

## Article

# The Dynamic Impact of Renewable Energy and Economic Growth on CO<sub>2</sub> Emissions in China: Do Remittances and Technological Innovations Matter?

Chafic Bassam Saliba <sup>1,2</sup>, Fida Ragheb Hassanein <sup>3</sup> , Seyed Alireza Athari <sup>4</sup> , Hazar Dördüncü <sup>5</sup>,  
Ephraim Bonah Agyekum <sup>6,\*</sup>  and Parise Adadi <sup>7</sup> 

<sup>1</sup> School of Business, Lebanese International University, Beirut 14604, Lebanon

<sup>2</sup> Department of Finance, Holy Spirit University of Kaslik, Kaslik 14604, Lebanon

<sup>3</sup> Business Department, Lebanese International University, Beirut 14604, Lebanon

<sup>4</sup> Department of Business Administration, Faculty of Economics and Administrative Sciences, Cyprus International University, Nicosia 99258, Northern Cyprus, Turkey

<sup>5</sup> Department of International Trade and Logistics, Faculty of Economics, Administrative and Social Sciences, Nisantasi University, Istanbul 34020, Turkey

<sup>6</sup> Department of Nuclear and Renewable Energy, Ural Federal University Named after the First President of Russia Boris Yeltsin, 19 Mira Street, 620002 Ekaterinburg, Russia

<sup>7</sup> Department of Food Science, University of Otago, Dunedin 9054, New Zealand

\* Correspondence: agyekumephraim@yahoo.com or agyekum@urfu.ru



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**Abstract:** Several investigations show that remittances, renewable energy, and innovation promote the socioeconomic advancement of a nation. Nevertheless, the impacts of remittances and renewable energy on ecological quality are yet to be evaluated thoroughly. Therefore, the current investigation assesses the effects of remittances and renewable energy on CO<sub>2</sub> emissions while taking into account the roles of technological innovation, globalization, and economic growth. Toward this end, this paper depends on yearly data between 1990 and 2019. The study employed bounds testing and its results disclosed long-term connections between CO<sub>2</sub> and the regressors. Moreover, unlike prior studies that employ time-domain causality, we employed frequency domain causality, which considers causality at different frequencies. Furthermore, the ARDL long- and short-run results showed that economic growth amplified CO<sub>2</sub> emissions, while green energy, remittances, and globalization lessened CO<sub>2</sub> emissions. Lastly, the frequency domain causality approach revealed that globalization, renewable energy, economic growth, technological innovation, and remittances could predict CO<sub>2</sub> emissions in the long-term. These findings' sturdiness was established utilizing DOLS and FMOLS regression. Several policy recommendations are suggested in light of these ground-breaking discoveries.

**Keywords:** remittances; renewable energy; technological innovation; globalisation

## 1. Introduction

As negative consequences of environmental degradation such as global warming and climate change become more palpable, the global awareness of environmental issues is raised. This awareness compels nations to take action to lessen environmental devastation. The inevitability of ecological deterioration due to economic expansion makes it difficult for policymakers to put ecological regulations into action rapidly [1,2]. As a result, economists focus their studies on the connection between ecological and macroeconomic indicators. At the same time, policymakers continue to look for environmental measures that have the fewest negative effects on economic indicators [3,4]. Over the years, both developing and developed nations have tried to find solutions to the environmental degradation that has resulted into global warming and climate change. The recent floods in Nigeria, Pakistan, and India, among others, are examples of the damaging effects of global warming and climate change. Though several nations have signed the Paris accord, meeting their carbon

reduction targets is in jeopardy due to their pro-growth agendas. Several studies have highlighted the environmental effects of economic growth, globalization, financial development, political risk, green energy, and technological innovation, among others [5–14].

The earlier-mentioned studies are often feasible in terms of the effective use of energy and the transition from fossil sources to green sources, since the structure of production and, consequently, the economy is centered mostly on fossil fuels for several nations. From this perspective, the financial sector is viewed as a crucial resource in supporting environmental expenditures due to the high costs of technology investments in enhancing energy efficacy and diversifying the portfolio of green energy [15–19]. Consequently, the amount of research studying drivers of ecological deterioration has significantly grown in recent years [20–23]. The most important resource in supporting environmental expenditures is official development aid for nations that wish to finance ecological investments using public funds [24]. Nevertheless, critics of public finance claim that these subsidies drain the economy and bloat the private sector. Remittances are being underutilized as essential resources. They help lessen ecological damage since they are one of the foundational components of the international financial sector [24,25].

Investigations of the connection between the environment and remittance (REM) are few since REM's effects on the ecosystem indirectly happen through several pathways. The first route is made possible by the rise in remittances, which raises individuals' disposable income. On the one hand, an increase in personal affluence results in an increase in energy-focused consumption, which worsens the ecosystem [26]. On the other hand, the rise in personal income encourages the creation of new employment and boosts output. Depending on the economy's degree of development, the effects of increased output brought on by economic expansion on the environment varies. The second route is concerned with how remittances affect consumption. Consumption growth is accompanied by output expansion, and as a result, energy demands rise sharply. Environmental contamination is thus anticipated to increase as energy demands rise. As more money is transferred from remittances to the banking sector, a third route opens up [8,9,27].

A complicated transmission mechanism that may be extended with different environmental and remittance-related elements was created [28,29]. For example, it is said that people can boost their educational attainment by utilizing the profits from their personal income increases brought on by their remittances [24–26]. Thus, a higher educational level can result in a greater understanding of the environment, which could aid in reducing behaviors that cause pollution. Foreign commerce is a potential outlet that is frequently disregarded. Due to remittances, an overvalued currency can harm a nation's ability to compete internationally by exposing a public moral hazard issue and restricting the variety of exports by adversely influencing internal commerce [26,29]. Lastly, since remittances are a crucial source of finance for nations with underdeveloped banking systems, sponsoring projects from this source with a focus on advancing renewable energy technology alters their ecological consequences. We must characterize the impact of remittances on the environment because of the intricate effects of all these pathways on environmental variables.

China intends to start along the route of a low-carbon transition by striving to become carbon neutral by 2060 in the face of escalating CO<sub>2</sub> emissions [8,9]. In addition to its domestic pledge to reduce CO<sub>2</sub> emissions, China has accepted the Paris Agreement under the UNFCCC and pledged to implement CO<sub>2</sub> emission-reduction measures to meet the relevant objectives for combating climate change [30,31]. China has changed its pledge to phase out coal use at the recently completed 26 Conference of the Parties (COP26) in Glasgow by promising to reduce coal usage. This change reflects China's position that CO<sub>2</sub> emissions will peak by 2030 and gradually decline to achieve the 2060 carbon neutrality goal. However, the country is committed to diversifying its national energy portfolio by growing and lowering the percentages of renewables (wind and solar) and coal, respectively, to achieve carbon neutrality by 2060 [32]. The world's economies are working to find pertinent elements that can separate CO<sub>2</sub> from economic growth to achieve carbon neutrality [31]. China also intends to implement specific decoupling techniques in this

area [30]. It is dedicated to accomplishing these linked goals by implementing novel efforts to create a modern energy system, particularly in respect to the 2060 carbon-neutrality aim. Modernization is anticipated to take the shape of a shift to clean energy, increased energy efficiency, and reduced reliance on fossil fuels for power production. Innovation may be considered a modern approach in attaining low-carbon growth in China by applying the most recent technologies, not just inside the energy sector but in the whole economy [31].

