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Nonlinear Simultaneous Equations for Individual-Tree Diameter Growth and Mortality Model of Natural Mongolian Oak Forests in Northeast China

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Academic Editors: M. Altaf Arain and Eric J. Jokela

Received: 27 March 2015 / Accepted: 12 June 2015 / Published: 23 June 2015

Abstract: A nonlinear equation system for individual tree diameter growth and mortality of natural Mongolian oak forests was developed based on 13,360 observations from 195 permanent sample plots in Northeast China. Weighted regression was used in a distance-independent diameter growth equation for dealing with heterocedasticity. Since diameter growth and mortality models have common predictors including the diameter at breast height (*DBH*), stand basal area (*BA*), basal-area-in-larger trees (*BAL*), and site index (*SI*), parameters were estimated using nonlinear three-stage least squares (N3SLS) and seemingly unrelated regression (SUR) which accounts for correlations of errors across models. The system equation provided better projection than individual fitting of the equation based on maximum likelihood estimation. Compared with the separate tree growth model, the simultaneous equations using N3SLS and SUR produced more efficient parameter estimation and smaller bias. Furthermore, N3SLS had more accurate projection. Overall, the simultaneous model will facilitate the growth and yield projection for better management of Mongolian oak forests in the region.

Keywords: Individual tree diameter growth model; mortality model; weighted regression; nonlinear three-stage least squares; seemingly unrelated regression

1. Introduction

Oak is the largest tree species in China [1]. Mongolian oak (*Ouercus mongolica* Fisch.) is a main tree species of the secondary broad-leaved forest that grows mainly throughout Northeast China. It not only has economically important value for timber production, but also has the capacity to improve environment, soil and water conservation. Unfortunately, the wood quality of most Mongolian oak in Northeast China is poor because of short clear bole lengths, frequent occurrence of water sprouts and crooked stems. However, the wood quality of Mongolian oak growing in other regions, such as Japan, is excellent, and its wood has been widely used and also exported internationally [2]. These facts indicate that the Mongolian oak growing in Northeast China has a potential to grow well like those species in Japan if adequate silvicultural treatments are applied. Forest management requires growth and vield models that enable prediction of the development of forest stands under different silvicultural systems. To overcome the negative effects of clearcutting, uneven-aged selection management has been practiced in Northeast China [3]. The selection management needs the individual tree growth and yield models because they provide good simulation of growth and mortality for short-term projections, produce detailed information about stand structure and allow consideration of a wide variety of silvicultural treatments. However, there are no individual tree growth models of Mongolian oak forests for good forest management with the exception of the integrated stand growth model developed by using system equations [4].

Diameter growth and mortality models are basic components of individual tree growth modeling. Diameter at breast height (*DBH*) is the most important and the easiest tree factor to measure. Based on *DBH*, we can directly calculate stand basal area. In addition, because there is a strong relationship between *DBH* and tree height and volume, tree height and volume can be calculated through tree height-*DBH* curve and volume equation. Therefore, the establishment of a diameter growth model for forest growth forecast is of great significance. On the other hand, mortality is an important growth process in maintaining the population dynamic changes, species diversity, and structure diversity. A mortality model is important to understand forest evolution, and thus an essential component of the growth and yield models. Both individual tree diameter growth equations and mortality models have been largely used to predict annual or periodic growth and mortality of trees [5–18]. Furthermore, from a biological perspective, the diameter growth and mortality equations of individual-tree growth model are related. Due to the highly internal correlation of these two equations, we can obtain more accurate and compatible parameter estimation through the establishment of joint estimation of simultaneous equations using nonlinear three stage least squares (N3SLS) and seemingly unrelated regression (SUR), which has been widely applied in forest growth and yield modeling [11,19–24].

Therefore, the objective of this study was to develop nonlinear equations for individual tree diameter growth and mortality through simultaneous equations, in order to obtain more accurate and compatible parameter estimation for reasonable management of natural Mongolian oak forest in Northeast China. N3SLS and SUR were examined and compared with the traditional individual tree diameter growth and mortality equations based on maximum likelihood estimation (MLE).

2. Materials and Methods

2.1. Data

The data used to develop individual-tree diameter growth and mortality models for predicting natural Mongolian oak were collected from 195 re-measured permanent sample plots (PSP) in Wangqing Forestry Bureau, Jilin Province, Northeast China (Figure 1). The area of each plot was 0.06 hm² and all of the plots were natural pure forests. These PSPs were selected using the following criteria: tree- and stand-level characteristics of these sample plots have been remeasured periodically with an interval of ten years from 1997 to 2007, and plots are located in forests without any evidence of silvicultural treatments or any other forms of human interference, such as cutting and artificial regeneration. One hundred and sixty plots were selected for model calibration, and the rest 35 plots were for model validation. On each sample plot, physical site attributes were collected including geographic coordinates, slope, aspect, and elevation. Trees with *DBH* above 5 cm were measured with species, diameter, and status (recruitment, live, or dead). In total, we received 13,360 observations. For each plot, we calculated the following stand or tree factors: stand basal area (*BA*), number of trees per hectare (*N*), stand density index (*SDI*), stand quadratic diameter and its increment (ΔD_q), and competition index.

