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Spatial Effects Analysis on Individual-Tree Aboveground Biomass in a Tropical *Pinus kesiya* var. *langbianensis* Natural Forest in Yunnan, Southwestern China

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Abstract: It is essential to analyze the spatial autocorrelation and heterogeneity of aboveground biomass (AGB). But it is difficult to accurately describe due to the lack of data in clear-cutting plots. Thus, measuring the AGB directly in a clear-cutting plot can provide a reference for accurately describing the spatial variation. Therefore, a 0.3-hectare clear-cutting sample plot of *Pinus kesiya* var. *langbianensis* natural forest was selected, and the AGB was calculated by each component. The intra-group variance was quantitatively described in terms of spatial heterogeneity, and the spatial autocorrelation was explored by global and local Moran's *I*. The results indicated that (1) there was different spatial heterogeneity for the different trees and organs. The intra-group variance tended to be stable after 20 m for *P. kesiya* var. *langbianensis* (PK) and other upper trees (UPs) and after 10 m for the other lower trees (LTs). (2) The spatial autocorrelation of AGB and wood biomass was similar, while the bark biomass and foliage biomass were consistent. PK and other UPs also exhibited strong spatial autocorrelation, with maximum Moran's *I* values of 0.1537 and 0.1644, respectively. (3) There was spatial heterogeneity in the different components except for the bark of PK. The lowest spatial heterogeneity was found for LT.

Keywords: aboveground biomass; individual tree; spatial structure; spatial heterogeneity; spatial autocorrelation; *Pinus kesiya* var. *langbianensis*

1. Introduction

Many factors affect forest growth, such as competition with other trees, tree distribution, forest disturbance, etc. [1–3]. The interaction among these trees produces spatial effects. Forestry data generally have spatial effects, making them relevant and independent. Without considering spatial effects, it may lead to deviations or errors in the analysis of biomass change rules and estimation [4–6]. Generally, the spatial effects are usually described in terms of spatial autocorrelation and spatial heterogeneity [7,8].

Spatial autocorrelation analysis is one of the most commonly used methods for spatial analysis, such as spatial structure, clustering test analysis, recognition of spatial distribution characteristics, and selection of appropriate spatial scales [9]. Generally, it can be divided into global and local spatial autocorrelation [10]. Global spatial autocorrelation mainly describes the average level of spatial objects over the whole research area. Moran index [11], Geary's C index [12], or Getis'G statistics [13] are commonly used to detect global spatial autocorrelation [8,14,15]. Local spatial autocorrelation can test whether there is similar or different observation aggregation in the local area, then determine the specific location of the hot spot, cold point, and outlier. The local indicators of spatial association (LISA), Getis'G



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Copyright: © 2023 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (https:// creativecommons.org/licenses/by/ 4.0/). statistics, and Moran scatter plots are the main indices for local spatial autocorrelation analysis [16,17]. These have been widely applied to improve the understanding of forest distribution and estimate the accuracy of forest information in forestry [18–24].

Spatial heterogeneity refers to the presence of structural instability characterized by systematic variations in model parameters or different response functions within a spatial dataset [4,7]. Neglecting spatial heterogeneity can result in erroneous significance tests and suboptimal predictions [16]. The importance of spatial heterogeneity lies in its provision of crucial information and insights in various fields, such as ecology, resource management, disaster risk assessment, and urban planning. It aids in better understanding and addressing spatial changes and patterns. For instance, effective forest resource management requires considering the spatial heterogeneity of biomass and growth rates to develop appropriate forestry management plans [25]. For forestry data, it represents spatial complexity, variation of ecosystem characteristics, and spatial non-homogeneity or the differences caused by the non-random distribution of forest vegetation [26-33]. Fully understanding and utilizing spatial heterogeneity will help us better understand the forest vegetation growth process and even the forest ecosystem's evolution [34–37]. For different spatial data, the methods to evaluate and measure spatial heterogeneity include a random index, aggregated index, the nearest distance, trend surface, spectrum analysis, variation function, fractal component, analysis of variance ratio, correlation analysis, etc. [38–44]. Moreover, spatial autocorrelation and spatial regression models could also be used to analyze spatial heterogeneity [11–13,16]. Furthermore, spatial regressions, particularly the geographically weighted regression model (GWR) model, are still the most common method to reflect spatial heterogeneity [7,45–48].

Biomass is an essential attribute of forests, and there are some studies on spatial effects analysis and the spatial regression model construction of forest biomass at different scales. The application of remote sensing, geographic information systems, and global position system technologies make the spatial location information of data relatively easy to obtain and the spatial effect analysis at the regional level [49]. At the stand scale, spatial regression models, considering spatial effects, can also play a role in fitting related variables at the stand scale [50]. Due to the spatial effects of forest data, forests are influenced by their interactions with the surrounding environment during their growth process. Therefore, it is assumed that spatial regression models can meet the requirements of data independence [51]. The AGB of a forest was estimated using an ordinary least squares model (OLS), spatial lag model (SLM), spatial error model (SEM), and GWR model, and the GWR model has good performance and can solve the problems of high underestimation and low overestimation to some extent [52,53]. At the individual tree scale, Ou et al. [54] constructed the individual tree biomass estimation model of Pinus kesiya var. langbianensis by geographic weighting regression, and the fitting accuracy was significantly higher than that of the ordinary regression model. Some researchers also constructed the mixed-effect model of individual tree biomass with higher fitting and prediction accuracy by using the mixed-effect model technology. These mixed-effect models do not consider the location information of their data but mostly use some functions to represent spatial autocorrelation in the selection of variance and covariance matrix within the group [5,55]. It can be found that spatial effects are common in forest biomass data. Considering the spatial effects of individual tree biomass, it can more accurately describe the change law of biomass. Additionally, it can improve the accuracy of forest biomass estimation to a certain extent.

Regarding the spatial effect analysis of the growth of individual trees, most studies focus on the spatial effect analysis of individual tree diameters and the relationship between the diameter and tree height. For example, Zhang et al. [5] analyzed the multivariate relationship of spatial heterogeneity between tree growth and diameter; Lu and Zhang [56–58] used a spatial regression model to fit tree height and diameter growth under the influence of spatial effect. Gu and Zhao [59] established a GWR model for forest growth and achieved high fitting accuracy. It can be seen that the individual tree diameter at breast height (DBH) and tree height (H) has been confirmed to have spatial effects, and most of these studies focus on the changes in diameter and tree height related to volume [60–62]. Moreover, the forest structural heterogeneity and spatial patterns of trees associated with age structure have been reported previously [63,64], but spatial effects on individual tree biomass have rarely been reported.

It can be seen that the spatial effects of biomass and carbon storage at a regional scale and the spatial effects of DBH and the volume of trees at the individual or stand levels have been analyzed. Previous studies have focused on the change in diameter and height related to volume, which can reflect the corresponding wood or wood biomass to a certain extent. However, a tree comprises wood, bark, branches, foliage, roots, and other components. Pearce et al. [65] used a linear mixed-effect model and site as a random effect to estimate the biomass of a New Zealand shrubbery firewood forest. Chai [66] proposed a biomass ratio to describe the spatial structure of typical secondary forests in the mid-mountain zone of the Qinling Mountains based on the relationship between adjacent trees. However, the biomass data used in the quantitative evaluation were not measured but calculated using the allometric growth equation of biomass of predecessors. Unfortunately, it is difficult to obtain accurate measurement data for forest biomass, and most of the biomass data investigated in the past do not record spatial information. Therefore, the variation in spatial effects of each component regarding individual tree biomass within a stand and the differences in spatial effects of individual tree biomass among different stand types has rarely been reported. Therefore, the differences in the spatial effects among the components need to be further studied.

Then, the spatial heterogeneity and autocorrelation were analyzed at an individual tree scale. The significant contributions of this work are:

(1) To describe the spatial heterogeneity and autocorrelation of AGB at an individual tree scale using a clear-cutting plot;

(2) To explain the differences in spatial effects among the different components of the AGB;

(3) To explain the differences in spatial effects among the different tree species in a tropical natural forest.

