

Article

# Impacts of Change in Atmospheric CO<sub>2</sub> Concentration on *Larix gmelinii* Forest Growth in Northeast China from 1950 to 2010

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**Abstract:** Although CO<sub>2</sub> fertilization on plant growth has been repeatedly modeled to be the main reason for the current changes in the terrestrial carbon sink at the global scale, there have been controversial findings on the CO<sub>2</sub> fertilization effects on forests from tree-ring analyses. In this study, we employed conventional dendrochronological tree-ring datasets from Northeast China, to detect the effect of CO<sub>2</sub> fertilization on *Larix gmelinii* growth from 1950 to 2010. Among four sites, there were two sites exhibiting a significant residual growth enhancement at a 90% confidence level after removing the size, age and climate-related trends of tree-ring indices. In addition, we found consistency ( $R$  from 0.26 to 0.33,  $p < 0.1$ ) between the high frequency CO<sub>2</sub> fluctuation and residual growth indices at two of the four sites during the common period. A biogeochemical model was used to quantitatively predict the contribution of elevated atmospheric CO<sub>2</sub> on accumulated residual growth enhancement. As found in the tree-ring data, 14% of the residual growth was attributed to the CO<sub>2</sub> fertilization effect, while climate was responsible for approximately the remaining 86%.

**Keywords:** tree growth; CO<sub>2</sub> fertilization; tree-ring; dendrochronology; InTEC model

## 1. Introduction

The rapid increase in CO<sub>2</sub> concentration in the atmosphere as a greenhouse gas is thought to be responsible for the increase in earth's surface temperature [1]. The response of forest ecosystems to elevated atmospheric CO<sub>2</sub> concentration will affect their net uptake or loss of carbon and may, therefore, have large consequences on the global carbon cycle [2]. Numerous global-level experiments were conducted to investigate and understand how the terrestrial ecosystem carbon cycle responds to rising atmospheric CO<sub>2</sub>. Recent studies based on terrestrial carbon cycle models suggested that the strength of the terrestrial C sink was growing at the global scale, while CO<sub>2</sub> fertilization was the predominant driver of the growth in the terrestrial C sink [3–6]. Free air CO<sub>2</sub> enrichment (FACE) experiments [7–10] also showed considerable growth enhancement due to the CO<sub>2</sub> fertilization effects. In spite of the wealth of global assessments and experimental evidences on CO<sub>2</sub> fertilization on tree growth, results from tree-ring studies on the CO<sub>2</sub> fertilization effect are controversial and still under debate [11–17].

It is widely held that the postindustrial rise in the concentration of CO<sub>2</sub> in the atmosphere should have enhanced tree growth through a fertilization effect, that can be ascribed this CO<sub>2</sub> fertilization effect on growth to the following two main mechanisms: first, direct CO<sub>2</sub> fertilization may occur because the higher partial pressure of CO<sub>2</sub> increases the rate of CO<sub>2</sub> reactions with Rubisco during photosynthesis, and inhibits photorespiration [18]. Second, the increase in water-use efficiency takes place when

the stomatal conductance is reduced but, at the same time, also the assimilation is kept constant or increasing, which results in an increase in the ratio of carbon gain to water loss [19]. As expected, several tree-ring studies have reported an increase in radial growth of trees with the rise of atmospheric CO<sub>2</sub> concentration in natural forests [11–13]. In contrast, some studies did not find any evidence of this CO<sub>2</sub> fertilization effect as a cause for enhanced tree growth in mid- and high-latitude forests of the northern hemisphere [15–17]. It may be argued that the CO<sub>2</sub> fertilization effect on forests at high latitudes is either non-existent because of other growth-limiting conditions such as temperature and nitrogen or is too small to detect [20,21].

Here, the objective of our study is to apply tree-ring analysis approaches to detect possible effects of atmospheric CO<sub>2</sub> fertilization on natural *Larix gmelinii* forests in Northeast China over the period 1950–2010. We examined the time-related trend in the residual growth after removing size-, age- and climate-related trends through conventional dendrochronological analysis. Consequently, any CO<sub>2</sub>-induced increases in productivity could be evident as an increase in the residual tree-ring width. Meanwhile, an integrated terrestrial ecosystem C-budget (InTEC) model [22], which integrates stand development with the effects of environmental conditions on growth, was used to assess the contributions of climate factors, and increased CO<sub>2</sub> concentration on growth enhancements in stemwood from 1950 to 2010. Finally, we will directly evaluate whether the CO<sub>2</sub> fertilization effect is significant in forest stands at high latitudes.