Summary statistics of tree and plot level variables for model calibration are shown in Table 1. These plots are good representative for the research site. Mortality occurred in 152 plots out of 160 plots. The distribution of total trees and dead trees by diameter classes is shown in Figure 2, which indicates that mortality mainly occurred in the small diameter classes.



Figure 1. Map of Mongolian oak plots distribution in Wangqing Forestry Bureau, Northeast China.



Figure 2. Distribution of total trees and dead trees by diameter classes of Mongolian oak.

Table 1. Summary statistics of tree- and plot-level variables for model calibration of Mongolian oak in Wangqing Forestry Bureau, Northeast China (n = 10,698).

Variable	Mean	Max.	Min.	Std.
DBH (cm)	72.0	5.1	12.9	7.3
BA (m^2/hm^2)	42.15	5.81	23.68	5.42
SDI (stems/hm ²)	1362	237	848	207
N (stems/hm ²)	3067	533	1643	569.75
$\Delta D_q(\mathrm{cm})$	4.00	0.42	1.53	0.57

2.2. Model Development

The usually identified components of individual-tree growth models are diameter (or basal area) increment, mortality, and recruitment [25]. Recruitment was not considered in the present study. At first, diameter growth and mortality models are established separately by maximum likelihood estimation, then due to the internal correlation of these two equations, simultaneous equations are built to obtain more accurate and compatible parameter estimation by three-stage least squares (N3SLS) [22] and seemingly unrelated regression (SUR) [23].

2.2.1. Diameter Growth Model

The diameter growth model predicts increment as a function of different variables describing tree size, competition and site conditions [13]. After the test of a variety of forms of diameter growth

equations, the individual tree diameter growth adding one in ten years $(\Delta d+1)$ was selected as the predicted variable. The potential explanatory variables and their transformation were tested:

Tree size: tree diameter at breast height (DBH).

Distance-independent competition indices (COMP): basal area larger than subject trees (*BAL*), the ratio of objective tree's *DBH* and average *DBH* from stand basal area (*RD*), stand basal area per hectare (*BA*), number of trees per hectare (*N*), stand density index (*SDI*), the square sum of tree's diameter bigger than objective tree (*DL*), the ratio of objective tree's *DBH* and the biggest tree's *DBH* (*DDM*).

Site: site condition which is a composited function of slope, aspect, altitude [26] and site index which is the stand mean height at reference age and computed as Equation (1) [27] since no dominant height was available in both class A and B inventory in China.

$$SI = H \exp(29.677 / T - 29.677 / 40) \tag{1}$$

where H is stand mean height, T is stand mean age.

The model of periodical diameter growth within ten years $(\Delta d+1)$ was expressed as Equation (2):

$$(\Delta d + 1) = e^{(a+b\times SIZE + c\times COMP + d\times SITE)}$$
(2)

Non-significant variables were removed in the modeling process. In order to avoid multicollinearity, the value of the variance inflation factor (*VIF*) was computed for the groups of variables that best predicted. Those alternatives with a *VIF* larger than 10 were rejected, as suggested by Myers and Kutner *et al.* [28,29].

2.2.2. Mortality Model

The possibility of a tree dying in next 10 years was modeled by using logistic function Equation (3), which is widely applied for tree mortality [14–17]. The candidate variables for the mortality model were numerous and diverse. Hamilton classified such variables in four groups: measures of individual tree size, tree competition, tree vitality and measures of stand density [18]. Several possibilities for transformation of this variable exist and have been applied [15,18,30–32]. Only significant variables (p < 0.05) with *VIF* less than 10 were selected.

$$P_{i} = \frac{1}{1 + e^{-(b_{0} + b_{1} \times x_{1} + b_{2} \times x_{2} + \dots + b_{n} \times x_{n})}}$$
(3)

where P_i is the probability of tree mortality, b_0-b_n are parameters to be estimated and x_1-x_n are explanatory variables.

2.3. Simultaneous Systems of Individual Diameter Growth and Mortality Models

Because a strong concomitant relationship exists between $(\Delta d+1)$ and P_i , they used some same tree- and stand-level variables as predictors in both Equation (2) and Equation (3). It is likely to improve projection accuracy and ensure logical consistency. Treated in this manner, Equation (2) and Equation (3) were fitted as a nonlinear simultaneous system. Fitting this simultaneous system using maximum likelihood estimation would result in biased estimates of the coefficients, termed

simultaneity bias [33–35]. To remove simultaneity bias, the simultaneous system was firstly fitted using nonlinear three-stage least squares (N3SLS) regression, with the following three stages [24]:

- (1) Tree- and stand-level variables were used to estimate a first-stage estimate of $(\Delta d+1)$ $((\Delta d+1)_{first})$ with a nonlinear model. Then a first-stage estimate of P_i (P_{ifirst}) was developed by the same basic procedure with a binary logistic model.
- (2) Some same or different tree- and stand-level variables were taken as predictor variables in both Equation (2) and Equation (3) in the second stage. Simultaneity bias across simultaneous systems of equations was removed by this. Simultaneous correlation and the correlation of errors, which arise in the two equations of the system, were then estimated by the error terms of these second-stage equations.
- (3) In the third stage, the estimated simultaneous correlation was used in which a seemingly unrelated regressions approach was taken to fit the two equations as a system.