2. Materials and Methods

2.1. Study Site

The study area is located in the town of Yutang in Mojiang County, which belongs to Pu'er City, Yunnan Province, in southwest China (Figure 1). *Pinus kesiya* var. *langbianensis* forest is one of the area's main forest types [67]. Mojiang County is located between 101°08′–102°04′ E and 22°51′–23°59′ N. The lowest altitude is 440 meters, and the highest altitude is 2278 meters. It has a typical subtropical monsoon climate, and the seasons here are not so obvious, but there is a distinct rainy and dry season. In addition, there was abundant rainfall, with an average annual rainfall of 1338.0 mm. The rainy season happens from May to October. According to the meteorological data (1980–2010) of Mojiang station, the mean annual temperature was about 17.8 °C. The coldest month was January, with an average temperature of 11.7 °C, and the hottest month was June, with an average temperature of 21.9 °C. Soil type is mainly red soil or lateritic red soil [68].

Pinus kesiya var. *langbianensis* is a tree species of the *Pinus* genus belonging to the Pinaceae family, and it is a geographic variation of *Pinus kesiya* [68,69]. The *Pinus kesiya* var. *langbianensis* forests, one of the specific types in Yunnan Province, were mainly distributed in the southern subtropical zone of the west Ailao mountains in Yunnan Province. Its stand volume accounted for about 11% of the forest land of Yunnan Province. Moreover, it is also distributed in central and northern Vietnam and Laos, etc. [70]. Rapid growth is a critical feature because its trunk and annual branches grow two or more rounds per year compared to other coniferous species. Therefore, it has been an important afforestation species in Yunnan [67].



Figure 1. Location of the clear-cutting plot. (**a**,**b**) shows the location of the plot in Mujiang county for Pu'er City, Yunnan Province, and (**c**) is a satellite image of the study site (Red frame); (**d**) shows the numbering of trees in the sample plot before logging, and (**e**–**g**) the trees that were logged.

2.2. Data Investigation

A typical sample plot of 0.3 hm^2 (30 m × 100 m) was investigated in April and May 2015. Co-ordinate information on the location, elevation, slope degree, and slope aspect of the sample plot was recorded, and the center co-ordinates are $23^{\circ}09'58.2''$ N, $101^{\circ}29'14.4''$ E (Table 1). In the plot, the *Pinus kesiya* var. *langbianensis* natural forest is a natural mature mixed forest. The upper layer is dominated by *Pinus kesiya* var. *langbianensis* and some hard broadleaf species, such as *Castanopsis hystrix*, *C. indica*, and *Cyclobalanopsis* spp., and other soft broadleaf species, such as *Schima wallichii* and *Alnus nepalensis*. The lower layer consists of *Wendlandia* sp., *Phyllanthus emblica*, *Anneslea fragrans*, etc. The forest floor was mostly covered with fallen foliage and branches.

Table 1. The basic characteristics of the clear-cutting plot.

Location	Latitude	Longitude	Altitude (m)	Slope Degree (°)	Aspect of Slope (°)
Yutang	23°9′58.2″ N	101°29′14.4″ E	1530	22	298

2.3. Biomass Measure and Calculate

We numbered each tree starting from the southwest corner of the sample plot and recorded the position of all trees; then, all trees were felled. Then, the DBH of trees was measured using a diameter tape measure. Meanwhile, we also measured the height of the fallen sample trees. The location, height, and DBH of each tree were recorded for all live trees with DBH > 3 cm and height \geq 2 m (Figure 2). All trees could be divided into three groups, *Pinus kesiya* var. *langbianensis* (PK), other trees at the upper layer (UP), and the trees at the lower layer (LT). There were 132 PK, 119 UP, and 261 LT. UP includes 17 *Castanopsis hystrix*, 17 *C. indica*, 16 *Cyclobalanopsis* spp., 54 *Schima wallichii*, 12 *Alnus nepalensis*, and three other trees. In the LT, there are 93 *Wendlandia* sp., 128 *Phyllanthus emblica*, 32 *Anneslea fragrans*, and eight others species.



Figure 2. Location of the trees in the clear-cutting plot. PK is *Pinus kesiya* var. *langbianensis*, UP is the other trees at the upper layer, and LT is the other trees at the lower layer. The circle diameter is proportional to the tree DBH.

Biomass data for tree components were obtained using the weighing method and the bulk density method, including wood, bark, branches, foliage. For wood and bark, the discs of individual trees with different heights and bark thicknesses were intercepted and weighed for fresh weight to obtain moisture content and wood density. For foliage and branches, the branches and foliage were weighed in the field and in the sample to assess dry weight [67]. All samples were dried in the laboratory at 105 °C to a fixed mass. The density of the disc and the corresponding volume of each stem segment were used to calculate the dry weight of the wood and bark using the proportion of the total fresh mass and the moisture content of the sample to determine the dry weight of the branches and foliage [71,72]. The information is listed in Table 2.

Trees or Trees Group		A an (Vaara)	Hoight (m)	DBH (am)	Biomass (kg)				
		Age (Teals)	fieight (iii)	DBH (cm)	Wood	Bark	Branches	Foliage	Aboveground
Pinus kesiya var.	Mean	37	17.05	24.26	236.44	6.78	37.24	4.45	284.91
	Std. Err.	10	4.58	9.93	224.73	5.33	47.83	4.74	268.62
(n - 122)	Min.	16	6.80	7.00	3.29	0.14	0.51	0.06	4.07
(n = 132)	Max.	66	25.60	45.10	1075.79	34.49	307.23	23.39	1180.42
	Mean	25	9.97	14.76	55.24	15.16	13.62	3.45	87.47
Other upper trees ($n = 119$)	Std. Err.	10	4.21	7.76	72.61	19.09	17.78	4.97	102.08
	Min.	5	3.30	3.40	0.82	0.29	0.07	0.02	1.74
	Max.	48	21.50	36.00	368.49	88.85	103.17	31.70	482.72
	Mean	16	6.36	6.73	5.97	1.32	1.77	0.54	9.60
Other lower trees $(n = 261)$	Std. Err.	5	1.66	2.62	6.81	1.55	2.74	0.89	11.09
	Min.	5	2.20	4.00	0.83	0.07	0.03	0.00	1.22
	Max.	35	12.00	21.00	44.20	10.57	21.04	7.67	71.80
Total (n = 512)	Mean	24	9.95	13.12	76.84	5.95	13.67	2.22	98.68
	Std. Err.	12	5.53	9.81	153.31	11.14	29.71	3.88	184.67
	Min.	5	2.20	3.40	0.82	0.07	0.03	0.00	1.22
	Max.	66	25.60	45.10	1075.79	88.85	307.23	31.70	1180.42

Table 2. The basic information of the sampling trees in the clear-cutting plot.

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2.4. Spatial Effects Analysis

2.4.1. Spatial Heterogeneity of the Aboveground Biomass

The intra-group variance can be used to describe the spatial heterogeneity of study subjects quantitatively. This is a function of block size and usually increases as the size of the blocks increases [73], and it can be calculated by Equation (1).

$$S_{\text{intra}} = \frac{1}{B} \sum_{g=1}^{B} \frac{1}{n_g} \sum_{h=1}^{n_g} \left(X_{gh} - \overline{X}_g \right)^2 \tag{1}$$

where *B* is the number of groups at a particular grouping distance; X_{gh} is the *h*-th obser-

vation biomass value in the *g*-th block; X_g is the mean of observation value in the *g*-th block.

2.4.2. Spatial Autocorrelation of the Aboveground Biomass

Moran's *I* effectively reflect the differences and correlations in the spatial distribution of observations. It can also reflect the overall spatial clustering pattern of objects in the study area [11]. The value range of Moran's *I* is between (-1, 1). The closer the absolute value of the index value is to 1, the greater the degree of aggregation. When Moran's *I* is positive, the observations within the distance (h) tend to be similar.

On the contrary, they tend to be dissimilar when the coefficient is negative. If Moran's *I* index is close to or equal to 0, the observations are spatially random and independent. So, Moran's *I* was chosen to describe the spatial autocorrelation for each component biomass (Equation (2)).

$$I = \left(\frac{n}{\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}(d)}\right) \frac{\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}(d) (x_i - \overline{x}) (x_j - \overline{x})}{\sum_{i=1}^{n} (x_i - \overline{x})^2}$$
(2)

where *n* is the number of plots; x_i and x_j are the observation value at plot *i* and plot *j*, \overline{x} is the average of all attribute values; $w_{ij}(d)$ is the spatial weight matrix value that represents the relationship between plot *i* and plot *j* of the spatial object.