## 2. Materials and Methods

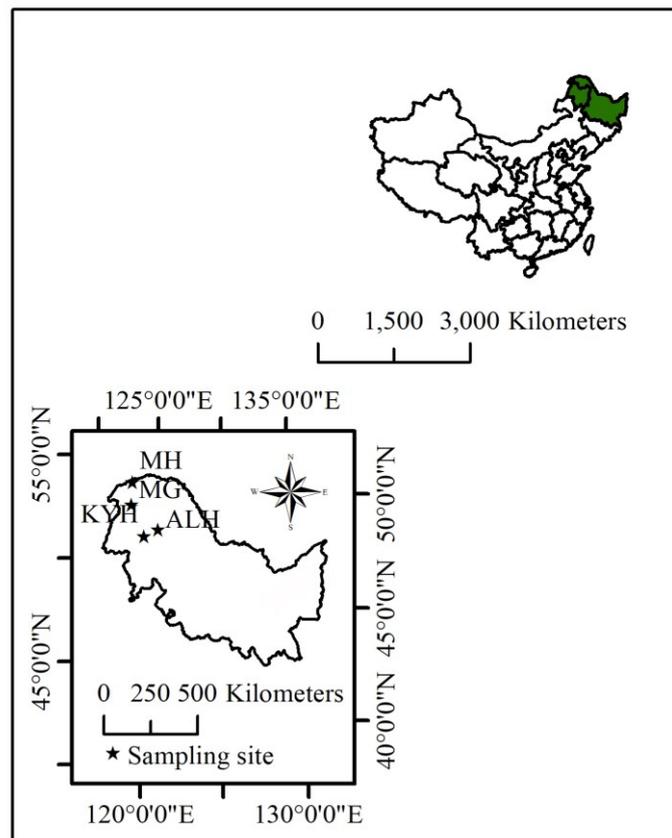
### 2.1. Tree-Ring Data

Increment cores of Chinese *Larix gmelinii* trees were collected from four sites. All sites were selected in stands with low levels of human disturbance and distributed in the main forest areas of Northeast China (Figure 1). These sites include four mature-tree sites (more than 200 years): Mohe, Heilongjiang, China (MH), Mangui, Inner Mongolia, China (MG), Alihe, Inner Mongolia, China (ALH), and Keyihe, Inner Mongolia, China (KYH). We measured annual growth increments from the pith to the outermost ring at a precision of 0.001 mm using a Velmax measuring system (sliding stage, Velmex Inc., Bloomfield, NY, USA), dating and measurement errors were further checked with COFECHA software [23]. Finally, a total of 171 cores (1 per tree) were collected, and ring-width measurements were recorded for a period extending from 1715 to 2010 (Table 1).

**Table 1.** Characteristics of four sampling sites.

Site	Longitude (E)	Latitude (N)	Elev. (m)	T (°C)	P (mm)	Length of Chronology	N
MH	122.13	53.30	617	−5.1	437	1802–2005	61
MG	121.80	52.23	695	−6.0	465	1760–2005	61
ALH	123.48	50.84	135	−3.9	471	1782–2010	27
KYH	122.38	50.64	779	−3.7	461	1715–2010	22

T: annual mean temperature; P: annual total precipitation; N: number of trees.



**Figure 1.** Locations of the sampling sites in Northeast China. Abbreviations for our sampling sites are MH (Mohe), MG (Mangui), ALH (Alihe) and KYH (Keyihe).

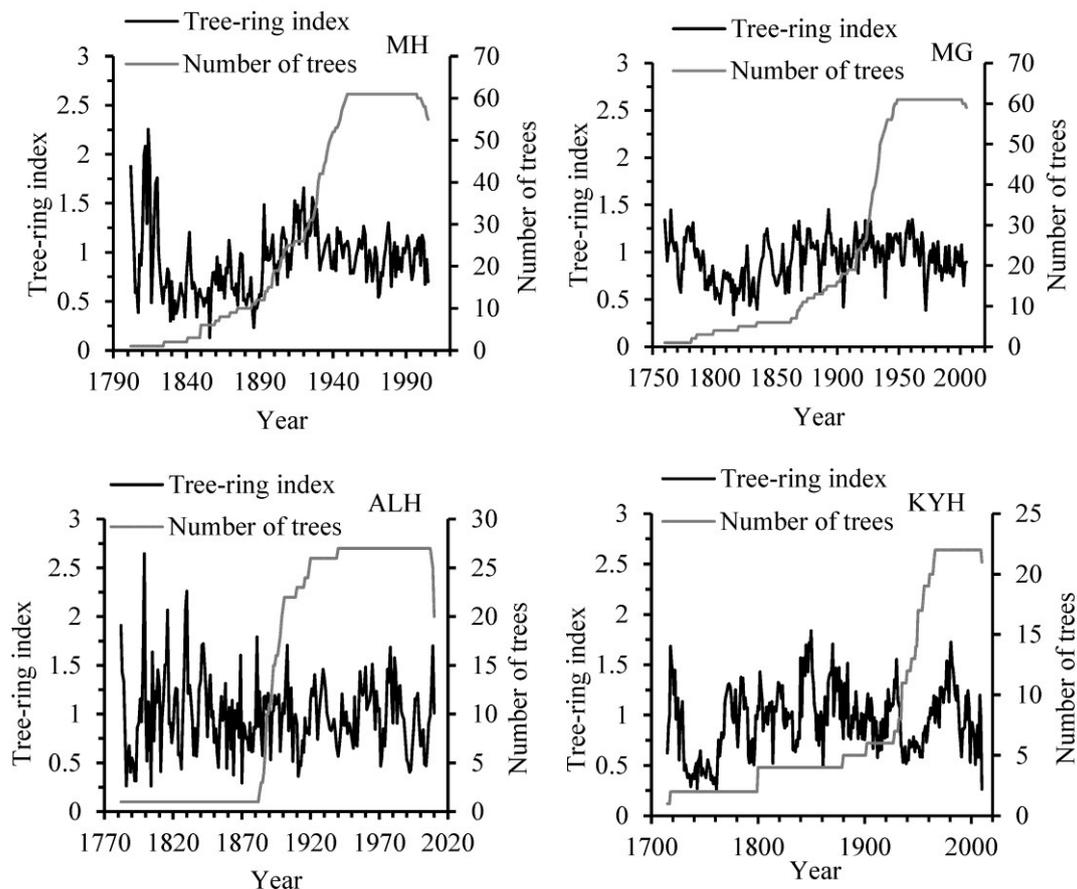
## 2.2. Tree-Ring Approach

### 2.2.1. Chronology Development

All of the ring-width measurements were detrended using linear lines or negative exponential curves. These conservative approaches to detrending should preserve virtually all of the low-frequency variability in the tree ring series [24]. The standard (STD) chronologies were then generated for each site. All series were created by the program ARSTAN [24] (Figure 2). The statistical characteristics of the STD chronologies for each site include (Table 2): the mean index and standard deviation (Std.) of the STD chronologies; the skewness and kurtosis are included to assess the effects on the probability distribution owing to the method of standardization [25]; the mean sensitivity, which quantifies the relative change in width among consecutive years [26]; the serial correlation is used to describe the correlation of each ring width with the preceding ring width [26]; the Rbar express average correlation among trees for the common overlap period among series [26]; the expressed population signal (EPS) (Equation (1)) is used to test the chronology confidence and strength of the common signal in the chronology. We employed 0.85 as an expressed population signal (EPS) threshold value to assess the robustness, which has been widely suggested as reliable [11,12,27].

$$\text{EPS} = \frac{N\bar{r}}{1 + (N-1)\bar{r}} \quad (1)$$

where  $N$  is the number of trees,  $\bar{r}$  is the mean inter-series correlation.