For comparison, a consistent and asymptotically efficient estimator also was obtained if the system of nonlinear equations was jointly estimated under seemingly unrelated regression SUR directly [23].

Initial model residual diagnosis showed that there was heterocedasticity for diameter growth equation. To deal with the problem, an additional weighting factor was used. Heterocedasticity residual variance was modeled as a power function of an independent variable [36,37]: $\hat{\sigma}^2 \propto DBH^q$. The *q* value was subsequently included in the weighting factor (*w*) as: $w = \sqrt{1/DBH^q}$.

2.4. Model Evaluation

For individual tree diameter growth model, the following statistics were calculated: mean bias (MB) and mean relative bias (MB%), which showed the bias of model prediction, root mean square error (RMSE), and relative root mean square error (RMSE%) which is expressed in the same units as the dependent variable, giving an idea of the mean error when using the model.

Mean bias:

$$MB = \sum_{i=1}^{n} \left(\frac{y_i - \hat{y}_i}{n} \right) \tag{4}$$

Mean relative bias:

$$MB\% = \left(\sum_{i=1}^{n} \frac{y_i - \hat{y}_i}{y_i}\right) / n \times 100$$
(5)

Root mean square error:

$$RMSE = \sqrt{\frac{\sum_{i=1}^{n} (y_i - \hat{y}_i)^2}{n - p}}$$
(6)

Relative root mean square error:

$$RMSE\% = \frac{\sqrt{\sum_{i=1}^{n} (y_i - \hat{y}_i)^2 / (n - p)}}{\overline{y}} \times 100$$
(7)

where \mathcal{Y}_i is observed value, $\hat{\mathcal{Y}}_i$ is predicted value, $\bar{\mathcal{Y}}$ is average value, *n* is the number of trees, and *p* is the number of parameters.

For individual tree mortality model, a probability threshold has to be determined to assign mortality. If the estimated probability of mortality exceeds the threshold then the tree is considered dead. The principle of using the maximum summary threshold (MST) was used in the study [38], namely taking the probability threshold when the sum of sensitivity and specificity is maximum (Table 2). Specificity is the ratio of the number of correctly predicted survived trees and observed survived trees. Sensitivity is the ratio of the number of correctly predicted dead trees and observed dead trees.

I de ant		Predic	Predicted Values		
Iter	n	Dead Trees	Survived Trees	1 otai	
Observed values	Dead trees	a_{11}	a_{12}	A_{1m}	
	Survived trees	a_{21}	a_{22}	A_{2m}	
Total		A_{n1}	A_{n2}	A_{nm}	
	Sensitivity = $a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/a_{11}/$	A _{1m} Specificity	$y = a_{22}/A_{2m}$		

Table 2. Contingency table of the observed values and predicted values.

Where a_{11} is the number of correctly predicted dead trees, a_{12} is the number of incorrectly predicted dead trees, a_{21} is the number of incorrectly predicted survived trees, a_{22} is the number of correctly predicted survived trees, A_{1m} is observed dead trees, A_{2m} is observed survived trees, A_{n1} is predicted dead trees, A_{n2} is predicted survived trees, A_{nm} is total trees.

The area under the receiver operating characteristic (ROC) curve was calculated for the mortality model. According to Hosmer and Lemeshow [39], the area under the ROC curve is a threshold independent measure of model discrimination, where a value of 0.5 suggests no discrimination, 0.7–0.8 suggests acceptable discrimination, and 0.8–0.9 suggests excellent discrimination. To carefully check the bias among diameter classes, the deviations between predicted and observed values were tested by means of Pearson Chi-Squared Statistics [40,41]. Chi-Squared values for classes within an attribute were calculated as:

$$\chi^{2} = \frac{\left(ND_{obs} - ND_{pred}\right)^{2}}{ND_{pred}}$$
(8)

where *ND*_{obs}, *ND*_{pred} are the number of observed and predicted dead trees in a diameter class, respectively.

Large chi-squared values provided evidence of lack of fit. When the chi-squared values we calculated are less than the critical chi-squared values with significant level p = 0.05, it means there is no significant difference between the probability of observed and predicted dead trees.

Model calibrations were carried out in R with the lme4 and systemfit package [42].

3. Results

3.1. Individual-Tree Diameter Growth Model

Equation (9) was the final model and its parameter estimates were listed in Table 3. The most important explanatory variables were the log of tree DBH ($\ln DBH$), stand basal area (*BA*), basal area larger than subject trees (*BAL*) and site index (*SI*). The standard errors of all explanatory variables are close to zero and the values of the variance inflation factor (*VIF*) are lower than 5.

$$(\Delta d + 1) = e^{(a_0 + a_1 \times \ln DBH + a_2 \times BA + a_3 \times BAL + a_4 \times SI)}$$
(9)

Table 3. The statistical analysis results for first-stage equations of individual-tree diameter growth model for Mongolian oak (n = 10,698).

Variable	Parameters	Standard Error	VIF
Intercept	0.4158 ***	0.0444	/
ln <i>DBH</i>	0.3092 ***	0.0108	3.2142
BA	-0.0137 ***	0.0008	2.0865
BAL	-0.0900 ***	0.0121	1.2412
SI	0.0009 *	0.0021	0.8546
	*(n < 0.05) **(n < 0.05)	(0.01) *** (n < 0.001)	

* (p < 0.05), ** (p < 0.01), *** (p < 0.001).

