

Technical Note

# Airborne LIDAR-Derived Aboveground Biomass Estimates Using a Hierarchical Bayesian Approach

Mengxi Wang <sup>1,2,†</sup>, Qingwang Liu <sup>1,†</sup>, Liyong Fu <sup>1</sup>, Guangxing Wang <sup>3</sup>  and Xiongqing Zhang <sup>2,\*</sup>

<sup>1</sup> Institute of Forest Resource Information Techniques, Chinese Academy of Forestry, Beijing 100091, China; 15501130818@163.com (M.W.); liuqw@ifrit.ac.cn (Q.L.); fuliyong@ifrit.ac.cn (L.F.)

<sup>2</sup> Key Laboratory of Tree Breeding and Cultivation of the State Forestry Administration, Research Institute of Forestry, Chinese Academy of Forestry, Beijing 100091, China

<sup>3</sup> Department of Geography and Environmental Resources, Southern Illinois University Carbondale, Carbondale, IL 62901, USA; gxwang@siu.edu

\* Correspondence: xqzhang85@caf.ac.cn; Tel.: +86-10-62888309

† The two authors contributed equally to this work.

Received: 16 April 2019; Accepted: 26 April 2019; Published: 3 May 2019



**Abstract:** Conventional ground survey data are very accurate, but expensive. Airborne lidar data can reduce the costs and effort required to conduct large-scale forest surveys. It is critical to improve biomass estimation and evaluate carbon stock when we use lidar data. Bayesian methods integrate prior information about unknown parameters, reduce the parameter estimation uncertainty, and improve model performance. This study focused on predicting the independent tree aboveground biomass (*AGB*) with a hierarchical Bayesian model using airborne LIDAR data and comparing the hierarchical Bayesian model with classical methods (nonlinear mixed effect model, NLME). Firstly, we chose the best diameter at breast height (*DBH*) model from several widely used models through a hierarchical Bayesian method. Secondly, we used the *DBH* predictions together with the tree height (*LH*) and canopy projection area (*CPA*) derived by airborne lidar as independent variables to develop the *AGB* model through a hierarchical Bayesian method with parameter priors from the NLME method. We then compared the hierarchical Bayesian method with the NLME method. The results showed that the two methods performed similarly when pooling the data, while for small sample sizes, the Bayesian method was much better than the classical method. The results of this study imply that the Bayesian method has the potential to improve the estimations of both *DBH* and *AGB* using LIDAR data, which reduces costs compared with conventional measurements.

**Keywords:** hierarchical Bayesian; classical method; airborne LIDAR data; diameter at breast height; aboveground biomass

## 1. Introduction

On the world's land surface, forests cover about 4200 million hectares, and the carbon stock of forests accounts for about 45% of the world's terrestrial carbon reserves [1]. Aboveground biomass is an important component of forest ecosystems and accounts for a large proportion of forest carbon stock; thus, quantifying forest aboveground biomass is important for forest managers [2].

Diameter at breast height (*DBH*) and aboveground biomass (*AGB*) are two important measurements of individual trees, widely used in yield estimations and forest growth [3]. Tree *DBH* is an easy measuring factor with high accuracy—all *DBH* data are measured. However, *AGB* is difficult to measure and has less accuracy than ground-observed *DBH*—only limited tree biomass can be measured in the sample [4]. Consequently, in the past few decades, many studies have proposed numerous equations to estimate aboveground biomass [5–10]. Developing an *AGB* model according to ground-based *DBH*

has been widely applied in forest investigation and yield modeling. It is important to note that *DBH* measurement in large-scale forest inventory is still time-consuming and costly.

LIDAR is a range detection system, which obtains the forest point cloud by LIDAR sensors. The LIDAR system has been used to measure the forest vegetation structure and has been applied to measure forest variables since the mid-1980s [11]. These variables, including tree height, stem volume, crown projection area, diameter at breast height, and other variables, are significantly connected to the forest biomass [3,11,12]. Using these LIDAR variables, individual tree *DBH* and aboveground biomass can be estimated by developing *DBH* and *AGB* models, respectively [3,12,13]. With some studies also using LIDAR data to predict timber volume and discern age class, LIDAR data have been widely used in forestry over the past two decades. The application of LIDAR technology has reduced the cost of forest management and provided a good opportunity for large-scale forest inventory [11,12].

The Bayesian method is used to integrate prior information about unknown parameters with the sample information; then, the posterior information is obtained according to the Bayes formula and the method of unknown parameters is further inferred according to the posterior information. Many studies have already used the Bayesian method to develop *DBH* and *AGB* models and have compared the Bayesian model with the classical model [14–17]. Zapata-Cuartas et al. [14] presented a comparison of aboveground biomass estimation in different sample sizes using the Bayesian method and the classical method. They obtained a result that the bias in the classical method was always bigger than that in the Bayesian approach.

Although LIDAR technology and Bayesian methods are used separately in individual *DBH* and *AGB* modeling [6,13–15,17–21], few studies have combined the two methods to estimate *DBH* and aboveground biomass. Tenneson [22] presented a combination of the Bayesian method and LIDAR to model LIDAR-derived forest inventory, combining the advantages of both *DBH* and *AGB* prediction; they used Bayesian model averaging (BMA), while we are using the hierarchical Bayesian method [23,24]. In this study, we used the hierarchical Bayesian method to choose the best *DBH* and *AGB* model to predict *DBH* and aboveground biomass on the basis of LIDAR point cloud data, integrating the advantages of the Bayesian approach and LIDAR data to reduce costs compared with conventional measurements [13,14]. Further, compared with the classical prediction method, using the Bayesian method can improve the accuracy to a certain degree.

## 2. Materials and Methods

### 2.1. Study Area and Data

The study area from which data were taken is located at the Xishui forest farm of Su'nan Yuguzu Autonomous County, China (38°290–38°350N, 100°120–100°200E) and lies within the temperate alpine cold semi-humid and semi-arid zone. The altitude is in the 2550 to 3680 m range, with a mean value of 2993 m in this area. This area is occupied by mountainous forests and steppes, and one of the main functions of these forests is to protect the water resources in the Dayekou Basin of the Qilian Mountains. Grass and natural mature secondary pure forests dominate the sunny and shady slopes, respectively. The dominant tree species in the forest in this area is *Picea crassifolia* Kom.