**Figure 2.** The standard (STD) chronologies and the number of trees for each site.

**Table 2.** General statistics of the standard chronologies for four sites.

Sites	MH	MG	ALH	KYH
Mean index	0.92	0.94	1.01	0.98
Std.	0.34	0.23	0.40	0.23
Skewness	0.89	0.28	1.02	0.03
Kurtosis	4.66	2.55	5.37	2.74
Mean Sensitivity	0.24	0.17	0.31	0.21
Serial correlation	0.65	0.68	0.47	0.39
Rbar	0.26	0.30	0.26	0.31
EPS	0.97	0.98	0.93	0.94

### 2.2.2. Dendroclimatic Analysis

Our dendroclimatic analysis was conducted by using the program PRECON [28] throughout 1950–2010 for four sites. In this work, a climatic response function analysis was performed with monthly mean temperatures and total precipitation records (from June of the previous year to October of the current year) as climatic predictors, and the STD chronologies as the dependent variables. Then the predicted ring-width indices were estimated through multiple regression, after extracting the principal components of the climatic predictors that explained the greatest variation in radial growth. This includes a bootstrap method to assess statistical significance of the regression coefficients in response functions. In this work, 1000 iterations were computed to estimate the confidence intervals of regression coefficients ( $p < 0.05$ ). If the confidence interval for the mean of a bootstrapped estimate of a regression coefficient does not include zero, the association between growth and that climate variable is deemed statistically significant. The final residual indices (STD residuals) could be calculated as

actual indices minus predicted indices; this process means that these climatically induced variabilities have been removed from the STD tree-ring chronologies.

### 2.3. CO<sub>2</sub> Fertilization Test

To test the CO<sub>2</sub> fertilization hypothesis, we used two methods. First, a nonparametric method was implemented in this study for detecting the trend of residual indices after removing the effects of age and climatic variability from 1950–2010. If a linear trend exists, the true slope can also be estimated by a nonparametric procedure developed by Sen (1968) [29], which is closely linked to the Mann–Kendall test [30]. The Sen’s method is not much affected by outliers and can be computed with missing data. Then the residual indices (STD residuals) were tested for an increasing trend in growth over time that cannot be attributed to climate by the Sen’s method. This test is premised on the assumption that once the effects of climatic variability have been removed from the chronologies, the residual indices should exhibit an increasingly positive bias over time due to CO<sub>2</sub> growth stimulations [14].

Second, to test the effect of CO<sub>2</sub> fluctuation on tree growth variation, we calculated the first order difference of the raw CO<sub>2</sub> concentration, then the linear trend was removed from the first order difference of the raw CO<sub>2</sub> concentration. In this way, a high frequency fluctuation in CO<sub>2</sub> concentration could be produced and correlated with interannual variations in tree growth [12]. Then, the comparisons were performed between the CO<sub>2</sub> fluctuation time series and the STD residuals for four sites, respectively.

### 2.4. Description of the InTEC Model

The InTEC model is a process-based biogeochemical model that mechanistically integrates the effects of non-disturbance factors (climate variables and atmospheric CO<sub>2</sub>) and disturbance factors (disturbance and regrowth) on the long-term C and N cycles in forest ecosystems. The model includes five core processes: (a) simulation of net primary production (NPP) in a recent reference year ( $NPP_{ref}$ ) using a two-leaf canopy photosynthesis model based on Farquhar’s leaf-level biochemical model [31]; (b) based on  $NPP_{ref}$  and past climate, the initial value of NPP in the starting year ( $NPP_0$ ) of simulation is reconstructed retrospectively through iterative adjustments of  $NPP_0$ , until the simulated NPP in the reference year agree with  $NPP_{ref}$  to within  $\pm 1\%$ ; (c) the NPP of a region for each pixel in any year is calculated from the  $NPP_0$  multiplied by a factor that integrates the effects of non-disturbance factors ( $\phi_{NPP}$ ) and forest stand age ( $F_{NPP(i)}$ ) (Equation (2)). In the InTEC model, NPP-age relationships are replaced by normalized productivity ( $F_{NPP}$ ) curves (Equation (3)) [32]. The curves are obtained by dividing NPP at a given age by their maximum NPP value in their forest life cycle with values ranging from 0 to 1. These relationships are used to simulate forest regrowth after disturbance and changes in the different C components; (d) a three-dimensional distributed hydrological model is used to simulate soil moisture and temperature [33]; and (e) a modified CENTURY model [34] and the net N mineralization model [35] are employed to simulate soil C and N cycles.

$$NPP(i) = NPP_0 \times \phi_{NPP}(i) \times F_{NPP}(i) \quad (2)$$

$$F_{NPP}(i) = NPP(i) / NPP_{max} \quad (3)$$

where  $NPP(i)$  represents the NPP at age  $i$ ,  $NPP_{max}$  represents a maximum of NPP.

#### 2.4.1. Model Inputs

In order to drive the InTEC model, several spatial datasets were created in this study, including monthly mean temperature, water vapor pressure, and total precipitation, which were collected from the UK Climate Research Unit [36], and monthly solar irradiance data was from the US National Center for Atmospheric Research [37]. A map of stand age in 2010 [38] and a forest type map in 2006 [39] were created from forest inventory data in this study. A reference NPP map in 2003 was produced by using the boreal ecosystem productivity simulators (BEPS) [40]. The annual atmospheric CO<sub>2</sub>

concentrations from 1950 to 2010 were taken from the dataset obtained at the Mauna Loa Observatory (20° N, 156° W) [41]. A maximum leaf area index (LAI) map was produced in 2003 by Deng et al. [42] using SPOT-VEGETATION data. The physical properties of the soil used in the InTEC, including the field capacity of soil water, wilting point, soil depth, and the fractions of clay, silt, and sand, were included. Field capacity and wilting point were derived from the International Geosphere-Biosphere Programme, Global Gridded Surfaces of Selected Soil Characteristics [43]. Soil depth was derived from the global soil texture dataset from Oak Ridge National Laboratory Distributed Active Archive Center, Tennessee, U.S. [44]. The fractions of clay, silt, and sand were obtained from the Harmonized World Soil Database (HWSD) constructed by the Food and Agriculture Organization of the United Nations (FAO) and the International Institute [45]. All of these spatial datasets were employed in the UTM WGS-84 coordinate system and interpolated to 1 km resolution. Detailed descriptions of datasets were reported elsewhere [46,47].