Given the facts above, the main goal of this essay is to ascertain the implications of remittances on the environment in China. In order to refrain from ignoring potential intermediary pathways in the remittances-environment interaction, the effects of GDP, REC, and INV are also explored. The following are the research's literary contributions: (i) The assessment is strengthened by accounting for the intricate relationships amongst remittance, the environment, and the prospective channels that constitute the relationship between CO<sub>2</sub> and remittances; (ii) The FMOLS and DOLS are employed for their sturdiness and to produce more thorough policy suggestions; (iii) Lastly, the frequency domain causality, which can evaluate causality at various frequencies suggested by [33], was employed. Thus, we can capture causality in the short-, medium-, and long-term.

The research's remaining sections are organized as follows: The literature is reviewed in Section 2. The methods and empirical models are described in Section 3. A discussion of the empirical findings is presented in Section 4. The findings' practical recommendations are highlighted in Section 5.

## 2. Literature Review

### 2.1. *Synopsis of Studies between CO<sub>2</sub> Emissions and Economic Growth and Remittances*

Between the years 1986 to 2017, [25] examined the relationship between remittances, growth, and emissions. In its econometric methodology, this research used second-generation unit root approaches. The results showed that remittances contribute to slowing down ecological deterioration because they have a detrimental impact on emissions, whereas economic growth increases CO<sub>2</sub> emissions. From 1980 to 2016, [24] investigated the emissions-remittances-growth nexus in Pakistan and Bangladesh. The researchers used a panel cointegration approach and a panel ARDL model to examine long-run correlations. The panel cointegration approach's predicted finding corroborated the presence of a long-term link between REM, GDP, and CO<sub>2</sub> emissions. Its conclusions showed that both economic development and an improvement in remittances received contributed to reducing emissions from the chosen panel nations. The short-run ARDL analysis, on the other hand, demonstrated that CO<sub>2</sub> emissions increased at a substantial rate in response to an increase in economic growth and remittance inflow. Additionally, this research utilized DOLS for reliable assessment and came to the same long-term conclusion as ARDL.

Moreover, Ref. [26] inspected the nexus of remittances and CO<sub>2</sub> emissions. For the period between 1980 and 2014, the study investigated the asymmetric connections between remittances, real growth, and CO<sub>2</sub> in India. Using time series data, the research adopted a nonlinear ARDL model to predict theoretical connections. The NARDL bound test's findings affirmed the indicators' long-run associations. The results demonstrated that, in contrast to a negative shock, positive surprise in remittances caused a rise in CO<sub>2</sub>. Real growth had a positive coefficient but lost statistical significance with time. According to empirical findings, the parameters had an asymmetric long-run interrelationship. Furthermore, [34] used FMOLS and DOLS models to estimate yearly data from 1990 to 2019 to compare CO<sub>2</sub> emissions, real growth, and remittance in a subset of G-20 nations. The CD was identified via LM tests, and the CADF test was used to determine if the variables were stationary. In the framework of the G-20 nations, the Panel Cointegration confirmed a long-run association. Two models showed a positive correlation between CO<sub>2</sub> emissions and the economic development of a subset of G-20 countries. Additionally, evidence found that remittances significantly reduced CO<sub>2</sub> emissions. Researchers, policymakers, environmentalists, and governments will benefit from this research's illumination on achieving an eco-friendly ecosystem through the prudent use of remittances and real growth.

## 2.2. Synopsis of Studies between CO<sub>2</sub> Emissions and Renewable Energy and Technological Innovation

In addition to being a reliable solution to guaranteeing green energy and security, renewable energy also aids in the transition to a low-carbon society and economy. The growth of renewable energy is influenced by a number of fundamental factors, including the advancement of technological innovation. Chinese province data from 2000 to 2015 were utilized by [35] to assess the causes of climate change and investigate how technical advancements in renewable energy were responding to high levels of CO<sub>2</sub>. The study revealed that the degree of technological innovation in China's regions varied significantly. Additionally, the research showed that the high levels of CO<sub>2</sub> emissions had boosted technical innovation in renewable energy sources; this indicated that the innovation process proactively reacts to climate change. Additionally, business and government R&D spending both contributed to raising the degree of innovation.

Moreover, ref. [36] utilized the yearly data from 1990 to 2018 to examine the influence of technological innovation and green energy in lowering the transport sector's CO<sub>2</sub> emissions in China. Implementing the QARDL technique revealed that clean energy and technological innovation significantly impacted CO<sub>2</sub> emissions in China's transportation industry. The use of green energy and green innovation both had a negative effect on CO<sub>2</sub> associated with transportation. In addition, the results showed that as innovation and clean energy grew, so too did the amount of CO<sub>2</sub> emitted in the transportation sector; yet, as a nation's GDP increased, so too did the amount of CO<sub>2</sub> emissions in that sector. To reduce CO<sub>2</sub> emissions, China should adopt new regulations that spur innovation in the transportation industry.

Based on the STIRPAT model, ref. [37] investigated the relationships between green technology, CO<sub>2</sub>, and renewable energy use in Turkey from 1990 to 2018. The research implemented tests similar to "unit-root" to confirm the integrative characteristics of the factors holding the information for structural fractures. The connection between the parameters was also examined using the BARDL-bound test procedure. A Granger causality test was used to examine the causative connection between the development of green technologies, renewable energy, and CO<sub>2</sub> emissions. The research findings uncovered the long-term integration of green renewable energy, emissions, and technological progress. Furthermore, it showed that innovation and clean energy reduced CO<sub>2</sub> emissions, but energy use increased CO<sub>2</sub> emissions.

From 1990 to 2017, ref. [38] examined the CO<sub>2</sub> dynamics of the N-11 nations. The research evaluated novel elements to determine CO<sub>2</sub>, such as technical innovation and the utilization of green energy. The empirical estimates were based on methodologies from the second generation. Consumption of renewable energy and technological innovation were negatively correlated with CO<sub>2</sub>. These results have significant ramifications, and prompt us to advise encouraging technical advancement and the usage of green energy sources. This will assist in accomplishing the COP21 objectives.