We established a single permanent sample plot (PSP) of 100 m × 100 m along the hillside and divided it into sixteen subplots of 25 m × 25 m. *DBH*, tree height (H), crown base height, and crown diameters in two perpendicular directions of all standing trees were measured in each subplot. We used a differential global positioning system unit to measure the corners and centers of the PSP in each subplot, and a total station to measure the positions of individual trees.

Airborne LIDAR data were acquired using a Lite Mapper 5600 system on 23 June 2008. The mean flight speed was 230 km h<sup>-1</sup> and the mean flight height was 3699 m above sea level. The point elevation varied from 2725 to 3193 m, and the LIDAR point density was 4.34 points m<sup>-2</sup>. We can obtain individual LIDAR tree heights and crown projection areas on the basis of the airborne LIDAR point cloud from a canopy height model (CHM), and the CHM was derived as the difference

between the digital surface model (DSM) and the digital elevation model (DEM) of the study area. The digital surface model grid, which contained elevation information including all ground objects, was interpolated from raw LIDAR point cloud data by a maximum height interpolation method and by filling null cells. For obtaining the absolute vegetation heights, we must remove the influence of the terrain. Thus, we obtained the digital elevation model from the ground-classified LIDAR point cloud data using a progressive morphological filter and, in this study, with a grid cell size of 0.5 m [3]. We also eliminated the noise using a Gaussian smoothing filter to obtain a more accurate crown projection area [3]. We used the local maxima algorithm to find the individual tree crown tops, and used the region growing method to obtain the tree crown boundary, the crown projection area was calculated according to the determined tree crown boundary. The details of the algorithms can be found in Fu et al. [3]. A total of 402 individual *P. crassifolia* tree crowns in the 16 subplots nested in the PSP were delineated; the individual tree LIDAR crown projection area and field measurement data were matched by individual tree locations measured by total station—Fu et al. [3] has drawn a figure to explain this. The biomasses of each component (stem, branch, foliage, and fruit) of the 402 *P. crassifolia* trees were estimated using the empirical allometric models developed by Wang et al. [25] using the field-measured individual *DBH* and tree height *H*. The total tree *AGB* was obtained by summing the component biomass values. For more information about the study sites and data collection, please see Fu et al. [3].

## 2.2. Random Sampling

In order to compare the LIDAR *AGB* estimation by the Bayesian method and the classical method in different sample sizes, five samples (Samples 1–4) of different sizes were randomly selected from the population. The independent variables in the population, including LIDAR crown projection area (*CPA*), LIDAR tree height (*LH*), ground-measured *DBH*, and the ground-estimated *AGB*, were assessed. Sample 1 randomly sampled 10% of the population, and Samples 2–4 randomly sampled 25%, 50%, and 75% of the population, respectively. The statistics of sample sizes 1–4 are presented in (Table 1), and we also put the statistics of the full dataset into this table.

**Table 1.** Summary statistics of different sample sizes (*DBH*, ground-measured diameter at breast height; *LH*, LIDAR tree height; *CPA*, crown projection area; *AGB*, ground-estimated aboveground biomass).

| Sample Size | Factor     | Maximum | Minimum | Mean   | SD     |
|-------------|------------|---------|---------|--------|--------|
| 10% (39)    | <i>DBH</i> | 36.40   | 2.50    | 21.26  | 7.84   |
|             | <i>LH</i>  | 9.79    | 1.96    | 6.81   | 2.17   |
|             | <i>AGB</i> | 395.83  | 1.45    | 153.08 | 102.39 |
|             | <i>CPA</i> | 11.06   | 2.94    | 7.42   | 2.20   |
| 25% (99)    | <i>DBH</i> | 40.50   | 8.20    | 23.38  | 7.77   |
|             | <i>LH</i>  | 10.10   | 2.21    | 6.80   | 1.88   |
|             | <i>AGB</i> | 508.02  | 15.68   | 178.08 | 119.52 |
|             | <i>CPA</i> | 12.80   | 3.56    | 7.32   | 1.97   |
| 50% (199)   | <i>DBH</i> | 54.50   | 2.50    | 22.73  | 7.41   |
|             | <i>LH</i>  | 11.29   | 1.96    | 6.96   | 1.85   |
|             | <i>AGB</i> | 959.49  | 1.45    | 169.51 | 118.90 |
|             | <i>CPA</i> | 13.50   | 2.94    | 7.48   | 2.03   |
| 75% (303)   | <i>DBH</i> | 54.50   | 2.50    | 23.09  | 7.78   |
|             | <i>LH</i>  | 11.30   | 1.96    | 6.92   | 1.89   |
|             | <i>AGB</i> | 959.49  | 1.45    | 176.03 | 125.55 |
|             | <i>CPA</i> | 13.50   | 2.94    | 7.41   | 2.06   |
| 100% (402)  | <i>DBH</i> | 81.10   | 2.50    | 23.40  | 8.30   |
|             | <i>LH</i>  | 11.20   | 1.90    | 6.90   | 1.90   |
|             | <i>AGB</i> | 1142.50 | 1.45    | 181.30 | 134.10 |
|             | <i>CPA</i> | 13.50   | 2.90    | 7.43   | 2.06   |