3.2. Mortality Model

Estimated parameters for mortality model Equation (10) are shown in Table 4. The most important explanatory variables were tree size: inverse of diameter in breast height (1/*DBH*); competition factors: basal area larger than subject trees (*BAL*), stand basal area per hectare (*BA*); Site factor: site index (*SI*). All the parameter estimates were highly significant and all *VIF*s were smaller than 5. The threshold of mortality probability is 0.3 according to MST principle.

$$P_{i} = \frac{1}{1 + e^{-(b_{0} + b_{1} \times 1/DBH + b_{2} \times BA + b_{3} \times BAL + b_{4} \times SI)}}$$
(10)

Table 4. The statistical analysis results for first-stage equations of individual tree mortality model (n = 10,698).

Independent Variable	Parameters Estimated Values	Standard Error	Odds Ratio	VIF		
Intercept	-3.9740 ***	0.2817	0.1745	/		
1/DBH	12.4642 ***	0.9982	4.8562	2.7452		
BA	-0.0254 ***	0.0103	0.95621	2.3256		
BAL	1.3093 ***	0.1609	10.3542	4.6253		
SI	-0.0177 **	0.0151	1.9854	2.8542		
* (

* (p < 0.05), ** (p < 0.01), *** (p < 0.001).

3.3. Simultaneous Systems of Equations Using N3SLS and SUR

Individual-tree diameter growth and mortality model are interrelated. For these above two models, the final instrumental variables used to obtain first stage estimates were *DBH*, *BA*, *BAL*, and *SI*. Therefore, the simultaneous equations are developed using seemingly unrelated regression N3SLS and SUR.

$$(\Delta d + 1) = e^{(a_0 + a_1 \times \ln DBH + a_2 \times BA + a_3 \times BAL + a_4 \times SI)}$$
(11)

$$\hat{P}_{i} = \frac{1}{1 + e^{-(b_{0} + b_{1} \times 1/DBH + b_{2} \times BA + b_{3} \times BAL + b_{4} \times SI)}}$$
(12)

where $(\Delta \hat{d} + 1)$ is the first-stage estimate of $(\Delta d+1)$, \hat{P}_i is the first-stage estimate of P_i , and a_0-a_4 and b_0-b_4 are sets of species-specific parameters for the first-stage equations of $(\Delta d+1)$ and P_i , respectively.

Parameter estimates for the system of equations using N3SLS and SUR are presented in Tables 5 and 6, respectively. Both methods produce similar parameter estimates, but SUR produced smaller standard error than N3SLS. For the individual-tree diameter growth model, higher values of ln*DBH* and *SI* led to greater increments, while higher values of any of the other variables (*BA*, *BAL*) resulted in lower increments. It is consistent with the separate individual-tree diameter growth model. In the case of the mortality model, higher values of *SI* resulted in lower mortality, while higher values of any of the other variables (*1/DBH*, *BAL*, *BA*) resulted in greater mortality. It is not consistent with the separate mortality model but biologically reasonable since small trees and dense stands with high competition normally resulted in high mortality. Thus, the simultaneous systems of equations removed simultaneity bias and had more accurate estimated parameters.

Table 5. The parameter estimates from N3SLS regression of $(\Delta d+1)$ and P_i in the system of equations for Mongolia oak (n = 10,698).

Model Predictor		Parameter Values	Standard Error	Pr > t
	Intercept	0.3778	0.0439	< 0.0001
	lnDBH	0.3344	0.0108	< 0.0001
Diameter growth model	BA	-0.0144	0.0079	< 0.0001
	BAL	-0.0855	0.0121	< 0.0001
	SI	0.0021	0.0021	0.0013
	Intercept	-0.2202	0.0249	< 0.0001
	1/DBH	1.4910	0.0856	< 0.0001
Mortality model	BA	0.0043	0.0006	< 0.0001
	BAL	0.0828	0.0088	< 0.0001
	SI	-0.0012	0.0015	0.0042

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Model Predictor		Parameter Values	Standard Error	Pr > t
	Intercept	0.4443	0.0433	< 0.0001
	lnDBH	0.3107	0.0105	< 0.0001
Diameter growth model	BA	-0.0141	0.0008	< 0.0001
	BAL	-0.0981	0.0120	< 0.0001
	SI	0.0022	0.0020	0.0016
	Intercept	-0.1904	0.0247	< 0.0001
	1/DBH	1.1283	0.0824	< 0.0001
Mortality model	BA	0.0042	0.0005	< 0.0001
	BAL	0.0954	0.0087	< 0.0001
	SI	-0.0009	0.0014	0.0029

Table 6. The parameter estimates from SUR of $(\Delta d+1)$ and P_i in the system of equations for Mongolia oak (n = 10,698).

