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Multi-Scale Parameter Identification of Lithium-Ion Battery Electric Models Using a PSO-LM Algorithm

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Abstract: This paper proposes a multi-scale parameter identification algorithm for the lithium-ion battery (LIB) electric model by using a combination of particle swarm optimization (PSO) and Levenberg-Marquardt (LM) algorithms. Two-dimensional Poisson equations with unknown parameters are used to describe the potential and current density distribution (PDD) of the positive and negative electrodes in the LIB electric model. The model parameters are difficult to determine in the simulation due to the nonlinear complexity of the model. In the proposed identification algorithm, PSO is used for the coarse-scale parameter identification and the LM algorithm is applied for the fine-scale parameter identification. The experiment results show that the multi-scale identification not only improves the convergence rate and effectively escapes from the stagnation of PSO, but also overcomes the local minimum entrapment drawback of the LM algorithm. The terminal voltage curves from the PDD model with the identified parameter values are in good agreement with those from the experiments at different discharge/charge rates.

Keywords: multi-scale parameter identification; lithium-ion battery (LIB); particle swarm optimization (PSO); Levenberg-Marquardt (LM) algorithm

1. Introduction

Lithium-ion batteries (LIBs) have been widely utilized as power sources of electrical vehicles (EVs) and hybrid electrical vehicles (HEVs) in recent years, due to their advantages of higher energy-to-weight ratios, longer cycle life and lower environmental pollution [1,2]. This demand has fueled the need for the improved safety and performance of LIBs. Elaborate models have been proposed for the prediction of battery performance [3], such as the empirical model [3], single particle model (SPM) [4], extended SPM [5], pseudo two-dimensional (P2D) model [6], two-dimensional potential and current density distribution (PDD) model [7], and temperature distribution model [8]. Among these [6–8] are nonlinear distributed parameter systems (DPSs) [9–11] and can describe the electrical and thermal performance more accurately. On the other hand, an accurate model of LIBs requires knowledge of a great number of physical properties; these are difficult to obtain directly. In this situation, unknown parameter values can be identified indirectly from measurements.

Several different techniques including the gradient method [4], the gradient-free method [12–14] and Kalman filter method [15,16] have been proposed for the parameter estimation or identification of LIBs. Among these methods, the gradient method and gradient-free method have been applied to identify the parameters of electrochemical models. The Levenberg-Marquardt (LM) algorithm is a gradient-based nonlinear regression method and can converge quickly to the optimum. Santhanagopalan et al. [4] employed the LM algorithm to identify the parameters of the SPM. However,

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this algorithm suffers from the problem of local minimum entrapment caused by inappropriate initial parameters [17]. Unlike the LM, some gradient-free methods have also been proposed in order to determine the global minimum for the objective function, e.g., genetic algorithm (GA) and particle swarm optimization (PSO). GA [12] was utilized to solve the parameter identification of the P2D model. The PSO algorithm was also applied in order to identify the parameters of the electrochemical model [13]. However, the stagnation problem of PSO and GA makes them extremely slow around the global optimum. Ref. [18] found that LM is better at finding the optimum with appropriate initial values than PSO. To overcome the drawbacks of LM and PSO, a hybrid PSO-LM algorithm was proposed in [19] so as to train the weights and threshold of a neural network for the nonlinear modeling of the fuel cell. Recently, an accelerated PSO algorithm based on LM was proposed in [20]. However, up to now, few PSO-LM based parameter identification algorithm has been proposed for the LIB models.

This study proposes a multi-scale parameter identification algorithm combining PSO and LM in order to globally optimize the parameter values for the two-dimensional uneven PDD battery electric model [7]. In the proposed identification algorithm, a hybrid multi-swarm PSO algorithm [21] is first applied for the coarse-scale searching in the global space to find near-optimal parameter values. Then LM is embedded for the fine-scale parameter identification in the vicinity of the optimum within the local space. Simulation results and the corresponding analysis are provided to demonstrate the effectiveness of the proposed multi-scale algorithm. The rest of this paper is structured as follows. Section 2 presents a two-dimensional PDD battery electric model description and the multi-scale parameter identification complexity for the nonlinear spatiotemporal electric model. The simulation results with optimized parameters are compared with the measurements in Section 4 which is followed by a detailed parameter analysis. Finally, Section 5 offers some concluding remarks.