### 2.3. DBH and AGB Models

In this paper, four frequently used candidate model forms [3] and another three models [26–28] were used to develop LIDAR DBH estimation models by regarding the LIDAR-derived tree height and the crown projection area as the predictor variables. These models are listed as follows:

$$DBH = \beta_1 + \beta_2 LH + \beta_3 CPA + \varepsilon_{DBH} \text{ (Linear)} \quad (I.1)$$

$$DBH = \beta_1 [1 - \exp(-\beta_2 LH - \beta_3 CPA)] + \varepsilon_{DBH} \text{ (Richards)} \quad (I.2)$$

$$DBH = \beta_1 / [1 + \beta_2 \exp(-\beta_3 LH - \beta_4 CPA)] + \varepsilon_{DBH} \text{ (Logistic)} \quad (I.3)$$

$$DBH = \beta_1 \exp(-\beta_2 LH - \beta_3 CPA) + \varepsilon_{DBH} \text{ (Exponential)} \quad (I.4)$$

$$DBH = -\beta_1^{-1} \ln[1 - (60^{-1}(LH - 1.3))^{1/\beta_2}] + \varepsilon_{DBH} \text{ (Richards)} \quad (I.5)$$

$$DBH = [-\beta_1^{-1} \ln(1 - 60^{-1}(LH - 1.3))]^{1/\beta_2} + \varepsilon_{DBH} \text{ ([27])} \quad (I.6)$$

$$DBH = \beta_1 (LH - 1.3)^{\beta_2} \exp(\beta_3 CPA) + \varepsilon_{DBH} \text{ ([28])} \quad (I.7)$$

where  $\beta_1$ – $\beta_4$  are parameters of the models,  $LH$  is the LIDAR tree height,  $CPA$  is the LIDAR crown projection area, and  $\varepsilon_{DBH}$  is the error term.

On the basis of the best DBH model, five commonly used AGB model forms were applied to estimate the aboveground biomass. They are listed as follows:

$$AGB = \alpha_1 DBH^{\alpha_2} + \varepsilon_{AGB} \quad (II.1)$$

$$AGB = \alpha_1 DBH^{\alpha_2} LH^{\alpha_3} + \varepsilon_{AGB} \quad (II.2)$$

$$AGB = \alpha_1 DBH^{\alpha_2} CPA^{\alpha_3} + \varepsilon_{AGB} \quad (II.3)$$

$$AGB = \alpha_1 LH^{\alpha_2} CPA^{\alpha_3} + \varepsilon_{AGB} \quad (II.4)$$

$$AGB = \alpha_1 DBH^{\alpha_2} LH^{\alpha_3} CPA^{\alpha_4} + \varepsilon_{AGB} \quad (II.5)$$

where  $\alpha_1$ – $\alpha_3$  are model parameters,  $DBH$  is the LIDAR estimated DBH, and  $\varepsilon_{AGB}$  is the error term.

### 2.4. Bayesian Method

Bayesian estimation is used to combine new evidence and prior distributions with the Bayesian theorem to get a new probability. Zhang et al. [16,29] presented the Bayesian rules in detail. Let  $y = (y_1, y_2, y_3, \dots)$  represent a vector of data and  $\theta = (\theta_1, \theta_2, \theta_3, \dots)$  be a vector of parameters to be estimated. Bayes' expression is

$$p(y, \theta) = p(y|\theta)p(\theta) = p(\theta|y)p(y)$$

where  $p(\theta)$  is the prior distribution for the parameters, which can be obtained by the classical method; and  $p(\theta|y)$  is the posterior distribution of the Bayesian frameworks.

An important feature of the Bayesian method is that the parameter will be given a prior distribution before sampling. The choice of prior distribution is critical for the Bayesian method [25]. Non-informative priors with large or infinite variance that reflect prior "ignorance" are generally chosen [29–31]. Alternatively, if prior information is available from external knowledge (reported parameters from the literature), the information can be used to construct a prior distribution. Zhang et al. [32] used the parameter estimates of height growth models from classical method as the prior information for the Bayesian method, consistent with Zhang et al. [32]. Here, we used the parameters of NLME estimation as the prior distribution for the Bayesian estimation.

### 2.5. Model Evaluation

The best *DBH* and *AGB* model was chosen from the value of the deviance information criterion (*DIC*) [16], when pooling the data:

$$DIC = \bar{D} + N$$

where  $\bar{D} = E_{\theta}\{-2\log [p(y|\theta)]\}$  represents the posterior mean of the deviance; and  $N = \bar{D} + 2\log [p(y|\bar{\theta})]$  is the effective number of parameters in the model and represents the complexity of the model. A smaller value of *DIC* for a model indicates a better fit to the data [31].

In addition, the stationarity test in Heidelberg–Welch Diagnostics [33,34] was regarded as a criterion to test if the model converged in our study. The stationary process is a random process in statistics, and its unconditional joint probability distribution does not change with time. Thus, the mean value and variance do not change with time either [35].

The determination coefficient ( $R^2$ ) and mean absolute deviation (*MAD*) were used to compare the Bayesian method with the classical method. Higher  $R^2$  and smaller *MAD* values indicate a better model fit to the data. They are calculated by

$$R^2 = \frac{\sum (y - \hat{y})^2}{\sum (y - \bar{y})^2}$$

$$MAD = \sum |y - \hat{y}| / n,$$

where  $y$  is the observed *AGB*,  $\hat{y}$  is the LIDAR predicted *DBH* or *AGB*,  $\bar{y}$  is the mean value of the observed *AGB*, and  $n$  is the number of samples.

This study developed several candidate *DBH* models and selected the best model. On the basis of the best *DBH* model, *DBH* predictions were introduced to the five candidate *AGB* models as one of the independent variables. Both the *DBH* and the *AGB* models were developed through the hierarchical Bayesian approach. In addition, five different random sample sizes were used for comparing the Bayesian method with the classical method.

Considering the random effects of plot, we added a random effect parameter to  $\beta_1$  in *DBH* models and to  $\alpha_1$  in *AGB* models using hierarchical Bayesian and classical methods (nonlinear mixed effects model, NLME). The Bayesian and NLME models were constructed by using SAS procedures MCMC and MODEL, respectively [36]. For the Bayesian estimation, a burn-in period of 10,000 steps and 100,000 iterations was used to estimate parameters. The thinning parameter was set to 5 to reduce the correlation between neighboring iterations.

## 3. Results

### 3.1. Model Selection

In this study, deviance information criterion (*DIC*) statistics and the stationarity test were used to evaluate the Bayesian *DBH* and *AGB* models. As shown in Table 2, all the *DBH* models other than model I.5 passed the stationarity test (Table 2). Model I.3 had the smallest *DIC* value, followed by I.4, I.7, I.2, I.1, and I.6. Thus, model I.3 was selected as the best model to predict *DBH* used as an independent variable for modeling *AGB*.