## 3. Data and Methods

### 3.1. Data

This study aims to explore the role of economic growth, remittances, innovation, and globalisation on CO<sub>2</sub> emissions in China using yearly data stretching between 1990 and 2019. The regressors are remittances (REM), renewable energy (REC), innovation (INV), globalisation (GLOB), and economic growth (GDP), while the dependent variable is Carbon emissions (CO<sub>2</sub>). The precise measurement, source, and variable name is presented in Table 1. All the variables are logged to ensure conformity to normality in line with the research of [39].

**Table 1.** Variables Description.

Sign	Variables	Measurement	Sources
GDP	Economic Growth	Constant 2010 US\$	World Bank [40]
GLOB	Globalisation	Index	KOF [41]
REC	Renewable Energy	Renewable energy consumption (% of total final energy consumption)	World Bank [40]
REM	Remittance	Personal remittances received (% of GDP)	World Bank [40]
INV	Innovation	Number of patent application	World Bank [40]
CO <sub>2</sub>	Carbon Emissions	Metric tonnes per capita	BP [42]

Table 2 presents the data description with the results presented as follows: the mean of GDP (29.07) is the highest, which ranges from 27.658 to 30.291; followed by LnINV (11.410), which ranges between 8.6712 and 14.194; LnGLOB (3.9889), which falls from 3.5866 to 4.1676; LnREC (2.9814), which ranges from 2.4283 to 3.5288; Ln CO<sub>2</sub> (1.3988), which ranges from 0.6494 to 2.0289; and LnREM (−1.8795), which ranges from −3.3983 to −0.7397. With the exemption of LnINV, the values of the other variables (LnCO<sub>2</sub>, LnGDP, LnGLOB, LnREC, and LnREM) are negatively skewed. Furthermore, the kurtosis values show that LnCO<sub>2</sub>, LnGDP, LnGLOB, LnREC, LnREM, and LnINV are platykurtic. Lastly, the JB Probability shows that all the variables align with normal distribution.

**Table 2.** Descriptive Statistics.

	LnCO <sub>2</sub>	LnGDP	LnGLOB	LnREC	LnREM	LnINV
Mean	1.3988	29.074	3.9889	2.9814	−1.8795	11.410
Median	1.4358	29.063	4.0725	2.9314	−1.7276	11.317
Maximum	2.0289	30.291	4.1676	3.5288	−0.7397	14.194
Minimum	0.6494	27.658	3.5866	2.4283	−3.3983	8.6712
Std. Dev.	0.5000	0.8134	0.1837	0.4227	0.6289	1.8841
Skewness	−0.0414	−0.1102	−0.7959	−0.0056	−0.7777	0.1143
Kurtosis	1.3562	1.7629	2.2707	1.2449	3.0753	1.5308
JB	3.3859	1.9735	3.8326	3.8505	3.0314	2.7633
Probability	0.1839	0.3727	0.1471	0.1458	0.2196	0.2511

### 3.2. Estimation Strategies

To achieve the goal of our study, the necessary data on the relevant variables were compiled from the outset. The empirical model was then specified in this research by referencing the literature. The timeframe of the study covered the period between 1990 and 2019. The defined empirical model was examined using a variety of advanced econometric methods to obtain results that can be relied upon for formulating policy. As a result, the investigation started by looking at the time series data' stationarity feature. The investigation validated the stated model's long-term association after defining the series' integration order. Since the Jarque Bera probability is insignificant, we refuted the null hypothesis of "nonlinearity". Therefore, the variables of investigation were linear, which warranted the study to use a linear approach. Based on this knowledge, we formulated the study's economic model as follows:

$$CO_{2t} = \beta_0 + \beta_1 GDP_t + \beta_2 INV_t + \beta_3 GLOB_t + \beta_4 REC_t + \beta_5 REM_t + \varepsilon_t \quad (1)$$

where CO<sub>2</sub>, GDP, INV, GLOB, REC, and REM represent carbon emissions, economic growth, technological innovation, globalisation, renewable energy, and remittances, respectively.

Furthermore,  $\beta_1 \dots \beta_5$  represents the coefficients of the regressors. Moreover,  $\varepsilon_t$  represents an error term.

Given its numerous advantages over the traditional cointegration methods, the analysis employed the approach suggested by [43], named the ARDL model, as an efficient estimation technique to reveal short- and long-term connections among the parameters of the specified model. Before using other cointegration strategies, it was crucial to uncover a series' integration property; nevertheless, this strategy did not call for pre-testing. The ARDL model can be used to address the endogeneity issue with the aid of the variable's lag length. Secondly, it can be used regardless of how the series under investigation is integrated [43,44]. Last but not least, the ARDL model can generate reliable results even when the sample size is small [45]. The ARDL bound testing approach for the provided econometric model in Equation (1) may be written in Equation (2) in the manner shown below. Consequently, it is simple to understand how the variables are connected in the long-run, as shown in Equation (2).

$$\begin{aligned} \Delta \ln \text{CO}_2 = & \theta_0 + \beta \text{DUM} + \sum_{i=1}^p \beta_{1i} \Delta \ln \text{CO}_{2t} + \sum_{i=1}^p \beta_{2i} \Delta \ln \text{GDP}_t + \sum_{i=1}^p \beta_{3i} \Delta \ln \text{INV}_t + \sum_{i=1}^p \beta_{4i} \Delta \ln \text{GLOB}_t \\ & + \sum_{i=1}^p \beta_{5i} \Delta \ln \text{REC}_t + \sum_{i=1}^p \beta_{6i} \Delta \ln \text{REM}_t + \gamma_1 \Delta \text{CO}_{2t-1} + \gamma_2 \Delta \text{GD}_{[t-1]} + \gamma_3 \Delta \text{INV}_{t-1} + \gamma_4 \Delta \text{GLOB}_{t-1} \\ & + \gamma_5 \Delta \text{REC}_{t-1} + \gamma_6 \Delta \text{REM}_{t-1} + \varepsilon_t \end{aligned} \quad (2)$$

After the long-term connection between series is defined, the short-run coefficient must be gathered. Thus, the error correction model must be evaluated to gather the short-run coefficients, as is shown in Equation (3).

$$\begin{aligned} \Delta \ln \text{CO}_2 = & \theta_0 + \beta \text{DUM} + \sum_{i=1}^p \beta_{1i} \Delta \ln \text{CO}_{2t} + \sum_{i=1}^p \beta_{2i} \Delta \ln \text{GDP}_t + \sum_{i=1}^p \beta_{3i} \Delta \ln \text{INV}_t + \sum_{i=1}^p \beta_{4i} \Delta \ln \text{GLOB}_t \\ & + \sum_{i=1}^p \beta_{5i} \Delta \ln \text{REC}_t + \sum_{i=1}^p \beta_{6i} \Delta \ln \text{REM}_t + \rho \text{ECT}_{t-1} + \varepsilon_t \end{aligned} \quad (3)$$