3.4. Model Comparison and Evaluation

The overall performance of Equations (11) and (12) as a system of equations using N3SLS and SUR was assessed through MB (cm), MB%, RMSE (cm) and RMSE%. Table 7 showed that these error statistics from simultaneous equations for both calibration data and validation data were lower than the separate individual tree diameter growth model. Overall, the simultaneous equations with N3SLS and SUR reduced the MB, MB%, RMSE, RMSE% by 4.56%, 13.37%, 3.38%, 8.29% and 3.00%, 9.18%, 2.63%, 4.93% for calibration data, and 3.31%, 15.61%, 3.90%, 6.19% and 2.44%, 6.47%, 2.67%, 4.19% for validation data from the individual-tree diameter growth model, respectively.

The plots of weighted residuals against predicted values for the final individual-tree diameter growth equations (Figure 3) showed no evidence of heterogeneous variance over the full range of the predicted values and no systematic pattern in the variation of the residuals after the correction for heterocedasticity and non-normality. Moreover, the plots of weighted residuals against predicted values for the system of simultaneous equations using N3SLS (Figure 4) and SUR (Figure 5) showed, more obviously, no evidence of heterogeneous variance over the full range of the predicted values. It indicated that simultaneous equations using N3SLS and SUR were more precise than estimates of the separate individual-tree diameter growth model, N3SLS had the most accurate projection.



Figure 3. Weighted residuals and predicted values of diameter growth model.



Figure 4. Weighted residuals and predicted values of simultaneous equations for N3SLS.

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Statistics	Calibration Data				Validation Data	
Statistics	Diameter Growth	N3SLS	SUR	Diameter Growth	N3SLS	SUR
MB	3.9421×10^{-5}	$3.7623 imes 10^{-5}$	3.8238×10^{-5}	$4.2245 imes 10^{-5}$	$4.0845 imes 10^{-5}$	4.1213×10^{-5}
MB%	0.1242	0.1076	0.1128	0.1345	0.1135	0.1258
RMSE	0.6358	0.6143	0.6191	0.6789	0.6524	0.6608
RMSE%	20.2546	18.5746	19.2546	17.9554	16.8437	17.2028

Table 7. Estimated errors of separate diameter growth model and simultaneous equation using N3SLS and SUR.



Figure 5. Weighted residuals and predicted values of simultaneous equations for SUR.

The ROC curve for both fittings were very similar, and the area under the curve indicated that the model discriminated well [39] between live and dead trees (Figures 6–8). The AUC values of simultaneous equations and mortality model were 0.776, 0.761, 0.740, respectively.

The chi-squared statistics in Table 8 gave no evidence of lack of fit between predicted and observed values for both models. Examining the predicted and observed values by diameter class showed that simultaneous equations using N3SLS and SUR produced more accurate projection than separate mortality model (Figure 9), but there are larger errors in big diameter classes. It indicated that estimates of simultaneous equations were more precise than estimates of the separate mortality model for the range of data tested, and N3SLS had the most precise projection.



Figure 6. ROC curve of separate mortality model.



Figure 7. ROC curve of simultaneous equations for N3SLS.



Figure 8. ROC curve of simultaneous equations for SUR.



Figure 9. Observed and predicted mortality trees for simultaneous equations using N3SLS and SUR and separate mortality model.

Data	Models	Total Trees	Observed Dead Trees	Predicted Dead Trees	Chi Squared
Calibration data	Mortality model	10,698	1420	1348	3.84 ^{ns}
	N3SLS	10,698	1420	1364	2.30 ^{ns}
	SUR	10,698	1420	1355	3.12 ^{ns}
	Mortality model	2,662	373	341	3.00 ^{ns}
Validation data	N3SLS	2,662	373	350	1.51 ^{ns}
	SUR	2,662	373	345	2.27 ^{ns}

Table 8. Test in predicting and observing the mortality probability for calibration and validation data sets.

ns is not significant.

4. Discussiony

Diameter growth and mortality models have long been recognized as a critical component in individual tree growth modeling. This study firstly presented a separate individual-tree diameter growth model and mortality model for Mongolia oak in Northeast China, based on permanent sample plots. Potential predictor variables were selected based on their biological importance for tree growth and mortality rather than simply fitting statistics because if a model does not make biological sense, it will not perform well for any data set other than that used for model development [18]. The predictor variables for both models are related to tree size, tree competition and site conditions. After testing different independent variables, the following were included in the two models: the initial diameter at

breast height (DBH), basal area of larger trees (BAL), stand basal area (BA), and site index (SI). Owing to same predictors and error correlation of both models, the simultaneous fitting was done through N3SLS and SUR. Possible correlated error components in the equations could be explained by the simultaneous fitting, because some tree- and stand-level variables depended on parameter estimates and were used in both equations. In contrast to individual fitting of the equations, this approach results in a decrease in the approximate standard errors of the coefficients and more accurate projections as shown in this study. Both stand- and tree-level system equations were reported to be superior to separate growth models [22,23,43–45]. However, the simultaneous fitting using N3SLS provided the more precise projection and similar parameter estimates than those obtained using SUR in the region. Because the bias from nonlinear N3SLS were smaller than those from SUR, which means N3SLS was more efficient than SUR in this region. Therefore, N3SLS is the most appropriate method for system equation estimate in the study. Recently, a system of equations was developed by Sattler and LeMay to predict crown length and crown radius for trees in structurally complex stands [24]. Compatible branch and foliage biomass simultaneous equations were constructed based on the crown biomass mixed models, in which the crown biomass was divided into the two components of branch and foliage biomass [46].