**Table 2.** Deviance information criterion (*DIC*) values and stationarity test of seven *DBH* candidate base models.

| Model | <i>DIC</i> | Stationarity Test |
|-------|------------|-------------------|
| I.1   | 2557.572   | passed            |
| I.2   | 2556.442   | passed            |
| I.3   | 2541.001   | passed            |
| I.4   | 2545.794   | passed            |
| I.5   | -          | failed            |
| I.6   | 2580.574   | passed            |
| I.7   | 2547.292   | passed            |

The parameter estimations of model I.3 for different sample sizes are listed in Table 3. In all sample sizes, the standard deviations of all the parameters estimated by the Bayesian method were smaller than those estimated by the classical method, meaning that the parameters estimated by the Bayesian method were more stable than those estimated by the classical method. Simultaneously, we can see from Table 3 that in Bayesian estimation, the range of parameters at sample size 39 was between  $-0.036$  and  $49.34$ , while it was between  $0.066$  and  $111.3$  at full data. The ranges of parameter estimates for small sample size were significantly narrower than those of the full data, which confirms that for a small sample size, the Bayesian method shows higher efficiency.

**Table 3.** Parameter estimates and standard deviation errors/deviations (values in brackets) of the LIDAR *DBH* model for different sample sizes (NLME, nonlinear mixed effects model).

| <i>n</i>  | 39               | 99            | 199           | 303           | 402           |
|-----------|------------------|---------------|---------------|---------------|---------------|
| Bayesian  |                  |               |               |               |               |
| $\beta_1$ | 49.34 (6.428)    | 99.62 (7.66)  | 99.58 (6.7)   | 99.44 (6.68)  | 111.3 (24.46) |
| $\beta_2$ | 8.326 (1.326)    | 11.92 (1.408) | 13.2 (1.223)  | 11.96 (1.213) | 13.6 (2.79)   |
| $\beta_3$ | $-0.036$ (0.045) | 0.108 (0.04)  | 0.148 (0.025) | 0.107 (0.03)  | 0.114 (0.024) |
| $\beta_4$ | 0.297 (0.055)    | 0.068 (0.037) | 0.043 (0.021) | 0.072 (0.026) | 0.066 (0.022) |
| NLME      |                  |               |               |               |               |
| $\beta_1$ | 49.48 (9.75)     | 99.57 (64.35) | 99.37 (50.25) | 99.2 (47.63)  | 99.36 (42.52) |
| $\beta_2$ | 8.246 (1.792)    | 11.75 (7.07)  | 13.04 (5.968) | 11.83 (5.282) | 12.29 (4.76)  |
| $\beta_3$ | $-0.036$ (0.07)  | 0.108 (0.043) | 0.146 (0.03)  | 0.109 (0.033) | 0.118 (0.026) |
| $\beta_4$ | 0.292 (0.092)    | 0.067 (0.042) | 0.044 (0.026) | 0.068 (0.03)  | 0.065 (0.025) |

In terms of the fitting statistics, *AGB* models II.1, II.3, and II.4 passed the stationarity test (Table 4), however, the other two models failed. Among the three models, II.1, II.3, and II.4, the *DIC* in II.3 (4754.906) was much smaller than those in the other two models. Therefore, model II.3 was selected as the best model to predict *AGB*. The parameter estimations of model II.3 are listed in Table 5. The standard deviation of parameter estimations by Bayesian method in all sample sizes are smaller than classical method.

**Table 4.** *DIC* values and stationarity test of the five *AGB* candidate models.

| Model No. | <i>DIC</i> | Stationarity Test |
|-----------|------------|-------------------|
| II.1      | 4763.795   | passed            |
| II.2      | -          | failed            |
| II.3      | 4754.906   | passed            |
| II.4      | 4777.023   | passed            |
| II.5      | -          | failed            |

**Table 5.** Parameter estimates and standard deviation errors/deviations (values in brackets) of the LIDAR AGB model for different sample sizes.

| <i>n</i>   | 39              | 99              | 199            | 303             | 402             |
|------------|-----------------|-----------------|----------------|-----------------|-----------------|
| Bayesian   |                 |                 |                |                 |                 |
| $\alpha_1$ | 0.133 (0.0320)  | 0.027 (0.0097)  | 0.281 (0.0604) | 0.281 (0.0604)  | 0.184 (0.0339)  |
| $\alpha_2$ | 2.334 (0.3959)  | 3.385 (0.2338)  | 1.370 (0.1881) | 1.37 (0.1881)   | 2.383 (0.1428)  |
| $\alpha_3$ | −0.246 (0.5815) | −1.332 (0.3225) | 0.990 (0.2447) | 0.990 (0.2447)  | −0.346 (0.1825) |
| NLME       |                 |                 |                |                 |                 |
| $\alpha_1$ | 0.014 (0.0461)  | 0.007 (0.0101)  | 0.167 (0.1068) | 0.052 (0.0371)  | 0.096 (0.0506)  |
| $\alpha_2$ | 3.522 (1.7199)  | 4.463 (0.8609)  | 1.670 (0.4611) | 2.973 (0.3612)  | 2.781 (0.3620)  |
| $\alpha_3$ | −0.967 (1.254)  | −2.019 (0.726)  | 0.765 (0.4559) | −0.617 (0.3463) | −0.652 (0.3438) |