After the Moran index value is calculated, the significance of the index is tested by a Z-score. The higher (or lower) the Z-score, the higher the degree of clustering. A positive Z-score indicates a cluster of high values, and a negative Z-score indicates clusters of low values. If the Z-score is close to zero, no significant clustering exists within this area, and the Z-score was calculated by Equation (3).

$$Z(I) = \frac{I - E(I)}{\sqrt{Var(I)}}, E(I) = \frac{-1}{n - 1}$$
(3)

where Z(I) is an index to measure the strength of spatial clustering patterns; E(I) and Var(I) are the expected value and variance of the index value *I*, respectively;

Moreover, local spatial autocorrelation can be used to describe the degree of spatial correlation between a research object and its adjacent units. In this study, the distance corresponding to the first peak after each component's spatial clustering pattern for biomass reached significance was used as the bandwidth, and ArcGIS was used to calculate local Moran's I_i . Furthermore, the local Moran's I_i was calculated by Equation (4):

$$I_{i} = \frac{n^{2}}{\sum_{i}^{n} \sum_{j}^{n} w_{ij}} \frac{(x_{i} - \overline{x}) \sum_{j}^{n} w_{ij}(x_{j} - \overline{x})}{\sum_{i}^{n} (x_{i} - \overline{x})^{2}}$$
(4)

where *n* is the number of plots; x_i and x_j are the observation value at plot *i* and plot *j*, *x* is the average of all attribute values; $w_{ij}(d)$ is the spatial weight matrix value. Different from the global Moran's *I*, the value range of the local Moran's *I_i* is not limited to (-1, 1); if the *I_i* value is positive, it shows that the location of the plot is a positive correlation and reflects the aggregation of similar values. But the *I_i* is negative, and it shows that the location of the plot is a negative correlation and reflects the aggregation of different values.

The global Moran's *I* and *Z* (*I*) values were calculated using the incremental spatial autocorrelation tool in ArcGIS. The most significant distance of the spatial clustering pattern obtained from the incremental spatial autocorrelation analysis was used as the scale parameter for local spatial autocorrelation analysis. On this basis, using ArcGIS clustering and outlier analysis tools, the spatial weight matrix was constructed in the form of fixed distance bandwidth, and the local Moran's I_i was calculated to quantitatively describe the spatial correlation between local, regional observations.

3. Results

3.1. Spatial Heterogeneity of the Aboveground Biomass

The spatial heterogeneity variation in biomass of each component showed the same trend in the range of 0–100 m (Figure 3). Aboveground, wood, branches, bark, and foliage were the order of the intra-group variance values for the different components. Although the value changed with the different organs, the spatial variation gradually increased with increasing distance; after reaching a certain distance, the variation tended to be stable. The variance values tended to be stable after 20 m for both PK and UP. In comparison, the variation of LT and all tree species tended to be stable after 10 m. Furthermore, the spatial variance values of PK were significantly larger than the other tree types for the AGB, wood, and branches. The spatial variance values of UP were the largest for both bark and foliage, and all the biomass components of the LT showed the smallest spatial variance values.



Figure 3. The intra-group variance of the different biomass components for different tree groups. (a): aboveground biomass, (b): wood biomass, (c): bark biomass, (d): foliage biomass, (e): branches biomass. PK is *Pinus kesiya* var. *langbianensis*, UP is the other trees at the upper layer, and LT is the other trees at the lower layer.

3.2. Spatial Autocorrelation of the Aboveground Biomass

3.2.1. Global Moran's I

The global Moran's *I* of each component biomass at the different lag distances were calculated, and the Moran's *I* correlogram and Z-score values of the biomass with lag distances are shown, respectively. For all trees (Figure 4), the distances at the first peak (DFP) with a significant Z-score larger than 1.96 were 21.2 m for the AGB and wood biomass. The distance value was 6.2 m, and the biomass for both the foliage and branch was 4.8 m. Moreover, at the corresponding lag distance, the spatial autocorrelation was significant because the Z-scores were larger than 1.96. Meanwhile, the Moran's *I* values of AGB and wood were lower than 0.01, and the values of foliage, bark, and branches were larger than 0.034.



Figure 4. Moran's *I* correlogram and *Z*-score values for the biomass of the different components for all trees. (**a**): aboveground biomass, (**b**): wood biomass, (**c**): bark biomass, (**d**): foliage biomass, and (**e**): branches biomass. The dots and rectangles represent the points of Moran's *I* and *Z*-score at the first peak (DFP) with a significant *Z*-score, and the values were shown in the figure, respectively.

For the different tree types, the PK had a similar change law for all trees; the DFP was 12.4 m for both AGB and wood biomass, 10.0 m for bark biomass, and 7.4 m for both foliage and branches. Furthermore, the Moran's *I* was larger than 0.13 for foliage and branches; the values for the AGB, wood, and bark were 0.0687, 0.0609, and 0.0389, respectively (Figure 5). For the UP, the spatial autocorrelation of both foliage and branches was not significant because the Z score was not significant along with the lag distance, and the absolute Z scores were not larger than 1.96. Meanwhile, the DFP of AGB, wood, and bark were 6.6 m, 6.6 m, and 9.0 m, respectively. The corresponding Moran's *I* values were 0.1277, 0.1372, and 0.1644 (Figure 6). Moreover, for LT, the branch biomass had no significant spatial autocorrelation, and the Moran's *I* of both AGB and wood were 0.0520 and 0.0617 at a DFP of 6.4 m. The bark biomass had the largest DFP at 17.5 m, and the Moran's *I* was 0.0233. Meanwhile, the foliage biomass had the lowest DFP at 6.0 m, and the Moran's *I* was 0.0771 (Figure 7).



Figure 5. Moran's *I* correlogram and *Z*-score values for the biomass of the different components for *Pinus kesiya* var. *langbianensis*. (a): aboveground biomass, (b): wood biomass, (c): bark biomass, (d): foliage biomass, and (e): branches biomass. The dots and rectangles represent the points of Moran's *I* and *Z*-score at the first peak (DFP) with a significant *Z*-score, and the values were shown in the figure, respectively.

Moreover, we can find that the AGB and wood biomass had the same DFP for each tree type for all trees. Additionally, they had the highest DFP for all trees and PK but the lowest UP, and most of the foliage biomass had a lower DFP, with the largest Moran's *I* for each tree type and all trees. Furthermore, a large portion of the PK and UP had the highest Moran's *I* when compared to that of LT.



Figure 6. Moran's *I* correlogram and *Z*-score values for the biomass of the different components of other upper trees. (**a**): aboveground biomass, (**b**): wood biomass, (**c**): bark biomass, (**d**): foliage biomass, and (**e**): branches biomass. The dots and rectangles represent the points of Moran's *I* and *Z*-score at the first peak (DFP) with a significant *Z*-score, and the values were shown in the figure, respectively.



Figure 7. Moran's *I* correlogram and *Z*-score values for the biomass of the different components of other lower trees. (**a**): aboveground biomass, (**b**): wood biomass, (**c**): bark biomass, (**d**): foliage biomass, and (**e**): branches biomass. The dots and rectangles represent the points of Moran's *I* and *Z*-score at the first peak (DFP) with a significant *Z*-score, and the values were shown in the figure, respectively.

3.2.2. Local Moran's I_i

The counting statistics of spatial autocorrelation were made and listed in Table 3. The PK had a lower percentage of the point with no significant spatial autocorrelation (NS) when compared with UP and LT. Especially for the LT, the percentage of the NS point was greater than 98%, and all points were the NS for the wood biomass and AGB. But for the PK, the percentage with NS for the wood biomass and AGB were the lowest, and the percentage was 77.27% and 78.03%, respectively. Moreover, for the point with significant spatial autocorrelation, there was no similar aggregation point for low-low clustering (LL) for all trees; there were similar aggregation points for high-high clustering (HH), and a dissimilarity aggregation point for high-low outliers (HLs) were not found in the LT. The point with significant spatial autocorrelation in the plot was primarily LH outliers and

HH clustering, and wood and AGB showed similar local spatial autocorrelation. HH points were mainly found in the wood, foliage, and AGB of PK, and bark biomass of UP. Meanwhile, HL points were mainly found in the wood and AGB of PK. For the dissimilarity aggregation points with a low-high outlier (LH), the percentage was lower; all UP was not found at the point with LH, and only 1 point for branches biomass for PK was found for LH.