In Equations (2) and (3),  $\Delta$  denotes the first difference operator,  $p$  represents the lag length of the series, and  $\varepsilon_t$  denotes the error term. Furthermore,  $\text{ECT}$  and  $\rho$  denote the error correction term and the coefficient of the ECT, respectively. Additionally, the above equation presents error correction dynamics and the series' long-term relationships. In Equation (2), we evaluate the existence of the long-run connection with the aid of the ARDL bounds test by proposing the following null and alternative hypotheses as shown in Equations (4) and (5):

$$H_0: \theta_1 = \theta_2 = \theta_3 = \theta_4 = \theta_5 = 0 \quad (4)$$

$$H_0: \theta_1 \neq \theta_2 \neq \theta_3 \neq \theta_4 \neq \theta_5 \neq 0 \quad (5)$$

The computed F-stat and its reference to the critical value provided by Pesaran et al. [43] determined the final verdict about the hypothesis (2001). Pesaran et al. [43] was chosen as the critical value because of its virtue in taking a longer duration into account. When the recorded statistical value of F exceeded the upper limit's crucial boundary value, the  $H_0$  was dismissed. When the computed F-statistic fell short of the lower critical limit, the  $H_0$  was supported. When the estimated number was between the top and lower boundary values, there was insufficient evidence against the  $H_0$ . The conclusion about the long-run interrelationship also relied on the ECT and its direction, i.e., the long-term connection required the negative and statistically significant ECT [46]. Furthermore, we used the AIC to evaluate the optimal lag length.

We also used the FMOLS (completely modified OLS) and dynamic OLS (DOLS) on the stated model to examine the long-term effects of explanatory factors on the  $\text{CO}_2$  as a robustness assessment for the proposed model. Two aspects drove the justification for utilizing these techniques. First, the DOLS and FMOLS can be employed if the cointegration requirement among the  $I(1)$  parameters are met. Second, these techniques address serial

correlation and endogeneity biases resulting from the cointegration connection. As a result, it provides asymptotic effectiveness in the results.

Lastly, the frequency domain causality, which can evaluate causality at various frequencies suggested by [33], was employed. Thus, we were able to capture causality in the short-, medium-, and long-term.

## 4. Results and Discussion

### 4.1. Pre-Estimation Results

The primary objective of this research was to examine the emergence of a long-term interrelationship between the series being discussed. The suggested model does not apply to  $I(2)$  indicators, even though ARDL bound testing does not need previous knowledge about the nature of the series associated with the integration order. Thus, to use the methods for creating a long-term interrelationship, unit root test evaluation gives vital knowledge about the integration characteristic of the parameters. As a result, several standard root tests, including the PP and ADF tests, were used to examine the series' integration properties. Table 3 summarizes the evolving results of the stationarity test. All variables have the unit root issue at level, as per the standard tests of unit root output, but it shifts to stationary at the first difference.

**Table 3.** Unit Root Test Results.

	ADF		PP		ZA			
	$I(0)$	$I(1)$	$I(0)$	$I(1)$	$I(0)$	BY	$I(1)$	BY
LnGDP	−1.016	−3.695 **	−1.788	−3.705 ***	−4.167	2010	5.937 *	2005
LnINV	−1.219	−4.018 **	−1.646	−4.114 **	−4.708	2015	−6.002 *	2010
LnGLOB	−0.928	−5.540 *	−0.928	−5.540 *	−4.349	2004	−8.544 *	2008
LnREC	−1.329	−4.468 *	−1.069	−4.526 *	−4.043	2010	−5.204 **	2003
LnREM	−1.839	−3.996 **	1.930	−4.045 **	−3.940	2003	−5.677 *	2003
InCO <sub>2</sub>	−1.675	−3.923 ***	−1.358	−4.082 **	−3.652	2009	−5.837 *	2009

Note: \*, \*\* and \*\*\* portrays 1%, 5% and 10% level of significance.

If the structural discontinuity in the time series is not accounted for during the estimating procedure, the computed traditional unit root test may yield misleading results. In order to comprehend the structure of the series while taking breaks into account, [47] test was used. According to Table 3, the results of the ZA showed that the investigated series were or  $I(1)$ . Consequently, it has been determined that the variables continue to follow the first difference stationarity despite the structural break. It has also been confirmed that none of the series are  $I(2)$ .

### 4.2. Cointegration Results

We have observed that the under-consideration series is  $I(1)$  series, based on unit root findings. Hence, we are required to assess the long- and short-term relationships between the series under consideration by utilizing the bounds test technique that was built on the ARDL framework proposed by Pesaran et al. [43]. The findings from ARDL are susceptible to the selection order of lag as an empirical method. Therefore, picking the right lag is crucial to solving the endogeneity issue in the model formulation. Based on our model's reliable findings from the different alternatives' information criterion, we decided to use the AIC for lag specification. By examining the test statistics for the cointegration investigation, the results shown in Table 4 demonstrate that the calculated F-stat is considerably above the upper CV of [45]. Thus, at the statistically significant level of 1%, the model shows an exceeding of the upper CV. As a result, we concluded that the regressors and dependent variable are cointegrated.

**Table 4.** ARDL Bounds Test Results.

Test Statistic	Value	Signif.	I(0)	I(1)
F-statistic	5.1397 *	10%	−2.57	−3.86
k	5	5%	−2.86	−4.19
		2.5%	−3.13	−4.46
		1%	−3.43	−4.79

Note: \* represents 1% level of significance.

#### 4.3. ARDL Long and Short-Term Results

This section elucidates the results of the ARDL both in long- and short-term (see Table 5). As observed in Table 5, the coefficient of GDP was positive and significant, implying that a 0.869% (long-term) and a 0.161% (short-term) upsurge in GDP contributed to an intensification of LnCO<sub>2</sub>. In summary, in both the short- and long-term, an increase in ecological deterioration was caused by economic progress in China. Moreover, the coefficient of LnINV was negative and significant, implying that a 0.869% (long-term) and 0.161% (short-term) upsurge in LnINV contributed to the mitigation of LnCO<sub>2</sub> emissions with an insignificant association in the short-term. Therefore, in the long-term, a lessening in ecological deterioration was caused by an upsurge in technological innovation in China.

**Table 5.** Long and Short-run Results.

Variable	Long-Term			Short-Term		
	Coefficient	t-Statistic	Prob	Coefficient	t-Statistic	Prob
LnGDP	0.869 ***	1.835	0.083	0.161 ***	1.786	0.090
LnINV	−0.370 *	3.753	0.001	−0.003	−0.136	−0.863
LnGLOB	−0.381 **	−2.603	0.018	−0.789 *	−4.258	0.000
LnREN	−0.789 *	−6.702	0.000	−1.027 **	2.683	0.015
LnREM	−1.208 *	−3.871	0.000	−1.947	−5.113	0.000
DUM	4.947	1.864	0.078	-	-	-
ECT (−1)	-	-	-	−0.396	−5.135	0.000
R <sup>2</sup>	0.99					
Adj-R <sup>2</sup>	0.98					
Diagnostic Test Results						
	Value	p-value				
Heteroskedasticity Test	0.374	0.545				
Serial Correlation	2.427	0.120				
Normality Test	1.336	0.517				
Ramsey RESET Test	1.127	0.275				

Note: \*, \*\* and \*\*\* portrays 1%, 5% and 10% level of significance.