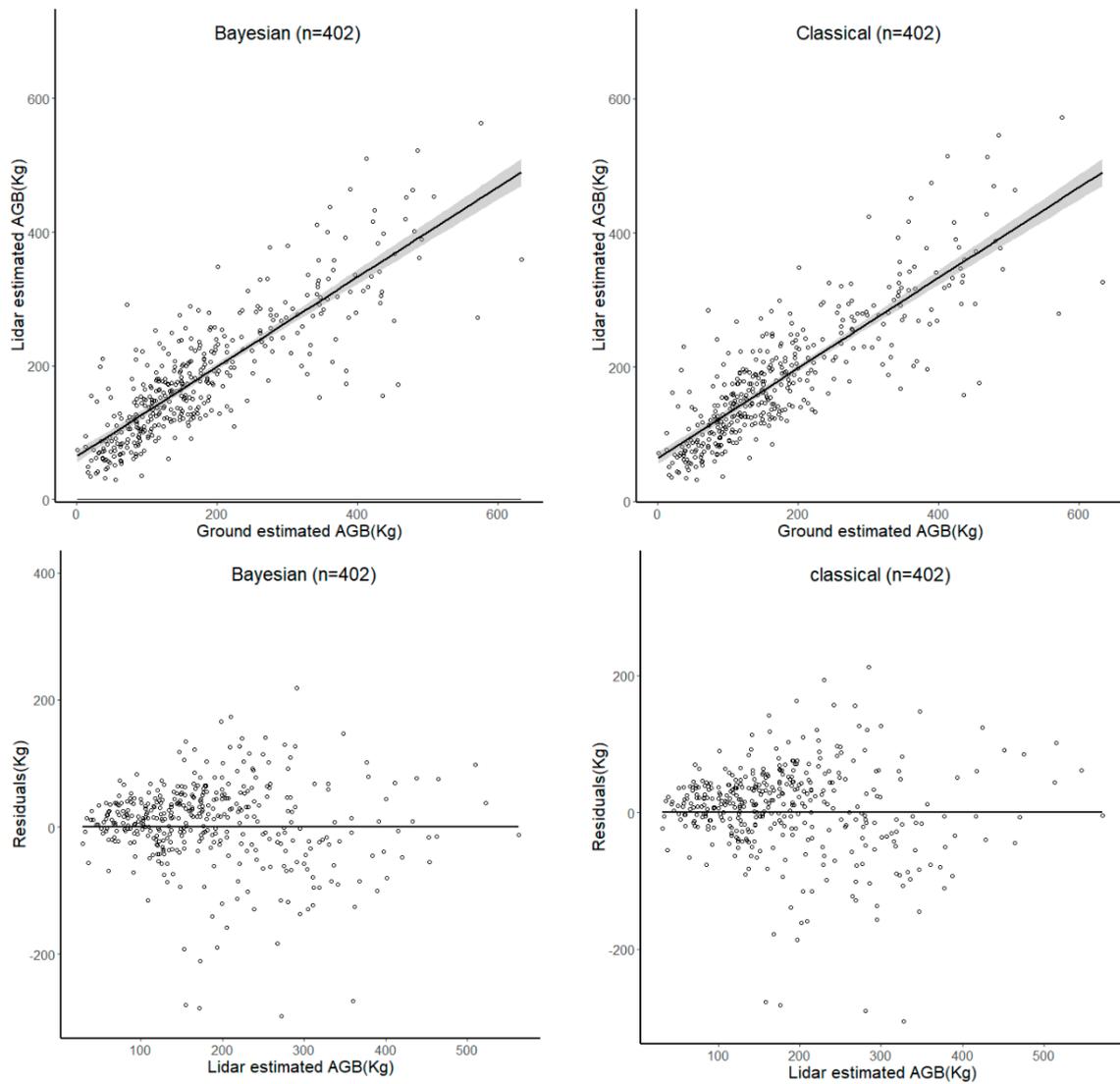
### 3.2. Comparing the Hierarchical Bayesian and NLME Methods

Figure 1 shows the correlations and residuals of AGB using the Bayesian method and the classical method when pooling the data. Both these plots show that the two methods performed well when modeling AGB. Also, the two plots from the two methods are very similar. However, as sample size decreased, the two methodologies produced large differences in  $R^2$  and MAD. For example, when pooling the data, the  $R^2$  of the Bayesian method was 0.5867, which was 2.587% larger than that of the classical method (0.5719). However, for a sample size of 39, the  $R^2$  of the Bayesian method was 11.238% larger than that of the classical method (Table 6). Similarly, although the MAD value from the Bayesian method was slightly smaller (2.172%) than that from the classical method using the whole data set, it was much smaller (16.62%) than that from the classical method for a sample size of 39 (Table 6).

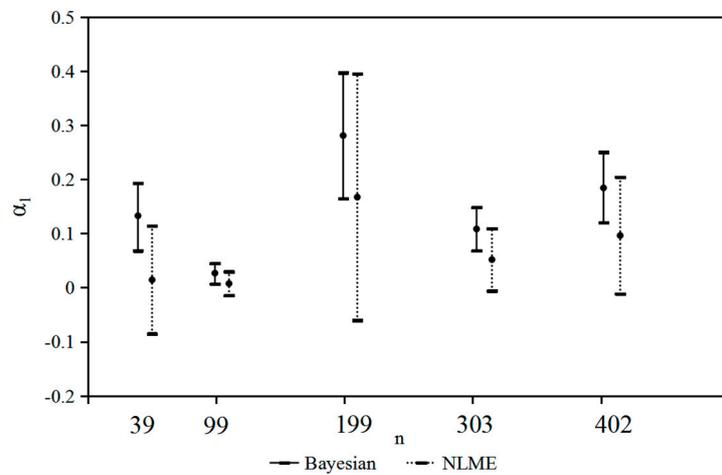
**Table 6.** The accuracy of model II.3 using different methods (hierarchical Bayesian and NLME) to estimate the value of AGB ( $R^2$ ; MAD, mean absolute deviation).

| Sample Size | Method    | $R^2$  | MAD     |
|-------------|-----------|--------|---------|
| 10% (39)    | Bayesian  | 0.8611 | 31.0571 |
|             | classical | 0.7741 | 34.2188 |
| 25% (99)    | Bayesian  | 0.7704 | 43.2709 |
|             | classical | 0.7529 | 43.1827 |
| 50% (199)   | Bayesian  | 0.7529 | 40.5010 |
|             | classical | 0.7503 | 40.2283 |
| 75% (303)   | Bayesian  | 0.7080 | 44.8375 |
|             | classical | 0.7046 | 44.6531 |
| 100% (402)  | Bayesian  | 0.5867 | 48.3125 |
|             | classical | 0.5719 | 49.3627 |