2. Battery Electric Model and Problem Formulation

This study considers a simple cell consisting of two parallel plate electrodes and assumes that the distance between the electrodes is extremely small. Figure 1 shows a schematic diagram of the current flow in the cell during discharge and charge. Since the distance between the electrodes is very small, the current flow between the electrodes is perpendicular to the electrodes. However, see Figure 1, the thickness along axis *z* is magnified to provide a clear reference. From the continuity of the current on the electrodes during discharge, the potential density distributions in the positive and negative electrodes are described by two Poisson equations in the two-dimensional spatial domain [7], respectively:

$$\nabla^2 \phi_p = -r_p J_c \text{ in } \Omega_p \tag{1}$$

$$\nabla^2 \phi_n = +r_n J_c \text{ in } \Omega_n \tag{2}$$

where ϕ_p and ϕ_n are the potentials (*V*) of positive and negative electrodes, respectively, r_p and r_n are the resistances (Ω) of the positive and negative electrodes, respectively, and J_c is the current density, which is current per unit area (A m⁻²). Ω_p and Ω_n are the domains of positive and negative electrodes, respectively. For the relevant boundary conditions of ϕ_p and ϕ_n , please refer to [7]. The difference between the governing equations for charge and those for discharge is that the signs in front of J_c in (1) and (2) are opposite according to the previous studies [22].



Figure 1. Schematic diagram of the current flow in the parallel plate electrodes. — for discharge; and — for charge.

The current density J_c is a function of the potential difference between positive and negative electrodes and is expressed as [23,24]:

$$J_c = Y_x(\phi_p - \phi_n - U) \tag{3}$$

where $\phi_p - \phi_n$ is the potential difference, Y_x and U are the fitting parameters, as suggested by Gu [25], these two parameters Y_x and U depend on the depth of discharge (*D*) and are expressed by the following expressions:

$$U = x_1 + x_2 D + x_3 D^2 + x_4 D^3 \tag{4}$$

$$Y_x = x_5 + x_6 D + x_7 D^2 \tag{5}$$

where x_1, x_2, \dots, x_7 are fitting parameters. When the battery is in a discharging process, the distribution of *D* on the electrodes is calculated from the integration of *J_c* as [7,26]:

$$D = \frac{1}{Q_T} \int_{t_0}^t J_c dt \tag{6}$$

where *t* is the discharge time (s), t_0 is discharge starting time (s), and Q_T is the theoretical capacity per unit area (Ah m⁻²) of the electrodes. When the battery is in a charging process, the distribution of *D* is calculated as follows [22]:

$$D = D_0 - \frac{1}{Q_T} \int_{t_0}^t J_c dt$$
 (7)

where D_0 is the initial value of D, t is the charge time (s), t_0 is charge starting time (s). Set $x \triangleq [x_1, x_2, \dots, x_7]^T$ for the sake of brevity. Due to the hysteresis behavior existed battery [27], the voltage curves during discharge and discharge process are fitted separately. Thus, two groups of parameters x and x^* will be obtained.

The described battery electric model is a time/space coupled high-dimensional nonlinear system [9]. The proper value of parameter vector x in the seven-dimensional hyperspace is usually difficult to obtain in the simulation [28]. The "art" of trial and error [29] was adopted to determine the value of x. However, it requires exhaustive enumeration to search the optimal value of x. Therefore,

it is valuable and interesting to investigate an efficient parameter identification algorithm to obtain the optimal value of *x*.

Although the battery electric model is extremely complicated with high-dimensional nonlinearity, it can still be approximately linearized in a neighborhood of any given operating point based on the theory of linear system. Thus, the identification can be considered as a multi-scale problem, for which a multi-scale approach could be developed as follows.

Multi-Scale Parameter Identification

The multi-scale properties can be described geometrically with a 3-dimensional example as in Figure 2. Due to the highly nonlinear properties, traditional optimization cannot be applied for the direct finding of the global optimum point *C* from a random start point *A*. However, the non-math-based coarse method can be easily used for coarse searching on a global scale. Once the proper point *B* near the optimum is found, then the model can be linearized at the local scale around the point *B*. Through this quantitative searching, the optimum point *C* can be found.



Figure 2. Geometric explanation for the multi-scale complexity.

3. Weighted PSO-LM Algorithm for Multi-Scale Parameter Identification

To address the problem of the multi-scale (global and local) complexity existing in the parameter identification, this section will provide a weighted PSO-LM combined algorithm for the multi-scale (coarse and fine) identification of parameter vector *x*. PSO is firstly utilized to implement the *coarse-scale* parameter identification to find the near-optimal point in the global space. Then the LM algorithm is embedded in PSO to conduct the *fine-scale* parameter identification in the vicinity of the optimum within the local space. The block diagram of the PSO-LM based identification approach is illustrated in Figure 3.



Figure 3. Block diagram of PSO-LM based identification approach.

3.1. Search Strategy

The search strategy determines whether PSO or LM is suitable for the next step identification. An objective/fitness function f is embedded in the search strategy as a criterion for the decision making. The strategy consists of two steps:

- Coarse searching: The identification procedure is assured to start from PSO by initializing a group of random parameter vectors in the global parameter space. After several generations of PSO optimization, a near-global optimum is acquired;
- (2) Fine searching: The identification procedure is then switched to LM by starting from the near-global optimum with the gradient search in a local space. The identified parameter vector is compared with the one obtained in PSO to select a better candidate to enter the next iteration.