In the simultaneous fitting using N3SLS, the individual tree diameter growth model showed that higher values of ln*DBH* and *SI* led to greater increments, while higher values of any of the other variables (*BA and BAL*) resulted in lower increments. The results confirmed that bigger trees and trees in better sites representing more vigor had higher diameter growth. The basal area of larger trees (*BAL*) confirmed that an increase in competition leads to a reduction in the diameter growth. *BAL* has been considered as a variable to capture competition for light [47]. Stand basal area (*BA*) also has negative effects, which showed a decrease in the growth rate as the stand becomes more crowded. It is consistent with the separate individual tree diameter growth model [13,48–50]. In addition, the crown characteristics were also included in some individual tree diameter growth models [9,51]. However, crown measurements were not used in this study, and no crown models were available for the species in the study area. Therefore, it will be interesting to test whether crown attributes would be beneficial to assess Mongolia oak stand dynamics in future research.

In the case of the mortality model from simultaneous fitting using N3SLS, higher values of *SI* representing good site condition resulted in lower mortality, while higher values of any of the other variables (1/*DBH*, *BAL*, *BA*) resulted in greater mortality. The results implied that the mortality probability of big trees is lower than that of small trees. Similar results have been reported for this species from Austria [15], Norway [41], and Finland [52]. However, Monserud and Sterba [15] hypothesized an increased mortality rate for very large diameters due to senescence of over-mature trees, included DBH² in the model, and found it significant in predicting mortality for Norway spruce. The basal area of larger trees (*BAL*), stand basal area (*BA*) indicated that an increase in competition of stand leads to a higher possibility of mortality. This means that the denser the stand and the more suppressed the tree, the greater is the mortality rate [41,53,54]. The size of the tree affects one-sided or asymmetric competition and occurs when larger trees have a disproportionate share of resources [55], so smaller trees are more likely to die due to stagnation. Light is the most common example of an asymmetric resource because of its directionality, but it is perhaps not the only one, since the degree of symmetry or asymmetry of mineral nutrients and water are not enough studied [47]. The basal area of

larger trees and the ratio of basal area of individual tree and stand have been considered as a variable to capture competition for light, and, hence, it is seen as a surrogate of one-sided competition [56], while stand basal area and number of trees is used to express two-side competition [53]. This makes sense as these trees grow in relatively sparse stands with high availability of light, where competition seems to be for belowground resources [57].

Owing to lack of stand dominant height, site productivity in the study was realized by the stand mean height at a given age, which is a practical measure of site productivity in the past [58]. It had significant effects on diameter and mortality in the both models as expected. There are still unexplained variances in both individual diameter growth and mortality models, and further model improvement will be considered in the future with the application of a mixed-effect model with plot as a random effect, for example. Height growth and recruitment models will also be needed.

5. Conclusions

A system of nonlinear simultaneous equations for individual tree growth and mortality in 10 years for natural Mongolian oak forests in Northeast China was developed. Parameters were estimated using nonlinear three-stage least squares (N3SLS) and seemingly unrelated regression (SUR), which provided better predictions than individual fitting of the equation, and N3SLS had the best prediction results. The model will facilitate the growth and yield prediction for better management of Mongolian oak forests in the region.

Acknowledgments

This work was supported by the National Twelfth-Five Year Research Program of China (Grant No. 2012BAD22B0201).

Author Contributions

Wu Ma contributed to the modelling and writing, and Xiangdong Lei initiated the research and led the writing.

Conflicts of Interest

The authors declare no conflict of interest.

References

- 1. SFA. *National Forest Resources Statistics (2009–2013)*; State Forestry Administration: Beijing, China, 2014; p. 233. (In Chinese)
- 2. Sano, J. Studies on regeneration patterns in natural oak (*Quercus mongolica* var. grosseserrata.) forests through community structure analysis. *Res. Bull. Exp. Hokkaido Univ.* **1988**, *45*, 221–266.
- 3. Hao, Q.Y.; Meng, F.R.; Zhou, Y.P.; Wang, J.X. Determining the optimal selective harvest strategy for mixed-species stands with a transition matrix growth model. *New For.* 2005, *29*, 207–219.