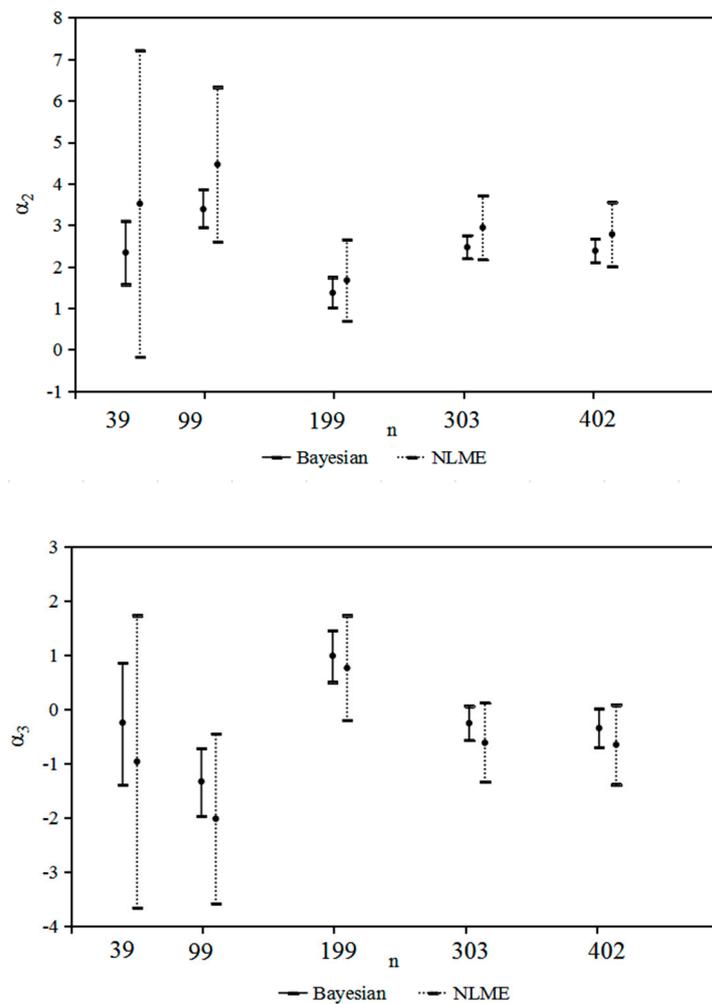
In addition, we found that parameter estimation through the Bayesian method was more stable than that through the classical method when using different sample sizes. For the parameter  $\alpha_1$ , the parameter range was 0.027–0.281 for the Bayesian method and 0.007–0.167 for the classical method. In particular, when the sample size was 39, the estimate changed largely compared with that for the pooled data set using the classical method, which could also be seen in parameters  $\alpha_2$  and  $\alpha_3$  (Figure 2). In addition, we found that the classical approach always produced estimates in a wider range than the Bayesian approach (Figure 2).



**Figure 1.** Correlation (95% interval) and residual plots for aboveground biomass (AGB) through the hierarchical Bayesian and classical (NLME, nonlinear mixed effects model) methods (From top to bottom: Correlation for AGB by hierarchical Bayesian and classical method; residual plots for AGB by hierarchical Bayesian and classical method).



**Figure 2.** Cont.



**Figure 2.** Parameter comparison with 95% intervals between the hierarchical Bayesian method and the classical method (NLME) using different sample sizes.

#### 4. Discussion and Conclusions

In this study, an individual tree *DBH* and *AGB* model was selected from seven *DBH* models and five *AGB* models using the Bayesian method on the basis of airborne LIDAR data. When selecting a model, we ensured that all models were compared at the same level and that the random effects of the parameters were considered in the same position. The criteria of model selection were the value of *DIC* and the results of the parameter stationarity test. When selecting the *AGB* model, the independent variable *DBH* was the LIDAR estimated *DBH*, which ensured that the selected *AGB* model took the correlation between *DBH* and *AGB* into account.

LIDAR has the ability to measure the forest structure, and it is a useful tool for obtaining canopy information [11]. This study used LIDAR to obtain the image data of our study area, then acquired individual tree information, including individual tree height and crown projection area, for model development. For aboveground biomass estimation, many studies have used LIDAR data. Næsset et al. [12] used LIDAR data for estimating the timber volume of forest stands. They obtained the LIDAR canopy height and canopy cover density for timber volume estimation, and they found that LIDAR data produced equal results to aerial photointerpretation, though different sites had large differences when estimating timber volume. Holmgren et al. [11] studied how to use LIDAR-derived data in regressive models for estimating the mean tree height and stem volume; the  $R^2$  of mean tree height linear regression functions was between 0.89 and 0.91 and that for average stem volumes was

between 0.82 and 0.9 when using the LIDAR mean tree height and LIDAR crown coverage area as variables. LIDAR is widely used in biomass prediction.

The study of forest biomass by the Bayesian method has been relatively less frequent than by the classical method. When the Bayesian method is compared to the classical method, an important difference is that the Bayesian method treats the parameters as random variables, while the classical method treats the parameters as fixed values; that is, the classical method treats the probability as the stable value of the frequency of a large number of repeated trials of events. Zapata-Cuartas et al. [14] used the existing biomass equation reported in the literature as the prior distribution, assuming that the information can be useful to probability distributions, and validated the larger  $R^2$  in the Bayesian method. This study used a different approach to the prior distribution—we used the parameters estimated by NLME for the prior distribution and obtained the same results.