The objective/fitness function f is constructed for the optimal identification of the parameter vector x in the battery electric model (1)–(7), which is a quadratic function of the difference between the experimental terminal voltage ϕ_k and the predicted voltage $\hat{\phi}_k$ with additional parameters $w_k > 0$, $k \in \{1, 2, \dots, N\}$:

$$f = \sum_{k=1}^{k=N} w_k (\phi_k - \hat{\phi}_k)^2 = (\boldsymbol{\phi} - \hat{\boldsymbol{\phi}})^T \boldsymbol{W} (\boldsymbol{\phi} - \hat{\boldsymbol{\phi}})$$
(8)

where *N* is the number of fitting curves, $\boldsymbol{\phi} \triangleq [\phi_1, \phi_2, \cdots, \phi_N]^T$ and $\hat{\boldsymbol{\phi}}$ is the estimation of $\boldsymbol{\phi}$, both are vectors consisting of ϕ_k and $\hat{\phi}_k$, $k \in \{1, 2, \cdots, N\}$, respectively, $\boldsymbol{W} \triangleq \text{diag}\{w_1, w_2, \cdots, w_N\}$ is the parametric matrix assigned to obtain a better fitness.

In what follows, we will define the specified form of the parameter matrix W, whose element w_k is applied to the *k*-th experimental terminal voltage curve:

$$w_k = \gamma_k \chi_k \tag{9}$$

where γ_k takes into account the experimental deviations in the *k*-th voltage curve and is defined as:

$$\gamma_k = \frac{1}{\frac{\sigma_k^2}{\sum\limits_{k=1}^N N_k} \left[\sum\limits_{k=1}^N \sum\limits_{j=1}^{N_k} \frac{1}{\sigma_j^2}\right]}$$
(10)

with $\sigma_k \triangleq \sqrt{\frac{1}{N_k - 1} \sum_{j=1}^{N_k} (\hat{\phi}_{k,j} - \phi_{k,j})^2}$, in which $\hat{\phi}_{k,j}$ is the predicted voltage of the *j*-th point in the *k*-th

curve and obtained by calculating the difference between $\hat{\phi}_{k,j}^p$ and $\hat{\phi}_{k,j}^n$, $\phi_{k,j}$ is the experimental voltage at the *j*-th point in the *k*-th curve, N_k is the number of experimental data points in the *k*-th voltage curve. The parameter χ_k is proportional to the difference between the previous voltage predictions and the experimental data in the *k*-th voltage curve and is defined as:

$$\chi_k = 1 + \frac{(w_{k,\max} - 1)\eta_k}{\eta_{k,\max}} \tag{11}$$

where $w_{k,\max} > 1$ is a given constant, $\eta_k \triangleq |\hat{\phi}_k - \phi_k|$, $\eta_{k,\max}$ is the maximum value of η_k for each $k \in \{1, 2, \dots, N\}$. Obviously, the parameter χ_k varies between 1 and $w_{k,\max}$, that is, $\chi_k = 1$ for the best fitting points and $\chi_k = w_{k,\max}$ for the worst fitting ones.

3.2. Implementation of PSO-LM Algorithm

The PSO algorithm adopted here is a variation of conventional PSO, named hybrid multi-swarm PSO (HMPSO). The HMPSO method divides the swarm into several sub-swarms, and adopts a parallel PSO search operator for the sub-swarms [30]. For a particle swarm with a population consisting of

M \overline{n} -dimensional particles, the velocity $v_{i,j}^m$ and position $x_{i,j}^m$, $i \in \{1, 2, \dots, M\}$, $j \in \{1, 2, \dots, \overline{n}\}$ of the *j*-th dimension of the *i*-th particle are updated as [30]:

$$\begin{cases} v_{i,j}^{m+1} = |\gamma_1|(\xi_{i,j}^m - \mathbf{x}_{i,j}^m) + |\gamma_2|(\boldsymbol{\psi}_{i,j}^m - \mathbf{x}_{i,j}^m) \\ x_{i,j}^{m+1} = x_{i,j}^m + v_{i,j}^{m+1} \end{cases}$$
(12)

where $\xi_{i,j}^m$ and $\psi_{i,j}^m$ are the best previous position of $x_{i,j}^m$ and the best position achieved with its neighbors at the *m*-th generation, respectively; γ_1 and γ_2 are two separately generated random numbers with uniform distribution in the range of [0, 1]. The privilege of HMPSO lies in that it can explore more promising regions of the search space by applying differential evolution (DE) to update the personal best of each particle. In conclusion, it is a more competitive and effective PSO method for solving optimizing problems. For more details on the HMPSO algorithm, please refer to ref. [30].