Table 3. Spatial distribution by Z-score using all trees. PK is *Pinus kesiya* var. *langbianensis*, UP is the other trees at the upper layer, and LT represents the other trees at the lower layer. NS is the point with no significant spatial autocorrelation, HH is the similar aggregation point with high-high clustering, LL is the similar aggregation point with low-low clustering, HL is the dissimilarity aggregation point with a high-low outlier, and LH is the dissimilarity aggregation point with a low-high outlier. N is the number of degrees of distribution for each component, and P is the proportion of distribution of each component.

Tree Tunes	Components –	NS		H	HH		HL		LH		LL	
free Types		Ν	P (%)	Ν	P (%)	Ν	P (%)	Ν	P (%)	Ν	P (%)	
	Wood	102	77.27	11	8.33	19	14.39	0	0.00	0	0.00	
	Bark	128	96.97	4	3.03	0	0.00	0	0.00	0	0.00	
PK	Branches	118	89.39	8	6.06	5	3.79	1	0.76	0	0.00	
	Foliage	116	87.88	15	11.36	1	0.76	0	0.00	0	0.00	
	Aboveground	103	78.03	12	9.09	17	12.88	0	0.00	0	0.00	
	Wood	117	98.32	2	1.68	0	0.00	0	0.00	0	0.00	
	Bark	99	83.19	15	12.61	5	4.20	0	0.00	0	0.00	
UP	Branches	119	100.00	0	0.00	0	0.00	0	0.00	0	0.00	
	Foliage	115	96.64	1	0.84	3	2.52	0	0.00	0	0.00	
	Aboveground	117	98.32	2	1.68	0	0.00	0	0.00	0	0.00	
	Wood	261	100.00	0	0.00	0	0.00	0	0.00	0	0.00	
	Bark	258	98.85	0	0.00	0	0.00	3	1.15	0	0.00	
LT	Branches	259	99.23	0	0.00	0	0.00	2	0.77	0	0.00	
	Foliage	260	99.62	0	0.00	0	0.00	1	0.38	0	0.00	
	Aboveground	261	100.00	0	0.00	0	0.00	0	0.00	0	0.00	
ALL	Wood	480	93.75	13	2.54	19	3.71	0	0.00	0	0.00	
	Bark	485	94.73	19	3.71	5	0.98	3	0.59	0	0.00	
	Branches	496	96.88	8	1.56	5	0.98	3	0.59	0	0.00	
	Foliage	491	95.90	16	3.13	4	0.78	1	0.20	0	0.00	
	Aboveground	481	93.95	14	2.73	17	3.32	0	0.00	0	0.00	

Moreover, the spatial autocorrelation for each biomass component of the different tree or tree groups was shown in the local Moran's I_i scatters. For the PK, the biomass of AGB, wood, and branches mainly showed significant HH clustering on the right side of the plot, while the spatial clustering pattern of biomass in the middle and left sides showed HL outliers. In addition, the biomass of bark and foliage mainly showed HH clustering within the plots. The biomass of some components was also accompanied by the LH outlier (Figure 8). For the UP, AGB and wood biomass showed HH clustering on the right side of the plot and NS on the left side. Similarly, foliage biomass also showed HH clustering on the right side of the plot, but an HL outlier was also present on the right and left sides. In contrast, bark biomass showed mainly HH clustering in the plot; HL outliers were found on the right center and near the left edge of the plot (Figure 9). But for the LT, the components mainly showed no significant aggregation (NS) and sporadic LH outliers in the plot (Figure 10).



Figure 8. Spatial distribution of local Moran's I_i for *Pinus kesiya* var. *langbianensis* (**a**) for aboveground biomass, (**b**) for wood biomass, (**c**) for bark biomass, (**d**) for foliage biomass, and (**e**) for branches biomass. NS is the point of no significant spatial autocorrelation, HH is the similar aggregation point with high-high clustering, HL is the dissimilarity aggregation point with a high-low outlier, and LH is the dissimilarity aggregation point with a low-high outlier.



Figure 9. Spatial distribution of local Moran's I_i for other upper trees (**a**) for aboveground biomass, (**b**) wood biomass, (**c**) bark biomass, (**d**) foliage biomass, and (**e**) for branches biomass. NS is the point of no significant spatial autocorrelation, HH is the similar aggregation point with a high-high clustering, and HL is the dissimilarity aggregation point with a high-low outlier.



Figure 10. Spatial distribution of local Moran's I_i for lower trees. The capital letters of the figure codes represent the different components (**a**) for aboveground biomass, (**b**) wood biomass, (**c**) bark biomass, (**d**) foliage biomass, and for (**e**) branches biomass. NS is the point of no significant spatial autocorrelation, and LH is the dissimilarity aggregation point with a low-high outlier.

Furthermore, for different tree species and components, we calculated the local Moran's I_i of each tree using all trees. The statistics are listed in Table 4. According to the local autocorrelation analysis of biomass in different tree species and different components, it can be seen that, for the mean local autocorrelation value of wood biomass,

the autocorrelation coefficient of PK showed a negative value, and the value (-0.45) was smaller than other trees (upper trees, lower trees, and total), with a negative autocorrelation. The minimum value (-10.32) and maximum value (8.43) were consistent with the total, and the minimum value was smaller than the other upper and lower trees, while the maximum value was larger than them. However, the standard error (0.23) was the largest among all tree species. For the local autocorrelation analysis of bark biomass, it can be seen that all tree species showed a positive correlation, and the highest mean correlation value (0.82) was found in other upper trees, and the highest value (20.07) was the same as that of the total. For the mean autocorrelation of foliage and branches biomass, except for the negative correlation of other upper trees, all showed a positive correlation, and the correlation of PK was the strongest; for foliage (0.39) and branches (0.17), the maximum value was the same as the total: foliage (12.73) and branches (14.35) and was higher than the other upper and lower trees. The mean, maximum, and minimum values of the local autocorrelation coefficients of AGB exhibited the same trend as wood biomass.

Table 4. Local Moran's I_i for the biomass of the different components.

Tree Species	Indices	Wood	Bark	Foliage	Branches	Aboveground
	Mean	-0.45	0.10	0.39	0.17	-0.42
Dinus kasing	Std. Err.	0.23	0.07	0.20	0.21	0.23
Pinus kesiyu	Skewness	0.01	2.52	2.87	3.25	-0.03
Var. Ianohianancio	Kurtosis	2.72	13.59	11.83	15.04	2.18
ungounensis	Min.	-10.32	-2.06	-3.87	-4.58	-9.67
	Max.	8.43	4.94	12.73	14.35	8.12
	Mean	0.15	0.82	-0.07	-0.04	0.07
	Std. Err.	0.09	0.31	0.11	0.06	0.10
Other upper trees	Skewness	1.78	3.81	0.16	-2.42	0.50
Other upper trees	Kurtosis	12.72	17.68	9.87	10.13	4.56
	Min.	-3.02	-3.71	-6.06	-3.76	-3.06
	Max.	6.49	20.07	6.14	1.26	5.40
	Mean	0.26	0.08	0.09	0.11	0.28
	Std. Err.	0.05	0.05	0.05	0.04	0.05
	Skewness	-0.72	-2.05	-2.59	-4.19	-0.86
Other lower trees	Kurtosis	0.18	7.46	11.12	28.11	0.79
	Min.	-2.60	-4.64	-4.23	-5.00	-3.24
	Max.	1.53	1.25	1.44	1.00	1.70
	Mean	0.05	0.26	0.13	0.09	0.05
	Std. Err.	0.07	0.08	0.06	0.06	0.07
T (1	Skewness	-0.56	6.64	3.15	4.67	-0.63
Iotal	Kurtosis	9.76	64.72	26.56	46.34	7.90
	Min.	-10.32	-4.64	-6.06	-5.00	-9.67
	Max.	8.43	20.07	12.73	14.35	8.12

4. Discussion

4.1. Spatial Heterogeneity and Autocorrelation in the Tropical Natural Forest

The forest community has high spatial heterogeneity [74]. Under heterogeneous conditions, community differentiation led to self-thinning, weakened inter-specific competition, and enhanced individual stress resistance, which was beneficial for the stability of various groups in the community and the overall development of the community. In this study, in the range of 0–100 m, the change in spatial heterogeneity tended to be stable after a distance of 20 m. The spatial distribution of wood, bark, branches, foliage, and AGB in the natural forests of *Pinus kesiya* var. *langbianensiss* was not random but was positively correlated and showed some spatial aggregation. The semi-variance model structure could indicate the spatial heterogeneity of AGB [75]. As the distance increases, the correlation between the spatial variables and their surrounding values decreases [62], and this pattern was similar to the biomass in the broadleaved forests [76].