- 4. Hong, L.X.; Lei, X.D.; Li, Y.C. Integrated stand growth model of Mongolian oak and its application. *For. Res.* **2012**, *25*, 201–206. (In Chinese)
- 5. Stage, A.R. Prognosis model for stand development. In *US Forest Service Research Paper INT-137*; Intermountain Forest and Range Experiment Station: Ogden, UT, USA, 1973; p. 32.
- Burkhart, H.E.; Farrar, K.D.; Amateis, R.L.; Daniels, R.F. Simulation of Individual Tree Growth and Stand Development in Loblolly Pine Plantations on Cutover, Site-Prepared Areas; Publication No. FWS-I-87; School of Forestry and Wildlife Research; Virginia Polytechnic Institute and State University: Blacksburg, VA, USA, 1987; p. 47.
- 7. Zhang, S.; Amateis, R.L.; Burkhart, H.E. Constraining individual tree diameter increment and survival models for loblolly pine plantations. *For. Sci.* **1997**, *43*, 414–423.
- 8. Cao, Q.V. Prediction of annual diameter growth and survival for individual trees from periodic measurements. *For. Sci.* **2000**, *46*, 127–131.
- 9. Cao, Q.V.; Li, S.; McDill, M.E. Developing a system of annual tree growth equations for the loblolly pine-shortleaf pine type in Louisiana. *Can. J. For. Res.* **2002**, *32*, 2051–2059.
- 10. Cao, Q.V. Predicting parameters of a Weibull function for modeling diameter distribution. *For. Sci.* **2004**, *50*, 682–685.
- 11. Cao, Q.V.; Strub, M.R. Simultaneous estimation of parameters of an annual tree survival and diameter growth model. *For. Sci.* **2008**, *54*, 617–624.
- 12. Coble, D.W.; Cao, Q.V.; Jordan, L. An annual tree survival and diameter growth model for loblolly and slash pine plantations in east Texas. *South. J. Appl. For.* **2012**, *36*, 79–84.
- 13. Adame, P.; Hynynen, J.; Cañellas, I.; del Río, M. Individual-tree diameter growth model for rebollo oak (*Quercus pyrenaica* Willd.) coppices. *For. Ecol. Manag.* **2008**, *255*, 1011–1022.
- 14. Monserud, R.A. Simulation of forest tree mortality. For. Sci. 1976, 22, 438-444.
- 15. Monserud, R.A.; Sterba, H. Modeling individual tree mortality for Austrian forest species. *For. Ecol. Manag.* **1999**, *113*, 109–123.
- Diéguez-Aranda, U.; Castedo-Dorado, F.; Álvarez-González, J.G.; Rodríguez-Soalleiro, R. Modelling mortality of Scots pine (*Pinus sylvestris* L.) plantations in the northwest of Spain. *Eur. J. For. Res.* 2005, *124*, 143–153.
- Crecente-Campo, F.; Marshall, P.; Rodríguez-Soalleiro, R. Modeling non-catastrophic individual-tree mortality for *Pinus radiata* plantations in northwestern Spain. *For. Ecol. Manag.* 2009, 257, 1542–1550.
- 18. Hamilton, D.A. A logistic model of mortality in thinned and unthinned mixed conifer stands of northern Idaho. *For. Sci.* **1986**, *32*, 989–1000.
- 19. Clutter, J.T. Compatible growth and yield model for 1oblolly pine. For. Sci. 1963, 9, 354–371.
- 20. Borders, B.E.; Bailey, R.L. A compatible system of growth and yield equations for slash pine fitted with restricted three-stage least squares. *For. Sci.* **1986**, *32*, 185–201.
- 21. Borders, B.E. Systems of equations in forest stand modeling. For. Sci. 1989, 35, 548-556.
- 22. Hasenauer, H.; Monserud, R.A. A crown ratio model for Austrian forests. *For. Ecol. Manag.* **1996**, *84*, 49–60.
- 23. Rose, C.E.; Lynch, T.B. Estimating parameters for tree basal area growth with a system of equations and seemingly unrelated regressions. *For. Ecol. Manag.* **2001**, *148*, 51–61.

- 24. Sattler, D.F.; LeMay, V. A system of nonlinear simultaneous equations for crown length and crown radius for the forest dynamics model SORTIE-ND. *Can. J. For. Res.* **2011**, *41*, 1567–1576.
- 25. Vanclay, J.K. *Modelling Forest Growth and Yield: Application to Mixed Tropical Forests*; CAB International: Wallingford, UK, 1994; p. 312.
- Stage, A.R.; Wykoff, W.R. Inventory Implications of Using Sampling Variances in Estimation of Growth Model Coefficients. In *Integrated Tools for Natural Resources Inventories in the 21st Century*; General Technical Report NC-212; United States Department of Agriculture Forest Service: St. Paul, MN, USA, 2000; pp. 587–595.
- 27. Du, J.S.; Tang, S.Z.; Wang, H.L. Natural forest growth models for subcompartment data updating. *For. Sci.* **2000**, *36*, 26–32. (In Chinese)
- 28. Myers, R.H. *Classical and Modern Regression with Applications*, 2nd ed.; Duxbury Press: Belmont, CA, USA, 1990; p. 488.
- 29. Kutner, M.H.; Nachtsheim, C.J.; Neter, J.; Li, W. *Applied Linear Statistical Models*, 5th ed.; McGraw-Hill: Lisbon, Portugal, 2005; p. 1396.
- Ojansuu, R.; Hynynene, J.; Koivunen, J.; Luoma, P. Luonnonprosessit Metsalaskelmassa. *METSA* 2000, 385, 1–59.
- 31. Vanclay, J.K. Mortality functions for north Queensland rain forests. J. Trop. For. Sci. 1991, 4, 15–36.
- 32. Tuhus, E. Natural mortality of trees; Norwegian Forest and Landscape Institute: Ås, Norway, 1997.
- 33. Zellner, A.; Theil, H. Three-stage least squares: Simultaneous estimation of simultaneous equations. *Econometrica* **1962**, *30*, 54–78.
- 34. Gallant, A.R. Seemingly unrelated nonlinear regressions. J. Econ. 1975, 3, 35–50.
- 35. Judge, G.G.; Griffiths, R.C.; Lutkepohl, H.; Lee, T.C. *The Theory and Practice of Econometrics*, 2nd ed.; John Wiley and Sons: New York, NY, USA, 1985.
- 36. Hann, D.W. An adjustable predictor of crown profile for stand-grown Douglas fir trees. *For. Sci.* **1999**, *45*, 217–225.
- Crecente-Campo, F.; Marshall, P.; LeMay, V.; Diéguez-Aranda, U. A crown profile model for *Pinus radiata* D. Don in northwestern Spain. *For. Ecol. Manag.* 2009, 257, 2370–2379.
- 38. Liu, C.; Berry, P.M.; Dawson, T.P.; Pearson, R.G. Selecting thresholds of occurrence in the prediction of species distributions. *Ecography* **2005**, *28*, 385–393.
- 39. Hosmer, D.W.; Lemeshow, S. *Applied Logistic Regression*, 2nd ed.; John Wiley & Sons: New York, NY, USA, 2000; p. 375.
- 40. Agresti, A. *An Introduction to Categorical Data Analysis*; Wiley: New York, NY, USA, 1996; p. 290.
- 41. Eid, T.; Tuhus, E. Models for individual tree mortality in Norway. *For. Ecol. Manag.* **2001**, *154*, 69–84.
- R Development Core Team. R: A Language and Environment for Statistical Computing; R Foundation for Statistical Computing: Vienna, Austria, 2013. Available online: http://www. R-project.org (accessed on 16 May 2013).