#### 2.4.2. Stemwood Biomass Growth

We used InTEC to estimate the stemwood biomass growth ( $G_{SW}$ ) of *Larix gmelinii* forests at four sites from 1950–2010. In this study, we assumed that stemwood growth consists of 31% of NPP in InTEC simulations [48].

#### 2.4.3. Baseline and Residual Stemwood Biomass Growth

The yield table for *Larix gmelinii* of Heilongjiang Province, China was developed by the Forest Survey Scheme Designing Institute in 2010. It was collected to provide information on the stand development, such as mean age, stand density ( $S$ ), volume ( $V$ ) and mean volume growth ( $V_g$ ). The age ranges in the yield table are 0–150, and five-year intervals were used in the yield tables. Using this information, the baseline stemwood growth ( $B_{SW}$ ) can be calculated based on the stand biomass equation.

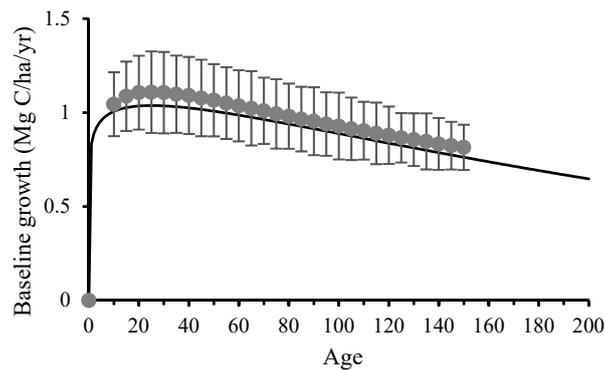
We used Equation (4) to calculate the  $B_{SW}$  for *Larix gmelinii* forests at different ages by combining stand mean volume growth and age information in the yield table. The regression coefficients in Equation (4) were taken from Dong (2015) [49]. The age response of  $B_{SW}$  ( $B_{SW}(\text{age})$ ) was then fitted using the Weibull distribution function as Equation (5) [32] (Figure 3). Finally, the residual stemwood biomass growth ( $R_{SW}$ ), which includes only the total signal of external forcings (e.g., climate,  $\text{CO}_2$ , and  $N$ ) was determined as the difference between  $G_{SW}$  and  $B_{SW}$ .

$$B_{SW} = \exp(a + b \times \log V_g) \times C_s \quad (4)$$

$$B_{SW}(A) = a_1 \left( 1 + \frac{a_2 \left( \frac{A}{a_3} \right)^{a_4} - 1}{\exp \left( \frac{A}{a_3} \right)} \right) \quad (5)$$

where  $\log$  denotes natural logarithm;  $a$  and  $b$  are regression coefficients;  $B_{SW}$  is baseline stemwood biomass growth (Mg/ha/yr);  $V_g$  is mean stand volume growth ( $\text{m}^3/\text{ha}/\text{yr}$ ), which is available in the yield table;  $A$  represents stand age (years);  $a_1$ ,  $a_2$ ,  $a_3$ , and  $a_4$  are fitted parameters that were listed in Table 3. We set the C content in stemwood as 0.47 in this study [50].

In this study, we did not integrate the age and size effects on stemwood growth from the observed tree-ring index into the InTEC model because of lacked of sufficient information to establish the NPP-age relationships for *Larix gmelinii* at these sites. Although, the  $B_{SW}$ -age relationship we established by using the yield table could represent the long-term average state of age response of  $B_{SW}$  at these sites. In addition, the stemwood growth component in the NPP-age relationship which integrated in the InTEC model was also calculated by using the same yield table and biomass equation as in this study [48], so that the age effect could be completely removed from simulated stemwood growth.



**Figure 3.** Variations of baseline stemwood growth with age for *Larix gmelinii*. Dots represent stemwood biomass growth calculated by yield tables; error bars represent standard deviation.

**Table 3.** Coefficient estimates and goodness of fit statistics (Mg C/ha/yr) of Equations (3) and (4).

Parameters	$a$	$b$	$a_1$	$a_2$	$a_3$	$a_4$	$R^2$	RMSE	Samples
Values	-0.726	1.038	0.019	91.639	710236	0.073	0.997	0.051	29