The LM algorithm has a strong ability to find a local and more optimistic result. The parameter correction vector Δx is obtained based on the objective Function (8) and the Marquardt method as follows:

$$\Delta \mathbf{x} = \left(\mathbf{J}^T \mathbf{W} \mathbf{J} + \lambda \mathbf{I}\right)^{-1} \mathbf{J}^T \mathbf{W} (\boldsymbol{\phi} - \hat{\boldsymbol{\phi}})$$
(13)

where *J* is a matrix of partial derivatives of the terminal voltage with respect to the fitting parameter vector *x* evaluated at all the experimental voltage data. LM adaptively alters the algorithmic parameter value λ updates between the gradient descent method and the Gauss-Newton method. The parameter λ determines how the LM algorithm works and is initialized to be large. If the iteration results in a better approximation, then the parameter λ is decreased to 0.1λ and LM is more like a Gauss-Newton update one. If the iteration provides a worse approximation, then the parameter λ is increased to 10λ and LM approaches that of the gradient descent update. The details of LM can refer to [17].

The proposed PSO-LM algorithm is a hybrid multi-scale approach to determine the value of x in the battery model (1)–(7) based on the fitness Function (8). It not only improves the convergence rate and effectively escapes from the stagnation of PSO, but also overcomes the drawback of the local minimum entrapment of LM. The pseudo code of the suggested weighted PSO-LM algorithm is shown in Table 1. The coarse-searching procedure of PSO can be found from Step 1 to Step 6, and the fine-searching procedure of LM is embedded from Step 7 to Step 16.

Step 1	Initialize the PSO population size, dimensions, and termination conditions, i.e., randomly generate an initial swarm P^0 consisting of M seven-dimensional particles, $\mathbf{x}_i^0 = \begin{bmatrix} x_{i,1}^0 & x_{i,2}^0 & \cdots & x_{i,7}^0 \end{bmatrix}^T$, $i \in \{1, 2, \cdots, M\}$, optimization error (f_{set}); Initialize the LM parameters, i.e., λ and the maximum number of iteration R_{max} ;
Step 2	Calculate the potential $\hat{\phi}_i$, parameter matrix W , and the fitness value $f(x_i^0)$ given in (7) for each particle, $i \in \{1, 2, \dots, M\}$;
Step 3	Record particle's personal bests, i.e., $\xi_1^0, \xi_2^0, \cdots, \xi_M^0$;
Step 4	while (fitness value $< f_{set}$) do
Step 5	Update the personal best ξ_i^m at <i>m</i> -th generation using the differential evolution (DE), and evaluate the fitness value $f(\xi_i^m)$ for the personal best;
Step 6	Split the swarm P^m into several sub-swarms, and each sub-swarm evolves in parallel according to the governing equations of the particles' velocity and position: $v_i^{m+1} = \gamma_1 (\xi_i^m - x_i^m) + \gamma_2 (\psi_i^m - x_i^m)$ $x_i^{m+1} = x_i^m + v_i^{m+1}$ where γ_1 and γ_2 are two scalars generated randomly in the range of (0, 1), and ψ_i^m is the best position achieved with its neighbors;

Table 1. Implementation of PSO-LM algorithm.

Step 7	Calculate the potential $\hat{\phi}_i$, parameter matrix W , and the fitness value $f(x_i^{m+1})$ for each particle $x_i^{m+1}, i \in \{1, 2, \dots, M\}$; Let $x_i^* \leftarrow x_i^{m+1}$ and $f_i^* \leftarrow f(x_i^{m+1})$; Set x_i^* be the initial values of LM algorithm;
Step 8	for $i = 1 : M$ do
Step 9	while (<i>iter</i> < R_{max} and $f_i^* > f_{\text{set}}$) do
Step 10	Calculate Jacobin matrix $J_i = [\partial \hat{\phi}_i / \partial x_0, \partial \hat{\phi}_i / \partial x_1, \dots, \partial \hat{\phi}_i / \partial x_6]$, parameter matrix W , Hessian matrix $H_i = J_i^T W_i J_i$; Calculate the potential $\hat{\phi}_i$ and f_i^* at current particle x_i^* ;
Step 11	Update the Hessian matrix $H_i = H_i + \lambda I$, the parameter correction vector $\Delta x_i = H_i^{-1} J_i^T (\hat{\phi}_i - \phi_i)$, the particle position $x_i^{lm} = x_i^* + \Delta x_i$, and the potential $\hat{\phi}_i$, the fitness value f_i^{lm} at the current particle x_i^{lm} ;
Step 12	if $f_i^{lm} < f_i^*$ do $\lambda = \lambda/10$, $x_i^* \leftarrow x_i^{lm}$, $f_i^* \leftarrow f_i^{lm}$ and go back to Step 10; else $\lambda = \lambda \times 10$, go back to Step 11; end if
Step 13	iter = iter + 1;
Step 14	end while
Step 15	$f(\mathbf{x}_i^{m+1}) \leftarrow f_i^*, \ \mathbf{x}_i^{m+1} \leftarrow \mathbf{x}_i^*;$ i = i + 1;
Step 16	end for
Step 17	m = m + 1;
Step 18	end while
Step 19	Output the identified parameter vector <i>x</i> .

Table 1. Cont.