- Hasenauer, H.; Monserud, R.A.; Gregoire, T.G. Cross-Correlations among Single Tree Growth Models; General Technical Report; United States Department of Agriculture, Forest Service, Rocky Mountain: Fort Collins, CO, USA, 1996; pp. 667–675.
- 44. Huang, S.; Titus, S.J. Estimating a system of nonlinear simultaneous individual tree models for white spruce in boreal mixed-species stands. *Can. J. For. Res.* **1999**, *29*, 1805–1811.
- 45. Crecente-Campo, F.; Soares, P.; Tomé, M.; Diéguez-Aranda, U. Modelling annual individual-tree growth and mortality of Scots pine with data obtained at irregular measurement intervals and containing missing observations. *For. Ecol. Manag.* **2010**, *260*, 1965–1974.
- 46. Zou, W.T.; Zeng, W.S.; Zhang, L.J.; Zeng, M. Modeling crown biomass for four pine species in China. *Forests* **2015**, *6*, 433–449.
- 47. Schwinning, S.; Weiner, J. Mechanisms determining the degree of size asymmetry in competition among plants. *Oecologia* **1998**, *113*, 447–455.
- Crecente-Campo, F.; Tomé, M.; Soares, P.; Diéguez-Aranda, U. A generalized nonlinear mixed-effects height-diameter model for *Eucalyptus globulus* L. in northwestern Spain. *For. Ecol. Manag.* 2010, 259, 943–952.
- 49. Subedi, N.; Sharma, M. Individual-tree diameter growth models for black spruce and jack pine plantations in northern Ontario. *For. Ecol. Manag.* **2011**, *261*, 2140–2148.
- 50. Lhotka, J.M.; Loewenstein, E.F. An individual-tree diameter growth model for managed uneven-aged oak-shortleaf pine stands in the Ozark Highlands of Missouri, USA. *For. Ecol. Manag.* **2011**, *261*, 770–778.
- 51. Weiskittel, A.R.; Garber, S.M.; Johnson, G.P.; Maguire, D.A.; Monserud, R.A. Annualized diameter and height growth equations for Pacific Northwest plantation-grown Douglas-fir, western hemlock, and red alder. *For. Ecol. Manag.* **2007**, *250*, 266–278.
- 52. Alenius, V.; Hökkä, H.; Salminen, H.; Jutras, S. Evaluating estimation methods for logistic regression in modelling individual-tree mortality. *Model. For. Syst.* **2003**, 225–236.
- 53. Yang, Y.; Titus, S.J.; Huang, S. Modeling individual tree mortality for white spruce in Alberta. *Ecol. Model.* **2003**, *163*, 209–222.
- 54. Mehmet, M.; Nuray, M.; Hakki, Y. Modeling individual tree mortality for Crimean pine plantations. *J. Environ. Biol.* **2007**, *28*, 167–172.
- 55. Weiner, J. Asymmetric competition in plant populations. *Trends Ecol. E* 1990, *5*, 360–364.
- Yao, X.; Titus, S.J.; MacDonald, S.E. A generalized logistic model of individual tree mortality for aspen, white spruce, and Lodgepole pine in Alberta mixedwood forests. *Can. J. For. Res.* 2001, *31*, 283–291.
- 57. Bravo-Oviedo, A.; Sterba, H.; del Río, M.; Bravo, F. Competition-induced mortality for Mediterranean *Pinus pinaster* Ait. and *P. sylvestris* L. *For. Ecol. Manag.* **2006**, *222*, 88–98.
- 58. Skovsgaard, J.P.; Vanclay, J.K. Forest site productivity: A review of the evolution of dendrometric concepts for even-aged stands. *Forestry* **2008**, *81*, 13–31.

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