We compared the classical method and the Bayesian method in estimating *DBH* and *AGB*; the results show that parameter estimation using the Bayesian method had higher stability than that using the classical method, and the  $R^2$  of the Bayesian method was higher than that of the classical method. Zianis et al. [17] also compared the two methods—they obtained a contrary result that the classical method was superior to the Bayesian method [17], but they did not compare the two methods on different sample sizes. In order to show that the Bayesian method is more accurate in estimating *AGB* for different sample sizes, we randomly sampled from the complete data set (Table 6). The data showed that the Bayesian method was more efficient than the classical method for different sizes of samples, which is the same as the result for the full data set. Zapata-Cuartas et al. [14] compared these two methods in the same way, and their results showed that for small sample sizes, the Bayesian method outperforms classical methods of least-square regression. Huang et al. [37] used the same method and showed that when the sample size is less than 50, we should use the Bayesian method to estimate aboveground biomass. Yao et al. [38] also presented a comparison of the Bayesian and classical methods—they found a more stable parameter value when using the Bayesian method. Thus, using the Bayesian method to estimate the aboveground biomass from small sample sizes is a good direction in biomass estimation. If classical methods are applied, especially with a small sample size, violation of statistical assumptions of error can lead to biased point estimates [20,39]. In addition, in this study, we have not taken into account the error propagation, thus a joint model is perhaps a good method, and further studies are needed to solve the joint model under the Bayesian framework for reducing the model uncertainties.

This study developed LIDAR *DBH* and LIDAR *AGB* models using the Bayesian method based on airborne LIDAR data. Previous studies have used the classical method to develop models and applied the classical method to airborne LIDAR data [9,10]. Fu et al. [1] developed a system to predict individual tree diameter and aboveground biomass using the classical method and compared two other widely used model structures (the classical method to estimate *AGB* depending on *DBH* and the classical method to estimate *AGB* not depending on *DBH*). The current article compared the classical method to the hierarchical Bayesian method, and the results of this study showed that the Bayesian method has higher  $R^2$  and smaller *MAD* than the classical method for all sample sizes; thus, the Bayesian method was better than the classical method in estimating LIDAR *DBH* and LIDAR *AGB* for *Picea crassifolia* Kom. in northwestern China.

Overall, this study combined the advantages of the Bayesian method and airborne LIDAR data; that is, airborne LIDAR data are easier to obtain than conventional aboveground measurements, and the Bayesian method was proven to have greater accuracy than the classical method. In addition, the present study also provides a reference for modeling with a small sample size.

**Author Contributions:** X.Z., L.F., Q.L., and G.W. conceived the study. M.W., X.Z., and Q.L. performed the analysis and wrote the draft of the paper. All authors contributed to reviewing and improving the paper.

**Funding:** This work was funded by the Central Public Interest Scientific Institution, Basal Research Fund (No. CAFYBB2019QD003), the Young Elite Scientists Sponsorship Program by CAST (No. 2017QNRC001), and the Central Public-Interest Scientific Institution Basal Research Fund of China under (Grant NO. CAFYBB2018GC005).

**Acknowledgments:** We appreciate the valuable and constructive suggestions from the Academic Editor and three anonymous referees that helped improve the manuscript.

**Conflicts of Interest:** The authors declare no conflict of interest.

## References

- Zhu, J.; Hu, H.; Tao, S.; Chi, X.; Li, P.; Jiang, L.; Ji, C.; Zhu, J.; Tang, Z.; Pan, Y.; et al. Carbon stocks and changes of dead organic matter in China's forests. *Nat. Commun.* **2017**, *8*. [[CrossRef](#)]
- He, Q.; Chen, E.; An, R.; Li, Y. Above-ground biomass and biomass components estimation using Lidar data in a coniferous forest. *Forests* **2013**, *4*, 984–1002. [[CrossRef](#)]
- Fu, L.; Liu, Q.; Sun, H.; Wang, Q.; Li, Z.; Chen, E.; Pang, Y.; Song, X.; Wang, G. Development of a system of compatible individual tree diameter and aboveground biomass prediction models using error-in-variable regression and airborne Lidar data. *Remote Sens.* **2018**, *10*, 325. [[CrossRef](#)]
- Fu, L.; Lei, Y.; Wang, G.; Bi, H.; Tang, S.; Song, X. Comparison of seemingly unrelated regressions with error-in-variable models for developing a system of nonlinear additive biomass equations. *Trees* **2016**, *30*, 839–857. [[CrossRef](#)]
- Nelson, R.; Krabill, W.; Tonelli, J. Estimating forest biomass and volume using airborne laser data. *Remote Sens. Environ.* **1988**, *24*, 247–267. [[CrossRef](#)]
- Neumann, M.; Saatchi, S.S.; Ulander, L.M.H.; Fransson, J.E.S. Assessing performance of l- and p-band polarimetric interferometric SAR data in estimating boreal forest above-ground biomass. *IEEE Trans. Geosci. Remote Sens.* **2012**, *50*, 714–726. [[CrossRef](#)]
- Tian, X.; Su, Z.; Chen, E.; Li, Z.; van der Tol, C.; Guo, J.; He, Q. Estimation of forest above-ground biomass using multi-parameter remote sensing data over a cold and arid area. *Int. J. Appl. Earth Obs.* **2012**, *14*, 160–168. [[CrossRef](#)]
- Zheng, C.; Mason, E.G.; Jia, L.; Wei, S.; Sun, C.; Duan, J. A single-tree additive biomass model of *Quercus variabilis* Blume forests in north China. *Trees* **2015**, *29*, 705–716. [[CrossRef](#)]
- Zianis, D.; Muukkonen, P.; Mäkipää, R.; Mencuccini, M. Biomass and stem volume equations for tree species in Europe. *Silva Fenn. Monogr.* **2005**, *4*, 1–63.
- Zianis, D.; Mencuccini, M. On simplifying allometric analyses of forest biomass. *For. Ecol. Manag.* **2004**, *187*, 311–332. [[CrossRef](#)]
- Holmgren, J.; Nilsson, M.; Olsson, H. Estimation of tree height and stem volume on plots using airborne laser scanning. *For. Sci.* **2003**, *49*, 419–428.
- Næsset, E. Estimating timber volume of forest stands using airborne laser scanner data. *Remote Sens. Environ.* **1997**, *61*, 246–253. [[CrossRef](#)]
- Hansen, E.; Ene, L.; Mauya, E.; Patocka, Z.; Mikita, T.; Gobakken, T.; Næsset, E. Comparing empirical and semi-empirical approaches to forest biomass modelling in different biomes using airborne laser scanner data. *Forests* **2017**, *8*, 170. [[CrossRef](#)]
- Zapata-Cuartas, M.; Sierra, C.A.; Alleman, L. Probability distribution of allometric coefficients and Bayesian estimation of aboveground tree biomass. *For. Ecol. Manag.* **2012**, *277*, 173–179. [[CrossRef](#)]
- Zell, J.; Bösch, B.; Kändler, G. Estimating above-ground biomass of trees: Comparing Bayesian calibration with regression technique. *Eur. J. For. Res.* **2014**, *133*, 649–660. [[CrossRef](#)]
- Zhang, X.; Duan, A.; Zhang, J.; Xiang, C. Estimating tree height-diameter models with the Bayesian method. *Sci. World J.* **2014**, *2014*, 683–691. [[CrossRef](#)] [[PubMed](#)]
- Zianis, D.; Spyroglou, G.; Tiakas, E.; Radoglou, K. Bayesian and classical models to predict above ground tree biomass allometry. *For. Sci.* **2016**, *62*, 247–259.
- Farid, A.; Goodrich, D.; Sorooshian, S. Using airborne Lidar to discern Age classes of cottonwood trees in a riparian Area. *West. J. Appl. For.* **2006**, *21*, 149–158.
- Hauglin, M.; Gobakken, T.; Astrup, R.; Ene, L.; Næsset, E. Estimating single-tree crown biomass of Norway spruce by airborne laser scanning: A comparison of methods with and without the use of terrestrial laser scanning to obtain the ground reference data. *Forests* **2014**, *5*, 384–403. [[CrossRef](#)]
- Kuyah, S.; Sileshi, G.; Rosenstock, T. Allometric models based on Bayesian frameworks give better estimates of aboveground biomass in the Miombo woodlands. *Forests* **2016**, *7*, 13. [[CrossRef](#)]