4. Experimental Validation

4.1. Experimental Setup

To illustrate the validity of the proposed PSO-LM algorithm, discharge experiments are carried out for a 60 Ah LIB (LiFePO₄) at a constant temperature of 25 °C on the experimental platform shown in Figure 4a. This platform consists of a thermal chamber, lithium-ion battery, battery testing system (BTS-300 A/60 V, Shenzhen Neware Technology, Shenzhen, China), battery management system (BMS) and computer. The schematic diagram of the signal and data flow for the experimental platform is also shown in Figure 4a. Figure 4b shows the dimensions of electrodes and positions of the tabs of a 60 Ah LIB. The discharge and charge tests of the battery are completed on this platform. The battery is first fully charged and then discharged at various discharging rates namely 1C (60 A), 2C (120 A) and 3C (180 A) until the cut-off voltage. The experimental terminal voltage data is collected per second during its discharge. Subsequently, the battery is fully discharged and then charged at a salso collected during the charging process. Due to the hysteresis behavior existed in battery, the voltage curves during discharge do not agree with that during charge process and thus are fitted separately [27]. There will be two sets of parameters $x_1 - x_7$ and $x_1^* - x_7^*$ for these two processes, respectively.



Figure 4. (a) Experimental platform and its schematic diagram of the signal and data flow; and (b) dimensions of electrodes and positions of the tabs of a 60 Ah LIB.

4.2. Numerical Calculation and Parameter Setup

The numerical solutions of the battery electric model (1)–(7) subjected to the associated boundary conditions are obtained by COMSOL (COMSOL Inc., Stockholm, Sweden), which is a commercial software package for accurate numerical simulation of partial differential equations (PDEs) using the finite element method. The simulation results and identified parameters are exchanged through the interface of COMSOL with MATLAB (The MathWorks Inc., Natick, MA, USA). In the proposed PSO-LM algorithm, set $\lambda = 0.1$, $R_{\text{max}} = 20$, M = 16, and for the parameter matrix W, set N = 3, $w_{1,\text{max}} = 4$, $w_{2,\text{max}} = 8$ and $w_{3,\text{max}} = 10$.

4.3. Results and Discussion

The value of the parameter vector x in the battery electric model (1)–(7) is identified using the proposed algorithm through fitting the three experimental terminal voltage curves simultaneously. Table 2 gives the identified value of the parameter vector x. As shown in Figure 5a, the predictions with identified parameter vector x and measurements of the terminal voltage curves are well matched

with each other, and this demonstrates the effectiveness of the proposed method. Figure 5b depicts the corresponding trajectory of the weight matrix $W = \text{diag}\{w_1, w_2, w_3\}$.



Table 2. The identified value of the parameter vector *x* during discharge.

Figure 5. Comparison between experimental and predicted discharge voltage curves at discharge rates of 1C, 2C, and 3C using three different identified algorithms: (a) weighted PSO-LM; (b) trajectory of the weight matrix $W = \text{diag}\{w_1, w_2, w_3\}$; (c) PSO-LM with $W = \text{diag}\{1, 1, 1\}$; (d) PSO; (e) LM and (f) method of trial and error.

The potential distributions on the positive and negative electrodes during discharge are obtained as a function of time for various discharge rates of 1C, 2C, and 3C. For example, Figure 6 indicates the potential distributions on the positive and negative electrodes with discharge rates of 1C, 2C, and 3C at discharge time t = 14 min, respectively. Because all the current flows into the tab from the entire electrode plate, the potential gradient on the positive electrode shown in Figure 6 is seen to be most severe in the region near to the tab. While the potential gradient on the negative electrode is also the highest in the region near tab. This is because all the current has to flow from the tab through the entire electrode plate.



Figure 6. Potential distributions of positive electrode and negative electrode for the discharge rate of (a) 1C; (b) 2C; (c) 3C at the discharging time t = 14 min.

To demonstrate the advantage of the proposed weighted PSO-LM algorithm, a comparison study is also conducted among PSO, PSO-LM with $W = \text{diag}\{1, 1, 1\}$, weighted PSO-LM, LM and the method of trial and error. Figure 5 also compares the experimental and predicted terminal voltage curves at discharge rates of 1C, 2C, and 3C through the simultaneous fit by PSO-LM with $W = \text{diag}\{1, 1, 1\}$ (Figure 5c), PSO (Figure 5d), LM (Figure 5e) and method of trial and error adopted in [29] (Figure 5f). It is clear from Figure 5 that the proposed weighted PSO-LM algorithm provides a more accurate discharging curve fitness compared to the other algorithms. Moreover, although the fitting result of PSO-LM with $W = \text{diag}\{1, 1, 1\}$ is better than that of PSO, it is worse than that of weighted PSO-LM. This is because the same weights are assigned to different curves but without considering the difference of the data number of each voltage curve (e.g., there are 2860 data points for the 1C-discharging voltage curve, 1360 data points for 2C and 850 for 3C). To evaluate the parameter importance, an a-priori sensitivity analysis is carried out for the identified parameters [27]. Each parameter is tuned to 0.5 times and 1.5 times the identified value. The results, shown in Figure 7, indicate that for the same range of parameter values, x_6 and x_7 are less sensitive than other parameters and hence can be held constant, while the other five parameters $x_1 - x_5$ are very sensitive and deserve more attention during the identification.