Moreover, the high biomass values were mainly concentrated on the right side of the sample plot, and the low values were mainly concentrated on the left and middle of the sample plot. The spatial clustering characteristics objectively existed in all biomass components [77]. Local spatial correlations were more robust for aboveground, wood, and bark biomass and mostly showed high-high aggregation and high-low outliers, with the lowest correlations for branch biomass. In addition, the findings indicated that the local spatial autocorrelation of the biomass components of lower tree species was more significant compared to that of other tree species. It could cause spatial patches within the stand by the specific dynamic processes that vary by spatial correlations among the tree species and local disturbances [64] and lead to spatial heterogeneity in the stand structure as well as within and among the tree species [78,79]. Therefore, it may be responsible for the spatial effects on individual tree biomass. Based on this, it can guide the study of the spatial structure, spatial heterogeneity, spatial autocorrelation, and biomass of *Pinus kesiya* var. langbianensiss natural forest in large regions. Rödig et al. [80] demonstrated how this linkage allowed for quantifying the spatial variation in forest structure caused by tree-level to regional-scale disturbances. In addition, the study of heterogeneity can also combine with remote sensing methods. Han et al. [81] estimated and explored the aboveground carbon storage (AGC) of bamboo forests by using Landsat images and five dates measured in the field. The results showed that the AGC of the moso bamboo increased gradually in the five periods, and its spatial heterogeneity had moderate autocorrelation.

The results of this study show that the AGB and the biomass of each component have spatial heterogeneity and spatial autocorrelation, which can be estimated by considering spatial regression models, such as GWR, in future studies. The GWR model has performed well in estimating biomass in many studies [51–53,61]. However, when estimating biomass or other tree characteristics, the applicability of the model needs to be fully considered before it can be calculated. The use of the SEM model enables a direct explanation of spatial autocorrelation. Meanwhile, the incorporation of spatial dependencies and variability in the modeling process through LMM (linear mixed model) and GWR techniques allows for improved data fitting and a more precise prediction of response variables [61].

4.2. Spatial Effect of the Different Components

The following results were obtained from the study of spatial heterogeneity regarding the biomass of different components. The components, which were ranked in terms of intragroup variance values, were aboveground, wood, branches, bark, and foliage, respectively. There was an allometric relationship for plant growth, and the plants preferentially allocated biomass to organ harvesting according to the biomass allocation theory [82]. Moreover, biomass allocation may vary with stand age, species competition, etc. [83], and therefore affect the spatial effect in the different components of biomass. Competition may have an especially meaningful impact on the biomass allocation of trees. Plant competition and herbivory interactions were crucial for reducing plant production performance and changing biomass allocation patterns [84]. The biomass allocation would prefer tree crowns with little light and tree roots with limited nutrients or water [82]. Under different light environments, there was no difference in the biomass allocation of stems and roots and leaf mass per area between invasive and native species [85]. In this study, the AGB had a much more significant spatial variation in the whole stand. This was consistent with the results of Du et al. [75] research on spatial heterogeneity and the carbon contribution of AGB of moso bamboo in Anji County, Zhejiang Province; in addition, their research also indicated that spatial autocorrelation had a determinative influence on the spatial heterogeneity and spatial distribution pattern of moso bamboo's AGC. Therefore, the study of biomass cannot ignore the impact of spatial heterogeneity and spatial autocorrelation.

In addition, it is important to carry out a spatial autocorrelation analysis of the biomass of each component. In this study, global and local spatial autocorrelation analyses were carried out for the different components of different tree species so that the spatial autocorrelation between trees in the whole stand could be more easily seen. The global spatial autocorrelation trends for each component of *Pinus kesiya* var. *langbianensis* were consistent with those of all trees in the sample plot, while the trends for other upper trees and other lower trees differed from those of all trees in the sample plot. Zhou et al. [83] conducted biomass modeling of the moso bamboo foliage as a way to accurately estimate the dynamics of foliage biomass, thus contributing to the development of accurate simulations. The spatial autocorrelation was used by Wu et al. [86] to conduct a spatio-temporal analysis of forest stock in Sichuan Province, and a mapping exercise was conducted for global spatial autocorrelation and local spatial autocorrelation.

The local spatial autocorrelation of biomass is also essential, as the distribution of habitat factors, such as light, water, and soil, are heterogeneous in natural environments [87]. Stand volume and biomass are closely related, as habitat variation can lead to the territorial and specific structural characteristics of forest volume [88]. The local spatial autocorrelation analysis revealed distinct patterns of aggregation in the sample plots for the different components studied. Specifically, the HH, HL, and LH aggregations exhibited varying location and aggregation patterns. Stand volume was a stochastic function that related to random variables and location [86]; therefore, forest biomass estimation needs to take full account of the spatial location and distribution of trees. Spatial autocorrelation to interpret and aid model outputs can identify areas that are systematically sensitive to specific model inputs and can also guide the collection of additional data to refine model predictions [89].

4.3. Spatial Effect of the Different Tree Species

The spatial variance values of PK were significantly larger than the other tree types for the aboveground, wood, and branches biomass. The spatial variance values of UP were the largest for both bark and foliage, and the biomass components of the LT showed the smallest spatial variance values. Then, significant or highly significant spatial autocorrelations were found for the biomass of all the components except for the branches of LT and the foliage and branches for UP. The results showed that there was similar aggregation (high-high clustering or low-low clustering) or dissimilarity aggregation (high-low outlier or low-high outlier) regarding biomass in the natural stand of *Pinus kesiya* var. *langbianensis* but the correlation was not strong. Therefore, the spatial autocorrelation variation patterns of AGB and wood biomass were similar, and the patterns of foliage and branches biomass were comparatively similar (Figures 4–7).

The species characteristics and specific dynamic process make the biomass allocation among the tree species different [64]; this then leads to the forest structure and spatial heterogeneity of different components and various tree species [78,79], which may be the reasons that affect the spatial effect of individual tree biomass. In this study, the stand was dominated by the masculine species of *Pinus kesiya* var. *Langbianensis*. *Pinus kesiya* var. *Langbianensis* is a typical fast-growing heliophilous species; the spatial autocorrelation of *Pinus kesiya* var. *langbianensis* natural forest was dominated by positive spatial autocorrelation for each component's biomass, and there was significant positive spatial autocorrelation. Therefore, the spatial correlation was more prominent in *Pinus kesiya* var. *langbianensis* in terms of different tree species.

Moreover, the local spatial autocorrelation of each biomass component for the lower tree species was stronger than that of other tree species. The foliage biomass of the lower trees and all trees, the bark biomass of other upper trees, and the branches biomass of *Pinus kesiya* var. *langbianensis* were strongly correlated. However, it was also found that the biomass of some components had insignificant spatial autocorrelation, such as the foliage biomass of the upper tree species and the foliage biomass of the lower tree species. Subsequently, competition may be one of the main reasons.

The competition strongly affected tree productivity and reproductive capacity [84,85]. Light is one of the most critical factors controlling tree survival and growth, and the canopy formed by the upper layer of trees can limit the use of light resources by the species in the lower layer [90]. When compared to the dominant and upper-layer species, the lower-layer species had the most negligible spatial variation in the biomass of each component. The

upper tree species allocated more of the obtained resources to the physiological activities of the foliage [47,61], coupled with the photosynthetic capacity of the bark being more sensitive to temperature [91]. The foliage and bark in the upper tree species got more light compared to the other types, and competition was more pronounced.