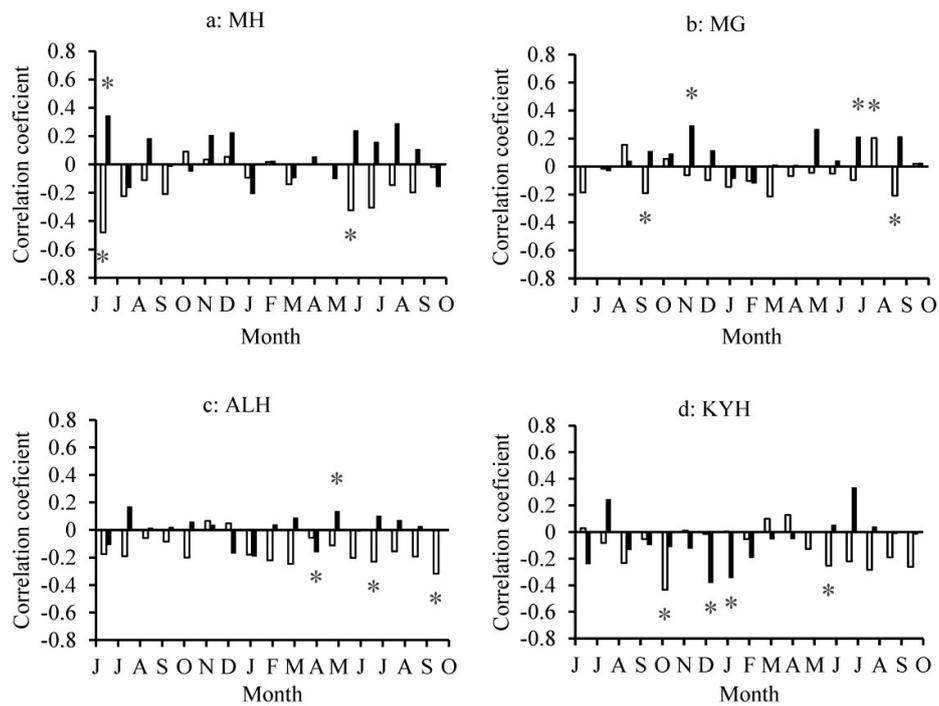
#### 2.4.4. The Partitioning Approach to Attribute Growth Variation to Various Factors

To examine the relative individual effects on growth of climate factors and CO<sub>2</sub> concentration on stemwood growth enhancement, we designed a series of modeling experiments by setting one factor in question constant at a time while using realistic historical values for all other factors in the simulation. For example, to investigate the effect of CO<sub>2</sub> on growth, we first set CO<sub>2</sub> to be a constant as the average over the 1901–1910 period, while the historical values of other factors were used in the simulation. We then detect the CO<sub>2</sub> effect as the difference between the full historical simulation and the constant CO<sub>2</sub> simulation.

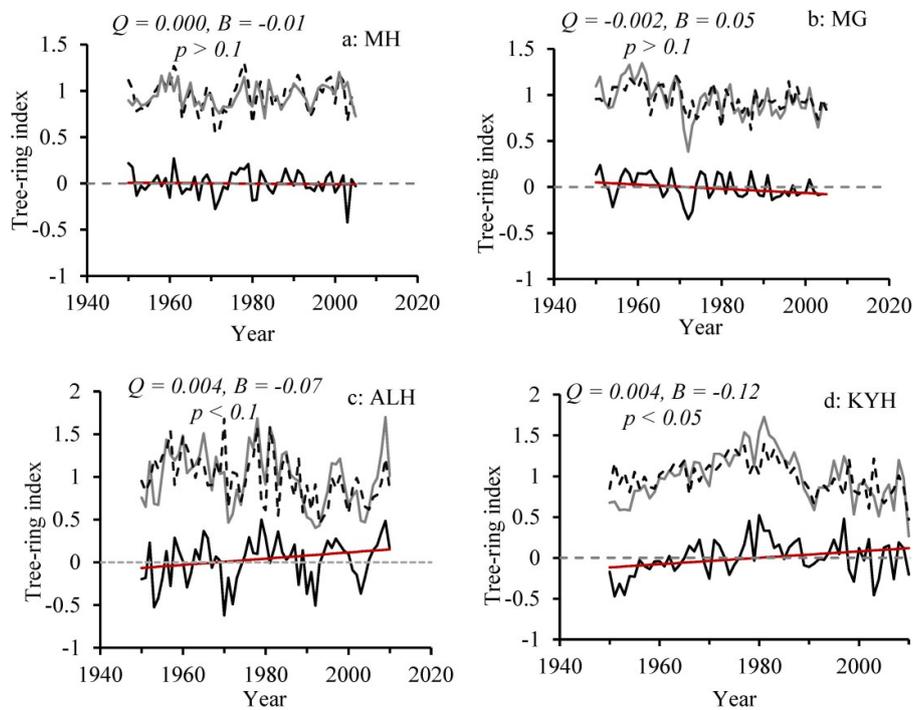
### 3. Results

Our climate-growth response function analysis showed that no clear common patterns were observed in the relationships between climate and radial growth across the four mature-tree sites (Figure 4). The high temperature during the growing season was generally unfavorable for tree growth at these sites, especially in June and July of the current year. The growing season precipitation was positively correlated with the tree growth, while negative correlations could be found in the winter months (from December of the previous year to February of the current year). Additionally, 58%, 63%, 52% and 56% of the variance in the raw growth indices of MH, MG, ALH, and KYH was explained by mean temperature and precipitation for the period 1950–2010. Based on these responses to climate variations, the predicted growth indices could be produced for each site; then we calculated residual indices by subtracting predicted from raw tree-ring indices.

Among the four sites, ALH and KYH exhibited a positive trend in their STD residuals at a 90% and 95% confidence level, respectively (Figure 5a,b), but the other two sites (MH and MG) did not have a significant positive trend in their residuals (Figure 5c,d).