Furthermore, the effect of LnGLOB on LnCO<sub>2</sub> emissions was negative, as shown by the negative and significant coefficients in both short- and long-term. This demonstrates that a 0.381% (long-term) and 0.789% (short-term) lessening of LnCO<sub>2</sub> was caused by a 1% upsurge in LnGLOB. These outcomes further imply that decarbonisation in China can be achieved via an increase in globalization. As expected, the coefficient of LnREC was significant and negative, suggesting that a 0.789% (long-term) and 1.027% (short-term) upsurge in LnREC promoted ecological quality. Lastly, the coefficient of LnREM was significant and negative, demonstrating that a 1.208% (long-term) and 1.947% (short-term) upsurge in LnREM contributed to the mitigation of LnCO<sub>2</sub>. In summary, in both the short

and long-term, the increase in ecological quality was caused by the inflow of remittances in China.

We have used 2009 as a dummy (DUM) variable in the estimate procedure to reflect the one structural break in CO<sub>2</sub>. Nevertheless, the DUM was negligible in both the short- and long-term results, demonstrating that structural discontinuities do not significantly affect CO<sub>2</sub>. Additionally, it was discovered that the ECT was significantly negative. This estimate, of 0.396, illustrated changes in the short-run equilibrium as it progressed toward achieving a long-run stable equilibrium, adjusted annually at a rate of 39%. It highlighted the feedback mechanism's usefulness in stabilizing China's CO<sub>2</sub> emissions. Lastly, the post-estimation results uncovered no serial correlation issue or heteroskedasticity; the residuals were distributed normally and without misspecification. The stability tests (see Figure 1a,b) show that the model was stable at 5%.

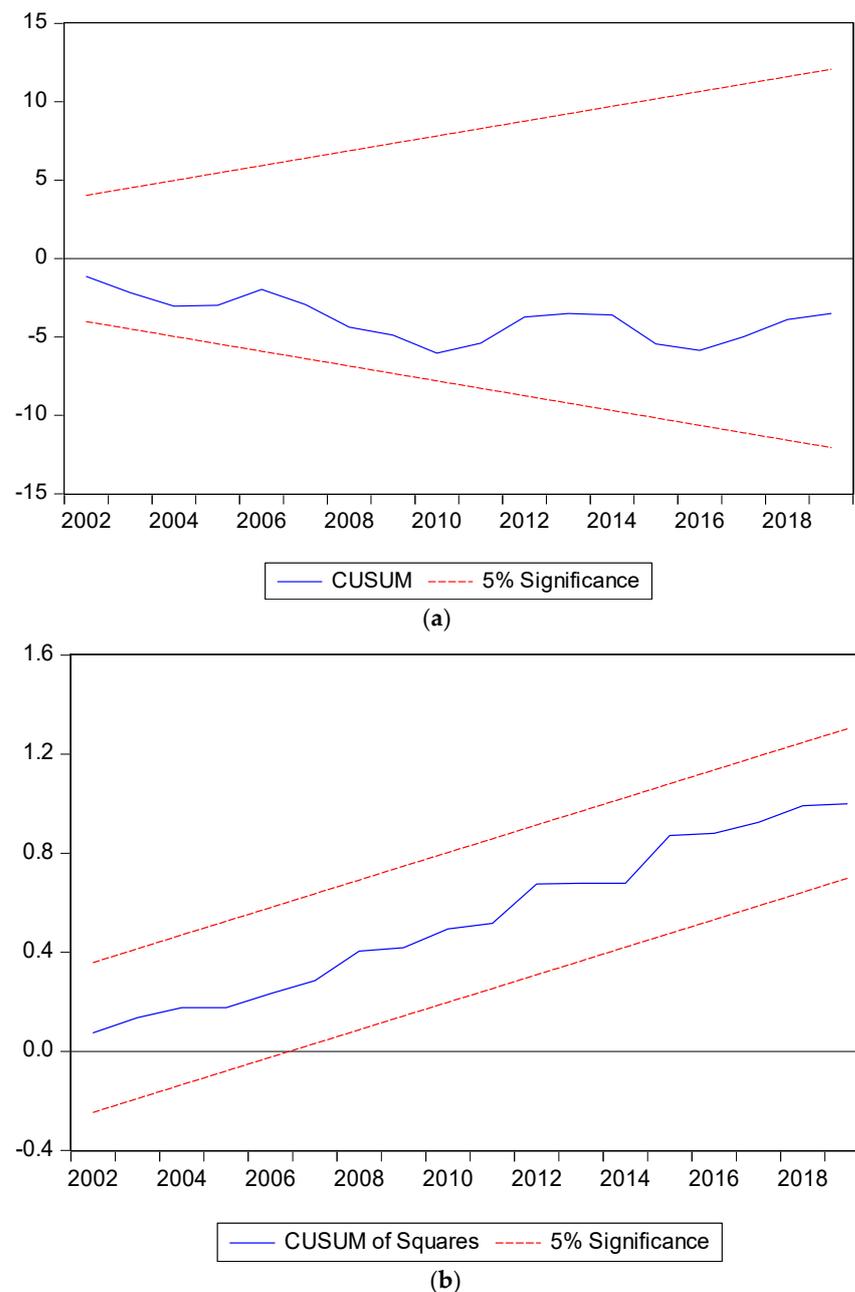


Figure 1. (a): CUSUM of the model. (b): CUSUM of Squares of the model.

#### 4.4. Robustness Check Results

Additionally, FMOLS and DOLS were utilized in this research to strengthen the validity of the ARDL framework's long-term results. While DOLS outperformed the FMOLS (Shahbaz et al., 2016) in some circumstances, two options were employed because we could not assess how the outcomes progressed consistently. The predicted outputs of FMOLS and DOLS are shown in Table 6. The FMOLS and DOLS results showed that they are reliable and consistent in sign and significance. As a result, they replicate the ARDL long-run outcomes in Table 5. The results uncovered that GDP intensified CO<sub>2</sub> while LnGLOB, LnREC, LnINV, and LnREM lessen CO<sub>2</sub> emissions. Consequently, the results may be utilized to formulate strategies with some degree of certainty.