A further quantized research needs to be carried to determine the importance order of parameters $x_1 - x_5$. On the other hand, Figure 7 shows that x_1 and x_5 have the similar effect toward model output, while x_2 , x_5 and x_4 have the similar effect.



Figure 7. Cont.





Figure 7. Sensitivity of the PDD model to parameter (**a**) *x*₁; (**b**) *x*₂; (**c**) *x*₃; (**d**) *x*₄; (**e**) *x*₅; (**f**) *x*₆ and (**g**) *x*₇.

4.3.1. 95% Confidence Interval

From a statistical point of view, it is very useful to obtain the confidence interval instead of the point estimation for the fitting parameter vector x [31]. In this paper, the 95% confidence interval of the parameter vector x is calculated as follows:

$$x_i^* - \mathcal{T}_{(1-0.05/2)} S_E \sqrt{a_{ii}} \le x_i \le x_i^* + \mathcal{T}_{(1-0.05/2)} S_E \sqrt{a_{ii}}$$
(14)

where x_i^* is the point estimation of the parameter x_i , $i \in \{1, 2, \dots, 7\}$, $\mathcal{T}_{(1-0.05/2)}$ is a value of \mathcal{T} -distribution with $(\overline{N} - 7)$ degrees of freedom, S_E is the unbiased estimate of unknown variance σ and is calculated by $S_E = \sqrt{\frac{1}{\overline{N} - 7} \sum_{\zeta=1}^{N} \sum_{j=1}^{N} (\phi_{\zeta,j} - \hat{\phi}_{\zeta,j})^2}$, \overline{N} is the number of experimental data points and derived by $\overline{N} = \sum_{\zeta=1}^{N} N_{\zeta}$, and a_{ii} is the *i*-th main diagonal element of $(J^T J)^{-1}$.

To demonstrate the goodness of the simultaneous fit, the model (1)–(7) is also to fit each experimental curve independently for comparison. The 95% confidence intervals of all seven parameters obtained from both the simultaneous fit and three independent fits are presented in Table 3. Although Table 3 shows that each independent fit leads to a smaller S_E compared to the simultaneous fit, the value of S_E for the simultaneous fit is also acceptable.

Parameter	Simultaneous Fit	Independent Fit (1C)	Independent Fit (2C)	Independent Fit (3C)	
x_1	3.125 ± 0.0012	3.094 ± 0.02895	2.844 ± 0.1022	2.619 ± 6.178	
<i>x</i> ₂	$(-5.077 \pm 0.1509) \times 10^{-1}$	$(-2.929 \pm 1.834) imes 10^{-1}$	$(-3.305 \pm 4.461) imes 10^{-1}$	-0.3632 ± 2.268	
<i>x</i> ₃	$(1.492 \pm 0.4326) imes 10^{-1}$	1.092 ± 0.4250	1.373 ± 0.05091	1.444 ± 1.481	
x_4	$(-1.768 \pm 0.3626) \times 10^{-1}$	-1.465 ± 0.2131	-1.748 ± 0.04607	-1.723 ± 0.4383	
<i>x</i> ₅	12860 ± 13.14	$(1.184 \pm 0.3005) imes 10^4$	$(1.017\pm 0.3847)\times 10^4$	$(1.043\pm 16.28)\times 10^4$	
<i>x</i> ₆	$(-0.6177 \pm 4.021) imes 10^2$	$(6.047 \pm 21.10) imes 10^3$	$(4.049 \pm 2.459) imes 10^3$	$(4.429 \pm 77.02) imes 10^3$	
<i>x</i> ₇	$(-2.377 \pm 0.7177) \times 10^3$	$(2.298 \pm 24.41) imes 10^3$	$(-5.160 \pm 3.377) \times 10^2$	$(-2.390 \pm 55.73) \times 10^3$	
S_E	0.0079	0.0019	0.0037	0.0041	

Table 3. Comparison of the 95% confidence intervals of *x* estimated from the simultaneous fit and the independent fit.