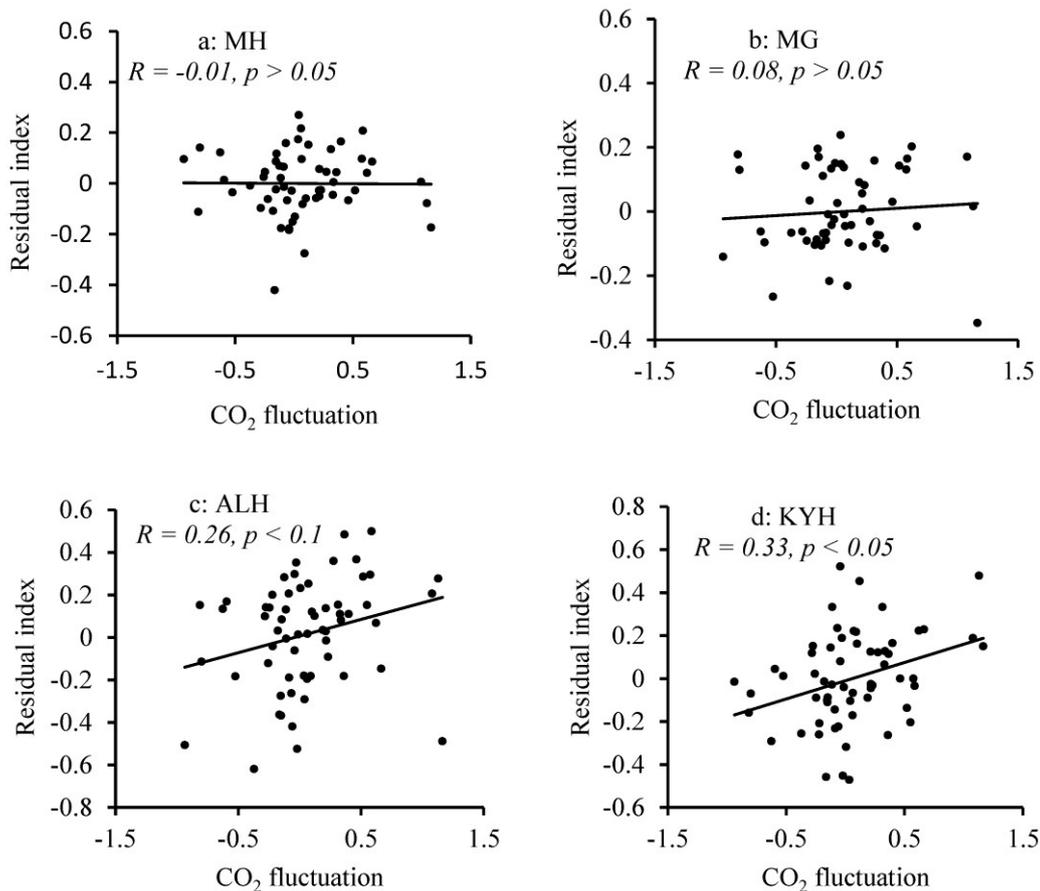


**Figure 4.** Correlation coefficients between the standard chronologies and total precipitation (■) and mean temperature (□) from June in the previous year to October in the current year for four sites over the period 1950–2010. The \* denotes months with significant response function coefficients determined from the bootstrapped response model.



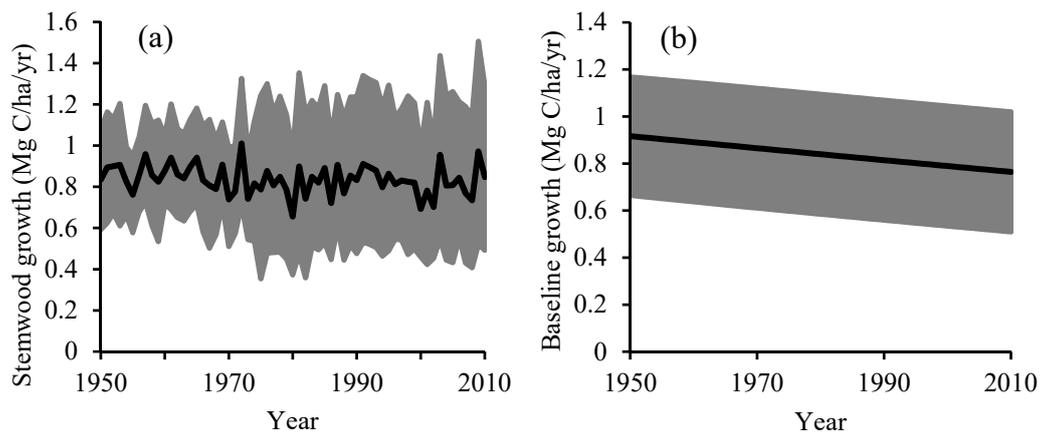
**Figure 5.** Raw growth indices (gray line), predicted growth indices (black dash line), and residual growth indices (black line), with Sen's regression line (red line) superimposed for four sites;  $Q$  is Sen's slope and  $B$  is a constant.

We then compared the residual indices with high frequency CO<sub>2</sub> interannual fluctuation (the linear trend was removed from the first order annual difference of the raw CO<sub>2</sub> concentration) (Figure 6). Two sites showing a statistically significant positive correlation between the CO<sub>2</sub> fluctuation and residual indices (Pearson's correlation coefficient ( $R$ ) from 0.26 to 0.33,  $p < 0.1$ ), except for MH ( $R = -0.01$ ,  $p > 0.05$ ) and MG sites ( $R = 0.08$ ,  $p > 0.05$ ), which also failed to detect a significant upward trend in their residual indices.