**Table 6.** Robustness Results.

Variable	FMOLS			DOLS		
	Coefficient	t-Statistic	Prob.	Coefficient	t-Statistic	Prob.
LnGDP	0.921 *	4.397	0.000	0.869 **	2.513	0.021
LnGLOB	−0.403 *	−6.217	0.000	−0.381 *	−3.566	0.002
LnINV	−0.101 *	−4.137	0.000	−0.403 *	−4.186	0.000
LnREC	−0.809 *	−15.52	0.000	−0.789 *	−9.180	0.000
LnREM	−0.008 **	−2.636	0.017	−0.008 ***	−1.742	0.098
DUM	0.014	0.705	0.488	−2.973	−1.528	0.142
R <sup>2</sup>		0.96			0.97	
Adj R <sup>2</sup>		0.94			0.96	

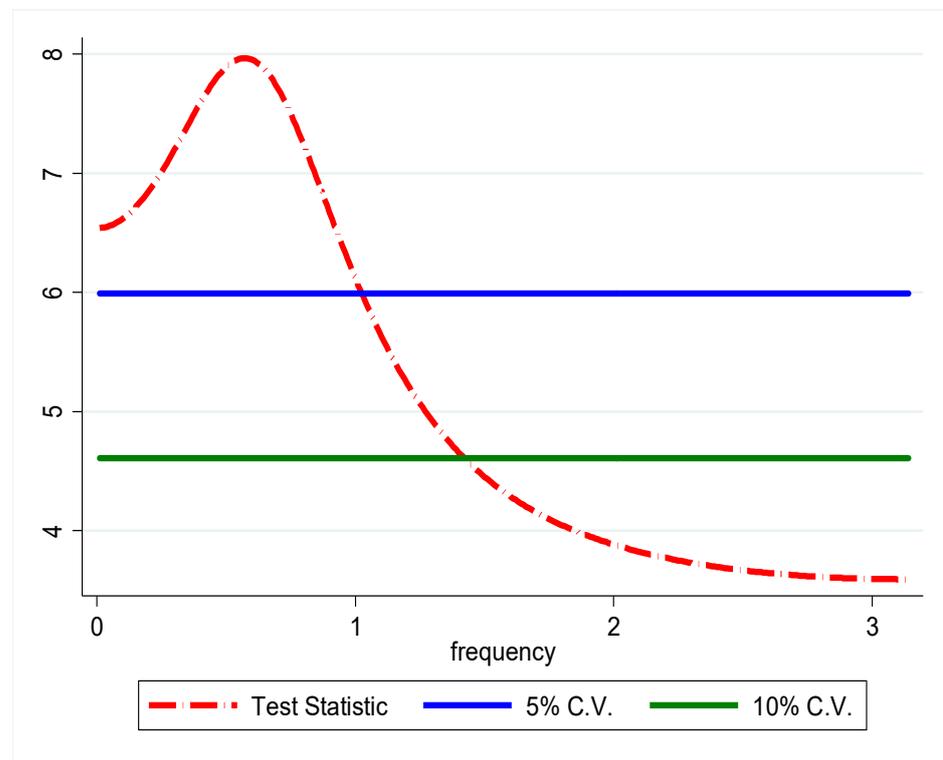
Note: \*, \*\* and \*\*\* portrays 1%, 5% and 10% level of significance.

#### 4.5. Causality Results

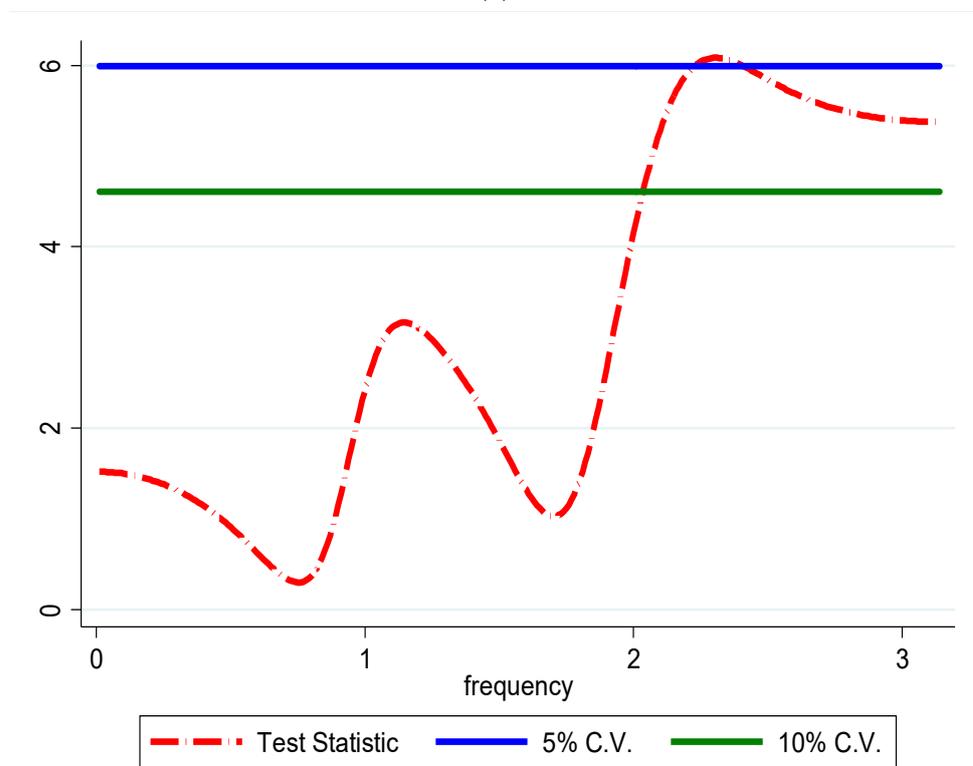
The causality test results are shown in Figure 2. The outcomes show that in the long-term, LnREC can forecast LnCO<sub>2</sub> (see Figure 2a) at 5% and 10% significance. Furthermore, in the short-term, LnREM can predict LnCO<sub>2</sub> (see Figure 2b) at a 5% and 10% level of significance with no causality in the long- and medium-term. Likewise, in the short- and long-term, LnGDP can forecast LnCO<sub>2</sub> (see Figure 2c) at a 10% level of significance. Similarly, in all frequencies, LnGLOB can forecast LnCO<sub>2</sub> (see Figure 2d) at a 10% level of significance. Lastly, in all frequencies, LnINV can predict LnCO<sub>2</sub> (see Figure 2e) at a 5% and 10% level of significance in the long-term.

#### 4.6. Discussion of Results

As reported by ARDL, FMOLS, and DOLS, LnREM impacts LnCO<sub>2</sub> negatively. The findings of [25,29,34] are in direct opposition to our findings. Conversely, our results are in line with the results of [24,28]. In this case, an upsurge in remittances has reduced CO<sub>2</sub> emissions by producing minor and significant scaled impacts in a production process, which entails shifting to energy-efficient inputs at the level of SMEs and industries. Remittances are a source of revenue that directly raises individual income. For a household, the increase in income opens up two possibilities. The first relates to spending, while the other relates to saving [34]. The demand for goods rises when the expenditure trend is in effect, directly impacting the industrial sector and indirectly affecting energy usage. In times of active saving, people spend a portion of their higher income on interest payments from financial institutions. Again, indirect energy consumption arises due to these companies lending the money they have raised to industrial groups to earn interest revenue [48]. Considering all the evidence, our findings demonstrate that remittances tend to shift toward renewable energy sources.