4.3.2. 95% Joint Confidence Region

We know from [17] that the individual confidence interval cannot reflect the correlations between the parameters in the nonlinear system. Hence it is necessary to construct the joint confidence region. The 95% joint confidence regions of the parameter vector x are calculated by Equation (15) [17]:

$$\frac{(\boldsymbol{x}-\hat{\boldsymbol{x}})^{T}(\boldsymbol{J}^{T}\boldsymbol{J})(\boldsymbol{x}-\hat{\boldsymbol{x}})}{7{S_{E}}^{2}} \le F_{1-\alpha}(7,\overline{N}-7)$$
(15)

where *x* is any point in the confidence region, \hat{x} is the point estimation value, $F_{1-\alpha}(7, \overline{N} - 7)$ is the quantile function of the F-distribution with 7 and $\overline{N} - 7$ degree of freedom, α is the statistical significance level. Set $\alpha = 0.05$, we get from the *F*-distribution table that $F_{0.95}(7, \overline{N} - 7) = 2.0986$. Without loss of generality, set \hat{x} to be the estimate of *x* from weighted PSO-LM, i.e., $\hat{x} = x^*$, we have:

$$\boldsymbol{x} - \hat{\boldsymbol{x}} = \boldsymbol{x} - \boldsymbol{x}^* = \begin{bmatrix} x_1 - 3.125 \\ x_2 + 5.077 \times 10^{-1} \\ x_3 - 1.492 \\ x_4 + 1.768 \\ x_5 - 1.286 \times 10^4 \\ x_6 + 6.177 \times 10^1 \\ x_7 + 2.377 \times 10^3 \end{bmatrix}$$
(16)

The confidence regions are obtained from the following equation:

$$(x_i - x_i^*)^2 (J^T J)_{ii} \le F_{0.95}(7, \overline{N} - 7)7S_E^2$$
(17)

The confidence region of x_1 is obtained from Equation (16) by fixing the parameters x_2, x_3, \dots, x_7 to their estimation value. The 95% joint confidence regions of all seven parameters obtained from the simultaneous fit are given in Table 4. The comparison of 95% joint confidence regions and 95% conference intervals of three simulated voltage curves is plotted in Figure 8. It can be observed from Table 4 and Figure 8 that the values of the parameter vector x defined by the confidence region lead to less uncertainty in model predictions than those defined by the confidence interval.

Table 4. The 95% joint confidence regions (CR) obtained from the simultaneous fit.

Parameter	<i>x</i> ₁	<i>x</i> ₂	<i>x</i> ₃	<i>x</i> ₄	
CR	$3.125\pm 6.762\times 10^{-4}$	$-(5.077\pm0.015)\times10^{-1}$	$1.492 \pm 2.355 \times 10^{-3}$	$-1.768 \pm 3.442 \times 10^{-3}$	
Parameter	<i>x</i> ₅	<i>x</i> ₆	<i>x</i> ₇		
CR	$1.286 \times 10^{4} \pm 13.83$	$(-6.177\pm 6.78)\times 10^{1}$	$(-2.377\pm0.1052)\times10^{3}$		



Figure 8. Comparison of the predicted voltage curves at different discharge rates (**a**) 1C; (**b**) 2C; (**c**) 3C, using different limits of the parameter x_1 from the 95% confidence interval and 95% joint confidence region. Point estimates obtained from the simultaneous fit were used for the rest parameters. LCIL and UCIL represent the lower and upper confidence interval limits, respectively; LJCR and UJCR represent the lower and upper joint confidence region limits, respectively.

As shown in Tables 3 and 4, the confidence interval and confidence region of parameter x_6 are much larger than any of other six parameters. This phenomenon is caused by the parameter correlations as explained by Evans and White [32]. Thus, a thorough correlation analysis of these seven parameters is conducted in what follows.

4.3.3. Correlation Analysis

The correlation coefficient matrix R of the parameters x_i , $i \in \{1, 2, \dots, 7\}$ is a symmetric matrix. Its elements r_{ij} have all their values in the range $-1 \le r_{ij} \le 1$, are calculated as follows:

$$r_{ij} = \frac{a_{ij}}{\sqrt{a_{ii}a_{jj}}}, \ i, j \in \{1, 2, \cdots, 7\},\tag{18}$$

Let matrix $H = J^T J$, then a_{ij} , $i, j \in \{1, 2, \dots, 7\}$ are the elements of H^{-1} after the simultaneous fit, and a_{ii} is the *i*-th main diagonal element of H^{-1} . The matrix H is calculated as:

$$H = J^{T}J = \begin{bmatrix} 2.485 \times 10^{3} & 9.715 \times 10^{2} & 5.069 \times 10^{2} \\ 9.715 \times 10^{2} & 5.322 \times 10^{2} & 3.193 \times 10^{2} \\ 5.069 \times 10^{2} & 3.193 \times 10^{2} & 2.049 \times 10^{2} \\ 2.988 \times 10^{2} & 2.056 \times 10^{2} & 1.379 \times 10^{2} \\ -5.855 \times 10^{-2} & -3.445 \times 10^{-2} & -2.294 \times 10^{-2} \\ -1.896 \times 10^{-2} & -1.080 \times 10^{-2} & -6.631 \times 10^{-3} \\ -1.026 \times 10^{-2} & -6.685 \times 10^{-3} & -4.372 \times 10^{-3} \\ 2.988 \times 10^{2} & -5.855 \times 10^{-2} & -1.896 \times 10^{-2} & -1.026 \times 10^{-2} \\ 2.056 \times 10^{2} & -3.445 \times 10^{-2} & -1.080 \times 10^{-2} & -6.685 \times 10^{-3} \\ 1.379 \times 10^{2} & -2.294 \times 10^{-2} & -6.631 \times 10^{-3} & -4.372 \times 10^{-3} \\ 9.592 \times 10^{1} & -1.664 \times 10^{-2} & -4.346 \times 10^{-3} & -2.990 \times 10^{-3} \\ -1.664 \times 10^{-2} & 5.942 \times 10^{-6} & 8.343 \times 10^{-7} & 5.756 \times 10^{-7} \\ -4.346 \times 10^{-3} & 8.343 \times 10^{-7} & 2.471 \times 10^{-7} & 1.545 \times 10^{-7} \\ -2.990 \times 10^{-3} & 5.756 \times 10^{-7} & 1.545 \times 10^{-7} & 1.027 \times 10^{-7} \end{bmatrix}$$
(19)