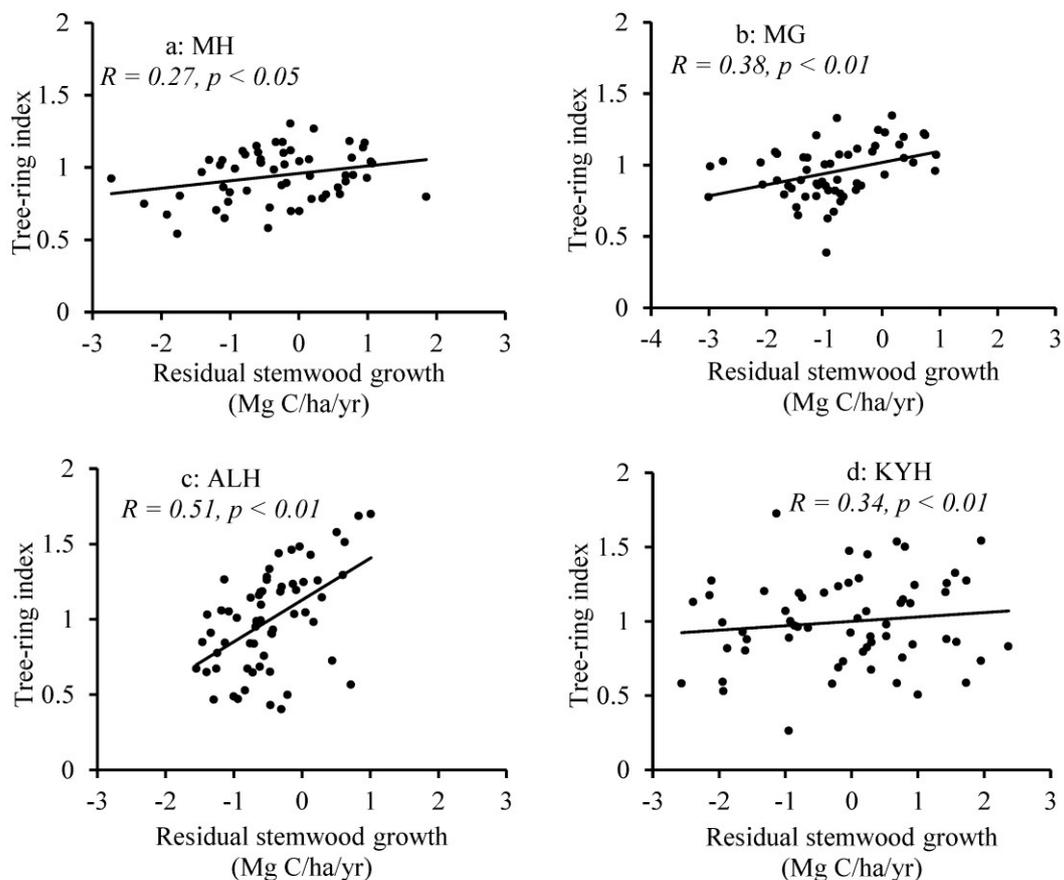
**Figure 6.** Comparison of the tree-ring residual indices and CO<sub>2</sub> interannual fluctuation (the first order annual difference of the CO<sub>2</sub> concentration) for four sites during the common period.

The relationship between stand age and baseline stemwood growth (B\_SW) for *Larix gmelinii* was produced by using yield tables (Figure 3). The B\_SW increased with stand age initially, reached a peak, then declined slowly at an approximately constant rate. For all sites, the curve presented a trend of slow decrease from 1950 to 2010 (Figure 7b), and it indicated that the stand age had a negative effect on stemwood growth since 1950.



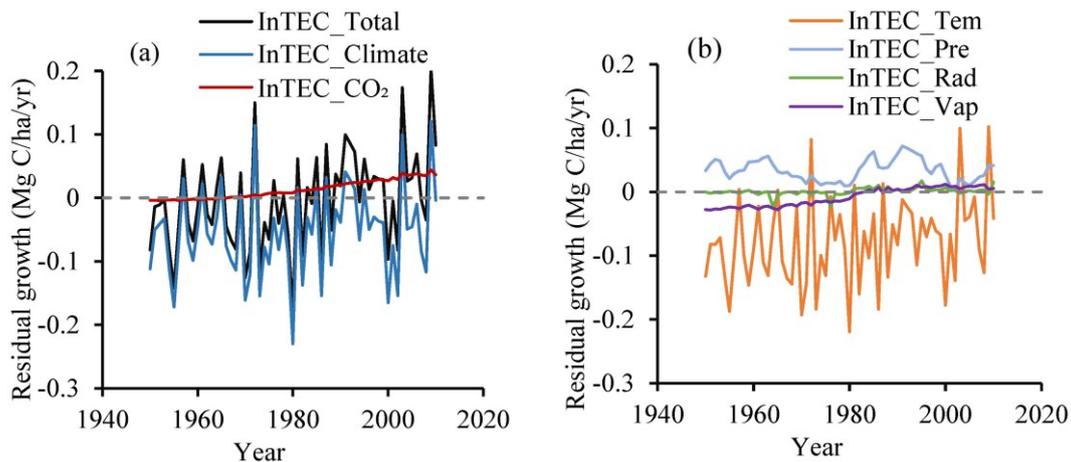
**Figure 7.** Temporal changes in stemwood growth ( $G_{SW}$ ) (a) and baseline stemwood growth ( $B_{SW}$ ) (b) from 1950 to 2010. Shaded area delineates the 90% confidence interval for the  $G_{SW}$  and  $B_{SW}$  estimates.

We removed the age effect from stemwood growth ( $G_{SW}$ ) (Figure 7a) to calculate the residual stemwood growth ( $R_{SW}$ ), and then we compared the  $R_{SW}$ s with the STD chronologies for each site, respectively. The  $R_{SW}$ s agreed well with those STD chronologies with a correlation coefficient from 0.27 to 0.51 ( $p < 0.05$ ) as shown in Figure 8. Through these site-level comparisons, there were reasons to believe that InTEC was a reliable model to link historical climate data and reconstruct the interannual variability of tree growth that was induced by external factors.



**Figure 8.** Comparison of the standard (STD) chronologies (black line) and the residual stemwood growth ( $R_{SW}$ ) simulated by the InTEC model (gray line) for four sites during the common period.

The contributions of climate and CO<sub>2</sub> on R\_SW were partitioned using the approach described in the methods section. Overall, 86% of accumulated R\_SW was mainly attributed to climate (Figure 9a). The CO<sub>2</sub> concentration increased from 346 ppmv in 1950 to 387 ppmv in 2010. The elevated CO<sub>2</sub> contributed to accumulated R\_SW by 14%. We also further separated the total climate effect into those from temperature, precipitation, radiation, and water vapor pressure (Figure 9b). Among the four climate factors, temperature was the dominating factor contributing 73% to the changes in the accumulated R\_SW. The second was precipitation, at 20%, followed by water vapor pressure at 4% and radiation at 3%.



**Figure 9.** Partition of the growth enhancement into contributions from climate and CO<sub>2</sub> fertilization (a), and the climate effect is partitioned to factors of temperature, precipitation, radiation, and water vapor pressure (b).