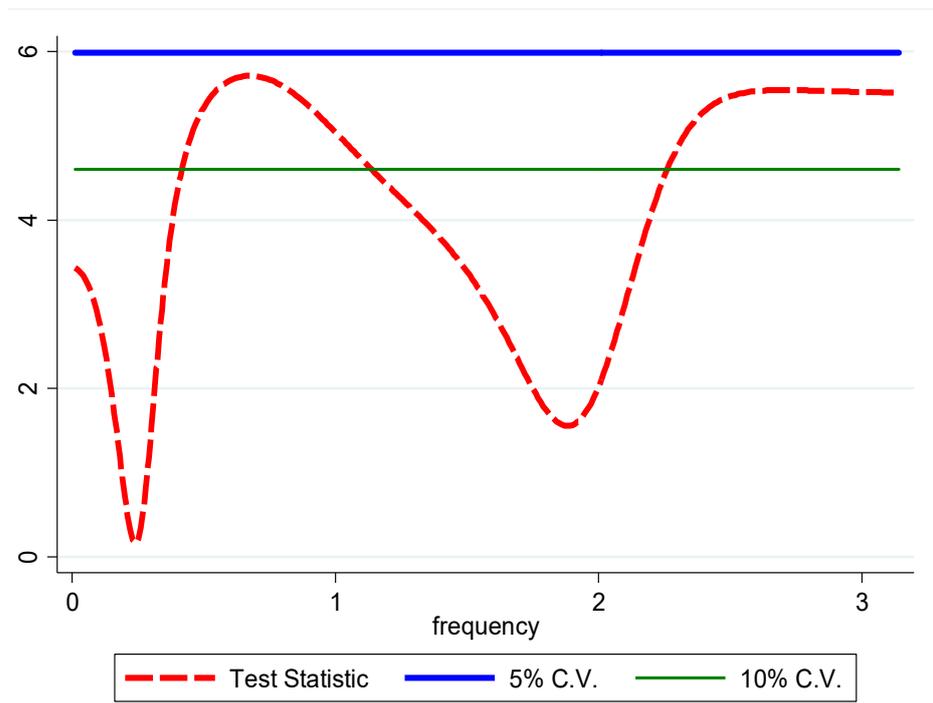


(a)

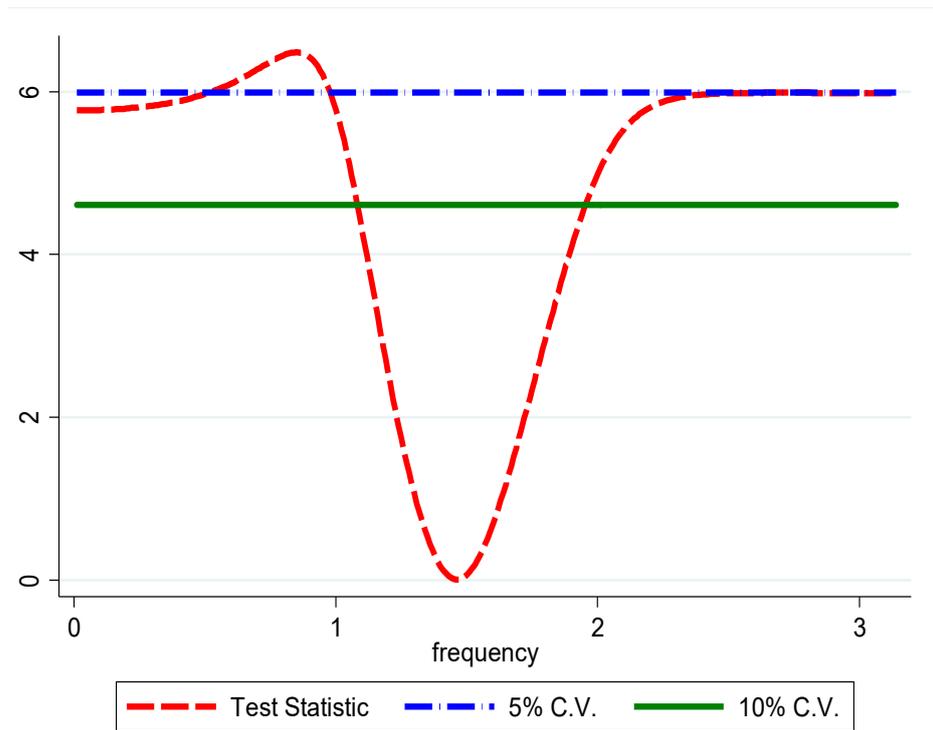


(b)

Figure 2. Cont.

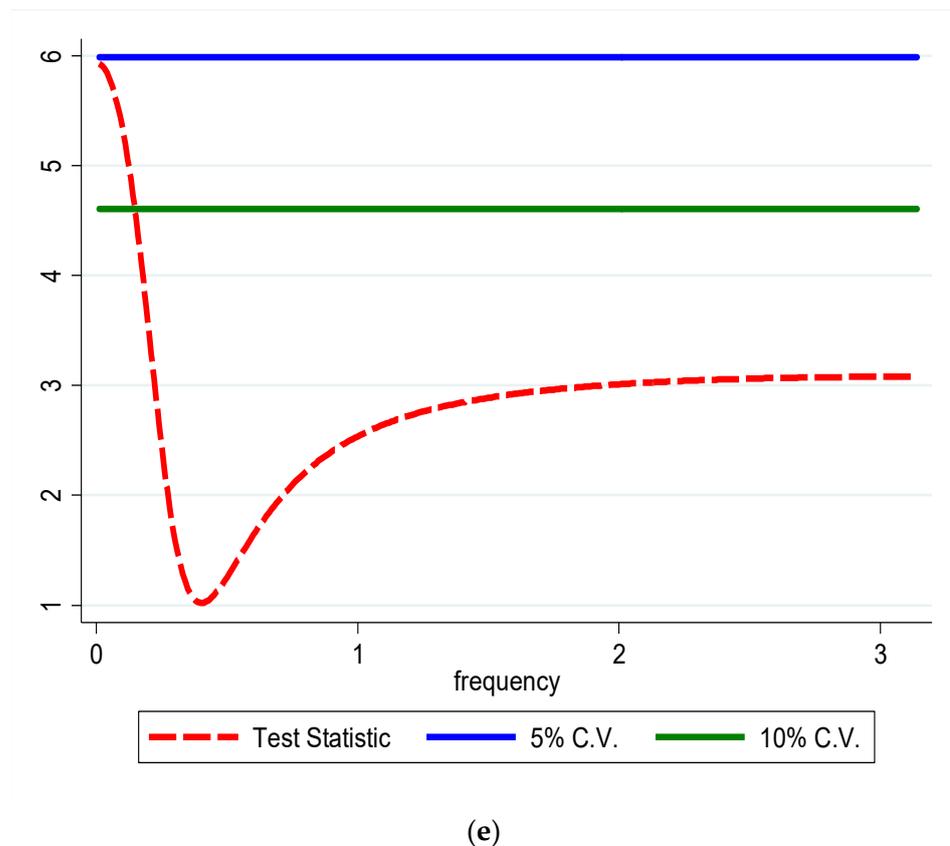


(c)



(d)

Figure 2. Cont.