4.4. The Limits of the Study

The clear-cutting sample plot of *Pinus kesiya* var. *langbianensis* natural forest is 0.3 hectares; this study carried out spatial effect analysis on the data by calculating the variance of each component biomass and Moran's index. It concluded that there was spatial heterogeneity in the biomass of an individual tree within a certain distance. This study explored spatial autocorrelation and spatial heterogeneity among individual trees. The study site is located in the main distribution area of *Pinus kesiya* var. *langbianensis* in Yunnan Province, namely the semi-humid areas of the northern tropical and southern subtropical regions of Yunnan Province. The stand structure is basically the same, with relatively small differences in site conditions and stand age [72]. Therefore, the sample plots were representative of the specific characteristics of the region. However, the number of sample plots could be increased in the next study; thus, spatial autocorrelation and spatial heterogeneity at the stand scale could be more easily investigated. Du et al. [75] investigated 55 sample plots with a sample size of 30 m \times 30 m to study the spatial heterogeneity and spatial patterns of the above-ground biomass of moso bamboo. Meanwhile, the differences in physiological, biochemical characteristics, and stand structure between heliophilous species and shadetolerant trees and the spatial effects of forest biomass dominated by other shade-tolerant species must be further studied.

Moreover, little research about the quantitative analysis of the spatial effects of individual tree AGB at a plot scale has been conducted because the measurement of biomass data is destructive and labor-intensive [80]. In this study, we discussed the reasons for the spatial effect of the natural forest of *Pinus kesiya* var. *langbianensis*. Moreover, only one sample plot was harvested using a clear-cutting plot; to some extent, it also reflects the organ and inter-specific differences of AGB at an individual tree scale for the *Pinus kesiya* var. *langbianensis* forest. However, there is a lack of validation or evidence from the relevant research. Therefore, in future research, the techniques employed in this study can be extended to investigate the spatial effects of biomass across a larger number of sample plots and more forest types to further confirm the spatial effect of the tree biomass in a forest at an individual tree scale. Additionally, the utilization of spatial regression models can be explored as a means to explore the spatial effect analysis for biomass.

5. Conclusions

In this study, all trees in a plot with 0.3 hectares (30 m \times 100 m) for *Pinus kesiya* var. *langbianensis* natural forest were entirely cut down in Mojiang County, Yunnan Province, southwestern China. Then, the AGB of each tree was measured according to different organs, such as wood, bark, foliage, and branches. The intra-group variance was applied to describe the spatial heterogeneity of biomass in each component, and the global and local Moran indices were used to analyze the spatial autocorrelation for the three types of tree species, such as *Pinus kesiya* var. *langbianensis* (PK), the other upper trees (UPs), and the other lower trees (LTs). Therefore, these are the following conclusions.

(1) There was spatial heterogeneity for the biomass of each component in the *P. kesiya* var. *langbianensis* natural forest. The spatial heterogeneity of each biomass component showed an increase with distance and stabilization at distances of about 20 m for the UP and PK, but for the LT, the distance values were 10 m.

(2) The spatial autocorrelation differences among the organs were obvious, and some components showed a similar change trend. The spatial autocorrelation trend of AGB and wood biomass were similar, while bark biomass and foliage biomass had a similar spatial variation trend.

(3) The spatial effect was various for the different trees or tree groups. The weakest spatial heterogeneity was found in LT, while PK showed strong spatial heterogeneity in all components except bark biomass. PK and UP also showed significant spatial autocorrelation, with a global Moran's I of 0.0389–0.1537 and 0.1277–0.1644, respectively.

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References

- 1. Pretzsch, H. Trees grow modulated by the ecological memory of their past growth. Consequences for monitoring, modelling, and silvicultural treatment. *For. Ecol. Manag.* **2021**, *487*, 118982. [CrossRef]
- Bhandari, S.K.; Veneklaas, E.J.; McCaw, L.; Mazanec, R.; Whitford, K.; Renton, M. Individual tree growth in jarrah (*Eucalyptus marginata*) forest is explained by size and distance of neighbouring trees in thinned and non-thinned plots. *For. Ecol. Manag.* 2021, 494, 119364. [CrossRef]
- 3. Aussenac, R.; Bergeron, Y.; Gravel, D.; Drobyshev, I. Interactions among trees: A key element in the stabilising effect of species diversity on forest growth. *Funct. Ecol.* **2019**, *33*, 360–367. [CrossRef]
- 4. Anselin, L.; Griffith, D.A. Do spatial effects really matter in regression analysis? Pap. Reg. Sci. 1988, 65, 11–34. [CrossRef]
- 5. Zhang, L.; Shi, H. Local modeling of tree growth by geographically weighted regression. For. Sci. 2004, 50, 225–244.
- 6. Stojanova, D.; Ceci, M.; Appice, A.; Malerba, D.; Džeroski, S. Dealing with spatial autocorrelation when learning predictive clustering trees. *Ecol. Inform.* 2013, *13*, 22–39. [CrossRef]
- Anselin, L. Lagrange multiplier test diagnostics for spatial dependence and spatial heterogeneity. *Geogr. Anal.* 1988, 20, 1–17. [CrossRef]
- 8. Chen, Y. New approaches for calculating Moran's index of spatial autocorrelation. PLoS ONE 2013, 8, e68336. [CrossRef]
- 9. Legendre, P. Spatial autocorrelation: Trouble or new paradigm? Ecology 1993, 74, 1659–1673. [CrossRef]
- 10. Kashlak, A.B.; Yuan, W. Computation-free nonparametric testing for local spatial association with application to the US and Canadian electorate. *Spat. Stat.* **2022**, *48*, 100617. [CrossRef]
- 11. Moran, P.A. Notes on continuous stochastic phenomena. Biometrika 1950, 37, 17–23. [CrossRef] [PubMed]
- 12. Geary, R.C. The contiguity ratio and statistical mapping. Inc. Stat. 1954, 5, 115–146. [CrossRef]
- 13. Getis, A.; Ord, J.K. The analysis of spatial association by use of distance statistics. *Geogr. Anal.* 1992, 24, 189–206. [CrossRef]
- 14. Darand, M.; Dostkamyan, M.; Rehmani, M. Spatial autocorrelation analysis of extreme precipitation in Iran. *Russ. Meteorol. Hydrol.* **2017**, *42*, 415–424. [CrossRef]
- 15. Sari, F.; Frananda, H.; Fransiska, S. Identification of Spatial Autocorrelation in the Poverty Level in West Pasaman Regency with Moran Index. *J. Phys. Conf. Ser.* **2020**, *1554*, 012052. [CrossRef]
- 16. Anselin, L. Local indicators of spatial association—LISA. Geogr. Anal. 1995, 27, 93–115. [CrossRef]
- 17. Dalposso, G.H.; Uribe-Opazo, M.A.; Mercante, E.; Lamparelli, R.A. Spatial autocorrelation of NDVI and GVI indices derived from Landsat/TM images for soybean crops in the western of the state of Paraná in 2004/2005 crop season. *Eng. Agrícola* **2013**, *33*, 525–537. [CrossRef]
- 18. Shi, H.; Zhang, L. Local analysis of tree competition and growth. *Forest Sci.* 2003, *49*, 938–955.
- Chas-AMil, M.L.; PresTeMon, J.P.; MccleAn, C.J.; TouzA, J. Human-ignited wildfire patterns and responses to policy shifts. *Appl. Geogr.* 2015, 56, 164–176. [CrossRef]
- Anselin, L.; Kelejian, H.H. Testing for spatial error autocorrelation in the presence of endogenous regressors. *Int. Reg. Sci. Rev.* 1997, 20, 153–182. [CrossRef]