Hence, the matrix *R* is obtained according to Equation (18):

$$\boldsymbol{R} = \begin{bmatrix} 1 & -0.7211 & 0.6311 & -0.5489 & -0.1471 & 0.0731 & -0.0592 \\ -0.7211 & 1 & -0.9352 & 0.7559 & 0.5082 & 0.3639 & -0.3952 \\ 0.6311 & -0.9352 & 1 & -0.9278 & -0.6087 & -0.1230 & 0.1899 \\ -0.5489 & 0.7559 & -0.9278 & 1 & 0.5769 & -0.2034 & 0.1479 \\ -0.1471 & 0.5082 & -0.6087 & 0.5769 & 1 & 0.1564 & -0.2760 \\ 0.0731 & -0.3639 & -0.1230 & -0.2034 & 0.1564 & 1 & -0.9640 \\ -0.0592 & -0.3952 & 0.1889 & 0.1479 & -0.2760 & -0.9640 & 1 \end{bmatrix}$$
(20)

The element r_{ij} stands for the correlation between the *i*-th parameter and the *j*-th parameter, where *i* stands for x_i , $i \in \{1, 2, \dots, 7\}$. As pointed out in [17], the absolute value of the elements of R is closer to 1, the correlation between two parameters is higher. It is observed from the matrix R given in Equation (20) that the values of all the main diagonal elements of R are equal to 1. This indicates that each parameter is highly correlated with itself. We also observe from Equation (20) that the highest correlation between two different parameters occurs to the $x_6 - x_7$ pair. A positive correlation coefficient between two parameters indicates that the errors causing the estimate of one parameter to be high also cause the other to be high, and vice versa. It is not difficult to conclude from $r_{67} = -0.9640$ that an underestimation of x_6 will cause an overestimation of x_7 . Moreover, the eigenvectors and eigenvalues of R can be applied to construct a hyperellipsoidal confidence region of the parameter space in the vicinity of the solution, where the model can be approximated linearized [17]. On the other hand, the high correlation between parameters also implies that it is difficult to obtain separate estimates of these parameters with the available data. The regression must be repeated and restarted from different initial parameter values many times to bypass the local minimum. While the multi-scale (global and local) approach provides a possible solution to find the global optima in this study.

4.4. Fitting to Charge Curves

The weighted PSO-LM algorithm is also applied to identify the parameters of $x_1^* - x_7^*$ during the charge process. The identified parameter values are in Table 5.

Parameter	x_1^*	<i>x</i> [*] ₂	<i>x</i> ₃ *	x_4^*	x_5^*	<i>x</i> ₆ *	x_{7}^{*}
Identified value	3.370	-1.033	2.142	1.459	$1.078 imes 10^4$	$-3.101 imes10^3$	$2.617 imes 10^3$

Table 5. The identified value of the parameter vector x^* during charge.

Figure 9 compares the experimental and predicted terminal voltages curves at charge rates of 1C, 2C, and 3C by the proposed PSO-LM algorithm and the method of trial and error adopted in [22]. It is concluded that the proposed PSO-LM algorithm provides a more accurate prediction than the method of trial and error. This result further demonstrates the effectiveness of the proposed multi-scale method. The analysis of the charge parameter values can refer to those during the discharge process, and thus is omitted.



Figure 9. Comparison between experimental and predicted charge voltage curves at charge rates of 1C, 2C, and 3C by (**a**) weighted PSO-LM; and (**b**) method of trial and error.

5. Conclusions

A multi-scale parameter identification approach has been developed to identify the proper parameter value of the PDD model for the LIB during charge and discharge. This multi-scale approach is a combination of the PSO algorithm and the LM algorithm. PSO has the advantage of the coarse searching in large scale and LM has the advantage of the fine searching in small scale. Integration of them can effectively solve the difficult identification of the model with multi-scale complexity. To obtain a better fitness, a weighted objective/fitness function is implemented in this algorithm to reduce the difference among the multiple curves. The experimental results demonstrate the effectiveness of the proposed multi-scale identification approach.

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