**Figure 2.** (a) Causality from LnREC to LnCO<sub>2</sub>; (b) Causality from LnREM to LnCO<sub>2</sub>; (c) Causality from LnREM to LnCO<sub>2</sub>; (d) Causality from LnGLOB to LnCO<sub>2</sub>; (e) Causality from LnINV to LnCO<sub>2</sub>.

The findings also demonstrate that GDP-based economic expansion correlates favorably with promoting larger CO<sub>2</sub> emissions. This suggests that the strategies intended to boost economic activities in China are not ecologically friendly and that these unsuitable measures do not support China's goal of becoming carbon neutral. This finding aligns with the views put forward by [9,17,49,50], in which the researchers affirmed that economic progress lessens environmental quality.

Energy is also a major factor in any economy's ability to advance economically. In essence, it is a crucial input for both consumption and production. People's daily lives are significantly impacted by energy. As such, we added renewable energy to the model to see if it affects CO<sub>2</sub> emissions. The results showed that changes in China's usage of renewable energy decreased the nation's CO<sub>2</sub> emissions. Our results support the claim made by [51,52] that REC can enhance ecological quality. Other research by [53] also demonstrated the negative effect of REC on CO<sub>2</sub> emissions, which matches our discoveries. They did, nevertheless, discover a short-term positive connection, which is distinct from our findings. This difference might be due to the fact that their study used data from the United States, while our analysis was conducted using data from China.

Our findings demonstrate the emissions-reducing effect of globalization. The principal defense is that lowering the obstacles created by international boundaries encourages the exchange of commodities and more environmentally friendly new technologies [54,55]. Parallel to this, there is a widespread assumption backed up by some proof that information will enhance concern and advocacy for ecological quality. Thus, environmental degradation should be reduced as society becomes more knowledgeable [14,56].

The results of this study showed a negative correlation between INV and CO<sub>2</sub>. These findings highlight the importance of implementing new technologies in the nation's industrial sector to foster an environment conducive to industrial transformation. These results offer crucial information for formulating policies that comply with the decisions made at

the worldwide climate change conference (COP21). Our outcomes are consistent with (REF) research, which showed how innovation could help China reduce its CO<sub>2</sub> emissions. The premise of our analysis is in accordance with [35]'s suggestion of the important contribution of R&D spending to ecological deterioration. Ullah et al.'s [2] findings, in contrast, diverge from our conclusions in that they discovered that while technological development does lower CO<sub>2</sub>, its impact on the environment is negligible. In contrast, our results demonstrate that the outcomes are significant. Additionally, [57,58] discovered outcomes that differed from our investigation. This research was carried out because the conclusions of earlier studies were ambiguous. As a result, our research's findings offer important information for formulating policy and are reliable because they were obtained using cutting-edge methods [59,60]. These results provide a useful starting point for the nation's innovation and ecological policy development.

Furthermore, in the long-term, there is evidence of causality from remittances, globalization, renewable energy use, economic growth, and technological innovation to CO<sub>2</sub> emissions in China. These interesting findings show that in the long-term, policies formulated to reduce CO<sub>2</sub> emissions must take into account the role of remittances, globalization, renewable energy use, economic growth, and technological innovation.

## 5. Conclusions and Policy Implications

### 5.1. Conclusions

This study examined the effect of remittance and renewable energy on CO<sub>2</sub> emissions in China by considering other drivers such as economic growth and technological innovation. The study uses data spanning between 1990 and 2019. The study used an autoregressive distributed lag model (ARDL) as well as the FMOLS and DOLS long-run estimators as a robustness check. Furthermore, unlike prior studies that employ the time domain causality, we employed frequency domain causality, which considers causality at different frequencies. The results of the ARDL, FMOLS, and DOLS revealed that economic growth amplifies CO<sub>2</sub> emissions while green energy, remittances, and globalization lessens them. Furthermore, the causality results show that remittances, globalization, renewable energy use, economic growth, and technological innovation can predict CO<sub>2</sub> emissions in the long-term.

### 5.2. Policy Recommendations

The research's findings suggest potentially significant policy ramifications. The research encourages policymakers to develop strategies to prevent the harmful impacts of remittances on the ecosystem by revealing a negative correlation between emissions and remittances. For example, individuals can be given incentives to use remittance money to buy environmentally friendly home appliances. Parallel to this, China may consider making its current emission control rules more stringent. Ultimately, the report recommends that ecological quality improvement activities be scaled following an upsurge in remittance inflows.

In terms of its ramifications for policy, this research suggests that China incorporate R&D and technological innovation linked with its 2060 carbon neutrality objective to make it more feasible. The Chinese government should concurrently increase R&D spending levels and encourage the private sector to fund R&D for the creation of technologies that promote ecological welfare. The government should focus on creating green technologies in this respect so that China can quickly attain low-carbon growth. In order to gradually eliminate the detrimental effects of economic expansion on the ecosystem, China should also look forward to promoting green economic activities. For example, it is advised that China invest in the construction of renewable energy facilities rather than building new coal power plants.

Thirdly, our analysis shows overwhelming evidence that renewable energy lowers emissions. Irrespective of a country's emission situation, renewable energy sources reduce emissions in all circumstances. Renewable energy, nevertheless, has significant infrastruc-

tural needs. However, once they are installed, these technologies gradually recuperate their expenses. China needs to stabilize its energy supply, beginning with less expensive renewable energy sources. Therefore, investment in renewable energy is needed in China. Furthermore, since renewable energy is driven by technological innovations in the tides of biomass, geothermal, wind, and solar, investment in green technologies will boost renewable energy and as a result improve ecological quality.

Lastly, environmental degradation has grown along with economic progress, showing that green growth has not yet started in China. Currently, a shift to a green economy is essential for sustainable development. As a result, policymakers in China should not be reluctant to employ some instruments in this regard: (i) Tax reductions should be reduced or stopped to boost remittances; (ii) The financial industry should be encouraged and assisted in creating internet applications; (iii) All private and public sectors, particularly educational institutions, should obtain training on ecological consciousness to promote ecological consciousness.

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## Abbreviations

ADF	Augmented Dickey Fuller
CO <sub>2</sub>	Carbon Emissions
DUM	Dummy
GDP	Economic Growth
GLOB	Globalisation
INV	Technological Innovation
PP	Phillip Perron
REC	Renewable Energy Consumption
REM	Remittances
SDGs	Sustainable Development Goals

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