21. Picard, N.; Henry, M.; Mortier, F.; Trotta, C.; Saint-Andre, L. Using Bayesian model averaging to predict tree aboveground biomass in tropical moist forests. *For. Sci.* **2011**, *58*, 15–23. [[CrossRef](#)]
22. Hoeting, J.A.; Madigan, D.; Raftery, A.E.; Volinsky, C.T. Bayesian model averaging: A tutorial. *Stat. Sci.* **1999**, *14*, 382–417.
23. Olubusoye, O.; Okonkwo, O. Choice of priors and variable selection in Bayesian regression. *J. Asian Sci. Res.* **2018**, *2*, 354–377.
24. Tenneson, K. Development of a regional Lidar-derived forest inventory model with Bayesian model averaging for use in ponderosa pine and mixed conifer forests in Arizona and new Mexico, USA. *Remote Sens.* **2018**, *10*, 442. [[CrossRef](#)]
25. Wang, J.; Che, K.; Fu, H.; Chang, X.; Song, C.; He, H. Study on biomass of water conservation forest on north Slope of Qilian mountains. *J. Fujian Coll. For.* **1998**, *18*, 319–323. (In Chinese)
26. Bi, H.; Fox, J.; Li, Y.; Lei, Y.; Pang, Y. Evaluation of nonlinear equations for predicting diameter from tree height. *Can. J. For. Res.* **2012**, *42*, 789–806. [[CrossRef](#)]
27. Yang, R.C.; Kozak, A.; Smith, J.H.G. The potential of Weibull-type functions as a flexible growth curve. *Can. J. For. Res.* **1978**, *8*, 424–431. [[CrossRef](#)]
28. Ruark, G.A.; Martin, G.L.; Bockheim, J.G. Comparison of constant and variable allometric ratios for estimating *Populus tremuloides* biomass. *For. Sci.* **1987**, *33*, 294–300.
29. Zhang, X.; Duan, A.; Zhang, J. Tree biomass estimation of Chinese fir (*Cunninghamia lanceolata*) based on Bayesian method. *PLoS ONE* **2013**, *8*, e79868. [[CrossRef](#)]
30. Ellison, A.M. Bayesian inference in ecology. *Ecol. Lett.* **2004**, *7*, 509–520. [[CrossRef](#)]
31. Zhang, X.; Zhang, J.; Duan, A. A hierarchical Bayesian model to predict self-thinning line for Chinese Fir in southern China. *PLoS ONE* **2015**, *10*. [[CrossRef](#)] [[PubMed](#)]
32. Zhang, X.; Zhang, J.; Duan, A. Tree-height growth model for Chinese fir plantation based on Bayesian method. *Sci. Silvae Sin.* **2014**, *50*, 69–75. (In Chinese)
33. Heidelberger, P.; Welch, P.D. Simulation run length control in the presence of an initial transient. *Opns. Res.* **1983**, *31*, 1109–1144. [[CrossRef](#)]
34. Heidelberger, P.; Welch, P.D. A spectral method for confidence interval generation and run length control in simulations. *Commun. ACM* **1981**, *24*, 233–245. [[CrossRef](#)]
35. Kipiński, L.; König, R.; Sieluzyczny, C.; Kordecki, W. Application of modern tests for stationarity to single-trial meg: Data transferring powerful statistical tools from econometrics to neuroscience. *Biol. Cybern* **2011**, *105*, 183–195. [[CrossRef](#)] [[PubMed](#)]
36. SAS Institute, Inc. *SAS/STAT 9.3 User's Guide*; SAS Institute, Inc.: Cary, NC, USA, 2011; 3316p.
37. Huang, X.; Chen, D.; Sun, X.; Zhang, S. Estimation on above-ground tree biomass based on probability distribution of allometric parameters. *Sci. Silvae Sin.* **2014**, *50*, 34–41. (In Chinese)
38. Yao, D. Individual Tree Growth Model for Mongolia Oak Forest with Bayesian Statistical Inference. Master's Thesis, Chinese Academy of Forestry, Beijing, China, 2015. (In Chinese).
39. Sileshi, G.W. A critical review of forest biomass estimation models, common mistakes and corrective measures. *For. Ecol. Manag.* **2014**, *329*, 237–254. [[CrossRef](#)]