#### 4. Discussion

Trees from cold environments may show either positive or negative growth responses to warming depending on both inherent acclimation potential and other potentially limiting factors [51]. Our results indicate that high temperature generally has a negative effect on the growth of *Larix gmelinii* at four sites, particularly in June or July (Figure 4). This result agrees with similar studies about coniferous trees in northern China [12,52]. High temperature can cause an increase in transpiration, evaporation, moisture loss, and the internal nutrient consumption of the trees, and since dark respiration is a temperature-dependent process, warmer temperatures increase the respiration rates of plants, which results in decreases in the accumulation of carbohydrate content in plants, and consequently limits growth [53].

In our tree-ring analysis, two sites exhibited evidence of residual indices (after removing the effects of age and climatic variability) to increase in response to increasing CO<sub>2</sub> (Figure 5c,d). These results indicate that tree growth has a “CO<sub>2</sub> acclimation” response to CO<sub>2</sub> enhancement [18–20]. We failed to detect a statistically significant positive trend in the residual indices at the MH and MG sites (Figure 5a,b). These two sites are located at higher latitudes, and the CO<sub>2</sub> fertilization effect may be severely limited by lower temperatures (annual mean temperature was  $-5.1$  °C for MH site and  $-6.0$  °C for MG site, respectively) there. Many studies reported that low temperature in some cases might impair Rubisco activity and diminish the positive influence of higher CO<sub>2</sub> on photosynthesis [54,55]. In addition, lower temperature makes the growing season shorter, giving little opportunity for trees to allocate carbon to the cambial production of stem tissue even if limited growth enhancement has occurred [56].

Some consistency was found between the time series of CO<sub>2</sub> high-frequency fluctuation and residual indices at two sites ( $R$  from 0.26 to 0.33,  $p < 0.05$ ) (Figure 6c,d). This result indicates that interannual variations in tree growth was linked to the high-frequency fluctuation of CO<sub>2</sub> concentration.

Our result was similar to that of Chen et al. [12], who reported that there was a significant correlation between CO<sub>2</sub> high-frequency fluctuation and tree-ring indices in *Pinus tabulaeformis* in the northeast of China over 1950–2010. These results reinforce our findings that conventional dendrochronological techniques can detect the effect of CO<sub>2</sub> fluctuations on yearly tree growth in coniferous forests.

It is not easy to quantitatively separate the CO<sub>2</sub> fertilization effect on tree growth from those of external environment factors. The InTEC model can not only capture key ecosystem processes but also describe interactions of terrestrial ecosystems with environmental forcing factors. Also, systematic tree growth-age relationships were integrated into the InTEC model, making it possible to remove the influences of the intrinsic age factors [57]. Our model results indicate that 14% of accumulated R\_SW can be attributed to atmospheric CO<sub>2</sub> increase (Figure 9a). Although increased CO<sub>2</sub> contributed only a small fraction to the overall growth enhancement, it may be important to fully explain the carbon balance in mid to high northern latitude forests.

The CO<sub>2</sub> fertilization effects reported were not uniform across the forest sites investigated [11–16]. Stand characteristics might be partly responsible for this cross-site variability [7,58]. Our results agree with most reports for Canadian, American and Chinese forests in terms of tree-ring analysis or terrestrial carbon model [12,57,59], and these results collectively provided clear evidence for greater efficiency of CO<sub>2</sub> uptake in plants from mid and high altitudes of the northern hemisphere [60,61]. Our results also support another view that the CO<sub>2</sub> fertilization would be more evident in the environments with annual precipitation between 300 mm and 500 mm [62,63].

Based on the increasing CO<sub>2</sub> fertilization evidences from many short-term CO<sub>2</sub>-enriched experiments and dendrochronological tree-ring studies, we speculated the failures of some studies [15–17] in detecting the CO<sub>2</sub> fertilization effect to be due to the following possible reasons: (1) temperature-limiting and N-limiting regions could preclude a direct CO<sub>2</sub> growth response by trees [20]; (2) many biased sampling of trees could also produce spurious trends in growth rates (i.e., slow-grower survivorship bias and big-tree selection bias), consequently, a small increase in biomass increment due to the CO<sub>2</sub> fertilization effect may be diminished [64]; and (3) carbon partitioning and growth within a plant that is species-specific [20].

Our results also point to weaknesses and additional work in the future: First, the NPP allocation coefficient to stemwood growth was set to a constant for *Larix gmelinii* based on the calculation results by using forest inventory data and yield tables [48]. However, the allocation coefficient varies with forest types, stand age, and environment. Unfortunately, due to lack of enough information, we could not calculate NPP allocation coefficients for each site. The use of forest inventory allowed us to check whether the allocation coefficients was reasonable for *Larix gmelinii*. We observed a 17% underestimate in simulated stemwood growth from InTEC compared with the forest inventory data for the period 2000–2005. Clearly, significant improvements can still be made in the relative allocation to stemwood component once more data become available. Second, since it is difficult to exclude the impact of atmospheric deposition of N on tree-ring growth; we could not be sure that the residual growth in the tree ring data was entirely due to the CO<sub>2</sub> fertilization effect. Although many studies suggest the N deposition effect was small with a very weak trend over the study period [65–67], these estimates could be refined with additional N data. Third, due to limited availability of high quality tree ring data, we have access to data only for four sites. Although we have confidence in the detected CO<sub>2</sub> fertilization signals, data at more sites are needed to assess the latitudinal gradient of these signals.

## 5. Conclusions

In this study, we employed a series of approaches for detecting the effect of CO<sub>2</sub> fertilization on *Larix gmelinii* tree growth in Northeast China from 1950–2010. We conclude from tree-ring analysis that elevated atmospheric CO<sub>2</sub> had a positive effect on tree growth. Meanwhile, the change in annual tree growth was found to be significantly correlated to CO<sub>2</sub> interannual fluctuations. These results indicate that the CO<sub>2</sub> fluctuations influence the high-frequency growth variability, but not necessarily the overall tree growth. Our InTEC modeling results also suggest that the growth change over 61 years

cannot be explained by climate alone, 14% of accumulated R\_SW can be attributed to CO<sub>2</sub> increase, while the remaining approximately 86% can be attributed to climate.

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