- 21. Yin, C.; Yuan, M.; Lu, Y.; Huang, Y.; Liu, Y. Effects of urban form on the urban heat island effect based on spatial regression model. *Sci. Total Environ.* **2018**, 634, 696–704. [CrossRef] [PubMed]
- 22. Liu, Z.; Jiang, F.; Zhu, Y.; Li, F.; Jin, G. Spatial heterogeneity of leaf area index in a temperate old-growth forest: Spatial autocorrelation dominates over biotic and abiotic factors. *Sci. Total Environ.* **2018**, *634*, 287–295. [CrossRef] [PubMed]
- 23. Junttila, V.; Laine, M. Bayesian Principal Component Regression model with spatial effects for forest inventory under small field sample size. *arXiv* **2016**, arXiv:1605.07439. [CrossRef]
- 24. Ferré, C.; Castrignanò, A.; Comolli, R. Comparison between spatial and non-spatial regression models for investigating tree–soil relationships in a polycyclic tree plantation of Northern Italy and implications for management. *Agrofor. Syst.* **2019**, *93*, 2181–2196. [CrossRef]
- 25. Klooster, D.J. Toward adaptive community forest management: Integrating local forest knowledge with scientific forestry. *Econ. Geogr.* 2002, *78*, 43–70. [CrossRef]
- Pickett, S.T.; Cadenasso, M.L. Landscape ecology: Spatial heterogeneity in ecological systems. *Science* 1995, 269, 331–334. [CrossRef]
- 27. Carey, A. Biocomplexity and restoration of biodiversity in temperate coniferous forest: Inducing spatial heterogeneity with variable-density thinning. *Forestry* **2003**, *76*, 127–136. [CrossRef]
- 28. Assal, T.J.; Anderson, P.J.; Sibold, J. Spatial and temporal trends of drought effects in a heterogeneous semi-arid forest ecosystem. *For. Ecol. Manag.* **2016**, *365*, 137–151. [CrossRef]
- 29. Beckage, B.; Clark, J.S. Seedling survival and growth of three forest tree species: The role of spatial heterogeneity. *Ecology* **2003**, *84*, 1849–1861. [CrossRef]
- Ngao, J.; Epron, D.; Delpierre, N.; Bréda, N.; Granier, A.; Longdoz, B. Spatial variability of soil CO2 efflux linked to soil parameters and ecosystem characteristics in a temperate beech forest. *Agric. For. Meteorol.* 2012, 154, 136–146. [CrossRef]
- 31. Ward, J.S.; Parker, G.R.; Ferrandino, F.J. Long-term spatial dynamics in an old-growth deciduous forest. *For. Ecol. Manag.* **1996**, *83*, 189–202. [CrossRef]
- 32. Brazhnik, K.; Shugart, H. Model sensitivity to spatial resolution and explicit light representation for simulation of boreal forests in complex terrain. *Ecol. Model.* 2017, 352, 90–107. [CrossRef]
- Gundale, M.J.; Metlen, K.L.; Fiedler, C.E.; DeLuca, T.H. Nitrogen spatial heterogeneity influences diversity following restoration in a ponderosa pine forest, Montana. *Ecol. Appl.* 2006, 16, 479–489. [CrossRef] [PubMed]
- 34. Gossner, M.M.; Getzin, S.; Lange, M.; Pašalić, E.; Türke, M.; Wiegand, K.; Weisser, W.W. The importance of heterogeneity revisited from a multiscale and multitaxa approach. *Biol. Conserv.* **2013**, *166*, 212–220. [CrossRef]
- 35. Hewitt, J.E.; Thrush, S.F.; Dayton, P.K.; Bonsdorff, E. The effect of spatial and temporal heterogeneity on the design and analysis of empirical studies of scale-dependent systems. *Am. Nat.* 2007, *169*, 398–408. [CrossRef]
- 36. Detto, M.; Asner, G.P.; Muller-Landau, H.C.; Sonnentag, O. Spatial variability in tropical forest leaf area density from multireturn lidar and modeling. *J. Geophys. Res. Biogeosci.* 2015, 120, 294–309. [CrossRef]
- Getzin, S.; Fischer, R.; Knapp, N.; Huth, A. Using airborne LiDAR to assess spatial heterogeneity in forest structure on Mount Kilimanjaro. *Landsc. Ecol.* 2017, 32, 1881–1894. [CrossRef]
- Clark, P.J.; Evans, F.C. Distance to nearest neighbor as a measure of spatial relationships in populations. *Ecology* 1954, 35, 445–453. [CrossRef]
- Madden, L.; Hughes, G.; Ellis, M. Spatial heterogeneity of the incidence of grape downy mildew. *Phytopathology* 1995, 85, 269–275. [CrossRef]
- Perry, G.L.; Miller, B.P.; Enright, N.J. A comparison of methods for the statistical analysis of spatial point patterns in plant ecology. *Plant Ecol.* 2006, 187, 59–82. [CrossRef]
- 41. Fotheringham, A.S. Trends in quantitative methods I: Stressing the local. Prog. Hum. Geogr. 1997, 21, 88–96. [CrossRef]
- 42. Fotheringham, A.S. "The problem of spatial autocorrelation" and local spatial statistics. Geogr. Anal. 2009, 41, 398–403. [CrossRef]
- 43. Yang, X.; Han, Y. Spatial heterogeneity of soil nitrogen in six natural secondary forests in mountainous region of northern China. *Sci. Soil Water Conserv.* **2010**, *8*, 95–102.
- Lamsal, S.; Rizzo, D.; Meentemeyer, R. Spatial variation and prediction of forest biomass in a heterogeneous landscape. *J. For. Res.* 2012, 23, 13–22. [CrossRef]
- 45. Fotheringham, A.S.; Charlton, M.E.; Brunsdon, C. Geographically weighted regression: A natural evolution of the expansion method for spatial data analysis. *Environ. Plan. A* **1998**, *30*, 1905–1927. [CrossRef]
- Wang, Q.; Ni, J.; Tenhunen, J. Application of a geographically-weighted regression analysis to estimate net primary production of Chinese forest ecosystems. *Glob. Ecol. Biogeogr.* 2005, 14, 379–393. [CrossRef]
- 47. Zhang, J.; Cheng, G.; Yu, F.; Kräuchi, N.; Li, M.-H. Interspecific variations in responses of Festuca rubra and Trifolium pratense to a severe clipping under environmental changes. *Biologia* 2009, *64*, 292–298. [CrossRef]
- 48. Nazeer, M.; Bilal, M. Evaluation of ordinary least square (OLS) and geographically weighted regression (GWR) for water quality monitoring: A case study for the estimation of salinity. *J. Ocean. Univ. China* **2018**, *17*, 305–310. [CrossRef]
- 49. Pradhan, B.; Youssef, A.M. Manifestation of remote sensing data and GIS on landslide hazard analysis using spatial-based statistical models. *Arab. J. Geosci.* 2010, *3*, 319–326. [CrossRef]
- 50. Dale, M.R.; Fortin, M.-J. Spatial Analysis: A Guide for Ecologists; Cambridge University Press: Cambridge, UK, 2014.

