of mothers attended a group meeting during the previous three months, in a study examining the impact of women's groups and home-based counseling on child growth in Jharkhand and Odisha, India.

Differences in study populations, CHW/lay counselor training/performance, exposure intensity (frequency of home visits and groups), and intervention duration also likely affected the coverage rates achieved. We were not able to assess the effects of length of intervention exposure due to the intervention period not being clearly specified in many studies. Dose–response analyses examining outcomes in relation to intervention duration and/or coverage would be useful and should be considered in future reviews.

All included studies in our review utilized an experimental design, though the overall quality of evidence was low for the majority of meta-analysed outcomes. A key limitation of our review was our inability to perform subgroup analyses to further investigate the high degree of heterogeneity, due to an insufficient number of studies. Though the inclusion of non-experimental studies in future reviews would generate a larger body of evidence, the introduction of additional biases would need to be weighed. Despite concerted efforts to select similar age/follow-up periods for anthropometric outcomes, variability across studies may have introduced some bias in the pooled analyses. However, this was not a concern for the infant and young child feeding (IYCF) outcomes, as time periods for ascertaining breastfeeding (0–6 months) and complementary feeding (6–23 months) practices were consistent across studies.

Our review strengthens the evidence through quantifying the effects of community home visits and peer groups and highlighting their potential value as platforms for improving IYCF practices in LMICs. As these tend to be established platforms within community-based child survival programmes in many countries, leveraging them for integrating nutrition-focused interventions that can be effectively provided by CHWs and through peer health promotion groups offers sizeable opportunities for nutrition behaviour change in these settings.

Therefore, the findings present an evidence-based rationale for continued investments in these community initiatives. Future research should account for platform-specific barriers affecting coverage to more broadly evaluate their effectiveness.

5. Conclusions

Knowledge of effective delivery modalities is needed to achieve maximum gains from evidencesupported interventions. In a landscape of multiple platforms and options, our findings indicate that community home visits and peer groups are important for improving nutrition-related behaviours across diverse contexts.

Supplementary Materials: The following are available online at www.mdpi.com/xxx/s1, Table S1: MEDLINE search strategy. Ovid MEDLINE(R) Epub Ahead of Print, In-Process & Other Non-Indexed Citations, Ovid MEDLINE(R) Daily and Ovid MEDLINE(R) 1946 to Present, Table S2: Characteristics of included studies for home visit and group platforms

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