- 51. Zhang, B.; Ou, G.; Sun, X.; Xu, T.; Xu, H. Application of Spatial Effect and Regression Model on Forestry Research. J. Southwest For. Univ. 2016, 36, 144–152.
- 52. Liu, C. Spatial Distribution of Forest Carbon Storage in Heilongjiang Province; Northeast Forestry University: Harbin, China, 2014.
- 53. Zhou, L.; Ou, G.; Wang, J.; Xu, H. Light Saturation Point Determination and Biomass Remote Sensing Estimation of *Pinus kesiya* var. langbianensis Forest Based on Spatial Regression Models. *Sci. Silvae Sin.* **2020**, *56*, 38–47.
- 54. Ou, G.L.; Wang, J.F.; Xiao, Y.F.; Xu, H. Modeling Individual Biomass of *Pinus kesiya* var.langbianensis Natural Forests by Geographically Weighted Regression. *For. Res.* 2014, 27, 213–218. [CrossRef]
- Ou, G.; Wang, J.; Xu, H.; Chen, K.; Zheng, H.; Zhang, B.; Sun, X.; Xu, T.; Xiao, Y. Incorporating topographic factors in nonlinear mixed-effects models for aboveground biomass of natural Simao pine in Yunnan, China. *J. For. Res.* 2016, *27*, 119–131. [CrossRef]
 Lu, L: Zhang, L. Evaluation of parameter estimation methods for fitting spatial regression models. *For. Sci.* 2010, *56*, 505–514.
- Lu, J.; Zhang, L. Evaluation of parameter estimation methods for fitting spatial regression models. *For. Sci.* 2010, *56*, 505–514.
 Lu, J.; Zhang, L. Modeling and prediction of tree height–diameter relationships using spatial autoregressive models. *For. Sci.* 2011, *57*, 252–264. [CrossRef]
- 58. Lu, J.; Zhang, L. Geographically local linear mixed models for tree height-diameter relationship. *For. Sci.* **2012**, *58*, 75–84. [CrossRef]
- 59. Gu, F.; Zhao, Q. Geographically weighted regression model for expressing tree growth relationships. J. Northeast For. Univ. 2012, 40, 129–140.
- 60. Zhang, L.; Gove, J.H.; Heath, L.S. Spatial residual analysis of six modeling techniques. Ecol. Model. 2005, 186, 154–177. [CrossRef]
- 61. Zhang, L.; Ma, Z.; Guo, L. An evaluation of spatial autocorrelation and heterogeneity in the residuals of six regression models. *For. Sci.* **2009**, *55*, 533–548.
- 62. Meng, Q.; Cieszewski, C.J.; Strub, M.R.; Borders, B.E. Spatial regression modeling of tree height–diameter relationships. *Can. J. For. Res.* 2009, *39*, 2283–2293. [CrossRef]
- 63. Soto, D.P.; Salas, C.; Donoso, P.J.; Uteau, D. Structural and spatial heterogeneity of a mixed Nothofagus donibeyi-dominate forest stand after a partial disturbance. *Rev. Chil. Hist. Nat.* **2010**, *83*, 335–347.
- 64. Rozas, V. Structural heterogeneity and tree spatial patterns in an old-growth deciduous lowland forest in Cantabria, northern Spain. *Plant Ecol.* **2006**, *185*, 57–72. [CrossRef]
- 65. Pearce, H.; Anderson, W.; Fogarty, L.; Todoroki, C.; Anderson, S. Linear mixed-effects models for estimating biomass and fuel loads in shrublands. *Can. J. For. Res.* **2010**, *40*, 2015–2026. [CrossRef]
- Chai, Z. Quantitative Evaluation and R Programming of Forest Spatial Structure Based on the Relationship of Neighborhood Trees: A Case Study of Typical Secondary Forest in the Mid-Altitude Zone of the Qinling Mountains; Northwest A&F University: Yangling, China, 2016.
- 67. Nong, M.; Leng, Y.; Xu, H.; Li, C.; Ou, G. Incorporating competition factors in a mixed-effect model with random effects of site quality for individual tree above-ground biomass growth of *Pinus kesiya* var. *langbianensis*. *N. Z. J. For. Sci.* **2019**, 49. [CrossRef]
- 68. Chen, Q.; Zheng, Z.; Feng, Z.; Ma, Y.; Sha, L.; Xu, H.; Nong, P.; Li, Z. Biomass and carbon storage of *Pinus kesiya* var. *langbianensis in Puer, Yunnan. J. Yunnan Univ.-Nat. Sci. Ed.* **2014**, *36*, 439–445.
- 69. Fan, Y.; Zhang, S.; Lan, Z.; Lan, Q. Possible causes for the differentiation of Pinus yunnanensis and P. Kesiya var. *Langbianensis in Yunnan, China: Evidence from seed germination. For. Ecol. Manag.* **2021**, 494, 119321.
- Flora of China Editorial Committee. Flora of China. 2018. Available online: http://www.efloras.org/flora_page.asp (accessed on 30 August 2018).
- Chen, G.; Zhang, X.; Liu, C.; Liu, C.; Xu, H.; Ou, G. Error Analysis on the Five Stand Biomass Growth Estimation Methods for a Sub-Alpine Natural Pine Forest in Yunnan, Southwestern China. *Forests* 2022, 13, 1637. [CrossRef]
- 72. Nong, M. Comparative Analysis on the Spatial Effects of Individualtree Biomass in Typical Subtropical Forests; Southwest Forestry University: Kunming, China, 2020.
- 73. Lieshout, M.v. AJ-function for marked point patterns. Ann. Inst. Stat. Math. 2006, 58, 235–259. [CrossRef]
- 74. Turner, M.G.; Donato, D.C.; Romme, W.H. Consequences of spatial heterogeneity for ecosystem services in changing forest landscapes: Priorities for future research. *Landsc. Ecol.* **2013**, *28*, 1081–1097. [CrossRef]
- 75. Du, H.; Zhou, G.; Fan, W.; Ge, H.; Xu, X.; Shi, Y.; Fan, W. Spatial heterogeneity and carbon contribution of aboveground biomass of moso bamboo by using geostatistical theory. *Plant Ecol.* **2010**, 207, 131–139. [CrossRef]
- Wang, W.; Dong, X.; Dong, X.; Lv, D.; Su, T.; Zheng, A. Study on spatial autocorrelation of forest biomass. For. Eng. 2018, 34, 35–39.
- 77. Liu, K.; Jiang, S.; Zhu, W. Estimation of carbon sequestration value and analysis of space effect of forests in Guangdong Province. *Chin. J. Agric. Resour. Reg. Plan.* **2015**, *36*, 120.
- Frelich, L.E.; Calcote, R.R.; Davis, M.B.; Pastor, J. Patch formation and maintenance in an old-growth hemlock-hardwood forest. *Ecology* 1993, 74, 513–527. [CrossRef]
- 79. Park, A.; Kneeshaw, D.; Bergeron, Y.; Leduc, A. Spatial relationships and tree species associations across a 236-year boreal mixedwood chronosequence. *Can. J. For. Res.* **2005**, *35*, 750–761. [CrossRef]
- Rödig, E.; Cuntz, M.; Heinke, J.; Rammig, A.; Huth, A. Spatial heterogeneity of biomass and forest structure of the Amazon rain forest: Linking remote sensing, forest modelling and field inventory. *Glob. Ecol. Biogeogr.* 2017, 26, 1292–1302. [CrossRef]
- Han, N.; Du, H.; Zhou, G.; Xu, X.; Cui, R.; Gu, C. Spatiotemporal heterogeneity of Moso bamboo aboveground carbon storage with Landsat Thematic Mapper images: A case study from Anji County, China. Int. J. Remote Sens. 2013, 34, 4917–4932. [CrossRef]

- Xu, Q.; Li, B.; McRoberts, R.E.; Li, Z.; Hou, Z. Harnessing data assimilation and spatial autocorrelation for forest inventory. *Remote Sens. Environ.* 2023, 288, 113488. [CrossRef]
- 83. Zhou, Z.; Tang, Y.; Xu, H.; Wang, J.; Hu, L.; Xu, X. Dynamic changes in leaf biomass and the modeling of individual Moso Bamboo (*Phyllostachys edulis* (Carrière) J. Houz) under intensive management. *Forests* **2022**, *13*, 693. [CrossRef]
- Nötzold, R.; Blossey, B.; Newton, E. The influence of below ground herbivory and plant competition on growth and biomass allocation of purple loosestrife. *Oecologia* 1997, 113, 82–93. [CrossRef]
- 85. Pattison, R.; Goldstein, G.; Ares, A. Growth, biomass allocation and photosynthesis of invasive and native Hawaiian rainforest species. *Oecologia* **1998**, *117*, 449–459. [CrossRef]
- Wu, H.; Xu, H.; Tian, X.; Zhang, W.; Lu, C. Multistage Sampling and Optimization for Forest Volume Inventory Based on Spatial Autocorrelation Analysis. *Forests* 2023, 14, 250. [CrossRef]
- 87. Wang, J.; Haining, R.; Cao, Z. Sample surveying to estimate the mean of a heterogeneous surface: Reducing the error variance through zoning. *Int. J. Geogr. Inf. Sci.* **2010**, *24*, 523–543. [CrossRef]
- Holmberg, H.; Lundevaller, E.H. A test for robust detection of residual spatial autocorrelation with application to mortality rates in Sweden. Spat. Stat. 2015, 14, 365–381. [CrossRef]
- Wulder, M.A.; White, J.C.; Coops, N.C.; Nelson, T.; Boots, B. Using local spatial autocorrelation to compare outputs from a forest growth model. *Ecol. Model.* 2007, 209, 264–276. [CrossRef]
- Bebre, I.; Annighöfer, P.; Ammer, C.; Seidel, D. Growth, morphology, and biomass allocation of recently planted seedlings of seven European tree species along a light gradient. *iFor.-Biogeosci. For.* 2020, 13, 261–269. [CrossRef]
- Ren, J.; Wang, H.; Wang, W.; Qu, D.; Wang, Q.; Zhong, Z. Responses of photosynthesis, chlorophyll fluorescence of poplar leaf and bark chlorenchyma to elevated temperature. *Bull. Bot. Res.* 2014, 34, 758–764.

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